# Online Appendix: Optimal Insurance: Dual Utility, Random Losses and Adverse Selection

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## **Online Appendix A: Stochastic Mechanisms**

We provide an example showing that a stochastic mechanism may be more profitable than the optimal deterministic mechanism. Consider an agent with risk preference represented by  $g(x) = x^2$  and assume that  $\theta \sim U[0, 1]$ . Suppose that there is a single deterministic loss level l, and let the agent's type  $\theta$  be the probability that a loss occurs. The optimal deterministic mechanism consists of full insurance to types  $\theta \geq \frac{2}{3}$  at a price  $\frac{8}{9}l$ , and no insurance for lower types. Using this mechanism, the insurer's profit is  $\frac{1}{54}l \approx 0.185l$ .

Consider now a stochastic direct mechanism of the form  $(t(\theta), p(\theta), l)$  such that: type  $\theta$  pays a premium  $t(\theta)$ ; in exchange, when a loss occurs, the insurer fully reimburses the agent's loss with (conditional) probability  $1 - p(\theta)$ . Note that the above class of mechanisms includes the optimal deterministic mechanism.

If type  $\theta$  reports to be type  $\theta'$  he receives  $-t(\theta') - l$  with probability  $p(\theta')\theta$  and receives  $-t(\theta')$  otherwise. Thus, in the proposed mechanism, this type of agent has a payoff of

$$\tilde{U}(\theta, \theta') = -l - t(\theta') + g \left(1 - p(\theta')\theta\right)l$$

One can verify that  $(t(\theta), p(\theta), l)$  is incentive compatible if and only if p is nonincreasing and

$$\tilde{U}(\theta) = \tilde{U}(\underline{\theta}) - l \int_{\underline{\theta}}^{\theta} p(z)g(1 - p(z)z) dz,$$

where we write  $\tilde{U}(\theta) = \tilde{U}(\theta, \theta)$  for short. The above conditions imply that

$$t(\theta) = -l - \tilde{U}(\underline{\theta}) + g\left(1 - p(\theta)\theta\right)l + l\int_{\underline{\theta}}^{\theta} p(z)g'\left(1 - p(z)z\right)dz$$

By using similar arguments to that of Lemma 1, it can be shown that the individual

rationality constraint holds if and only if

$$\tilde{U}(\underline{\theta}) \ge -l(1-g(1-\underline{\theta})) = 0.$$

From now onward, we only consider mechanism for which  $\tilde{U}(\underline{\theta}) = 0$ . The insurer's profit is

$$\begin{split} \pi(p,t) &= \int_{\underline{\theta}}^{\theta} \left[ t(\theta) - (1-p(\theta))\theta l \right] f(\theta) d\theta \\ &= -l + \int_{\underline{\theta}}^{\overline{\theta}} \left[ g \left( 1 - p(\theta)\theta \right) l + l \int_{\underline{\theta}}^{\theta} p(z)g' \left( 1 - p(z)z \right) dz - (1-p(\theta))\theta l \right] f(\theta) d\theta \\ &= l \int_{\underline{\theta}}^{\overline{\theta}} \left[ g \left( 1 - p(\theta)\theta \right) - (1-p(\theta))\theta + \frac{1 - F(\theta)}{f(\theta)} p(\theta)g' \left( 1 - p(\theta)\theta \right) \right] f(\theta) d\theta - l \\ &= l \int_{\underline{\theta}}^{\overline{\theta}} \left[ \theta(3\theta - 2)p^2 - (3\theta - 2)p + 1 - \theta \right] d\theta - l. \end{split}$$

To obtain the this equality, we used integration by parts:

$$\int_{\underline{\theta}}^{\overline{\theta}} \left[1 - F(\theta)\right] p(\theta)g'\left(1 - p(\theta)\theta\right) d\theta - \int_{\underline{\theta}}^{\overline{\theta}} f(\theta) \left[\int_{\underline{\theta}}^{\theta} p(z)g'\left(1 - p(z)z\right) dz\right] d\theta = 0.$$

