## Online Appendix

Can Information Reduce Ethnic Discrimination？<br>Evidence from Airbnb<br>Morgane Laouénan，Roland Rathelot

## A Additional Figures

Figure A1：Example of a listing＇s dashboard，with the most salient information


Information on listings＇ratings

7 Reviews

Summary

| Accuracy | t t t t $\boldsymbol{*}$ | Location | ＊ |
| :---: | :---: | :---: | :---: |
| Communication | t $k$ t $k$ t | Check In | thttt |
| Cleanliness | 大＊大 大 | Value | 大 大 大 |

[^0]
## Information on listings' amenities



For illustrative purposes only, screenshot of the Airbnb platform captured by the author on May 2016.

## Peer-reviewing system

## Describe Your Experience (required)

Your review will be public on your profile and your host's listing page. If you have additional feedback that you don't want to make public, you can share it with Airbnb on the next page.
How did your host make you feel welcome? Was the listing
description accurate? What was the neighborhood like?

500 words left

## Private Host Feedback

We won't make it public and your feedback will only be shared with your host, Airbnb employees and its service providers

What did you love about staying at this listing?


How can your host improve?


Overall Experience (required)


## Next

For illustrative purposes only, screenshot of the Airbnb platform captured by the author on May 2016.

## Information on listings' locations



In this example, the listing is shown to be located in the neighbourhood of Pimlico, in London, and the area of the .6 -mile-radius circle is almost entirely in that neighbourhood. For illustrative purposes only, screenshot of the Airbnb platform captured by the author on September 2018.

Figure A2: Number of observations by listing


Notes: This figure shows the number of observations by listing depending on the number of waves (x-axis). It starts at 2 waves as we restrict the sample to listings that have gained at least one review over the observation period.

Figure A3: Distribution of the number of reviews (left) and of the longitudinal variation in the number of reviews within a property (right)


Notes: The left figure shows the distribution of the number of reviews. The right figure shows the distribution of the longitudinal variation in the number of reviews within a property. Both figures are right truncated with a maximum of 50 reviews. The sample is restricted to listings that have gained at least one review over the observation period.

Figure A4: Illustration of the conceptual framework: Prices with the number of reviews, by unobservable quality


Note: This illustrative graph displays $(K v-\rho / 5) /(K+\rho)$ as function of $K$, where $v$ takes values in $\{-2,1,0,1,2\}$ and $\rho=8$.

## B Additional tables

Table A1: Number of observations \& listings by city

| City | Observations |  | Listings |  |
| :--- | :---: | :---: | :---: | :---: |
|  | $\#$ | share | $\#$ | share |
| Amsterdam | 51,189 | 2.07 | 6,122 | 2.77 |
| Barcelona | 173,180 | 7.00 | 14,529 | 6.58 |
| Berlin | 151,887 | 6.14 | 13,948 | 6.31 |
| Boston | 43,637 | 1.76 | 4,330 | 1.96 |
| Chicago | 42,990 | 1.74 | 4,408 | 2.00 |
| Florence | 67,106 | 2.71 | 4,967 | 2.25 |
| London | 264,705 | 10.70 | 23,889 | 10.81 |
| Los Angeles | 159,228 | 6.43 | 15,182 | 6.87 |
| Madrid | 65,753 | 2.66 | 5,359 | 2.43 |
| Marseille | 55,643 | 2.25 | 4,921 | 2.23 |
| Miami | 67,373 | 2.72 | 6,383 | 2.89 |
| Milan | 85,365 | 3.45 | 8,360 | 3.78 |
| Montreal | 69,331 | 2.80 | 6,525 | 2.95 |
| New York City | 349,471 | 14.12 | 31,717 | 14.36 |
| Paris | 464,493 | 18.77 | 39,026 | 17.66 |
| Rome | 152,644 | 6.17 | 11,547 | 5.23 |
| San Francisco | 108,144 | 4.37 | 10,148 | 4.59 |
| Toronto | 56,843 | 2.30 | 5,359 | 2.43 |
| Vancouver | 45,569 | 1.84 | 4,219 | 1.91 |

Notes: The table shows the number of observations (column 1), its share (column 2) and the number of listings (column 3), and its share (column 4) for each of the 19 cities included in our dataset. The sample is restricted to listings that have gained at least one review over the observation period. The total number of observations is $2,474,551$ and the total number of listings is 220,939 .

