Quantile Local Projections: Identification, Smooth Estimation, and Inference

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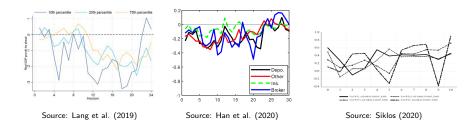
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Introduction

- Dynamic causal effects of shocks impulse response functions
- Local Projections (LP)
 - Simple and flexible alternative to VAR
 - Only capture the mean lose information
 - Usually estimated by least squares not robust
- Quantile Regression (QR)
 - Describes the entire conditional distribution
 - Special case median regression robust
- LP estimated by QR quantile local projections (QLP)
 - Increasingly popular in empirical macroeconomics and finance

 - Researchers face three important challenges in practice.

Open questions and contribution



- QLP open questions
 - How to identify quantile treatment effects?
 - How to construct closed-form confidence intervals?
 - How to smooth the impulse response functions?
- Contribution: coherent econometric framework for QLP
 - Solves the open questions identification, inference, smoothing
 - Financial conditions affect the entire GDP growth distribution.
 - Heterogeneous effects of monetary policy

Local projections: very simple examples (OLS vs QR)

- Univariate time series $\{y_t\}$, forecast horizon $h \in \{1, 2, 3, \dots\}$
- Standard local projections

$$IR(h) = \frac{\mathbb{E}[y_{t+h}|y_t, y_{t-1}, \dots] = \alpha_h + \beta_h y_t}{\partial y_t}$$

$$= \frac{\partial \mathbb{E}[y_{t+h}|y_t, y_{t-1}, \dots]}{\partial y_t} = \beta_h$$

• Quantile local projections, quantile $au \in (0,1)$

$$IR(h,\tau) = \frac{Q_{\tau}[y_{t+h}|y_t, y_{t-1}, \dots] = \alpha_h(\tau) + \beta_h(\tau)y_t}{\partial y_t} = \beta_h(\tau)$$

Theoretical results – overview of contribution

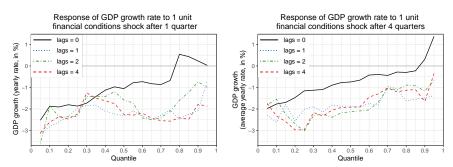
- Causality in dynamic quantile models
 - Lemma: Structural quantile form of a stochastic process
 - Connects multivariate time series and IV-QR (Chernozhukov & Hansen, 2005, 2006, 2008)
 - Every multivariate higher-order Markov process can be represented in a nonlinear simultaneous equations model with iid disturbances.
 - Definition: Impulse response function
 - Quantile treatment effect of an exogenous impulse
 - Can be identified by recursive short-run restrictions or external instruments
- Recursive short-run restrictions identification, estimation, inference
 - Proposition: Identification
 - Definition: SQLP estimator (QR + roughness penalties)
 - Propositions: Consistency, asymptotic normality
 - Information criteria for optimal smoothing
- External instruments identification, estimation, inference
 - Proposition: Identification
 - Definition: SQLPI estimator (IV-QR + roughness penalties)
 - Proposition: Inference with weak instruments
 - · Quasi-information criteria for optimal smoothing

- Adrian et al. (2019, AER) study the effect of financial conditions on GDP in the US.
- Their main specification involves two time series:
 - Real GDP growth rate, denoted gGDP_t
 - National Financial Conditions Index, denoted NFCI_t
- Sample: 1973Q1 to 2015Q4 (172 observations)
- Their main specification is:

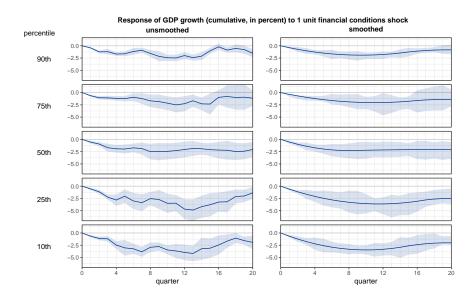
$$Q_{\tau}(\textit{gGDP}_{t+h}|\textit{gGDP}_t,\textit{NFCI}_t) = \beta_{0,0,h}(\tau) + \beta_{1,0,h}(\tau)\textit{gGDP}_t + \beta_{2,0,h}(\tau)\textit{NFCI}_t$$

- Contrasting it with our Proposition 2, we notice that:
 - Their specification entails the following identification restriction:
 NFCI has no contemporaneous effect on GDP growth distribution.
 - Also, their specification does not control for any lags.

