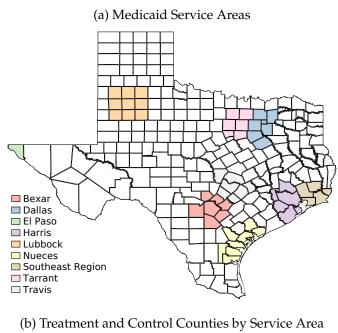
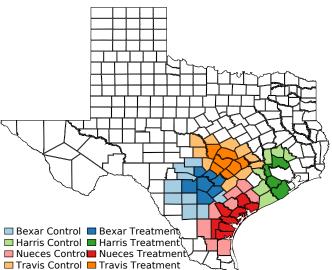
# Healthcare Rationing in Public Insurance Programs: Evidence from Medicaid Tim Layton Nicole Maestas Daniel Prinz Boris Vabson

Online Appendix

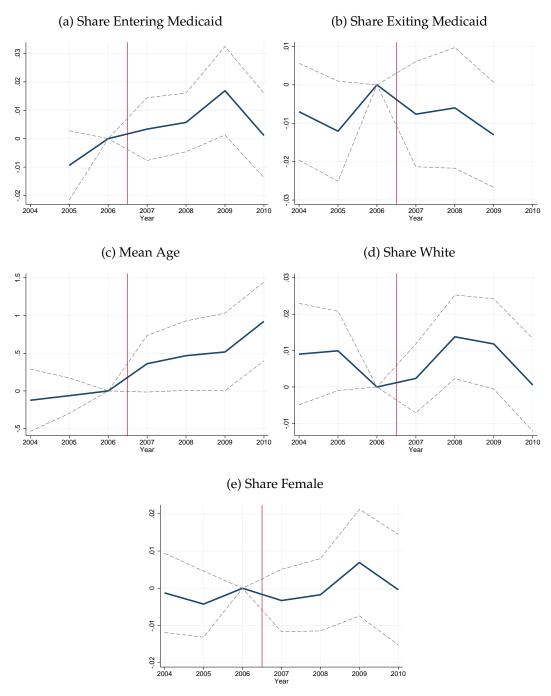
# A Additional Figures and Tables

Appendix Figure A1: Texas Counties



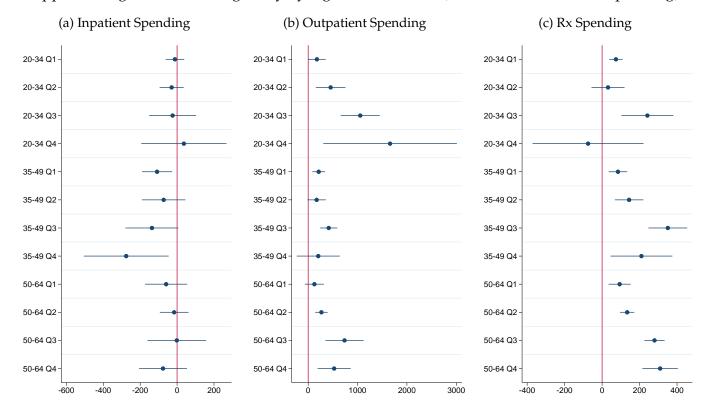


**Note:** Figure shows Medicaid service areas and the treatment and control counties we define based on these service areas. Panel (a) shows all ten of the Medicaid service areas designated by the Texas Health and Human Services Commission in April 2004. Nine service areas are marked by colors, while the 10th service area comprises much of the state and is shown in white. Panel (b) shows treatment and control counties by service area. For more details, see Section 4.



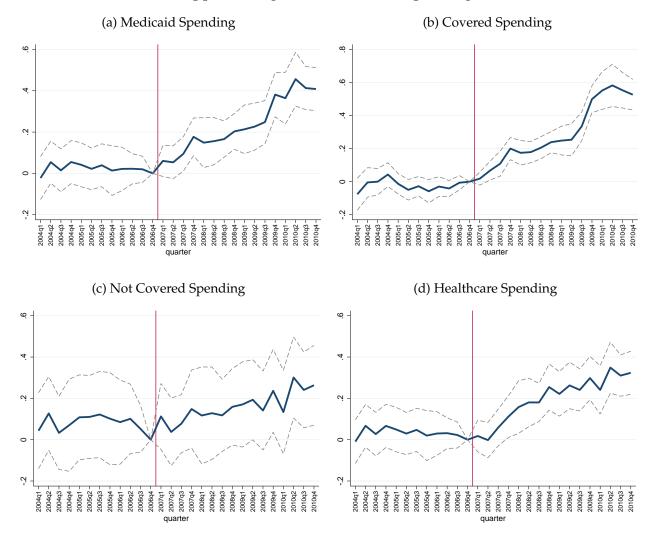
Appendix Figure A2: Composition

**Note:** Figure shows the impact of Medicaid managed care on sample composition measures, including Medicaid entry and exit, mean age, share white, and share female. These coefficients are from estimating the event study difference-in-differences specification in Equation (1). For more details, see Section 4. (N = 168, 658 beneficiary-years.)



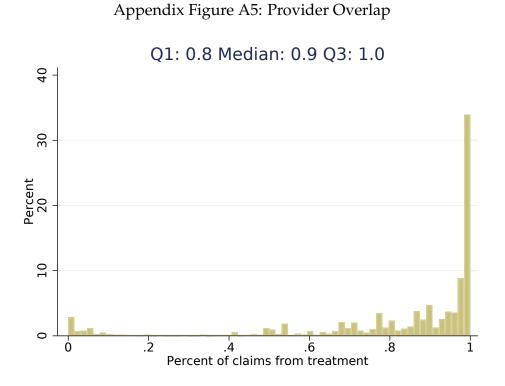
#### Appendix Figure A3: Heterogeneity by Age Health Status (Quartile of Pre-Period Spending)

**Note:** Figure shows the impact of Medicaid managed care on inpatient spending, outpatient spending, and prescription drug spending by age and health status. Health status is measured as quartile of average spending during the pre-period, limiting our sample to beneficiaries for whom this measure can be generated. These coefficients are from estimating our instrumental variable specification separately for each age (20-34, 35-49, 50-64) by quartile of pre-period spending group. For more details, see Section 4. (N = 478,938 beneficiary-quarters.)



#### Appendix Figure A4: Medicaid Spending

**Note:** Figure shows control-treatment differences in Medicaid spending outcomes in percent terms relative to the treatment mean in the pre-period. These coefficients are from estimating the event study difference-in-differences specification in Equation (1), including individual fixed effects. For more details, see Section 4. (N = 643,751 beneficiary-quarters.)



**Note:** Figure shows the 2004 distribution of the percent of claims at a provider that come from patients who live in treatment counties, weighted by each provider's volume. For more details, see Section 5. (N = 14,445 unique provider identifiers.)

	(1)	(2)	(3)	(4)	(5)
	Enter	Exit	Age	Female	White
Treatment $\times$ Post	.011	002	.593	.003	.001
	(.004)	(.004)	(.248)	(.006)	(.006)
Baseline Mean	.108	.054	44.3	.521	.441

Appendix Table A1: Composition

**Note:** Table shows estimates for sample composition measures, including Medicaid entry and exit, mean age, share white, and share female from estimating the pooled version of the reduced form specification in Equation (1). We control for service area by year fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 168, 658 beneficiary-years.)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
	Spending	Any Prescriptions	Days Supply	Spending Branded Drugs	Days Supply Branded Drugs	Spending Generic Drugs	Days Supply Generic Drugs	Spending High Value Drugs	Days Supply High Value Drugs
Treatment $\times$ Post (2007-2008)	80	001	38.8	52.7	13.1	23.6	23.5	29.1	11.6
	(15.2)	(.004)	(5.21)	(13.8)	(2.08)	(2.94)	(3.28)	(8.12)	(2.28)
Treatment $\times$ Post (2009-2010)	210	.002	78.9	154	25.1	46.3	49.1	67.4	25.1
	(25.3)	(.005)	(8.7)	(20.1)	(3.02)	(6.41)	(5.87)	(10.3)	(3.51)
IV Coefficient	199	.001	79.7	142	25.8	47.3	49.2	65.8	24.9
	(22.3)	(.005)	(6.70)	(19)	(2.64)	(5.11)	(4.46)	(10.3)	(3.01)
Baseline Mean	623	.676	187	524	95	84	84	260	77
Percent Change	.32	.001	.427	.272	.271	.564	.583	.253	.325
-	(.036)	(.007)	(.036)	(.036)	(.028)	(.061)	(.053)	(.04)	(.039)
Individual Fixed Effects	Х	Х	Х	Х	Х	Х	Х	Х	Х

Appendix Table A2: Rx Outcomes

**Note:** Table shows reduced form and instrumental variable estimates for prescription drug outcomes. The first row shows estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and the second row shows estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, when the post-period is broken into two separate periods, 2007-2008 and 2009-2010. We estimate the IV coefficient using the two separate treatment × post interaction terms as instruments. We control for service area by quarter fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 643,751 beneficiary-quarters.)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	Anti- Infec- tive Agents	Anti- neo- plastic Agents	Auto- nomic Drugs	Blood Form/ Coagul Agents	Cardio- vascular Agents	Central Nervous System	Gastro- intestinal Drugs	Hormones & Synthetic Subst	Immuno- suppres- sants	Misc Thera- peutic Agents
Treatment $\times$ Post	21.1	2.57	9.61	2.44	20.7	46.1	15.3	9.13	948	-4.33
	(5.77)	(3.51)	(2.57)	(7.15)	(2.92)	(9.02)	(3.52)	(4.02)	(4.11)	(4.25)
IV Coefficient	28.1	3.43	12.8	3.26	27.6	61.6	20.4	12.2	-1.27	-5.79
	(7.76)	(4.39)	(2.93)	(8.94)	(3.79)	(11.5)	(4.01)	(4.77)	(5.16)	(5.33)
Baseline Mean	47.8	11.2	18.1	19.5	66.4	305	34.8	52	19.0	10.6
Percent Change	.588	.306	.71	.167	.416	.202	.587	.234	067	547
Ũ	(.162)	(.391)	(.162)	(.459)	(.057)	(.038)	(.115)	(.091)	(.272)	(.504)
Individual Fixed Effects	Х	Х	Х	Х	Х	Х	Х	Х	Х	Х