The optimal p is then given by by

$$p^*(\theta) = \begin{cases} 1 & \text{if } \theta \leq \frac{1}{2} \\ \frac{1}{2\theta} & \text{if } \frac{1}{2} < \theta < \frac{2}{3} \\ 0 & \text{if } \theta \geq \frac{2}{3}. \end{cases}$$

That is, within the above described class of potentially stochastic mechanisms, it is optimal to offer no insurance to agents with type below  $\frac{1}{2}$ , to offer unconditional full insurance to those with type above  $\frac{2}{3}$ , and to offer to reimburse the loss with (conditional) probability  $1 - \frac{1}{2\theta}$  to intermediate types in  $(\frac{1}{2}, \frac{2}{3})$ . This mechanism yields, approximately, an expected profit of 0.188l > 0.185l, and is thus superior to the optimal deterministic mechanism.

# **Online Appendix B: Finite Number of Losses**

**Proof of Proposition 3.** It holds that

$$H_{\theta}(z) = \begin{cases} 1-\theta & \text{if } z < l_{1} \\ 1-\theta + \theta \sum_{i=1}^{k-1} p_{i} & \text{if } l_{k-1} \le z < l_{k} \text{ and } k \in \{2, ..., n\} \\ 1 & \text{if } z \ge l_{n} \end{cases}$$
  
and  $\frac{\partial H_{\theta}(z)}{\partial \theta} = \begin{cases} -1 & \text{if } z < l_{1} \\ -1 + \sum_{i=1}^{k-1} p_{i} & \text{if } l_{k-1} < z < l_{k} \text{ and } k \in \{2, ..., n\} \\ 0 & \text{if } z > l_{n} \end{cases}$ 

In any incentive compatible mechanism, the menu of deductibles  $D(\theta)$  is non-increasing in the probability of accident  $\theta$ . In particular,  $D(\theta)$  is continuous almost everywhere.

Fix such a non-increasing menu, and let  $\theta_0 = \overline{\theta}$ . Denote by  $\theta_1 = \inf\{\theta : D(\theta) \le l_1\}$ . If this set is empty, define  $\theta_1 = \theta_0 = \overline{\theta}$ . Similarly, for  $i \in \{2, ..., n\}$  define  $\theta_i = \inf\{\theta : D(\theta) \le l_i\}$  with  $\theta_i := \theta_{i-1}$  if the set is empty.

By the monotonicity of  $D(\theta)$ , it holds that  $\underline{\theta} = \theta_n \leq \theta_{n-1} \leq \ldots \leq \theta_1 \leq \theta_0 = \overline{\theta}$ . The insurer's profit becomes then