Table A2: Collection dates of waves

| Wave | Collection date |
| :---: | :---: |
| 1 | 15 June 2014 |
| 2 | 8 July 2014 |
| 3 | 28 July 2014 |
| 4 | 11 August 2014 |
| 5 | 25 August 2014 |
| 6 | 8 September 2014 |
| 7 | 25 September 2014 |
| 8 | 15 October 2014 |
| 9 | 5 November 2014 |
| 10 | 25 November 2014 |
| 11 | 15 December 2014 |
| 12 | 7 January 2015 |
| 13 | 13 January 2015 |
| 14 | 3 February 2015 |
| 15 | 4 March 2015 |
| 16 | 25 March 2015 |
| 17 | 13 April 2015 |
| 18 | 4 May 2015 |
| 19 | 26 May 2015 |
| 20 | 15 June 2015 |
| 21 | 11 November 2017 |

Table A3: Summary statistics: Property \& host characteristics

|  | Full <br> Sample | Listings that have gained at least <br> one review over the period |
| :--- | :---: | :---: |
| Type of property |  |  |
| Entire property | 0.665 | 0.705 |
| Flat | 0.802 | 0.843 |
| House | 0.064 | 0.106 |
| Loft | 0.016 | 0.019 |
| Size |  |  |
| Person capacity | 3.148 | 3.211 |
| Number of bedrooms | 1.252 | 1.244 |
| Number of bathrooms | 1.162 | 1.153 |
| Type of bed |  |  |
| Couch | 0.005 | 0.006 |
| Airbed | 0.003 | 0.003 |
| Sofa | 0.026 | 0.033 |
| Futon | 0.009 | 0.012 |
| Real bed | 0.958 | 0.946 |
| Amenities |  |  |
| Cable TV | 0.290 | 0.346 |
| Wireless | 0.901 | 0.899 |
| Heating | 0.876 | 0.887 |
| AC | 0.395 | 0.380 |
| Elevator | 0.341 | 0.340 |
| Wheelchair accessible | 0.077 | 0.098 |
| Doorman | 0.080 | 0.096 |
| Fireplace | 0.077 | 0.080 |
| Washer | 0.697 | 0.697 |
| Dryer | 0.402 | 0.388 |
| Parking | 0.200 | 0.179 |
| Gym | 0.072 | 0.064 |
| Pool | 0.063 | 0.054 |
| Buzzer | 0.293 | 0.386 |
| Hot Tub | 0.069 | 0.069 |
| Services | 0.111 |  |
| Breakfast served | 0.466 | 0.091 |
| Family/Kids friendly | 0.045 | 0.448 |
| Suitable for events |  | 0.052 |
| Rules \& Extras |  |  |
| Additional people |  |  |
| Price per additional people |  |  |
|  |  |  |

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Table A3: Summary statistics: Property \& host characteristics

| Smoking allowed | 0.133 | 0.144 |
| :--- | :---: | :---: |
| Pets allowed | 0.125 | 0.131 |
| Host Characteristics |  |  |
| Has multiple properties | 0.356 | 0.345 |
| Member since 2008 | 0.001 | 0.001 |
| Member since 2009 | 0.006 | 0.009 |
| Member since 2010 | 0.019 | 0.033 |
| Member since 2011 | 0.063 | 0.107 |
| Member since 2012 | 0.126 | 0.209 |
| Member since 2013 | 0.166 | 0.263 |
| Member since 2014 | 0.198 | 0.291 |
| Member since 2015 | 0.068 | 0.075 |
| Number of languages spoken | 0.851 | 1.408 |
| Superhost | 0.023 | 0.053 |
| Verified email | 0.620 | 0.960 |
| Verified offline | 0.320 | 0.525 |
| Verified phone | 0.281 | 0.428 |
| Number of Facebook friends | 153.567 | 237.714 |
| Number of words in description | 217.000 | 240.168 |
| Number of words in profile | 49.560 | 49.822 |
| Number of pictures | 13.058 | 13.921 |
| Number of pictures by professionals | 0.979 | 0.703 |
| $N$ | 663,090 | 220,939 |

Notes: The left column displays the mean of each characteristics in the full sample, while the right column focuses on the sub-sample of listings that have gained at least one review over the observation period (between the first and the last waves).