#### Appendix Table A3: Therapeutic Classes

#### (a) Spending

#### (b) Any Spending

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	Anti- Infec- tive Agents	Anti- neo- plastic Agents	Auto- nomic Drugs	Blood Form/ Coagul Agents	Cardio- vascular Agents	Central Nervous System	Gastro- intestinal Drugs	Hormones & Synthetic Subst	Immuno- suppres- sants	Misc Thera- peutic Agents
Treatment $\times$ Post	.053	.004	.050	.007	.035	.029	.046	.043	.001	.013
	(.007)	(.001)	(.006)	(.002)	(.007)	(.005)	(.007)	(.007)	(.001)	(.003)
IV Coefficient	.071	.006	.067	.009	.046	.039	.061	.058	.002	.017
	(.007)	(.001)	(.007)	(.003)	(.008)	(.006)	(.008)	(.008)	(.002)	(.003)
Baseline Mean	.213	.013	.156	.05	.266	.504	.144	.231	.008	.026
Percent Change	.332	.457	.426	.184	.174	.077	.423	.251	.201	.646
C	(.032)	(.103)	(.044)	(.061)	(.030)	(.011)	(.055)	(.033)	(.189)	(.115)
Individual Fixed Effects	Х	Х	Х	Х	Х	Х	Х	Х	Х	Х

**Note:** Table shows reduced form and instrumental variable estimates for the most common therapeutic classes. The first row shows estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and the second row shows estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. We control for service area by quarter fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 643,751 beneficiary-quarters.)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Spen	ding		nber of ient Days	Any	Use	ED V	/isits
Treatment $\times$ Post	366		.310		009		048	
	(45)		(.247)		(.004)		(.133)	
Treatment $\times$ Post (2007-2008)		200		.057		013		063
		(36.7)		(.214)		(.004)		(.083)
Treatment $\times$ Post (2009-2010)		501		.537		004		193
		(62.9)		(.330)		(.004)		(.223)
IV Coefficient	489	480	.414	.423	013	010	064	176
	(54.3)	(57.5)	(.312)	(.333)	(.005)	(.005)	(.168)	(.186)
Baseline Mean	1,551	1,551	8.20	8.20	.717	.717	2.17	2.17
Percent Change	.316	.309	.050	.052	018	014	03	081
2	(.035)	(.037)	(.038)	(.041)	(.007)	(.007)	(.077)	(.086)
Individual Fixed Effects	X	X	X	Х	X	X	Х	X

#### Appendix Table A4: Outpatient Outcomes

**Note:** Table shows reduced form and instrumental variable estimates for outpatient outcomes. For each outcome, the first column shows estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. The second column shows reduced form and instrumental variable estimates, when the post-period is broken into two separate periods, 2007-2008 and 2009-2010. In the second and fourth columns we estimate the IV coefficient using the two separate treatment × post interaction terms as instruments. We control for service area by quarter fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 643,751 beneficiary-quarters.)

(a) Ag	ge 20-34, N	lo Comor	bidities		(b) Age 20	-34, 1-3 C	omorbic	lites	(c) Age 20	-34, 4+ C	omorbid	ites
	(1)	(2)	(3)	(4)	(1)	(2)	(3)	(4)	(1)	(2)	(3)	(4)
	Any Inpatient Admissions	Outpatient Spending	Rx Spending	ED Visits	Any Inpatient Admissions	Outpatient Spending	Rx Spending	ED Visits	Any Inpatient Admissions	Outpatient Spending	Rx Spending	ED Visits
Treatment $\times$ Post	002	342	55.2	119	006	433	63.7	471	.017	1,280	-221	.369
<b>T</b> LO (01)	(.002)	(94.6)	(16.4)	(.107)	(.003)	(108)	(51.2)	(.217)	(.028)	(681)	(366)	(1.29)
IV Coefficient	003 (.003)	496 (136)	80.1 (22.4)	173 (.148)	010 (.005)	714 (195)	105 (79.1)	776 (.369)	.033 (.049)	2,393 (1,235)	-413 (624)	.691 (2.27)
Baseline Mean	.007	760	147	.399	.041	2,385	697	1.63	.266	4,651	1,347	7.58
Percent Change	369	.653	.544	432	251	.299	.151	475	.123	.515	306	.091
	(.432)	(.179)	(.152)	(.371)	(.121)	(.082)	(.114)	(.226)	(.185)	(.265)	(.463)	(.300)
(d) Aş	$\frac{35-49, N}{(1)}$			(4)	(e) Age 35	(2)	(3)	(4)	(f) Age 35	(2)		(4)
	Any Inpatient Admissions	Outpatient Spending	Rx Spending	ED Visits	Any Inpatient Admissions	Outpatient Spending	Rx Spending	ED Visits	Any Inpatient Admissions	Outpatient Spending	Rx Spending	ED Visits
Treatment $\times$ Post	008 (.006)	216 (84.4)	39.1 (23.2)	122 (.329)	005 (.004)	186 (54.1)	140 (32.2)	150 (.223)	039 (.014)	432 (190)	429 (121)	.944 (.511)
IV Coefficient	012 (.008)	302 (112)	54.5 (30.2)	170 (.429)	007 (.005)	239 (64)	181 (37.8)	193 (.275)	050 (.017)	552 (229)	549 (140)	1.21 (.626)
Baseline Mean	.006	312	108	.544	.05	1,334	708	2.01	.262	2,923	1,132	6.85
Percent Change	-1.95	.966	.507	313	136	.179	.255	096	189	.189	.485	.176
	(1.32)	(.360)	(.281)	(.788)	(.101)	(.048)	(.053)	(.137)	(.064)	(.078)	(.124)	(.091
(g) Ag	ge 50-64, N	Jo Comor	bidities		(h) Age 50	-64, 1-3 C	omorbic	lites	(i) Age 50	-64, 4+ Co	omorbid	ites
(g) Ag	ge 50-64, N	Jo Comor	bidities	(4)	(h) Age 50	-64, 1-3 C	Comorbic	lites	(i) Age 50	-64, 4+ Co	omorbid	ites
(g) Aş				(4) ED Visits						-		
(g) Ag	(1) Any Inpatient	(2) Outpatient	(3) Rx	ED	(1) Any Inpatient	(2) Outpatient	(3) Rx	(4) ED	(1) Any Inpatient	(2) Outpatient	(3) Rx	(4) ED

#### Appendix Table A5: Outcomes by Age and Pre-Period Health Status (Number of Comoborbidities)

**Note:** Table shows reduced form and instrumental variable estimates for the main outcomes broken down by age and health status. Health status is measured as the average number of comorbidities during the pre-period, limiting our sample to beneficiaries for whom this measure can be generated. The first row shows estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and the second row shows estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. We control for service area by quarter fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 478,938 beneficiary-quarters.)

401

(66.1)

1,081

.371

(.061)

210

(27.9)

606

.347

(.046)

.102

(.156)

1.56

.065

(.100)

-.013

(.011)

.228

-.057

(.046)

375

(63.7)

1,118

.335

(.057)

.169

(.366)

4.93

.034

(.074)

996

(236)

2,841

.350

(.083)

.002

(.005)

.051

.045

(.095)

IV Coefficient

Baseline Mean

Percent Change

157

(99.7)

184

.855

(.543)

-.006

(.006)

006

-1.05

(1.06)

125

(36.6)

98.0

1.28

(.374)

.101

(.185)

.338

.300

(.549)