Are the results of Adrian et al. (2019) sensitive to the number of lags?
 The black solid lines (no lags) are their main estimates:



- The figures reveal that if we add at least 1 lag, then:
 - the effect in the upper part of the distribution flips the sign from positive to negative, and
 - the heterogeneity across quantiles mostly disappears.



Conclusion

- Quantile regression local projections increasingly popular
 - More robust and more informative than standard LP
 - Researchers face 3 important challenges in practice.
- Contribution: framework that overcomes the challenges
 - Identification: popular identification schemes
 - Smoothing: roughness penalties computationally convenient
 - Inference: closed-form and bootstrap confidence intervals
- Financial conditions affect the entire GDP growth distribution not just the lower part as argued by Adrian et al. (2019).
- Statistically significant effect of monetary policy in line with Gertler & Karadi (2015), contrary to Stock & Watson (2018)
- Heterogeneous effects of conventional monetary policy:
 - More effective at curbing inflation than at preventing deflation
 - More effective at taming excessive economic growth than at supporting recovery

Annex – Related literature

- LP Jordà (2005)
- LP with instrumental variables Stock & Watson (2018)
- Smooth LP Barnichon & Brownlees (2019)
- Quantile Autoregression Koenker & Xiao (2006)
- IV-QR Chernozhukov & Hansen (2005, 2006, 2008)
- Quantile VAR White et al. (2015), Montes-Rojas (2018, 2019), Ruzicka (2021b), Chavleishvili & Manganelli (2024)
- Examples of research that uses QLP
 - Distante et al. (2013) technology shocks, business cycle reversals
 - Winkler & Linnemann (2015) fiscal multipliers
 - Adrian et al. (2019, AER) financial conditions and GDP
 - Boyarchenko et al. (2020) bank capital and GDP
 - Han et al. (2020) systemic risk of financial institutions
 - Lang & Forletta (2020) risks to bank profitability
 - Mano & Sgherri (2020) capital flow shocks
 - Siklos (2020) inflation dynamics

• The observable macroeconomic variables follow a K-variate stochastic process $\{Y_t\}$. Denote $X_t = (Y'_{t-1}, \dots, Y'_{t-P})'$

Assumption 1 (Stationary Higher Order Markov Process)

Let $\{Y_t\}_{t\in\mathbb{Z}}=\{(y_{1t},y_{2t},\ldots,y_{Kt})'\}_{t\in\mathbb{Z}}$ be a K-variate strictly stationary stochastic process with natural filtration $\{\mathcal{F}_t\}_{t\in\mathbb{Z}}$. For $P\geq 1$, let $\underline{Y}_{t-1}=(Y'_{t-1},\ldots,Y'_{t-P})'$. Assume that $\forall y\in\mathbb{R}^K$: $P(Y_t\leq y|\mathcal{F}_{t-1})=P(Y_t\leq y|\underline{Y}_{t-1})$.

- The process $\{Y_t\}$ covers as a special case strictly stationary vector autoregressions with independent errors.
- To allow structural identification, we represent $\{Y_t\}$ in a suitable way (structural quantile form).

Lemma 1 (Structural Quantile Form of a Stochastic Process)

Under Assumption 1 there exist functions $S_1, S_2, \ldots, S_K : \mathbb{R}^{K(P+1)} \mapsto \mathbb{R}$, non-decreasing in their first argument, and independent standard uniform random variables $\{u_{1t}, u_{2t}, \ldots, u_{Kt}\}_{t \in \mathbb{Z}}$ such that:

For all $\tau \in (0,1)$,

$$Q_{\tau}\left(y_{1t} - S_{1}(\tau|y_{2,t}, y_{3,t}, \dots, y_{K,t}, \underline{Y}_{t-1}) \middle| \{u_{1t}, \dots, u_{Kt}\} \setminus u_{1t}, \mathcal{F}_{t-1}\right) = 0,$$

$$Q_{\tau}\left(y_{2t} - S_{2}(\tau|y_{1,t}, y_{3,t}, \dots, y_{K,t}, \underline{Y}_{t-1}) \middle| \{u_{1t}, \dots, u_{Kt}\} \setminus u_{2t}, \mathcal{F}_{t-1}\right) = 0,$$

$$Q_{\tau}\Big(y_{\mathcal{K}t}-S_{\mathcal{K}}(\tau|y_{1,t},y_{2,t},\ldots,y_{\mathcal{K}-1,t},\underline{Y}_{t-1})\Big|\{u_{1t},\ldots,u_{\mathcal{K}t}\}\setminus u_{\mathcal{K}t},\mathcal{F}_{t-1}\Big)=0.$$

 Y_t is the a.s. unique solution of (1).