Table A4: Log daily rate

|  | (1) | (2) |
| :---: | :---: | :---: |
| Shared Flat | $\begin{gathered} \hline-0.828^{* * *} \\ (0.004) \end{gathered}$ | $\begin{gathered} \hline-0.715^{* * *} \\ (0.007) \end{gathered}$ |
| Person Capacity ( $>2$ ) | $\begin{gathered} 0.164^{* * *} \\ (0.004) \end{gathered}$ | $\begin{gathered} 0.175^{* * *} \\ (0.005) \end{gathered}$ |
| \# bedrooms | $\begin{gathered} 0.273 * * * \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.293^{* * *} \\ (0.004) \end{gathered}$ |
| \# bathrooms | $\begin{gathered} 0.167^{* * *} \\ (0.004) \end{gathered}$ | $\begin{gathered} 0.143^{* * *} \\ (0.005) \end{gathered}$ |
| Flat | $\begin{gathered} -0.154^{* * *} \\ (0.009) \end{gathered}$ | $\begin{gathered} -0.179^{* * *} \\ (0.013) \end{gathered}$ |
| House or Loft | $\begin{gathered} -0.159^{* * *} \\ (0.010) \end{gathered}$ | $\begin{gathered} -0.061^{* * *} \\ (0.014) \end{gathered}$ |
| Couch | $\begin{gathered} -0.193^{* * *} \\ (0.014) \end{gathered}$ | $\begin{gathered} -0.165^{* * *} \\ (0.015) \end{gathered}$ |
| Airbed | $\begin{gathered} -0.192^{* * *} \\ (0.027) \end{gathered}$ | $\begin{gathered} -0.125^{* * *} \\ (0.025) \end{gathered}$ |
| Sofa | $\begin{gathered} -0.175^{* * *} \\ (0.006) \end{gathered}$ | $\begin{gathered} -0.166^{* * *} \\ (0.009) \end{gathered}$ |
| Futon | $\begin{gathered} -0.158^{* * *} \\ (0.011) \end{gathered}$ | $\begin{gathered} -0.116^{* * *} \\ (0.010) \end{gathered}$ |
| Cable TV | $\begin{gathered} 0.141^{* * *} \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.098^{* * *} \\ (0.004) \end{gathered}$ |
| Wireless | $\begin{gathered} 0.033^{* * *} \\ (0.005) \end{gathered}$ | $\begin{gathered} 0.021^{* * *} \\ (0.006) \end{gathered}$ |
| Heating | $\begin{gathered} -0.019^{* * *} \\ (0.005) \end{gathered}$ | $\begin{gathered} 0.001 \\ (0.007) \end{gathered}$ |
| AC | $\begin{gathered} 0.134^{* * *} \\ (0.004) \end{gathered}$ | $\begin{gathered} 0.113^{* * *} \\ (0.006) \end{gathered}$ |
| Elevator | $\begin{gathered} 0.093^{* * *} \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.084^{* * *} \\ (0.005) \end{gathered}$ |
| Wheelchair Accessible | -0.039*** | -0.007 |

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Table A4: Log daily rate

|  | $(0.004)$ | $(0.005)$ |
| :--- | :---: | :---: |
| Doorman | $0.102^{* * *}$ | $0.036^{* * *}$ |
| Fireplace | $(0.005)$ | $(0.007)$ |
|  | $0.158^{* * *}$ | $0.117^{* * *}$ |
| Washer | $(0.005)$ | $(0.005)$ |
|  | $-0.021^{* * *}$ | $0.020^{* * *}$ |
| Dryer | $(0.004)$ | $(0.006)$ |
|  | $0.146^{* * *}$ | $0.094^{* * *}$ |
| Parking | $(0.003)$ | $(0.004)$ |
|  | $-0.133^{* * *}$ | $0.028^{* * *}$ |
| Gym | $(0.004)$ | $(0.005)$ |
| Pool | $0.062^{* * *}$ | $0.042^{* * *}$ |
|  | $(0.007)$ | $(0.009)$ |
| Buzzer | $0.083^{* * *}$ | $0.082^{* * *}$ |
| Hot Tub | $(0.007)$ | $(0.012)$ |
|  | $0.050^{* * *}$ | $0.008^{* *}$ |
| Breakfast served | $(0.003)$ | $(0.003)$ |
| Family /Kids Friendly | $0.012^{* *}$ | 0.010 |
| Suitable for events | $(0.005)$ | $(0.006)$ |
| Price per Additional People | 0.005 | $0.033^{* * *}$ |
| Additional People | $(0.004)$ | $(0.005)$ |
| Cancellation Policy | $0.014^{* * *}$ | $0.033^{* * *}$ |
|  | $(0.003)$ | $(0.003)$ |
|  | $0.072^{* * *}$ | $0.062^{* * *}$ |
|  | $(0.006)$ | $(0.008)$ |
|  | $-0.034^{* * *}$ | $-0.013^{* * *}$ |
|  | $(0.002)$ | $(0.002)$ |
|  | $0.001^{* * *}$ | $-0.001^{* * *}$ |
|  | $(0.000)$ | $(0.000)$ |
|  | $0.040^{* * *}$ | $0.015^{* * *}$ |
|  | $(0.002)$ | $(0.002)$ |
|  | $-0.123^{* * *}$ | $-0.093^{* * *}$ |
|  | $(0.004)$ | $(0.004)$ |