$ \begin{array}{ c c c c c c c c c c c c c c c c c c c$	visits           7        093           )         (.369)           4        277           )         (1.06)           2         3.27           2        085           a)         (.324)
$\frac{\ln patient}{Admissions} = \frac{Virgatient}{Spending} = \frac{Virgatient}{V$	ing Visits 7093 ) (.369) 4277 ) (1.06) 2 3.27 2085 2) (.324) le 4
$\frac{\text{Admissions}}{\text{Admissions}} \frac{\text{Spending}}{\text{Admissions}} \frac{\text{Spending}}{\text{C}} \frac{\text{Visits}}{\text{C}} \frac{\text{Inpatient}}{\text{Admissions}} \frac{\text{Spending}}{\text{Spending}} \frac{\text{Visits}}{\text{Spending}} \frac{\text{Inpatient}}{\text{Admissions}} \frac{\text{Spending}}{\text{Spending}} \frac{\text{Visits}}{\text{Spending}} \frac{\text{Inpatient}}{\text{Admissions}} \frac{\text{Spending}}{\text{Spending}} \frac{\text{Visits}}{\text{Spending}} \frac{\text{Inpatient}}{\text{Admissions}} \frac{\text{Spending}}{\text{Spending}} \frac{\text{Visits}}{\text{Spending}} \frac{\text{Inpatient}}{\text{Admissions}} \frac{\text{Spending}}{\text{Spending}} \text$	ing Visits 7093 ) (.369) 4277 ) (1.06) 2 3.27 2085 2) (.324) le 4
$ \begin{array}{c c c c c c c c c c c c c c c c c c c $	) (.369) 4277 ) (1.06) 2 3.27 2085 4) (.324) le 4
$\frac{ V Coefficient001 176 73.7235}{(0003) (91.2) (18.4) (.130)} (005 (153) (45 (224) (009) (202) (7.7 (509) (019 190) -10 (009) (202) (7.7 (509) (019 190) -10 (019 190) -10 (019 190) (019 190) -10 (019 190) (019 190) -10 (019 190) (0100 190) (0$	4    277       )     (1.06)       2     3.27       2    085       3)     (.324)
$\frac{(003)  (91.2)  (18.4)  (130)}{\text{Baseline Mean}  .001  60.5  32.5  .299 \\ (2.8)  (1.51)  (2.66)  (1.51)  (2.66)  (1.53)  (1.54)  ($	) (1.06) 2 3.27 2085 4) (.324) le 4
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	085 (.324)
(2.88)       (1.51)       (.566)       (.436)       (.269)       (.448)       (.161)       (.194)       (.175)       (.237)       (.086)       (.273)       (.243)       (.134)       (.32         (e) Age 35-49, Quartile 1       (f) Age 35-49, Quartile 2       (f) Age 35-49, Quartile 2       (g) Age 35-49, Quartile 3       (h) Age 35-49, Quart         (a) Any Inpatient Admissions       Outpatient Spending       Rx Visits       ED Inpatient (.003)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1)       (2)       (3)       (4)       (1) <th< td=""><td>e) (.324) le 4</td></th<>	e) (.324) le 4
(e) Age 35-49, Quartile 1 $(f) Age 35-49, Quartile 2 $ $(f) Age 35-49, Quartile 2 $ $(g) Age 35-49, Quartile 3 $ $(h) Age 35-49, Quart 1 $ $(h)$	le 4
$ \begin{array}{c c c c c c c c c c c c c c c c c c c $	
$ \begin{array}{c c c c c c c c c c c c c c c c c c c $	(4)
Inpatient AdmissionsOutpatient SpendingKxED InpatientInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient AdmissionsOutpatient SpendingKxED SpendingInpatient Admissions	(1)
$ \begin{array}{c c c c c c c c c c c c c c c c c c c $	ED
(.003)         (51.7)         (20.1)         (.268)         (.00709)         (81)         (33.5)         (.265)         (.00881)         (76.1)         (49.6)         (.279)         (.0107)         (195)         (92           IV Coefficient        005         213         84.8        0033        00238         172         144        214        017         416         351         .0533        0219         290         17	
IV Coefficient005 213 84.803300238 172 144214017 416 351 .05330219 290 17	
(004) (66.1) (75.2) (328) (00847) (94.9) (38.9) (31.8) (0105) (89) (52.9) (32.8) (01.48) (26.5) (12	
Baseline Mean         .003         67.7         33.9         .352         .019         339         299         1.35         .074         845         786         2.692         .211         4,201         1,4	
Percent Change         -1.67         3.15         2.50        095        125         .506         .482        159        230         .492         .446         .020        104         .069         .11           (1.25)         (.976)         (.743)         (.933)         (.446)         (.280)         (.130)         (.236)         (.141)         (.105)         (.067)         (.122)         (.070)         (.063)         (.08)	
(1.25) (.976) (.743) (.933) (.446) (.280) (.130) (.236) (.141) (.105) (.067) (.122) (.070) (.063) (.067) (.122)	7) (.102)
(i) Age 50-64, Quartile 1 (j) Age 50-64, Quartile 2 (k) Age 50-64, Quartile 3 (l) Age 50-64, Quartile	e 4
(1)     (2)     (3)     (4)       (1)     (2)     (3)     (4)       (1)     (2)     (3)     (4)	(4)
Any Outpatient Rx ED Any Outpatient Rx ED Any Outpatient Rx ED Any Outpatient Rx	ED
Inpatient Suparchi RX ED Inpatient Suparchi RX	
Treatment × Pos00214 97.9 73 .143 .00243 229 11510700142 620 236 .10800853 555 300	.24
$- \underbrace{(.00475)}_{\text{IV Coefficient}} \underbrace{(82.9)}_{00274} \underbrace{(25.7)}_{126} \underbrace{(.127)}_{.184} \underbrace{(.00431)}_{.00284} \underbrace{(59.6)}_{.68} \underbrace{(18.1)}_{.164} \underbrace{(.164)}_{0164} \underbrace{(.00644)}_{.00644} \underbrace{(170)}_{.170} \underbrace{(25.1)}_{.189} \underbrace{(.00852)}_{.00852} \underbrace{(201)}_{.111} \underbrace{(67.1)}_{.124} \underbrace{(.00644)}_{.125} \underbrace{(170)}_{.125} \underbrace{(.127)}_{.125} (.127$	.313
IV Coefficient        00274         126         93.7         .184         .00284         268         134        125        00168         737         280         .129        0111         724         399           (.00565)         (98.1)         (30)         (.148)         (.00478)         (64.3)         (19.3)         (.183)         (.00731)         (197)         (27.6)         (.215)         (.0106)         (235)         (81.4)	
$\frac{-(30005)}{1000} (5017) (507) (140) (140) (1007) (1507) (1007) (1507) (1007) (1707)$	(425)
Percent Change914 1.832 2.66 .746 .167 .766 .407114023 .84 .373 .063047 .194 .31	
<u>(1.88)</u> (1.43) (.852) (.602) (.281) (.184) (.059) (.167) (.100) (.224) (.037) (.105) (.045) (.063) (.063)	4.95 .063

#### Appendix Table A6: Outcomes by Age and Pre-Period Health Status (Quartile of Pre-Period Spending)

**Note:** Table shows reduced form and instrumental variable estimates for the main outcomes broken down by age and health status. Health status is measured as quartile of average spending during the pre-period, limiting our sample to beneficiaries for whom this measure can be generated. The first row shows estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and the second row shows estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. We control for service area by quarter fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 478,938 beneficiary-quarters.)

#### Appendix Table A7: Main Outcomes (Balanced Panel)

 $(a) 2005_{2008}$ 

	(a) 200.	5-2008		
	(1)	(2)	(3)	(4)
	Healthcare	Rx	Outpatient	Inpatient
	Spending	Spending	Spending	Spending
Treatment $\times$ Post	305	100	251	-46
	(67)	(19)	(43)	(27)
IV Coefficient	400	132	329	-61
	(81)	(21)	(54)	(34)
Baseline Mean	2,803	757	1,680	367
Percent Change	.143	.174	.196	166
U U	(.029)	(.028)	(.032)	(.094)
Individual Fixed Effects	Х	Х	Х	Х

		(b) 2	2004-20	010				
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Healthca	re Spending	Rx Spe	ending	Outpatie	nt Spending	Inpatien	t Spending
Treatment $\times$ Post	493		130		395		-32	
	(76)		(25)		(55)		(18)	
Treatment $\times$ Post (2007-2008)		271		94		203		-26
		(68)		(20)		(50)		(20)
Treatment $\times$ Post (2009-2010)		660		160		522		-21
		(83)		(30)		(65)		(23)
IV Coefficient	643	628	169	169	516	489	-42	-30
	(92)	(89)	(29)	(28)	(71)	(70)	(24)	(23)

733

.231

(.040)

Х

2,685

.240

(.034)

Х

2,685

.234

(.033)

Х

733

.231

(.039)

Х

1,644

.314

(.043)

Х

1,644

.297

(.043)

Х

309

-.135

(.077)

Х

309

-.098

(.073)

Х

**Baseline** Mean

Percent Change

Individual Fixed Effects

**Note:** Table shows reduced form and instrumental variable estimates for the main outcomes in a balanced panel. Panel (a) shows a shorter panel, for 2005-2008 and Panel (b) shows all years, 2004-2010. In Panel (b), for each outcome, the first and third columns show estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. The second and fourth columns show reduced form and instrumental variable estimates, when the post-period is broken into two separate periods, 2007-2008 and 2009-2010. In the second and fourth columns we estimate the IV coefficient using the two separate treatment  $\times$  post interaction terms as instruments. We control for service area by quarter fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 172,496beneficiary-quarters in Panel (a). N = 187,684 beneficiary-quarters in Panel (b).)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)
	H	Iealthcare	e Spendir	ıg		Rx Spe	ending		0	utpatien	t Spendir	ıg	]	npatient	Spending	g
Treatment $\times$ Post	324		444		109		118		255		314		-39.7		11.4	
	(118)		(181)		(47.7)		(70.5)		(87.1)		(91.7)		(84.8)		(90.3)	
Treatment $\times$ Post (2007-2008)		236		278		53.2		40.7		184		222		541		15.4
		(121)		(127)		(36.7)		(28)		(79.2)		(60.8)		(94.2)		(78.9)
Treatment $\times$ Post (2009-2010)		411		566		182		174		269		361		-39.4		31.7
		(115)		(271)		(76.9)		(108)		(101)		(154)		(84.6)		(108)
IV Coefficient	409	407	651	638	137	150	173	167	322	283	461	436	-50.2	-26	16.8	35.7
	(131)	(127)	(274)	(285)	(55.7)	(61.8)	(100)	(97)	(94.6)	(97.1)	(145)	(155)	(98.5)	(95.8)	(129)	(123)
Baseline Mean	2,711	2,711	2,711	2,711	700	700	700	700	1,388	1,388	1,388	1,388	623	623	623	623
Percent Change	.151	.150	.240	.235	.196	.214	.248	.238	.232	.204	.332	.314	081	042	.027	.057
5	(.048)	(.047)	(.101)	(.105)	(.079)	(.088)	(.143)	(.138)	(.068)	(.070)	(.105)	(.112)	(.158)	(.154)	(.206)	(.198)
Individual Fixed Effects	X	X			X	X			X	X			X	X		

#### Appendix Table A8: Main Outcomes (Bexar Service Area)

**Note:** Table shows reduced form and instrumental variable estimates for the main outcomes in Texas's Bexar Service Area. For each outcome, the first and third columns show estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. The second and fourth columns show reduced form and instrumental variable estimates, when the post-period is broken into two separate periods, 2007-2008 and 2009-2010. In the second and fourth columns we estimate the IV coefficient using the two separate treatment × post interaction terms as instruments. Standard errors are clustered at the county level. For more details, see Section 4. (N = 96,678 beneficiary-quarters.)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)
	H	lealthcare	e Spendir	ng		Rx Spe	ending		С	utpatien	t Spendir	ng	I	Inpatient	Spending	g
Teatment $\times$ Post	502		464		138		146		436		380		-70.9		-62.6	
	(135)		(126)		(29.6)		(41.5)		(89.5)		(79.5)		(50.3)		(50.6)	
Treatment $\times$ Post (2007-2008)		203		238		80.7		102		228		220		-105		-83.5
		(119)		(125)		(21.3)		(37.6)		(85.2)		(85)		(58.6)		(57.9)
Treatment $\times$ Post (2009-2010)		865		640		211		180		672		499		-17.9		-40
		(165)		(151)		(39.9)		(52.5)		(121)		(91.4)		(73.1)		(59)
IV Coefficient	665	742	736	745	183	198	232	233	577	616	603	604	-93.9	-72.8	-99.4	-91.8
	(175)	(168)	(203)	(209)	(34.6)	(34.3)	(66.5)	(72.1)	(130)	(145)	(132)	(138)	(60.7)	(48)	(76)	(77.3)
Baseline Mean	3,042	3,042	3,042	3,042	637	637	637	637	1,631	1,631	1,631	1,631	773	773	773	773
Percent Change	.219	.244	.242	.245	.286	.311	.365	.366	.354	.378	.369	.37	121	094	129	119
U	(.057)	(.055)	(.067)	(.069)	(.054)	(.054)	(.104)	(.113)	(.08)	(.089)	(.081)	(.085)	(.079)	(.062)	(.098)	(.100)
Individual Fixed Effects	Х	X			X	X			X	X			X	X		