$$\begin{split} \pi &= \int_{\underline{\theta}}^{\overline{\theta}} \left[ -\mathbb{E}[L(\theta)] + \int_{0}^{D(\theta)} [g(H_{\theta}(z)) - H_{\theta}(z) - \frac{1 - F(\theta)}{f(\theta)} g'(H_{\theta}(z)) \frac{\partial H_{\theta}(z)}{\partial \theta}] dz \right] f(\theta) d\theta - U(\underline{\theta}) \\ &= -\int_{\underline{\theta}}^{\overline{\theta}} \mathbb{E}[L(\theta)] f(\theta) d\theta + \int_{\theta_{1}}^{\overline{\theta}} \left[ \int_{0}^{D(\theta)} [g(H_{\theta}(z)) - H_{\theta}(z) - \frac{1 - F(\theta)}{f(\theta)} g'(H_{\theta}(z)) \frac{\partial H_{\theta}(z)}{\partial \theta}] dz \right] f(\theta) d\theta \\ &+ \sum_{k=2}^{n} \int_{\theta_{k}}^{\theta_{k-1}} \left[ \int_{0}^{D(\theta)} [g(H_{\theta}(z)) - H_{\theta}(z) - \frac{1 - F(\theta)}{f(\theta)} g'(H_{\theta}(z)) \frac{\partial H_{\theta}(z)}{\partial \theta}] dz \right] f(\theta) d\theta \\ &+ \int_{\underline{\theta}}^{\theta_{n}} \left[ \int_{0}^{D(\theta)} [g(H_{\theta}(z)) - H_{\theta}(z) - \frac{1 - F(\theta)}{f(\theta)} g'(H_{\theta}(z)) \frac{\partial H_{\theta}(z)}{\partial \theta}] dz \right] f(\theta) d\theta - U(\underline{\theta}) \\ &= -\int_{\underline{\theta}}^{\overline{\theta}} \mathbb{E}[L(\theta)] f(\theta) d\theta - U(\underline{\theta}) + \int_{\theta_{1}}^{\overline{\theta}} \left[ \int_{0}^{D(\theta)} [g(1 - \theta) - (1 - \theta) + \frac{1 - F(\theta)}{f(\theta)} g'(1 - \theta)] dz \right] f(\theta) d\theta \\ &+ \sum_{k=2}^{n} \int_{\theta_{k}}^{\theta_{k-1}} \int_{0}^{D(\theta)} \left[ g(1 - \theta + \theta \sum_{i=1}^{k-1} p_{i}) - (1 - \theta + \theta \sum_{i=1}^{k-1} p_{i}) \right] dz f(\theta) d\theta \\ &= -\int_{\underline{\theta}}^{\overline{\theta}} \mathbb{E}[L(\theta)] f(\theta) d\theta - U(\underline{\theta}) + \int_{\theta_{1}}^{\overline{\theta}} D(\theta) [g(1 - \theta) - (1 - \theta) + \frac{1 - F(\theta)}{f(\theta)} g'(1 - \theta)] f(\theta) d\theta \\ &= -\int_{\underline{\theta}}^{\overline{\theta}} \mathbb{E}[L(\theta)] f(\theta) d\theta - U(\underline{\theta}) + \int_{\theta_{1}}^{\overline{\theta}} D(\theta) [g(1 - \theta) - (1 - \theta + \theta \sum_{i=1}^{k-1} p_{i}) \right] f(\theta) d\theta \\ &+ \sum_{k=2}^{n} \int_{\theta_{k}}^{\theta_{k-1}} D(\theta) \left[ g(1 - \theta + \theta \sum_{i=1}^{k-1} p_{i}) - (1 - \theta + \theta \sum_{i=1}^{k-1} p_{i}) \right] f(\theta) d\theta . \end{split}$$

By definition, in each interval  $[\theta_k, \theta_{k-1}]$ , the given deductible  $D(\theta)$  belongs to the interval  $[l_{k-1}, l_k]$ , where we denote  $l_0 = 0$ . Note that, on each interval  $[\theta_k, \theta_{k-1}]$ , the obtained expression for profit is linear in D:

$$\int_{\theta_k}^{\theta_{k-1}} D(\theta) \left[ \begin{array}{c} g(1-\theta+\theta\sum_{i=1}^{k-1}p_i) - \left(1-\theta+\theta\sum_{i=1}^{k-1}p_i\right) \\ + \frac{1-F(\theta)}{f(\theta)}g'(1-\theta+\theta\sum_{i=1}^{k-1}p_i) \left(1-\sum_{i=1}^{k-1}p_i\right) \end{array} \right] f(\theta)d\theta.$$

Depending on the sign of the integrand, the above expression is maximized with respect to D at an extreme point of the respective feasible set, i.e., either at  $D^*(\theta) = l_{k-1}$  or at  $D^*(\theta) = l_k$ . Thus, the profit from the given mechanism can be increased by changing all deductibles  $D(\theta)$  on the interval  $[\theta_k, \theta_{k-1}]$  to the value of  $D^*(\theta)$  that maximizes the above expression. The obtained  $D^*$  is non-increasing by construction, and thus also implementable. Hence, we have shown that the search for an optimal mechanism can be confined to menus consisting of at most n + 1 deductibles, where each deductible equals either zero, or one of the possible losses.