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Table A4: Log daily rate

| Pets Allowed | $\begin{gathered} -0.024^{* * *} \\ (0.004) \end{gathered}$ | $\begin{gathered} -0.027^{* * *} \\ (0.004) \end{gathered}$ |
| :---: | :---: | :---: |
| Host has multiple properties | $\begin{gathered} 0.050 * * * \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.024^{* * *} \\ (0.004) \end{gathered}$ |
| Member since 2009 | $\begin{gathered} 0.145 * * * \\ (0.019) \end{gathered}$ | $\begin{gathered} 0.118^{* * *} \\ (0.021) \end{gathered}$ |
| Member since 2010 | $\begin{gathered} 0.121^{* * *} \\ (0.015) \end{gathered}$ | $\begin{gathered} 0.097^{* * *} \\ (0.016) \end{gathered}$ |
| Member since 2011 | $\begin{gathered} 0.098^{* * *} \\ (0.014) \end{gathered}$ | $\begin{gathered} 0.087^{* * *} \\ (0.015) \end{gathered}$ |
| Member since 2012 | $\begin{gathered} 0.077 * * * \\ (0.014) \end{gathered}$ | $\begin{gathered} 0.070^{* * *} \\ (0.015) \end{gathered}$ |
| Member since 2013 | $\begin{gathered} 0.076^{* * *} \\ (0.014) \end{gathered}$ | $\begin{gathered} 0.066^{* * *} \\ (0.015) \end{gathered}$ |
| Member since 2014 | $\begin{gathered} 0.051^{* * *} \\ (0.014) \end{gathered}$ | $\begin{gathered} 0.048^{* * *} \\ (0.014) \end{gathered}$ |
| Member since 2015 | $\begin{gathered} 0.052^{* * *} \\ (0.014) \end{gathered}$ | $\begin{gathered} 0.047^{* * *} \\ (0.015) \end{gathered}$ |
| Superhost | $\begin{gathered} 0.023^{* * *} \\ (0.005) \end{gathered}$ | $\begin{gathered} 0.014^{* * *} \\ (0.005) \end{gathered}$ |
| Verified Email | $\begin{gathered} -0.022^{* * *} \\ (0.007) \end{gathered}$ | $\begin{gathered} -0.000 \\ (0.007) \end{gathered}$ |
| Verified Offline | $\begin{gathered} 0.013^{* * *} \\ (0.003) \end{gathered}$ | $\begin{aligned} & 0.005^{*} \\ & (0.003) \end{aligned}$ |
| Verified Phone | $\begin{gathered} 0.003 \\ (0.007) \end{gathered}$ | $\begin{gathered} 0.004 \\ (0.012) \end{gathered}$ |
| Nber of Facebook friends | $\begin{gathered} 0.000^{* * *} \\ (0.000) \end{gathered}$ | $\begin{gathered} 0.000 \\ (0.000) \end{gathered}$ |
| Nber of words in Description | $\begin{gathered} -0.000^{* * *} \\ (0.000) \end{gathered}$ | $\begin{gathered} -0.000^{* * *} \\ (0.000) \end{gathered}$ |
| Nber of words in Profile | $\begin{gathered} -0.000^{* * *} \\ (0.000) \end{gathered}$ | $\begin{gathered} -0.000^{* * *} \\ (0.000) \end{gathered}$ |
| Nber of Languages | $-0.005^{* * *}$ | -0.005*** |