#### Appendix Table A9: Main Outcomes (Harris Service Area)

**Note:** Table shows reduced form and instrumental variable estimates for the main outcomes in Texas's Harris Service Area. For each outcome, the first and third columns show estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. The second and fourth columns show reduced form and instrumental variable estimates, when the post-period is broken into two separate periods, 2007-2008 and 2009-2010. In the second and fourth columns we estimate the IV coefficient using the two separate treatment × post interaction terms as instruments. Standard errors are clustered at the county level. For more details, see Section 4. (N = 179,623 beneficiary-quarters.)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)
	H	Iealthcare	e Spendir	ng		Rx Sp	ending		C	Outpatien	t Spendir	ng	I	npatient	Spending	g
Treatment $\times$ Post	664		534		220		181		499		446		-53.8		-93.3	
	(170)		(151)		(46.9)		(42.8)		(124)		(119)		(42.9)		(51.1)	
Treatment $\times$ Post (2007-2008)		342		402		146		114		262		310		-65.7		-21.8
		(155)		(132)		(38.3)		(39.1)		(88.5)		(90)		(66.2)		(51)
Treatment $\times$ Post (2009-2010)		917		600		298		225		674		522		-55		-147
		(227)		(183)		(54.9)		(46.2)		(189)		(146)		(39.5)		(70)
IV Coefficient	766	719	715	678	253	253	243	232	575	535	597	565	-62.1	-68.6	-125	-119
	(180)	(184)	(194)	(183)	(49.9)	(47.5)	(54.7)	(53)	(131)	(139)	(153)	(146)	(45.7)	(50.8)	(66.3)	(62.2)
Baseline Mean	2,898	2,898	2,898	2,898	640	640	640	640	1,590	1,590	1,590	1,590	668	668	668	668
Percent Change	.264	.248	.247	.234	.396	.396	.379	.363	.362	.336	.376	.356	093	103	187	178
U U	(.062)	(.064)	(.067)	(.063)	(.078)	(.074)	(.085)	(.083)	(.082)	(.087)	(.096)	(.092)	(.068)	(.076)	(.099)	(.093)
Individual Fixed Effects	X	X			X	X			X	X			X	X		

Appendix Table A10: Main Outcomes (Nueces Service Area)
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**Note:** Table shows reduced form and instrumental variable estimates for the main outcomes in Texas's Nueces Service Area. For each outcome, the first and third columns show estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. The second and fourth columns show reduced form and instrumental variable estimates, when the post-period is broken into two separate periods, 2007-2008 and 2009-2010. In the second and fourth columns we estimate the IV coefficient using the two separate treatment × post interaction terms as instruments. Standard errors are clustered at the county level. For more details, see Section 4. (N = 115,986 beneficiary-quarters.)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)
	H	Healthcare Spending			Rx Spending			Outpatient Spending			Inpatient Spending					
Treatment $\times$ Post	391		335		119		119		320		208		-48		7.83	
	(87.1)		(65.5)		(22.1)		(19.2)		(55.3)		(36.9)		(36.2)		(22.2)	
Treatment $\times$ Post (2007-2008)		223		225		60.7		72.2		162		126		0785		27
		(55.2)		(80.8)		(18.3)		(17.8)		(55.6)		(51.5)		(21.6)		(26.4)
Treatment $\times$ Post (2009-2010)		550		372		183		149		439		222		-71		.543
		(109)		(85)		(27.3)		(20.9)		(64.1)		(49.2)		(43.1)		(30.8)
IV Coefficient	587	611	587	551	179	195	209	209	480	478	364	324	-72	-62.4	13.7	18.6
	(119)	(103)	(111)	(130)	(29.6)	(29.7)	(32.6)	(32.2)	(76)	(71.3)	(62.2)	(66.4)	(49.4)	(41.3)	(37.6)	(48.4)
Baseline Mean	2,621	2,621	2,621	2,621	573	573	573	573	1,517	1,517	1,517	1,517	530	530	530	530
Percent Change	.224	.233	.224	.21	.312	.34	.365	.365	.316	.315	.24	.213	136	118	.026	.035
U	(.045)	(.039)	(.043)	(.049)	(.052)	(.052)	(.057)	(.056)	(.05)	(.047)	(.041)	(.044)	(.093)	(.078)	(.071)	(.091)
Individual Fixed Effects	X	X			X	X			X	X			X	X		

Appendix Table A11: Main Outcomes (Travis Service Area)

Note: Table shows reduced form and instrumental variable estimates for the main outcomes in Texas's Travis Service Area. For each outcome, the first and third columns show estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. The second and fourth columns show reduced form and instrumental variable estimates, when the post-period is broken into two separate periods, 2007-2008 and 2009-2010. In the second and fourth columns we estimate the IV coefficient using the two separate treatment × post interaction terms as instruments. Standard errors are clustered at the county level. For more details, see Section 4. (N = 251,464 beneficiary-quarters.)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	Analg/ Antipyr, Nonstr/ Antiinflm	Analg/ Antipyr, Opiate Agonists	Analg/ Antipyr, NEC	Anti- convulsant, Benzo- diazepine	Anti- conv, Hydantoin Derivative	Anticonv, Misc	Psych other, Anti- depres- sants	Psychother, Tranq/ Anti- psychotic	ASH, Benzo- diazepines	ASH, NEC
Treatment $\times$ Post	2.15	2.62	1.1	.27	.351	7.51	5.72	15.8	1.12	2.91
	(.475)	(4.49)	(.291)	(.0928)	(.2)	(3.33)	(2.55)	(5.23)	(.542)	(.715)
IV Coefficient	2.88	3.5	1.47	.361	.469	10	7.64	21.1	1.5	3.89
	(.612)	(5.62)	(.379)	(.113)	(.252)	(4.34)	(3.05)	(6.66)	(.678)	(.896)
Baseline Mean	9.77	25.1	2.50	1.15	2.64	75.8	39.4	125.1	4.93	8.80
Percent Change	.295	.139	.587	.315	.177	.132	.194	.169	.303	.442
0	(.063)	(.224)	(.152)	(.099)	(.095)	(.057)	(.077)	(.053)	(.137)	(.102)

#### Appendix Table A12: Central Nervous System Classes

(a) Spending

#### (b) Any Spending

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	Analg/ Antipyr, Nonstr/ Antiinflm	Analg/ Antipyr, Opiate Agonists	Analg/ Antipyr, NEC	Anti- convulsant, Benzo- diazepine	Anticonv, Hydantoin Derivative	Anticonv, Misc	Psychother, Anti- depressants	Psychother, Tranq/ Anti- psychotic	ASH, Benzo- diazepines	ASH, NEC
Treatment $\times$ Post	.042	.025	.017	.004	.004	.015	.037	.018	.020	.025
	(.005)	(.007)	(.003)	(.001)	(.002)	(.004)	(.005)	(.003)	(.005)	(.002)
IV Coefficient	.056	.034	.023	.006	.006	.020	.049	.023	.027	.034
	(.005)	(.008)	(.004)	(.002)	(.002)	(.005)	(.006)	(.004)	(.005)	(.003)
Baseline Mean	.085	.19	.04	.031	.029	.131	.181	.125	.099	.054
Percent Change	.664	.178	.571	.18	.201	.151	.271	.187	.276	.629
-	(.064)	(.043)	(.096)	(.061)	(.079)	(.038)	(.031)	(.030)	(.055)	(.056

**Note:** Table shows reduced form and instrumental variable estimates for the most common subclasses of the central nervous system therapeutic class. The first row shows estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and the second row shows estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. We control for service area by quarter fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 643,751 beneficiary-quarters.)

	(4)	(2)	(2)	(1)	_	-		(0)	(0)	(1.0)
	(1) NEC	(2) ACE Inhibitors	(3) Cardiac Glycosides	(4) Anti- arrhythmic Agents	(5) Alpha-Beta Blockers	(6) Beta Blockers	(7) Calcium Channel	(8) Anti- hyperlipidemic Drugs,	(9) Hypotensive Agents, NEC	(10) Vasodilating Agents, NEC
Treatment $\times$ Post	1.32	1.43	.0957	.101	.0219	1.81	1.43	NEC 10.5	.961	3.03
fredulient × 1 05t	(.866)	(.295)	(.025)	(.079)	(.033)	(.490)	(.599)	(2.16)	(.371)	(2.17)
IV Coefficient	1.77	1.91	.128	.135	.0293	2.42	1.91	14.1	1.28	4.04
	(1.07)	(.388)	(.031)	(.0991)	(.0416)	(.596)	(.765)	(2.73)	(.45)	(2.75)
Baseline Mean	8.17	4.85	.169	.314	.153	5.79	10.6	30.3	2.02	3.44
Percent Change	.216	.394	.756	.431	.192	.418	.181	.464	.636	1.18
0	(.130)	(.080)	(.184)	(.316)	(.272)	(.103)	(.072)	(.090)	(.223)	(.800)
	(1)	(2)	(3)	(4)	) Any Spen (5)	(6)	(7)	(8)	(9)	(10)
	(1) NEC	(2) ACE Inhibitors	(3) Cardiac Glycosides	(4) Anti- arrhythmic Agents	(5) Alpha-Beta Blockers	(6) Beta Blockers	(7) Calcium Channel	Anti- hyperlipidemic Drugs,	(9) Hypotensive Agents, NEC	(10) Vasodilating Agents, NEC
			224	0				NEC		
Treatment $\times$ Post	.008	.027	.004	.0002	.0004	.024	.014	.025	.008	.008
	(.004)	(.004)	(.001)	(.0005)	(.0005)	(.004)	(.002)	(.007)	(.002)	(.001)
IV Coefficient	.010	.036	.005	.0003	.0006	.032	.019	.034	.010	.011
Baseline Mean	(.004)	(.004)	(.001)	(.0007)	(.0006)	(.005)	(.003)	(.009)	(.002)	(.002)
	.047	.079	.009	.002	.002	.076	.067	.109	.022	.017
Percent Change	.222 (.095)	.454 (.056)	.571 (.158)	.129 (.326)	.279 (.310)	.421 (.061)	.284 (.042)	.309 (.080)	.463 (.101)	.663 (.089)

