

Online Appendix

The Focal Luce Model

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Appendix B: Proof of Proposition 7

Suppose we observe $c_{a_i^l}^k \in \{0, 1\}$ for each $l \leq L$, $k \leq N^l$ such that $\sum_i c_{a_i^l}^k = 1$. Here $c_{a_i^l}^k = 1$ means that consumer k (or a consumer visits the menu the k -th time) chooses a_i^l from menu $A^l = \{a_0^l, a_1^l, \dots, a_{m_l}^l\}$. Here, N^l is the number of observations from menu A^l . By our model,

$$c_{a_i^l}^k = \mathbb{1} \left\{ a_i \in \arg \max_{a_j^l \in A^l} \left\{ \bar{u}(a_j^l) + \bar{\delta}(A^l) \mathbb{1}\{a_j^l \in F(A^l)\} + \epsilon_{a_j^l}^k \right\} \right\},$$

where $\epsilon_{a_j^l}^k$ is i.i.d and follows the standard type I extreme distribution. Since

$$\bar{c}_{a_i^l} \equiv \mathbb{E}[c_{a_i^l}^k] = \frac{\exp(\bar{u}(a_i^l) + \bar{\delta}(A^l) \mathbb{1}\{a_i^l \in F(A^l)\})}{\sum_{a_j^l \in A^l} \exp(\bar{u}(a_j^l) + \bar{\delta}(A^l) \mathbb{1}\{a_j^l \in F(A^l)\})},$$

we can write that

$$c_{a_i^l}^k = \frac{\exp(\bar{u}(a_i^l) + \bar{\delta}(A^l) \mathbb{1}\{a_i^l \in F(A^l)\})}{\sum_{a_j^l \in A^l} \exp(\bar{u}(a_j^l) + \bar{\delta}(A^l) \mathbb{1}\{a_j^l \in F(A^l)\})} + \bar{\epsilon}_{a_i^l}^k = \bar{c}_{a_i^l} + \bar{\epsilon}_{a_i^l}^k,$$

where $\bar{\epsilon}_{a_i^l}^k$ is i.i.d across k , $\mathbb{E}[\bar{\epsilon}_{a_i^l}^k] = 0$, and $\sum_{i=0}^{m_l} \bar{\epsilon}_{a_i^l}^k = 0$. We usually aggregate N^l separate choices to obtain choice frequencies as follows:

$$p_i^l = p(a_i^l, A^l) = \frac{\sum_{k=1}^{N^l} c_{a_i^l}^k}{N^l} = \frac{\exp(\bar{u}(a_i^l) + \bar{\delta}(A^l) \mathbb{1}\{a_i^l \in F(A^l)\})}{\sum_{a_j^l \in A^l} \exp(\bar{u}(a_j^l) + \bar{\delta}(A^l) \mathbb{1}\{a_j^l \in F(A^l)\})} + \frac{\sum_{k=1}^{N^l} \bar{\epsilon}_{a_i^l}^k}{N^l} = \bar{c}_{a_i^l} + \bar{\epsilon}_{a_i^l}.$$

Now let us explain the algebra behind LMP we have discussed in Appendix A.A4. Notice that

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$$\begin{aligned}
y_i^l &\equiv \log\left(\frac{p_i^l}{p_0^l}\right) = \log \left[\frac{\exp(\bar{u}(a_i^l) + \bar{\delta}(A^l) \mathbb{1}\{a_i^l \in F(A^l)\}) + \bar{\epsilon}_{a_i^l} \left(\sum_{a_j^l \in A^l} \exp(\bar{u}(a_j^l) + \bar{\delta}(A^l) \mathbb{1}\{a_j^l \in F(A^l)\}) \right)}{\exp(\bar{u}(a_0^l) + \bar{\delta}(A^l) \mathbb{1}\{a_0^l \in F(A^l)\}) + \bar{\epsilon}_{a_0^l} \left(\sum_{a_j^l \in A^l} \exp(\bar{u}(a_j^l) + \bar{\delta}(A^l) \mathbb{1}\{a_j^l \in F(A^l)\}) \right)} \right] \\
&= \log \left[\frac{\exp(v(a_i^l) + \delta_l d_i^l) + \bar{\epsilon}_{a_i^l} \left(\sum_{a_j^l \in A^l} \exp(v(a_j^l) + \delta_l d_j^l) \right)}{1 + \bar{\epsilon}_{a_0^l} \left(\sum_{a_j^l \in A^l} \exp(v(a_j^l) + \delta_l d_j^l) \right)} \right] \\
&= \log \left[\exp(v(a_i^l) + \delta_l d_i^l) + \frac{(\bar{\epsilon}_{a_i^l} - \exp(v(a_i^l) + \delta_l d_i^l) \bar{\epsilon}_{a_0^l}) \left(\sum_{a_j^l \in A^l} \exp(v(a_j^l) + \delta_l d_j^l) \right)}{1 + \bar{\epsilon}_{a_0^l} \left(\sum_{a_j^l \in A^l} \exp(v(a_j^l) + \delta_l d_j^l) \right)} \right] \\
&= \log \left[\exp(v(a_i^l) + \delta_l d_i^l) + \zeta_i^l \right] = v(a_i^l) + \delta_l d_i^l + \bar{\zeta}_i^l,
\end{aligned}$$