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Table A4: Log daily rate

|  | $(0.001)$ | $(0.001)$ |
| :--- | :---: | :---: |
| Nber of words in Rules | $-0.000^{* * *}$ | $-0.000^{* *}$ |
|  | $(0.000)$ | $(0.000)$ |
| Nber of pictures | $0.003^{* * *}$ | $0.003^{* * *}$ |
| Nber of pictures taken by professionals | $(0.000)$ | $(0.000)$ |
|  | $\left(0.001^{* * *}\right.$ | $0.002^{* * *}$ |
| Nber of picture changes | $-0.034^{* * *}$ | $-0.037^{* * *}$ |
|  | $(0.002)$ | $(0.002)$ |
| City-wave FE | Yes | Yes |
| Neighbourhood FE | No | Yes |
| Block FE | No | Yes |
| Adj $R^{2}$ | 0.627 | 0.733 |
| $N$ obs. | $2,474,551$ | $2,474,551$ |

Notes: OLS regression on the daily log-price. In column (2), neighbourhood and block fixed effects are included in the estimation. Robust standard errors clustered at the property level.

Table A5: Distribution of the last rating

|  | Obs | Share |
| :--- | :---: | :---: |
| 3.5 stars | 9,560 | $4.39 \%$ |
| 4 stars | 26,943 | $12.37 \%$ |
| 4.5 stars | 85,047 | $39.06 \%$ |
| 5 stars | 96,178 | $44.17 \%$ |

Notes : The sample corresponds to listings for which last rating is observed. Listings with less than 3.5 stars are included in the first row.

Table A6: Number of neighbourhoods \& blocks by city

| City | \# neighbourhoods | \# Blocks |
| :--- | :---: | :---: |
| Amsterdam | 45 | 101 |
| Barcelona | 70 | 82 |
| Berlin | 88 | 404 |
| Boston | 42 | 250 |
| Chicago | 75 | 242 |
| Florence | 18 | 102 |
| London | 150 | 838 |
| Los Angeles | 115 | 1267 |
| Madrid | 67 | 166 |
| Marseille | 61 | 615 |
| Miami | 80 | 430 |
| Milan | 25 | 155 |
| Montreal | 53 | 242 |
| New York City | 189 | 527 |
| Paris | 64 | 116 |
| Rome | 44 | 107 |
| San Francisco | 169 | 495 |
| Toronto | 115 | 286 |
| Vancouver | 34 | 307 |
| Total | 1,504 | 6,732 |
| Notes: The definition of neighbourhoods directly |  |  |
| comes from Airbnb while blocks are created via the |  |  |
| approximate coordinates of the listing. |  |  |

## C Ethnic differences in the exit rate

In this section, we look at the issue of differential selection in the sample across ethnic groups and find that minority hosts are not more likely to leave the market than the majority. We consider that a listing $i$ leaves the market at $t$ if it is present at $t$, and not present anytime after $t$, and define $q_{i t}=1$ and 0 for $s \neq t$. Within the period of observation, 65,358 majority hosts ( $31.6 \%$ ) and 4,777 minority hosts (33.6\%) leave the platform. We regress $q_{i t}$ on a minority dummy, and control for property characteristics, ratings, neighbourhood fixed-effects, block fixed-effects and price.

Table A7 shows that the exit rate is similar for both groups when controlling for property characteristics, ratings, neighbourhood and block fixed-effects, price of the listing and number of reviews.

Table A7: Probability to leave the market at wave $t$

|  | $(1)$ | $(2)$ | $(3)$ |
| :--- | :---: | :---: | :---: |
| Minority host | 0.0004 | 0.0003 | 0.0003 |
|  | $(0.0005)$ | $(0.0005)$ | $(0.0005)$ |
| Log-price |  | $-0.0043^{* * *}$ | $-0.0053^{* * *}$ |
|  |  | $(0.0003)$ | $(0.0003)$ |
| Number of reviews |  |  | $-0.0001^{* * *}$ |
|  |  |  | $(0.0000)$ |
| Adj $R^{2}$ | 0.04 | 0.04 | 0.04 |
| $N$ obs. | $2,474,551$ | $2,474,551$ | $2,474,551$ |

Notes: OLS regressions of the probability to leave the market at wave $t$. Covariates include, aside from the ones mentioned in the table, neighbourhood fixed effects, block fixed-effects, property characteristics and ratings. Robust standard errors clustered at the property level.