#### Appendix Table A13: Cardiovascular Classes

(a) Spending

**Note:** Table shows reduced form and instrumental variable estimates for the most common subclasses of the cardiovascular agents therapeutic class. The first row shows estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and the second row shows estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. We control for service area by quarter fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 643,751 beneficiary-quarters.)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	Adrenals & Comb, NEC	Contra- ceptive, Oral Comb, NEC	Estrogens & Comb, NEC	Anti- diabetic Agents, Insulins	Anti- diabetic Agents, Sulfo nylureas	Anti- diabetic Agents, Misc	Para- thyroid Hor- mones, NEC	Pituitary Hor- mones, NEC	Pro- gestins, NEC	Thy /Antithy, Thyroid/ Hor- mones
Treatment $\times$ Post	5.43	0425	.616	139	1.21	2.49	.106	-2.11	.0266	.696
	(1.37)	(.168)	(.198)	(2.42)	(.284)	(1.41)	(.098)	(.989)	(.0545)	(.0994)
IV Coefficient	7.26	0567	.823	185	1.61	3.33	.141	-2.82	.0355	.93
	(1.58)	(.211)	(.246)	(3.03)	(.343)	(1.76)	(.124)	(1.24)	(.0678)	(.116)
Baseline Mean	7.60	1.74	2.00	13.3	4.39	18.5	.356	1.77	.146	1.64
Percent Change	.955	033	.411	014	.367	.18	.397	-1.594	.243	.566
0	(.208)	(.121)	(.123)	(.227)	(.078)	(.095)	(.347)	(.700)	(.465)	(.071)

Appendix Table A14: Hormones Classes

(a) Spending

<b>/1</b> \		0 1.	
(h)	$\Delta \mathbf{n} \mathbf{v}$	Spondit	າຕ
101	<b>MIIV</b>	Spendir	12
( )	5		0

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	Adrenals & Comb, NEC	Contra- ceptive, Oral Comb, NEC	Estrogens & Comb, NEC	Anti- diabetic Agents, Insulins	Anti- diabetic Agents, Sulfo nylureas	Anti- diabetic Agents, Misc	Para- thyroid Hor- mones, NEC	Pituitary Hor- mones, NEC	Pro- gestins, NEC	Thy /Antithy, Thyroid/ Hor- mones
Treatment $\times$ Post	.0318	0007	.004	.001	.015	.019	.0007	0003	.0003	.018
	(.004)	(.001)	(.002)	(.003)	(.003)	(.003)	(.0005)	(.0003)	(.0007)	(.004)
IV Coefficient	.042	001	.005	.002	.020	.025	.001	0004	.0004	.024
	(.005)	(.002)	(.002)	(.004)	(.004)	(.004)	(.0006)	(.0004)	(.0009)	(.004)
Baseline Mean	.049	.018	.02	.047	.047	.07	.002	.002	.003	.048
Percent Change	.866	054	.244	.041	.418	.355	.501	191	.141	.503
	(.094)	(.101)	(.099)	(.075)	(.075)	(.056)	(.289)	(.192)	(.307)	(.085)

**Note:** Table shows reduced form and instrumental variable estimates for the most common subclasses of the hormones and synthetic substances therapeutic class. The first row shows estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and the second row shows estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. We control for service area by quarter fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 643,751 beneficiary-quarters.)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	Infectious and parasitic Diseases	Neo- plasms	Endocrine; nutritional; and metabolic Diseases and immunity disorders	Mental Illness	Diseases of the circu- latory system	Diseases of the respi- ratory system	Diseases of the diges- tive system	Diseases of the genito- urinary system	Diseases of the skin and subcutaneous tissue	In- jury and poiso- ning
Treatment $\times$ Post	-9.14	.116	-6.04	-27.9	7.21	-5.89	-4.14	17	-1.54	.417
	(5.55)	(5.52)	(2.9)	(6.84)	(8.14)	(8.9)	(5.16)	(3.79)	(2.05)	(11.2)
IV Coefficient	-12.2	.155	-8.07	-37.2	9.64	-7.87	-5.53	227	-2.06	.557
	(7.09)	(6.91)	(3.65)	(8.01)	(10.1)	(11.2)	(6.48)	(4.75)	(2.56)	(14)
Baseline Mean	41.5	60.9	31.1	47.2	130.6	79.8	71.7	22.7	19.1	67.6
Percent Change	294	.003	260	789	.074	099	077	01	108	.008
U	(.171)	(.113)	(.117)	(.170)	(.078)	(.141)	(.090)	(.209)	(.134)	(.207)

#### Appendix Table A15: Inpatient Spending on the Top 10 CCS Categories

(a) Spending

#### (b) Any Spending

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	Infectious and parasitic	Neo- plasms	Endocrine; nutritional; and metabolic	Mental Illness	Diseases of the circu-	Diseases of the respi-	Diseases of the diges-	Diseases of the genito-	Diseases of the skin and	In- jury and poiso-
	Diseases	plasifis	Diseases and immunity disorders	miless	latory system	ratory system	tive system	urinary system	subcutaneous tissue	ning
Treatment $\times$ Post	.0001	.0003	002	004	.0006	0007	0008	00003	0004	0.000008
	(.0006)	(.0003)	(.0006)	(.001)	(.0009)	(.001)	(.0006)	(.0005)	(.0004)	(.0006)
IV Coefficient	.0002	.0004	002	006	.0008	001	001	00004	0005	.00001
	(.0007)	(.0003)	(.0007)	(.001)	(.001)	(.001)	(.0008)	(.0007)	(.0005)	(.0008)
Baseline Mean	.004	.005	.006	.009	.016	.01	.01	.005	.004	.007
Percent Change	.047	.086	402	637	.048	098	105	009	134	.002
	(.173)	(.064)	(.123)	(.127)	(.070)	(.122)	(.077)	(.139)	(.126)	(.110)

**Note:** Table shows reduced form and instrumental variable estimates for the top 10 most common Clinical Classification Software (CCS) groups of diagnoses. The first row shows estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and the second row shows estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. We control for service area by quarter fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 643,751 beneficiary-quarters.)

	(1)	(2)	(3)	(4)	(5)	(6)
	Adjusted	Adjusted	Adjusted	Adjusted	Adjusted	Adjusted
	Spending	Spending	Spending	Spending	Spending	Spending
	Total	Total	No Rx Match	No Rx Match	Rx Match	Rx Match
Treatment $\times$ Post	265		172		93	
	(43)		(42)		(7)	
Treatment $\times$ Post (2007-2008)		123		71		52
		(36)		(34)		(5)
Treatment $\times$ Post (2009-2010)		374		240		134
		(60)		(59)		(10)
IV Coefficient	354	343	230	215	124	128
	(53)	(57)	(52)	(56)	(7)	(8)
Baseline Mean	1,549	1,549	1,396	1,396	153	153
Percent Change	.229	.221	.165	.154	.808	.834
Percent Change (SE)	(.034)	(.037)	(.037)	(.04)	(.046)	(.053)
Individual Fixed Effects	Х	Х	Х	Х	Х	Х

Appendix Table A16: Outpatient Spending and Prescriptions

**Notes:** Table shows reduced form and instrumental variable estimates for outpatient spending. Columns (1) and (2) show results for total outpatient days, columns (3) and (4) show results for outpatient days when no prescription is written, and columns (5) and (6) show results for outpatient days when a prescription is written. For each outcome, the first column shows estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. The second column shows reduced form and instrumental variable estimates, when the post-period is broken into two separate periods, 2007-2008 and 2009-2010. In the second and fourth columns we estimate the IV coefficient using the two separate treatment × post interaction terms as instruments. We control for service area by quarter fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 643,751 beneficiary-quarters.)

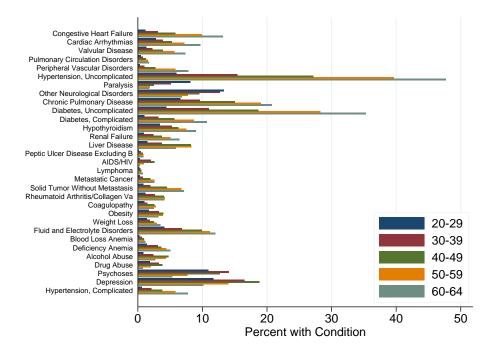
## **B** Medicaid and SSI

The majority of adults with disabilities enrolled in Medicaid are eligible for Medicaid due to their enrollment in the Supplemental Security Income (SSI) program. The SSI program is one of the largest welfare programs in the United States, providing monthly payments to more than 8.2 million disabled or elderly beneficiaries in December 2017. Of these, 4.8 million were adults with disabilities between the ages of 18 and 64, and the average monthly payment for this group was \$564.34 (Social Security Administration, 2018). For the non-elderly, eligibility for SSI is based on medical criteria as well as income and asset tests. SSI has the same medical eligibility criteria for adults as the Social Security Disability Insurance (SSDI) program, but does not share SSDI's work history requirements. Approximately one-third of SSI beneficiaries are also enrolled in the SSDI program because they have sufficient prior work history for SSDI but low enough income to quality for SSI as well.

SSI beneficiaries are categorically eligible for Medicaid in most states, meaning that they can enroll in Medicaid without having to apply separately.<sup>1</sup> SSDI beneficiaries are categorically eligible for Medicare, making those SSI beneficiaries who also qualify for SSDI dually eligible for both Medicaid and Medicare. In Texas (as well as in most other states where private provision has been rolled out to adults with disabilities), dually eligible beneficiaries were excluded from the shift to private managed care plans. Thus, our analysis focuses on the two-thirds of SSI beneficiaries who were not also eligible for SSDI.