- S_1, S_2, \dots, S_K are structural quantile functions (Chernozhukov & Hansen, 2008).
- The disturbances $u_{1t}, u_{2t}, \dots, u_{Kt}$ determine which quantiles occur.

Assumption 2 (Counterfactual Stochastic Process)

Let $\{Y_t\}_{t\in\mathbb{Z}}, \{u_{1t}, u_{2t}, \ldots, u_{Kt}\}_{t\in\mathbb{Z}}, S_1, S_2, \ldots, S_K$ be as in Lemma 1. Take $i\in\{1,2,\ldots,K\}$, $s\in\mathbb{R}$ such that $\{Y_t\}_{t\in\mathbb{Z}}$ has the identical support as $\{\tilde{Y}_t\}_{t\in\mathbb{Z}}=\{(\tilde{y}_{1t},\tilde{y}_{2t},\ldots,\tilde{y}_{Kt})'\}_{t\in\mathbb{Z}}$ that is given by

$$\begin{split} \tilde{y}_{1,t} &= \mathbb{I}(t<0)y_{1t} &+ \mathbb{I}(t\geq 0)S_{1}(u_{1,t}|\tilde{y}_{2,t},\tilde{y}_{3,t},\ldots,\tilde{y}_{K,t},\tilde{Y}_{t-1}), \\ &\vdots \\ \tilde{y}_{i-1,t} &= \mathbb{I}(t<0)y_{i-1,t} &+ \mathbb{I}(t\geq 0)S_{i-1}(u_{i-1,t}|\tilde{y}_{1,t},\ldots,\tilde{y}_{i-2,t},\tilde{y}_{i,t},\ldots,\tilde{y}_{K,t},\tilde{Y}_{t-1}), \\ \tilde{y}_{i,t} &= \mathbb{I}(t\leq 0)y_{it} + \mathbb{I}(t=0)s + \mathbb{I}(t>0)S_{i}(u_{i,t}|\tilde{y}_{1,t},\ldots,\tilde{y}_{i-2,t},\tilde{y}_{i,t},\ldots,\tilde{y}_{K,t},\tilde{Y}_{t-1}), \\ \tilde{y}_{i+1,t} &= \mathbb{I}(t<0)y_{i+1,t} &+ \mathbb{I}(t\geq 0)S_{i+1}(u_{i+1,t}|\tilde{y}_{1,t},\ldots,\tilde{y}_{i,t},\tilde{y}_{i+2,t},\ldots,\tilde{y}_{K,t},\tilde{Y}_{t-1}), \\ \vdots &\\ \tilde{y}_{K,t} &= \mathbb{I}(t<0)y_{Kt} &+ \mathbb{I}(t\geq 0)S_{K}(u_{K,t}|\tilde{y}_{1,t},\tilde{y}_{2,t},\ldots,\tilde{y}_{K-1,t},\tilde{Y}_{t-1}), \end{split}$$

where $\tilde{Y}_{t-1} = (\tilde{Y}'_{t-1}, \dots, \tilde{Y}'_{t-P})'$. We call $\{\tilde{Y}_t\}_{t \in \mathbb{Z}}$ the counterfactual process.

- QR models are random coefficient models, so responses to shocks are stochastic. We want to describe their distribution.
 - The QIR definition differs from other definitions in the literature valid for any point-identification scheme and reflects the QTE of a shock.