## D Pictures from which host ethnicity cannot be measured

Hosts can choose whether to post a picture of themselves on their host profiles. Popular alternative choices are pictures of their properties, pets, furniture, landscapes, etc. We identify pictures for which it was impossible to say anything about the ethnicity of anyone in the picture. In our data, there are $17 \%$ of such listings. If minorities are aware of the existence of discrimination on the platform, they might more often obfuscate their skin colour.

In this appendix, we try to understand the choice leading hosts to post or not their pictures. First, is the price set by no-person-picture hosts higher in neighbourhoods where the share of blacks is high? First, how do no-person-picture hosts set their price? Second, does the probability of having a no-person picture depend on the share of Blacks in the neighbourhood?

Table A8 first shows that, controlling for listing characteristics, hosts with a listing located in a neighbourhood with more Black hosts are not more or less likely to post a picture of themselves (Column 1). This result is at odds with a model of strategic hosts anticipating discrimination. Column 2 shows that, controlling for neighbourhoods and characteristics, hosts post very similar prices whether they choose to publish their pictures or not. Column 3 shows that the pattern does not seem to vary much with the ethnic composition of the neighbourhood. If anything, in areas with more Black hosts, the hosts that do not post their pictures have lower prices than those posting their pictures.

Table A8: Behaviour of hosts posting non-person pictures

|  | Non-person picture | Log-price |  |
| :---: | :---: | :---: | :---: |
|  | (1) | (2) | (3) |
| Local share of Blacks | $\begin{gathered} 0.007 \\ (0.018) \end{gathered}$ |  |  |
| Non-person picture |  | $\begin{gathered} 0.002 \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.004 \\ (0.004) \end{gathered}$ |
| Non-person picture $\times$ share Blacks |  |  | $\begin{aligned} & -0.078 \\ & (0.064) \end{aligned}$ |
| Neighbourhood FE | No | Yes | Yes |
| Adj $R^{2}$ | 0.036 | 0.713 | 0.713 |
| N obs. | 2,466,726 | 2,466,726 | 2,466,726 |

Notes: OLS regressions. Aside from those mentioned in the Table, controls include city-wave FE, and property characteristics (see Table A4). Specifications in Columns 2 and 3 include neighbourhood FE and block FE, not in Column 1. Robust standard errors clustered at the listing level.

## E Using a non-normal prior distribution of quality with a discrete signal

Assume that $v \sim \mathcal{B}\left(\alpha_{v}, \beta_{v}\right)$ (a Beta distribution). A Beta distribution looks more similar to the measures of quality that we have empirically: it is bounded and can be really skewed.

A single rating being a discrete signal, let's assume that we can model it as a draw in a Binomial $(n, v)$, where $n$ depends on how much information a single rating contains (to what extent it is discrete). A rating takes values in $0 \ldots n$.

The pdf of the posterior distribution, given the observation of a rating $r$ can be written as:

$$
f(v \mid r)=\frac{P(r \mid v) f(v)}{\int P(r \mid v) f(v) d v}
$$

Working on the numerator, we have:

$$
P(r \mid v) f(v)=\binom{n}{r} \frac{v^{r}(1-v)^{n-r} v^{\alpha_{v}-1}(1-v)^{\beta_{v}-1}}{B\left(\alpha_{v}, \beta_{v}\right)}
$$

where $B(.,$.$) is the beta function. This simplifies to:$

$$
P(r \mid v) f(v)=\binom{n}{r} \frac{v^{\alpha_{v}-1+r}(1-v)^{\beta_{v}-1+n-r}}{B\left(\alpha_{v}, \beta_{v}\right)}
$$

Because $f(v \mid r)$ is a density, we know it is of integral one and thus should be equal to the density of a $\mathcal{B}\left(\alpha_{v}+r, \beta_{v}+n-r\right)$. We can also prove it by computing the integral of $P(r \mid v) f(v)$ wrt $v$ and computing $f(v \mid r)$ explicitly.