Cash benefit payments for disabled SSI beneficiaries quadrupled between 1990 (\$12.2 billion) and 2017 (\$48.2 billion) (Social Security Administration, 2018); however, these expenditures are dwarfed by Medicaid expenditures for this population— \$187 billion in 2014 (Kaiser Family Foundation, 2014a). Adults with disabilities are the most expensive group in Medicaid, with per capita spending equal to \$16,859 in 2014, almost five times higher than per capita spending for adults without disabilities (\$3,278) (Kaiser Family Foundation, 2014b). One reason for this higher spending profile is that SSI beneficiaries disproportionately qualify for the program due to mental disorders: 57.4% of SSI beneficiaries qualified for SSI due to a mental disorder, with intellectual disabilities (19% of beneficiaries who qualify due to a mental disorder) being the largest sub-category, followed by mood disorders (16%), and schizophrenic and other psychotic disorders (8.9%). After mental disorders, the next largest categories are musculoskeletal disabilities (13%) and nervous system disabilities (7.7%) (Duggan, Kearney and Rennane, 2015). Thus, this population differs greatly from the average non-disabled Medicaid beneficiary and even from the typical SSDI beneficiary, in its high prevalence of mental illness, indicating a high level of need for mental healthcare. Also contributing to high costs is the fact that individuals in this population suffer from multiple serious health problems. This suggests that (1) the tools of managed care may be particularly effective for this group and (2) strict rationing in public FFS Medicaid programs (such as Texas's three-drug cap) is likely to be binding for this group and could potentially have detrimental (and observable) health effects. In Appendix Figure B1, we present the prevalence of a variety of clinical conditions by age.

<sup>&</sup>lt;sup>1</sup>10 states have stricter criteria, while 7 states require a separate application but have no additional criteria. In Texas, Medicaid eligibility is automatic for SSI beneficiaries



#### Appendix Figure B1: Comorbidites By Age

**Note:** Figure shows the distribution of the 31 components of the Elixhauser Comorbidity Index by age group. (N = 189,935 beneficiary-years.)

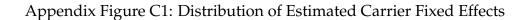
### C Price Variation Across Managed Care Plan Carriers

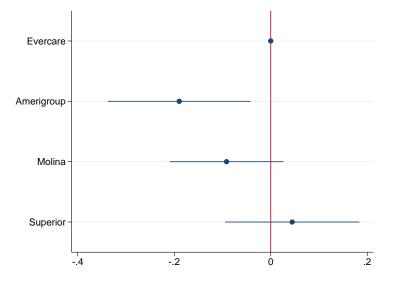
The Texas outpatient data includes information on actual cost amounts for both the public *and the private* programs. Specifically, the data contains the negotiated amounts actually paid to providers by the public or private plans at the claim-line level. These actual provider payment amounts are available for all public Medicaid claims, as well as for about 80% of all private Medicaid plan claims. In this section, we examine the variation in the observed rates across the four carriers (EverCare, Amerigroup, Molina, and Superior HealthPlan) that the state contracted with in 2009 and 2010, the years during which Medicaid managed care was already rolled out and for which we have a crosswalk from plan identifiers in the data to carrier names.

We examine the sample of outpatient managed care claims for 2009 and 2010 for which the payment from plans to providers is available, which comprises 80% of managed care claims. Furthermore, we restrict to the 99.9% of claims that are associated with a plan and carrier that we observe as a plan contracted by the state of Texas in the actuarial reports. To decrease noise in prices, we exclude claims that have a quantity of service provided different from 1 and claims that have a procedure modifier code. On this final sample of outpatient managed care claims, we estimate the following regression:

$$log(p_{ichpt}) = \gamma_c + \delta_h + \psi_p + \tau_t + \varepsilon_{ichpt}$$
(1)

where *i* indexes individuals, *c* indexes carriers, *h* indexes providers, *p* indexes procedures, and *t* indexes time;  $\gamma_c$  is a set of carrier fixed effects,  $\delta_h$  is a set of provider fixed effects,  $\psi_p$  is a set of procedure fixed effects, and  $\tau_t$  is a set of year fixed effects. We define procedures as unique combinations of procedure codes and place of service codes. We cluster standard errors at the carrier level. Figure C1 shows the estimated carrier fixed effects. Relative to the omitted carrier, Evercare, Amerigroup has on average 19% lower prices, Molina has on average 9% lower prices, and Superior has on average 4% higher prices, though these latter two differences are not statistically significant.





**Note:** Figure shows the distribution of estimated carrier fixed effects from estimating Equation (1). Standard errors are clustered at the carrier level. (N = 2,265,378 outpatient managed care claims.)

### D Border Zip Code Analysis

In this appendix we replicate our main results limiting to border zip codes. The motivation for this analysis is that one might be concerned that our treatment counties are more urban than control counties and urban and rural counties may have been differentially impacted by potential shocks that occurred around the time of our treatment (February 2007). Focusing on border zip codes may make control and treatment counties even more similar. Border zip codes are defined as zip codes in a control county that are within 25 miles of a treatment county and zip codes in a treatment county that are within 25 miles of a control county. Distance is measured as great-circle distance calculated using the Haversine formula based on internal points in zip codes.<sup>2</sup>

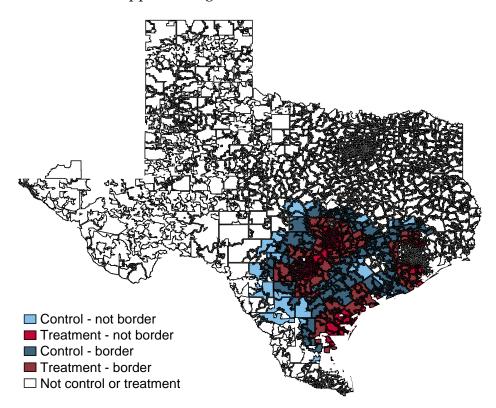
Appendix Figure D1 shows a map of zipcodes in Texas. Control and treatment counties are highlighted in shades of blue and shades of red, respectively, separating border and non-border zipcodes.

Appendix Table D1 replicates Table 1, limiting to the border zipcodes.

Appendix Table D2 replicates Table 2, limiting to the border zipcodes. For each primary outcome (healthcare spending, inpatient spending, drug spending, outpatient spending), we report coefficients from four regressions. The first two regressions include individual fixed effects while the second two regressions do not. The first and third regressions include an interaction between an indicator for residing in a treatment county ("Treatment") and an indicator for the quarter being after February 2007 ("Post"), the month in which mandated enrollment in private Medicaid plans began in Texas. The second and fourth columns break the "post" period into two periods, an "early-post" period (2007-2008) and a "late-post" period (2009-2010). For each regression specification we report both reduced form and IV coefficients. Reduced form coefficients should be interpreted as the effect of a county-level private-plan enrollment mandated enrollment. IV coefficients should be interpreted as the effect of a step as the difference in the outcome in the public Medicaid program vs. in a private plan for the average beneficiary who was induced by the mandate to enroll in a private plan. We highlight that our main results remain quite similar on this restricted sample.

<sup>&</sup>lt;sup>2</sup>Files with distances between zip codes are available at https://www.nber.org/data/ zip-code-distance-database.html.

Appendix Figure D1: Texas ZIP Codes



**Note:** Figure shows the map of zip codes in Texas. For our analysis of zip codes we classify zip codes within the control and treatment counties into border and non-border zip codes. Border zip codes are zip codes in control counties within 25 miles of a treatment zip code and zip codes in treatment counties within 25 miles of a control zip code. Not border zip codes are all the other zip codes in control and treatment counties. Distance is measured as great-circle distance calculated using the Haversine formula based on internal points in zip codes.

	Control	Treatment
Average quarterly healthcare spending 2004	2,662	2,912
Average quarterly inpatient spending 2004	722	745
Average quarterly outpatient spending 2004	1,360	1,531
Average quarterly Rx spending 2004	580	636
Age 20 to 24	.095	.111
Age 25 to 29	.078	.083
Age 30 to 34	.080	.080
Age 35 to 39	.084	.086
Age 40 to 44	.099	.111
Age 45 to 49	.130	.124
Age 50 to 54	.142	.130
Age 55 to 59	.162	.149
Age 60 to 64	.130	.126
Female	.578	.560
Male	.422	.441
Heart Disease	.339	.327
Diabetes	.198	.214
HIV/AIDS	.010	.090
Cancer	.052	.052
Rheumatoid Arthritis	.036	.036
Obesity	.028	.029
Substance Use	.052	.051
Mental Illness	.212	.201
N recipients Jan 2004	6,092	8,710
N recipients Dec 2010	7,191	11,548
N pre-period recipient months	234,355	339,409
N post-period recipient months	315,790	503,044

Appendix Table D1: Summary Statistics (Zipcodes)

**Note:** Table shows summary statistics for border zipcodes in control and treatment counties. For our analysis of zipcodes we classify zipcodes within the control and treatment counties into border and not border zipcodes. Border zipcodes are zipcodes in control counties within 25 miles of a treatment zipcode and zipcodes in treatment counties within 25 miles of a control zipcode. Not border zipcodes are all the other zipcodes in control and treatment counties. Distance is measured as great-circle distance calculated using the Haversine formula based on internal points in zipcodes.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)
	H	lealthcare	e Spendir	ıg		Rx Spe	ending		0	utpatient	t Spendir	ıg	]	npatient	Spending	g
Treatment $\times$ Post	477		516		151		145		391		388		-65.2		-17.2	
	(81.8)		(70.3)		(28.7)		(23.1)		(52)		(49.1)		(32.3)		(26.7)	
Treatment $\times$ Post (2007-2008)		241		327		89.9		82		200		254		-48.4		-8.5
		(72.1)		(58.6)		(23.1)		(24)		(50)		(41.6)		(35.6)		(27.2)
Treatment $\times$ Post (2009-2010)		718		645		219		187		556		469		-56.8		-10.9
		(93.6)		(93.5)		(38.9)		(27.3)		(65.9)		(63.7)		(33)		(34)
IV Coefficient	611	622	778	762	194	199	219	213	500	489	585	564	-83.3	-66.3	-25.9	-14.8
	(95.1)	(87.7)	(110)	(110)	(33.3)	(34.2)	(35.3)	(36.5)	(60.6)	(63)	(74.8)	(75.1)	(38.8)	(35.2)	(39.8)	(38.1)
Baseline Mean	2,781	2,781	2,781	2,781	630	630	630	630	1,477	1,477	1,477	1,477	673	673	673	673
Percent Change	.220	.224	.28	.274	.308	.316	.347	.337	.339	.331	.396	.382	124	098	038	022
U U	(.034)	(.032)	(.04)	(.04)	(.053)	(.054)	(.056)	(.058)	(.041)	(.043)	(.051)	(.051)	(.058)	(.052)	(.059)	(.057)
Individual Fixed Effects	X	X			X	X			X	X			X	X		