#### **Proof of Corollary 1.** Here

$$H_{\theta}(z) = \begin{cases} 1 - \theta \text{ if } z < l \\ 1 \quad \text{if } z \ge l \end{cases} \text{ and } \frac{\partial H_{\theta}(z)}{\partial \theta} = \begin{cases} -1 \text{ if } z \le l \\ 0 \text{ if } z \ge l \end{cases}.$$

The insurer's profit becomes:

$$\begin{aligned} \pi &= \int_{\underline{\theta}}^{\overline{\theta}} \left[ -\mathbb{E}[L(\theta)] + \int_{0}^{D(\theta)} [g(H_{\theta}(z)) - H_{\theta}(z) - \frac{1 - F(\theta)}{f(\theta)} g'(H_{\theta}(z)) \frac{\partial H_{\theta}(z)}{\partial \theta}] dz \right] f(\theta) d\theta - U(\underline{\theta}) \\ &= -l + \int_{\underline{\theta}}^{\overline{\theta}} \left[ \int_{0}^{D(\theta)} [g(1 - \theta) - (1 - \theta) + \frac{1 - F(\theta)}{f(\theta)} g'(1 - \theta)] dz \right] f(\theta) d\theta - U(\underline{\theta}) \\ &= -l + \int_{\underline{\theta}}^{\overline{\theta}} D(\theta) \left[ g(1 - \theta) - (1 - \theta) + \frac{1 - F(\theta)}{f(\theta)} g'(1 - \theta) \right] f(\theta) d\theta - U(\underline{\theta}). \end{aligned}$$

The above expression is linear in D, and hence the pointwise maximum in the above expression is attained at an extreme point of the feasible set: it can be either at D = l or at D = 0, depending on the sign of the virtual value.

### **Online Appendix C: Binary Lotteries**

In this Appendix we document several instances of well-known, non-expected utility formulations that coincide with Yaari's dual utility on the class of binary lotteries (e.g., in an insurance framework with a single, deterministic loss). 1. Gul's [1991] disappointment-averse preferences with linear utility over outcomes:<sup>1</sup>

$$\mathcal{U}(x) = \frac{\alpha}{1 + (1 - \alpha)\beta} \mathbb{E}[x|x \ge CE(x)] + \frac{(1 - \alpha)(1 + \beta)}{1 + (1 - \alpha)\beta} \mathbb{E}[x|x < CE(x)]$$

where CE(x) is a certainty equivalent of lottery  $x \in X$ ,  $\alpha$  is the probability that the outcome of the lottery is above its certainty equivalent, and  $\beta$  is a parameter. For binary lotteries, the above functional form is a special case of Yaari's dual utility with<sup>2</sup>

$$g(p) = \frac{p}{1 + (1 - p)\beta}.$$

2. Versions of the disappointment aversion theories due to Loomes and Sugden [1986], and Jia et al. [2001] with linear utility over outcomes:

$$\mathcal{U}(x) = \mathbb{E}(x) + (e - d)\mathbb{E}\left[\max\left\{x - \mathbb{E}(x), 0\right\}\right],\$$

where e > 0, d > 0. For binary lotteries, this is a special case of Yaari's dual utility with

$$g(p) = p(1 + e - d) + (d - e)p^{2}.$$

Risk aversion (either in the weak or strong sense) is obtained when e < d.

3. The modified Mean-Variance preferences (see Rockafellar et al. [2006]) with linear utility over outcomes are given by<sup>3</sup>:

$$\mathcal{U}(x) = \mathbb{E}(x) - \frac{1}{2}r\mathbb{E}\left[\mid x - \mathbb{E}(x) \mid\right],$$

where  $r \in [0, 1]$ . For binary lotteries this is again a special case of Yaari's preferences where

$$g(p) = p - rp(1-p).$$

<sup>&</sup>lt;sup>1</sup>This is implicit. See also Cereia-Voglio et al. [2020] for an explicit formulation.

<sup>&</sup>lt;sup>2</sup>(Weak) risk aversion corresponds then to  $\beta > \frac{1}{2}$  and a version to mean-preserving spreads corresponds to  $\beta > 0$ .

 $<sup>^{3}</sup>$ The modification relative to the standard mean-variance preferences is needed in order to ensure consistency with FOSD.

# References

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