The expectation of $v$ conditional on $r$ is therefore equal to:

$$
E(v \mid r)=\frac{\alpha_{v}+r}{\alpha_{v}+\beta_{v}+n}
$$

Now, suppose that we have $K$ signals instead of just one. I also rescale the signal between 0 and 1 (which is the range of $v$ ) and define $\bar{r}=\sum_{k} r_{k} /(n K), \tilde{\alpha}_{v}=\alpha_{v} / n$ and $\tilde{\beta} v=\beta v / n$. We can show that the expectation depends only on $\bar{r}$ :

$$
E(v \mid \bar{r}, K)=\frac{\hat{\alpha}_{v}+K \bar{r}}{\hat{\alpha}_{v}+\hat{\beta}_{v}+K}
$$

Dividing everything by $n$ rescales the signal between 0 and 1 (which is the range of $v$ ) and we obtain an expression that is exactly identical, up to a change in notations, to the one with normal distributions.

$$
\mathbb{E}(v \mid \bar{r}, K, m)=\frac{\rho \bar{v}+K \bar{r}}{\rho+K}
$$

with $\alpha_{v}=\rho \bar{v}$ and $\alpha_{v}+\beta_{v}=\rho$.

## F Proofs for the identification results

## F. 1 Accurate beliefs

We start from the equation (2), reproduced here:

$$
p=p_{0}-\lambda \gamma m+\lambda \alpha w+\lambda \beta \zeta+\lambda \beta \frac{K r+\rho \bar{v}_{m}}{K+\rho}
$$

Assuming that we know $\rho$, the regression line of $p_{i t}$ conditional on $\mathcal{I}_{i t}$, an information set made of $\frac{K_{i t}}{K_{i t}+\rho}, m_{i} \frac{K_{i t}}{K_{i t}+\rho}, \bar{r}_{i} \frac{K_{i t}}{K_{i t}+\rho}$, characteristics $X_{i t}$ and listing fixed effects $\mu_{i}$ :

$$
\mathbb{E}\left(p_{i t} \mid \mathcal{I}_{i t}\right)=\mathbb{E}\left(p_{0}-\lambda \gamma m+\lambda \alpha w_{i t}+\lambda \beta \zeta_{i t} \mid \mathcal{I}_{i t}\right)+\mathbb{E}\left(\left.\lambda \beta \frac{K_{i t} r_{i t}+\rho \bar{v}_{m}}{K_{i t}+\rho} \right\rvert\, \mathcal{I}_{i t}\right)
$$

By assumption, the first term $\mathbb{E}\left(p_{0}-\lambda \gamma m+\lambda \alpha w_{i t}+\lambda \beta \zeta_{i t} \mid \mathcal{I}_{i t}\right)$ is equal to a linear combination of the fixed effects and the observable characteristics.

$$
\mathbb{E}\left(p_{i t} \mid \mathcal{I}_{i t}\right)=\mu_{i}+X_{i t} \beta_{x}+\lambda \beta \mathbb{E}\left(\left.\frac{K_{i t} r_{i t}}{K_{i t}+\rho} \right\rvert\, \mathcal{I}_{i t}\right)+\lambda \beta \mathbb{E}\left(\left.\frac{\rho \bar{v}_{m}}{K_{i t}+\rho} \right\rvert\, \mathcal{I}_{i t}\right)
$$

At this stage, it is key that $\mathbb{E}\left(r_{i t} \mid \mathcal{I}_{i t}\right)=\mathbb{E}\left(r_{i t} \mid \bar{r}_{i}\right)$. In particular, $r_{i t}$ does not depend on ethnicity conditional on $\bar{r}_{i}$.

$$
\begin{aligned}
& \mathbb{E}\left(p_{i t} \mid \mathcal{I}_{i t}\right)=\mu_{i}+X_{i t} \beta_{x}+\lambda \beta \frac{K_{i t}}{K_{i t}+\rho} \mathbb{E}\left(r_{i t} \mid \bar{r}_{i}\right)+\lambda \beta \frac{\rho \bar{v}_{0}}{K_{i t}+\rho}+\lambda \beta \frac{\rho}{K_{i t}+\rho}\left(\bar{v}_{1}-\bar{v}_{0}\right) m_{i} \\
& \text { As } \frac{\rho}{K_{i t}+\rho}=1-\frac{K_{i t}}{K_{i t}+\rho}: \\
& \qquad \mathbb{E}\left(p_{i t} \mid \mathcal{I}_{i t}\right)=\mu_{i}+X_{i t} \beta_{x}+\lambda \beta \frac{K_{i t}}{K_{i t}+\rho}\left[\mathbb{E}\left(r_{i t} \mid \bar{r}_{i}\right)-\bar{v}_{0}\right]-\lambda \beta \frac{K_{i t}}{K_{i t}+\rho}\left(\bar{v}_{1}-\bar{v}_{0}\right) m_{i}
\end{aligned}
$$