where $v(a_i^l) = u(a_i^l) - u(a_0^l)$, $\delta_l = \bar{\delta}(A^l) \text{sign}(\mathbb{1}\{a_i^l \in F(A^l)\} - \mathbb{1}\{a_0^l \in F(A^l)\})$ and $d_i^l = |\mathbb{1}\{a_i^l \in F(A^l)\} - \mathbb{1}\{a_0^l \in F(A^l)\}|$.

We assume that $\delta_l \neq 0$ for each l . We shall prove Proposition 7 in three steps.

Step 1: for each $l \leq L$, $\bar{\zeta}_i^l \rightarrow^P 0$ as $N^l \rightarrow \infty$.

Proof of Step 1. By direct calculations,

$$\bar{\zeta}_i^l = \log \left(1 + \frac{\zeta_i^l}{\exp(v(a_i^l) + \delta_l d_i^l)} \right) = \frac{\zeta_i^l}{\exp(v(a_i^l) + \delta_l d_i^l)} + O((\zeta_i^l)^2).$$

Therefore, it is enough to prove that $\zeta_i^l \rightarrow^P 0$ as $N^l \rightarrow \infty$. By the Law of Large Number, for each l , $\bar{\epsilon}_{a_i^l} = \frac{\sum_{k=1}^{N^l} \bar{\epsilon}_{a_i^k}^k}{N^l} \rightarrow^P 0$ as $N^l \rightarrow \infty$. Therefore,

$$\zeta_i^l = \frac{(\bar{\epsilon}_{a_i^l} - \exp(v(a_i^l) + \delta_l d_i^l) \bar{\epsilon}_{a_0^l}) \left(\sum_{a_j^l \in A^l} \exp(v(a_j^l) + \delta_l d_j^l) \right)}{1 + \bar{\epsilon}_{a_0^l} \left(\sum_{a_j^l \in A^l} \exp(v(a_j^l) + \delta_l d_j^l) \right)} \rightarrow^P 0 \text{ as } N^l \rightarrow \infty.$$

Before we proceed to Step 2, let $x_i^l = a_i^l - a_0^l$. Then we solve the following least squares problem:

$$\min_{\alpha \in \mathbb{R}^n, \delta_l \in \mathbb{R}, d_i^l \in \{0,1\}} \sum_{1 \leq i \leq m^l, l \leq L} (y_i^l - \alpha x_i^l - \delta_l d_i^l)^2,$$

where $y_i^l = \alpha x_i^l + \delta_l d_i^l + \bar{\zeta}_i^l$. Let $(\hat{\alpha}, \hat{\delta} \equiv (\hat{\delta}_l)_l, \hat{d} \equiv (\hat{d}_i^l)_{i,l})$ be the solution to the above LMP.

Step 2: $\lim_{N^l \rightarrow \infty} \text{for all } l \text{ } Pr(\hat{d} = d) = 1$.