Appendix Table D2: Main Outcomes (Border Zipcodes)

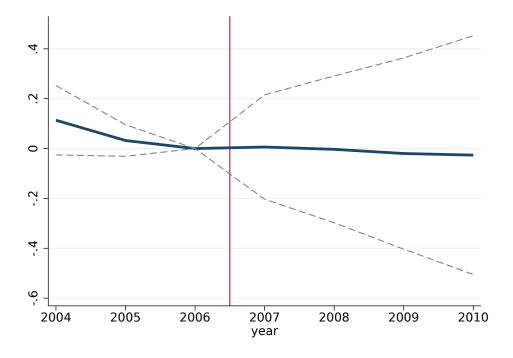
**Note:** Table shows reduced form and instrumental variable estimates for the main outcomes using only border zipcodes. For each outcome, the first and third columns show estimates of control-treatment differences from estimating the pooled reduced form specification in Equation (1) and estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. The second and fourth columns show reduced form and instrumental variable estimates, when the post-period is broken into two separate periods, 2007-2008 and 2009-2010. In the second and fourth columns we estimate the IV coefficient using the two separate treatment × post interaction terms as instruments. For more details, see Section 4. (N = 369, 823 beneficiary-quarters.)

# E Long-Term Care Spending

In this appendix, we present results related to spending on and utilization of institutional long-term care services. Use of these services is not uncommon among disabled Medicaid beneficiaries. However, beneficiaries using institutional long-term care are exempted from the STAR+Plus program, suggesting that the roll out of STAR+Plus should not have much effect on utilization of these services. Further, our data do not provide the same detailed information about utilization of these services that they provide for other types of services discussed in the paper. Indeed, we only observe annual summary measures of spending and utilization for institutional long-term care, preventing us from performing the same quarterly analyses we perform for other types of services. For these reasons, we present these results separately here in this appendix.

Figure E1 and Table E1 present results of our analyses of the effects of the reform on longterm care spending. We estimate the same regressions as we estimate for other outcomes but at the annual instead of the quarterly level. The event study presented in Figure E1 suggests no effect of the reform on long-term care spending, with the difference in spending between treatment and control counties remaining roughly constant over the period before and after the reform. However, standard errors are quite large, especially toward the end of the sample period. The results in Table E1 provide a similar picture: In the regressions with individual fixed effects (our preferred specification), coefficients are small and statistically insignificant. Regressions without individual fixed effects produce larger coefficients but are (1) subject to the caveat that changes in composition could potentially explain these results (though not the results with individual fixed effects) and (2) still statistically insignificant in all cases but one, and the single significant case (the "early-post" coefficient in column (4)) is only marginally significant at the 10% level. Further, in results not presented here, the coefficients on total healthcare spending from Table 2 change only slightly when including long-term care spending as part of total healthcare spending, indicating that our omission of long-term care spending has no material effect on the conclusions of the paper.

Taken together, these results suggest no effect of the reform on long-term care spending. This is not surprising, given that individuals using institutional long-term care are exempted from the STAR+Plus program.



Appendix Figure E1: Long-term Care Spending

**Note:** Figure shows control-treatment differences in long-term care spending in percent terms relative to the treatment mean in the pre-period. These coefficients are from estimating the event study difference-in-differences specification in Equation (1), including individual fixed effects. For more details, see Section 4. (N = 189,935 beneficiary-years.)