Therefore, regressing $p_{i t}$ on $\frac{K_{i t}}{K_{i t}+\rho} \mathbb{1}\left\{\bar{r}_{i}=\bar{r}\right\}$, for all values $\bar{r}$ in the support of $\bar{r}_{i}$, and $\frac{K_{i t}}{K_{i t}+\rho} m_{i}$, conditional on listing fixed effects and characteristics $X_{i t}$, will identify:

- $\beta_{\bar{r}}=\lambda \beta\left[\mathbb{E}\left(r_{i t} \mid \bar{r}\right)-\bar{v}_{0}\right]$ for each value $\bar{r}$ in the support of $\bar{r}_{i}$.
- $\beta_{m}=-\lambda \beta\left(\bar{v}_{1}-\bar{v}_{0}\right)$.

To finish the proof, note that $\rho$ is identified non-parametrically within listing conditional on $\beta_{\bar{r}}$ and $\beta_{m}$.

## F. 2 Inaccurate beliefs

The first part of the proof directly follows the one of the case with accurate beliefs. For the second part, we apply the same reasoning, except that we attempt to characterise the regression line of $p_{i t}$ conditional on $\mathcal{I}_{i t}^{\prime}$, an information set equal to $\mathcal{I}_{i t}$ minus $\bar{r}_{i} \frac{K_{i t}}{K_{i t}+\rho}$. The main difference is that Bayesian updating starts from the wrong bias $\tilde{\nu}_{1}$ instead of $\bar{\nu}_{1}$ for listings held by minority hosts. We obtain:

$$
\mathbb{E}\left(p_{i t} \mid \mathcal{I}_{i t}^{\prime}\right)=\mu_{i}+X_{i t} \beta_{x}+\lambda \beta \mathbb{E}\left(\left.\frac{K_{i t} r_{i t}}{K_{i t}+\rho} \right\rvert\, \mathcal{I}_{i t}^{\prime}\right)+\lambda \beta \mathbb{E}\left(\left.\frac{\rho\left(\bar{v}_{0}+m_{i}\left(\tilde{v}_{1}-\bar{v}_{0}\right)\right)}{K_{i t}+\rho} \right\rvert\, \mathcal{I}_{i t}^{\prime}\right)
$$

Now, note that $\mathbb{E}\left(r_{i t} \mid \mathcal{I}_{i t}^{\prime}\right)=\mathbb{E}\left(r_{i t} \mid m_{i}\right)=\bar{v}_{m_{i}}=\bar{v}_{0}+m_{i}\left(\bar{v}_{1}-\bar{v}_{0}\right)$.
$\mathbb{E}\left(p_{i t} \mid \mathcal{I}_{i t}^{\prime}\right)=\mu_{i}+X_{i t} \beta_{x}+\lambda \beta \frac{K_{i t}}{K_{i t}+\rho}\left(\bar{v}_{0}+m_{i}\left(\bar{v}_{1}-\bar{v}_{0}\right)\right)+\lambda \beta \frac{\rho}{K_{i t}+\rho}\left(\bar{v}_{0}+m_{i}\left(\tilde{v}_{1}-\bar{v}_{0}\right)\right)$
As $\frac{\rho}{K_{i t}+\rho}=1-\frac{K_{i t}}{K_{i t}+\rho}$ :

$$
\mathbb{E}\left(p_{i t} \mid \mathcal{I}_{i t}\right)=\mu_{i}+X_{i t} \beta_{x}+\lambda \beta\left(\tilde{v}_{1}-\bar{v}_{1}\right) \frac{K_{i t}}{K_{i t}+\rho} m_{i}
$$

Therefore, regressing $p_{i t}$ on $\frac{K_{i t}}{K_{i t}+\rho} m_{i}$, conditional on listing fixed effects and characteristics $X_{i t}$, will identify $\tilde{\beta}_{m}=\lambda \beta\left(\tilde{v}_{1}-\bar{v}_{1}\right)$.


[^0]:    For illustrative purposes only，screenshot of the Airbnb platform captured by the author on May 2016.