Proof of Step 2. For any $\tilde{d} \in \{0,1\}^{\sum_{l=1}^L m_l}$, let

$$X(\tilde{d}) = \begin{bmatrix} x_1^1 & \tilde{d}_1^1 & 0 & \dots & 0 & \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & 0 & \ddots & \vdots & \vdots \\ x_{m_1}^1 & \tilde{d}_{m_1}^1 & 0 & \dots & 0 & \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & 0 & \ddots & \vdots & \vdots \\ x_1^l & 0 & 0 & \dots & \tilde{d}_1^l & \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots & \ddots & \vdots & \vdots \\ x_{m_l}^l & 0 & 0 & \dots & \tilde{d}_{m_l}^l & \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots & \ddots & \vdots & \vdots \\ x_1^L & 0 & 0 & \dots & 0 & \dots & 0 & \tilde{d}_1^L \\ \vdots & \vdots & \vdots & \vdots & \ddots & \ddots & \vdots & \vdots \\ x_{m_L}^L & 0 & 0 & \dots & 0 & \dots & 0 & \tilde{d}_{m_L}^L \end{bmatrix}, \beta = \begin{bmatrix} \alpha_1 \\ \vdots \\ \alpha_n \\ \delta_1 \\ \vdots \\ \delta_L \end{bmatrix} \text{ and } \zeta(\tilde{d}) = \begin{bmatrix} \bar{\zeta}_1^1 + (d_1^1 - \tilde{d}_1^1)\delta_1 \\ \vdots \\ \bar{\zeta}_{m_1}^1 + (d_{m_1}^1 - \tilde{d}_{m_1}^1)\delta_1 \\ \vdots \\ \bar{\zeta}_1^l + (d_1^l - \tilde{d}_1^l)\delta_l \\ \vdots \\ \bar{\zeta}_{m_l}^l + (d_{m_l}^l - \tilde{d}_{m_l}^l)\delta_l \\ \vdots \\ \bar{\zeta}_1^L + (d_1^L - \tilde{d}_1^L)\delta_L \\ \vdots \\ \bar{\zeta}_{m_L}^L + (d_{m_L}^L - \tilde{d}_{m_L}^L)\delta_L \end{bmatrix}$$

For given \tilde{d} , let the vector $(\alpha(\tilde{d}), \delta(\tilde{d}))$ be the solution to the minimization problem:

$$(B1) \quad \min_{\alpha \in \mathbb{R}^n, \delta_l \in \mathbb{R}} \sum_{1 \leq i \leq m^l, l \leq L} (y_i^l - \alpha x_i^l - \delta_l \tilde{d}_i^l)^2.$$

Moreover, let

$$S(\tilde{d}) = \sum_{1 \leq i \leq m^l, l \leq L} (y_i^l - \alpha(\tilde{d}) x_i^l - \delta(\tilde{d})_l \tilde{d}_i^l)^2.$$

Indeed, by the definition of $(\hat{\alpha}, \hat{\delta} \equiv (\hat{\delta}_l)_l, \hat{d} \equiv (\hat{d}_{i,l}^l)_{i,l})$, we shall have $(\hat{\alpha}, \hat{\delta}) = (\alpha(\hat{d}), \delta(\hat{d}))$ and $S(\hat{d}) \leq S(\tilde{d})$ for any \tilde{d} .

Note that Equation (B1) is an OLS regression for $y = X(\tilde{d})\beta + \zeta(\tilde{d})$. Therefore, following standard calculations, we have

$$S(\tilde{d}) = \sum_{1 \leq i \leq m^l, l \leq L} (y_i^l - \alpha(\tilde{d}) x_i^l - \delta(\tilde{d})_l \tilde{d}_i^l)^2 = \zeta(\tilde{d})^T \left(I - X(\tilde{d})^T (X(\tilde{d})^T X(\tilde{d}))^{-1} X(\tilde{d}) \right) \zeta(\tilde{d}).$$

By way of contradiction, suppose that $\hat{d} \neq d$; i.e., $\Delta \equiv \zeta(\hat{d}) - \zeta(d) = \zeta(\hat{d}) - \bar{\zeta} \neq 0$. First, note that

$$\begin{aligned} S(\hat{d}) &= (\bar{\zeta} + \Delta)^T (I - X(\hat{d})^T (X(\hat{d})^T X(\hat{d}))^{-1} X(\hat{d})) (\bar{\zeta} + \Delta) = \\ &= \Delta^T (I - X(\hat{d})^T (X(\hat{d})^T X(\hat{d}))^{-1} X(\hat{d})) \Delta \\ &+ (\bar{\zeta} + \Delta)^T (I - X(\hat{d})^T (X(\hat{d})^T X(\hat{d}))^{-1} X(\hat{d})) \bar{\zeta} + \bar{\zeta}^T (I - X(\hat{d})^T (X(\hat{d})^T X(\hat{d}))^{-1} X(\hat{d})) \Delta. \end{aligned}$$