	(1)	(2)	(3)	(4)
	Long-term Care	Long-term Care	Long-term Care	Long-term Care
	Spending	Spending	Spending	Spending
Treatment $\times$ Post	-146		-421	
	(556)		(257)	
Treatment $\times$ Post (2007-2008)		-117		-292
· · · ·		(453)		(167)
Treatment $\times$ Post (2009-2010)		-193		-537
``````````````````````````````````````		(721)		(343)
IV Coefficient	-209	-219	-786	-807
	(626)	(653)	(485)	(501)
Baseline Mean	3,084	3,084	3,084	3,084
Percent Change	068	071	255	262
5	(.203)	(.212)	(.157)	(.162)
Individual Fixed Effects	X	X		

#### Appendix Table E1: Long-term Care Spending

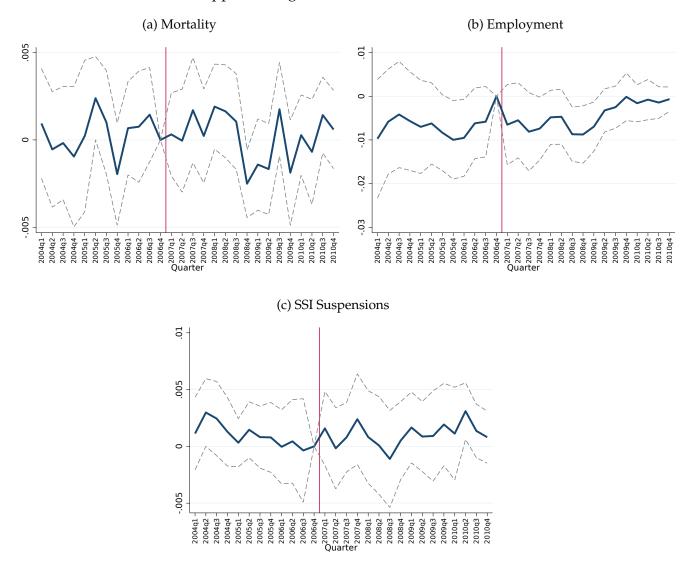
**Note:** Table shows reduced form and instrumental variable estimates for long-term care spending. The first row shows estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and the second row shows estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, when the post-period is broken into two separate periods, 2007-2008 and 2009-2010. We estimate the IV coefficient using the two separate treatment × post interaction terms as instruments. We control for service area by year fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 189,935 beneficiary-years.)

### F Mortality and Employment Outcomes

In this section, we examine indicators of beneficiary health and functional capacity, including death, employment, and the suspension of SSI benefits using the SSA's Disability Analysis File (DAF). The DAF contains monthly administrative records on the universe of SSI and SSDI beneficiaries. We isolate adults (21-64) enrolled only in the SSI program during our sample period. We only observe Medicaid and Medicare eligibility and county of residence but not private vs. public plan enrollment in the SSA data. Mortality is defined as a binary indicator for whether a beneficiary died in a given quarter. Employment is defined as a binary indicator for whether the beneficiary had positive earnings in a given quarter. SSI suspension is defined as a binary indicator for whether a beneficiary's SSI benefits were suspended due to work in a given quarter. Mortality provides a direct measure of beneficiary health. Employment and SSI suspensions provide indirect measures of functional capacity, with the assumption being that take-up of employment or the suspension of benefits due to work indicate improvements in functional capacity and overall well-being.

Regression specifications follow Equation (1) (intent-to-treat estimator), as we do not observe managed care plan enrollment in the SSA data and therefore cannot account for incomplete take-up of managed care as well as the accompanying relaxation of the drug cap in an instrumental variables framework.<sup>3</sup> Regression results are presented in Appendix Figure F1 and Appendix Table F1. Odd columns pool all years in the post-period, and even columns split the post-period into an early and a late period. Coefficients generally go in a direction consistent with overall improvements in health and functional capacity, with managed care plus relaxation of the drug cap leading to long-run reductions in mortality, increases in employment, and more suspensions of benefits due to work. However, none of the coefficients are statistically significantly different from zero, and confidence intervals are quite wide. For mortality, we get a point estimate of -0.06 percentage points, or a reduction of 6% relative to the baseline mean quarterly mortality rate of 1%. However, the 95% confidence interval ranges from a mortality reduction of 0.18 percentage points (18%) to a mortality increase of 0.05 percentage points (5%), implying that we can only rule out mortality increases larger than 5%. For employment, we can only rule out reductions larger than 0.32 percentage points (6%), and for suspensions we can only rule out reductions larger than 0.21 percentage points (15%). We thus conclude that while the signs on these coefficients are all consistent with improvements in health and functional capacity, they are too noisy to lead to any firm conclusions.

<sup>&</sup>lt;sup>3</sup>We also do not include individual fixed effects, as this is not appropriate with the mortality and suspension outcomes, which are absorbing states.



Appendix Figure F1: Other Outcomes

**Note:** Figure shows control-treatment differences in mortality, employment, and SSI suspensions. These coefficients are from estimating the event study difference-in-differences specification in Equation (1). For more details, see Section 4. (N = 1, 123, 706 beneficiary-quarters.)

	(1)	(2)	(3)	(4)	(5)	(6)
	Mor	tality	Emplo	yment	SSI Susp	ensions
Treatment $\times$ Post	-0.0003		0.003		0.00008	
	(0.0005)		(0.004)		(0.001)	
Treatment $\times$ Post (2007-2008)		0.00008		0.0003		-0.0004
		(0.0007)		(0.003)		(0.001)
Treatment $\times$ Post (2009-2010)		-0.0006		0.005		0.0005
		(0.0006)		(0.004)		(0.001)
Baseline Mean	0.010	0.010	0.051	0.051	0.014	0.014

Appendix Table F1: Other Outcomes

**Note:** Table shows reduced form estimates for mortality, employment, and SSI suspension. For each outcome, the first column shows estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1), pooling over the entire post-period. The second column shows reduced form estimates, when the post-period is broken into two separate periods, 2007-2008 and 2009-2010. We control for service area by quarter fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 1, 123, 706 beneficiary-quarters.)

# G Exploration of Alternative Explanations for Increase in Drug Utilization

In this appendix we explore two alternative explanations for the increase in drug utilization we observe after the introduction of the STAR+Plus program. First, we consider the possibility that the increase in drug utilization occurred because MMC plans who were not responsible for drug spending encouraged utilization of drugs with the hope of offsetting non-drug medical expenditures. Second, we consider the possibility that the increase in drug utilization was a side-effect of the roll out of managed care for non-drug services. The evidence we present here suggests that these were not the driving factors behind the increase in drug utilization. Instead, as we discuss in Section 7, the evidence suggests that the relaxation of the drug cap was the driving factor.

**Carve-Out of Prescription Drugs** Even though the relaxation of the drug cap appears to be the main mechanism through which the Texas reform impacted drug utilization, the fact that drugs were carved out of private managed care plan contracts could also play a role; recall that drugs were paid for by the public program for all beneficiaries in all years, even for beneficiaries enrolled in a private plan. With this carve-out, plans had no incentive to reduce drug spending, and may have instead been incentivized to drive up drug utilization, given potential drug-driven medical offsets (Chandra, Gruber and McKnight, 2010; Starc and Town, 2020), including the inpatient offsets we document in Sections 6 and 7. If drugs had been "carved-in" or included in managed care plan contracts, plans may have chosen to ration access to drugs more aggressively than they did in the presence of the carve-out, possibly limiting the effect of relaxing the public drug cap.

To investigate this possibility, we leverage the fact that drugs were carved in to managed care plan contracts in Texas starting in 2012. Our detailed claims and enrollment data ends in 2010, so we cannot use it to study the effects of the carve-in of prescription drugs. Instead, we follow Dranove, Ody and Starc (2021) and use publicly available aggregate data describing prescription drug utilization and spending in Texas's Medicaid program (both public and private plans) over time.<sup>4</sup> In Appendix Figure G1, we document per-enrollee prescription drug services into private Medicaid contracts. The figures show no meaningful change in any of these measures of drug use within Texas Medicaid following the carve-in. In the figure, we also show the same set of outcomes for Arkansas as a reference and control, as it is the neighboring state with the most similar pre-2012 trends in drug utilization.

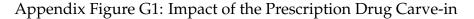
These results provide suggestive evidence that the prescription drug carve-out is relatively inconsequential for patterns of drug utilization in Texas. This is consistent with results from Dranove, Ody and Starc (2021) who show that, when a large set of states carve prescription drug benefits into managed care plan contracts, patterns of utilization change in ways that are similar to states which privatized medical and drug benefits at the same time. While they do find changes overall spending, these appear driven by changes in unit prices rather than by changes in utilization. The implication of this body of evidence is that there seems to be little consequence of including or excluding drugs from managed care plan contracts.

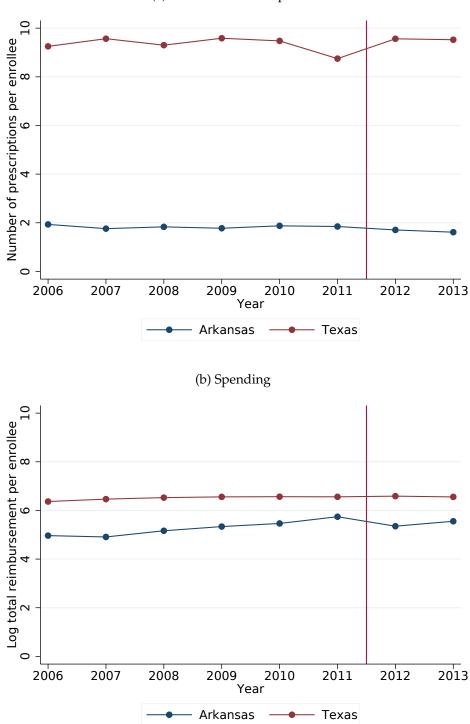
<sup>&</sup>lt;sup>4</sup>The Medicaid State Drug Utilization Data is available online from https://www.medicaid.gov/ medicaid/prescription-drugs/state-drug-utilization-data/index.html.

In other words, the evidence suggests that the effect of shifting to managed care on drug utilization would have been comparable irrespective of whether drugs were carved in or carved out, at the time of the shift.

**Shift to Private Provision of Medical Benefits** While the analysis so far suggests that the overall effect of the Texas reform on drug utilization came partly through the accompanying relaxation of drug caps, we cannot completely rule out the alternative mechanism of the drug effect instead coming through private provision's effect on patterns of medical care. For example, it is possible that the activities of the private Medicaid plans related to outpatient care (i.e. care management) naturally led to increased levels of drug utilization. Specifically, we showed that private provision led to increased use of outpatient care in Texas, and it is possible that just seeing the doctor more could lead to higher levels of drug utilization.

However, it seems unlikely that any care management activities would *only* affect utilization on the margin of taking three or more drugs, the margin we showed to be by far the most important for the drug effect we estimate. That said, our analysis cannot entirely rule out comparable drug effects, even absent the lifting of the public drug cap under privatization. To make this point, we must instead rely on the null result for drugs in our work studying the shift to managed care in New York, which had no drug cap (Layton et al., 2019).

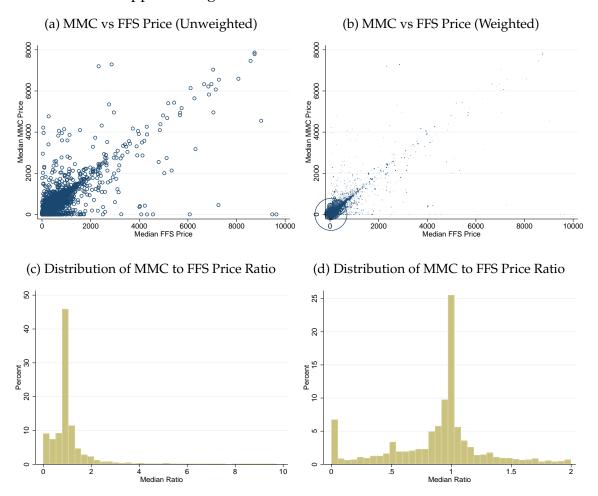




(a) Number of Prescriptions

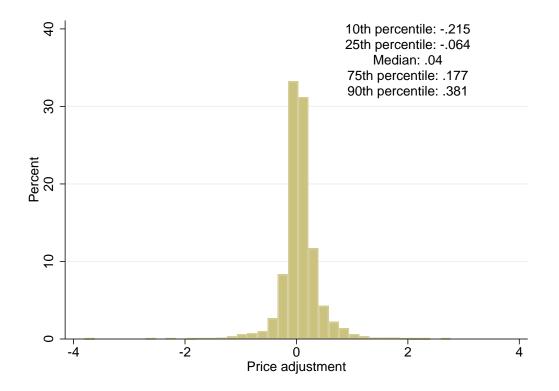
**Note:** Figure shows the number of prescriptions and the amount of spending per enrollee in Texas and Arkansas before and after Texas carved prescription drugs into its managed care contracts in 2012. The data displayed here come from the publicly available Medicaid State Drug Utilization Data. For more details, see Appendix G. (N = 64 state-years.)

### **H** Price Analysis



Appendix Figure H1: Prices Under FFS and MMC

**Note:** Figure shows how MMC and FFS prices compare in 2010, the final year in our data. For each procedure that we observe both under fee-for-service (FFS) Medicaid and Medicaid managed care (MMC), we compute the median price under FFS and under MMC. Panel (a) shows an unweighted scatterplot of the median MMC price vs the median FFS price, censored at \$10,000 for readability. Panel (b) shows a weighted (by frequency under FFS) scatterplot of the median MMC price vs the median MMC price vs the median MMC price to the median MMC price to the median MMC price to the median FFS price, censored at 10 for readability. Panel (d) shows a histogram of the distribution of the ratio of the ratio of the median MMC price to the median FFS price, censored at 2 for readability. For more details, see Section 7. (N = 6,053 unique procedure codes.)



Appendix Figure H2: Distribution of Procedure-Specific Price Differences

**Note:** Figure shows the distribution of procedure specific price differences. We estimate Equation (2) on the sample of procedures that we observe both under fee-for-service (FFS) Medicaid and Medicaid managed care, allowing the price difference to vary by procedure. We then plot the distribution of the estimated price differences. For more details, see Section 7. (N = 6,053 unique procedure codes.)

	(1)
	Log
	Medicaid
	Payment
Medicaid Managed Care	.085
	(.0002)

-

**Note:** Table shows the estimated difference in log Medicaid payments between Medicaid managed care and fee-for-service Medicaid. The results are from estimating Equation (2). For more details, see Section 7. (N = 94,472,521 claims.)

	(1)	(2)	(3)	(4)
	Spending	Spending	Adjusted	Adjusted
	openang	opending	Spending	Spending
Treatment $\times$ Post	366		265	
	(45)		(42.8)	
Treatment $\times$ Post (2007-2008)		200		123
· · · ·		(36.7)		(35.9)
Treatment $\times$ Post (2009-2010)		501		374
· · · ·		(62.9)		(60.2)
IV Coefficient	489	480	354	343
	(54.3)	(57.5)	(53.3)	(56.9)
Baseline Mean	1,551	1,551	1,549	1,549
Percent Change	.316	.309	.229	.221
č	(.035)	(.037)	(.034)	(.037)
Individual Fixed Effects	X	X	X	ŶX

Appendix Table H2: Price-Adjusted Outpatient Spending Outcomes

**Note:** Table shows reduced form and instrumental variable estimates for price-adjusted outpatient spending outcomes. For each outcome, the first column shows estimates of control-treatment differences from estimating the pooled version of the reduced form specification in Equation (1) and estimates of the impact of Medicaid managed care from estimating our instrumental variable specification, pooling over the entire post-period. The second column shows reduced form and instrumental variable estimates, when the post-period is broken into two separate periods, 2007-2008 and 2009-2010. In the second and fourth columns we estimate the IV coefficient using the two separate treatment × post interaction terms as instruments. We control for service area by quarter fixed effects. Standard errors are clustered at the county level. For more details, see Section 4. (N = 643,751 beneficiary-quarters.)

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