Notice that for any $\tilde{d} = (\tilde{d}_{i,l}^l)_{i,l} \in \{0, 1\}^{\sum_{l=1}^L m_l}$, the sum of the squares of residuals of an OLS regression for $z = X(\tilde{d})\gamma + \Delta$ is strictly positive; i.e.,

$$\Delta^T (I - X(\tilde{d})^T (X(\tilde{d})^T X(\tilde{d}))^{-1} X(\tilde{d})) \Delta = (z - X(\tilde{d})\hat{\gamma})^T (z - X(\tilde{d})\hat{\gamma}) > 0.$$

Moreover, since there are only finitely many different \tilde{d} , there is $\eta > 0$ such that

$$\Delta^T (I - X(\tilde{d})^T (X(\tilde{d})^T X(\tilde{d}))^{-1} X(\tilde{d})) \Delta > \eta \text{ for all } \tilde{d}.$$

However, in order to have $S(\hat{d}) \leq S(d)$, we shall have

$$-\eta > (\bar{\zeta} + \Delta)^T (I - X(\hat{d})^T (X(\hat{d})^T X(\hat{d}))^{-1} X(\hat{d})) \bar{\zeta} + \bar{\zeta}^T (I - X(\hat{d})^T (X(\hat{d})^T X(\hat{d}))^{-1} X(\hat{d})) \Delta \\ - \bar{\zeta}^T (I - X(d)^T (X(d)^T X(d))^{-1} X(d)) \bar{\zeta}.$$

However, since $\bar{\zeta} \rightarrow^P 0$ as $N^l \rightarrow \infty$ for all l by Step 1,

$$(\bar{\zeta} + \Delta)^T (I - X(\hat{d})^T (X(\hat{d})^T X(\hat{d}))^{-1} X(\hat{d})) \bar{\zeta} + \bar{\zeta}^T (I - X(\hat{d})^T (X(\hat{d})^T X(\hat{d}))^{-1} X(\hat{d})) \Delta \\ - \bar{\zeta}^T (I - X(d)^T (X(d)^T X(d))^{-1} X(d)) \bar{\zeta} \rightarrow^P 0 \text{ as } N^l \rightarrow \infty \text{ for all } l.$$

Therefore, the above is a contradiction to the assumption that $\Delta \neq 0$. In other words, $\hat{d} = d$ as $N^l \rightarrow \infty$ for all l ; i.e., the focal sets will be correctly identified.

Step 3: $(\hat{\alpha}, \hat{\delta}) \rightarrow^P (\alpha, \delta)$ as $N^l \rightarrow \infty$ for all l .

Given Step 2, by standard results the OLS regression returns consistent estimators for α and δ .

Appendix C: Maximum Likelihood Estimation

Consider a data that defined in Appendix B. Let $c_l^k = (c_{a_i^l}^k)_{i=1}^{m_l}$, $n_{a_i^l}^l = \sum_{k=1}^{N^l} c_{a_i^l}^k$, and $p_i^l = p(a_i^l, A^l) = \frac{n_{a_i^l}^l}{N^l}$. Let us first derive the likelihood function. By our model,

$$c_{a_i^l}^k = \mathbf{1} \left\{ a_i \in \arg \max_{a_j^l \in A^l} \{ \bar{u}(a_j^l) + \bar{\delta}(A^l) \mathbf{1}\{a_j^l \in F(A^l)\} + \epsilon_{a_j^l}^k \} \right\},$$

where $\epsilon_{a_j^l}^k$ is i.i.d and follows the standard type I extreme distribution. Since

$$Pr(c_{a_i^l}^k = 1) = \frac{\exp(\bar{u}(a_i^l) + \bar{\delta}(A^l) \mathbf{1}\{a_i^l \in F(A^l)\})}{\sum_{a_j^l \in A^l} \exp(\bar{u}(a_j^l) + \bar{\delta}(A^l) \mathbf{1}\{a_j^l \in F(A^l)\})},$$

$$Pr(c_l = c_l^k) = \prod_{i=1}^{m_l} \left(\frac{\exp(\bar{u}(a_i^l) + \bar{\delta}(A^l) \mathbf{1}\{a_i^l \in F(A^l)\})}{\sum_{a_j^l \in A^l} \exp(\bar{u}(a_j^l) + \bar{\delta}(A^l) \mathbf{1}\{a_j^l \in F(A^l)\})} \right)^{c_{a_i^l}^k}.$$

Therefore, the likelihood function for given $\bar{u}, \bar{\delta}, F$ is

$$\mathcal{L}(\{c_l^k\}_{k \leq N^l, l \leq L} | \bar{u}, \bar{\delta}, F) = \prod_{l=1}^L \prod_{i=1}^{m_l} \left(\frac{\exp(\bar{u}(a_i^l) + \bar{\delta}(A^l) \mathbf{1}\{a_i^l \in F(A^l)\})}{\sum_{a_j^l \in A^l} \exp(\bar{u}(a_j^l) + \bar{\delta}(A^l) \mathbf{1}\{a_j^l \in F(A^l)\})} \right)^{n_{a_i^l}^l}.$$

Then to find the maximum likelihood estimators, we shall solve

$$\max_{\bar{u}, \bar{\delta}, F} \mathcal{L} \left(\{c_i^k\}_{k \leq N^l, l \leq L} | \bar{u}, \bar{\delta}, F \right);$$

equivalently, we solve

$$\max_{\bar{u}, \bar{\delta}, F} \sum_{l=1}^L \left[\sum_{i=1}^{m^l} n_{a_i^l}^l (\bar{u}(a_i^l) + \bar{\delta}(A^l) \mathbf{1}\{a_i^l \in F(A^l)\}) - N^l \log \left(\sum_{j=1}^{m^l} \exp(\bar{u}(a_j^l) + \bar{\delta}(A^l) \mathbf{1}\{a_j^l \in F(A^l)\}) \right) \right].$$

Let us derive the first order conditions.

FOC for u : for each a ,

$$(C1) \quad 0 = \sum_{l: a=a_i^l} N^l \left(p_i^l - \frac{\exp(\bar{u}(a_i^l) + \bar{\delta}(A^l) \mathbf{1}\{a_i^l \in F(A^l)\})}{\sum_{a_j^l \in A^l} \exp(\bar{u}(a_j^l) + \bar{\delta}(A^l) \mathbf{1}\{a_j^l \in F(A^l)\})} \right).$$

FOC for $\bar{\delta}$: for each $l \leq L$,

$$(C2) \quad 0 = \sum_{i=1}^{m^l} \mathbf{1}\{a_i^l \in F(A^l)\} \left(p_i^l - \frac{\exp(\bar{u}(a_i^l) + \bar{\delta}(A^l) \mathbf{1}\{a_i^l \in F(A^l)\})}{\sum_{a_j^l \in A^l} \exp(\bar{u}(a_j^l) + \bar{\delta}(A^l) \mathbf{1}\{a_j^l \in F(A^l)\})} \right).$$

FOC for F : for l and $a_i^l \in A^l$, $a_i^l \in F(A^l)$ if and only if

$$(C3) \quad \frac{\exp(p_i^l \bar{\delta}(A^l)) - 1}{\exp(\bar{\delta}(A^l)) - 1} > \frac{\exp(\bar{u}(a_i^l))}{\exp(\bar{u}(a_i^l)) + \sum_{j \neq i} \exp(\bar{u}(a_j^l) + \bar{\delta}(A^l) \mathbf{1}\{a_j^l \in F(A^l)\})}.$$

For linear models, $u(x) = \alpha \cdot x = \sum_{k=1}^n \alpha_k x_k$, Equation (C1) should be replaced with FOCs for α : for any $k \leq n$,

$$0 = \sum_{l=1}^L N^l \sum_{i=1}^{m^l} a_{i,k}^l \left(p_i^l - \frac{\exp(\alpha \cdot a_i^l + \bar{\delta}(A^l) \mathbf{1}\{a_i^l \in F(A^l)\})}{\sum_{a_j^l \in A^l} \exp(\alpha \cdot a_j^l + \bar{\delta}(A^l) \mathbf{1}\{a_j^l \in F(A^l)\})} \right).$$