Definition 1 (Impulse Response Function)

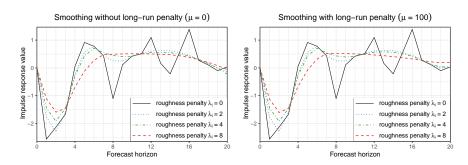
Let $\{Y_t\}_{t\in\mathbb{Z}}, \{u_{1t}, u_{2t}, \dots, u_{Kt}\}_{t\in\mathbb{Z}}, \{\tilde{Y}_t\}_{t\in\mathbb{Z}}, i \text{ and } s \text{ be as in Assumption 2. Let } j\in\{1,2,\dots,K\} \text{ and } \tau\in(0,1).$ Let $X_{it}=\underline{Y}_{t-1} \text{ or } X_{it}=(y_{1t},\dots,y_{it},\underline{Y}'_{t-1})'.$ Define the response of the jth variable to the ith impulse (or shock) s at quantile τ and horizon $h\in\mathbb{N}_0$ as

$$IR_{j,i}^{h,\tau,s}(Y_0,X_{i0}) = Q_{\tau}(Q_{\tau}(\tilde{y}_{jh}|u_{i0},X_{i0})|Y_0,\underline{Y}_{-1}) - Q_{\tau}(Q_{\tau}(y_{jh}|u_{i0},X_{i0})|Y_0,\underline{Y}_{-1}).$$
 (1)

- QIR is the QTE of the ith shock on variable j after h periods.
- Under appropriate identification restrictions (formulated in the paper),
 it is possible to estimate QIR consistently by quantile local projections.

Recursive short-run restrictions – smoothing illustration

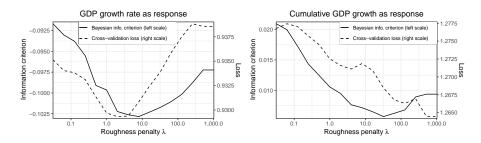
The data and specification are taken from an empirical application.



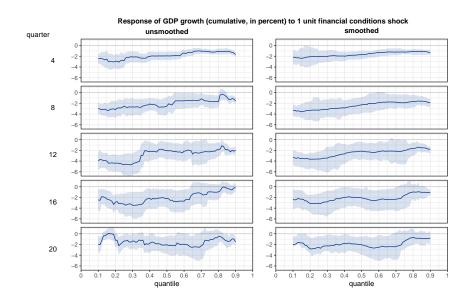
- Higher roughness penalty means more smoothing.
- Long-run penalty straightens the function at the end.

Recursive short-run restrictions – optimal smoothing

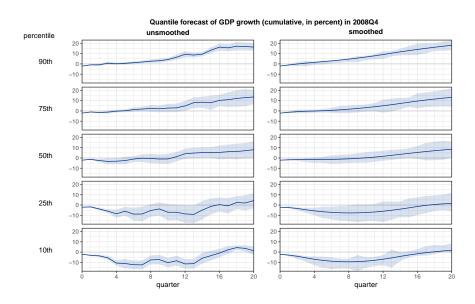
- The roughness penalty value can be chosen by information criteria.
- Let us illustrate the roughness penalty selection on a real example.



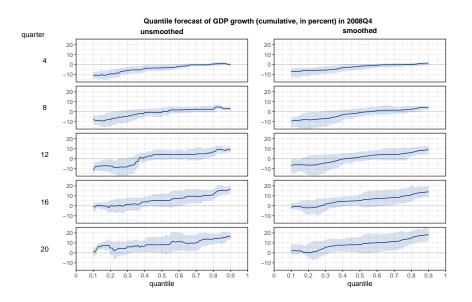
- Bias vs variance tradeoff: $\lambda \uparrow \Rightarrow \downarrow$ variance, \uparrow bias
- Intuitively, the smoothing makes the estimates more accurate by reducing their variance, at the cost of a small bias.



SQLP application: quantile forecasts



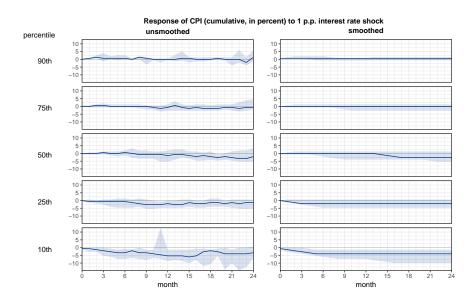
SQLP application: quantile forecasts



SQLPI application: monetary policy shocks

- Specification and data from Stock & Watson (2018) (except for the interest rate)
- Four time series:
 - Federal Funds rate (instead of one-year treasury rate)
 - Growth of industrial production index
 - Growth of consumer price index (CPI)
 - Excess bond premium
- Identification by an external instrument changes in Federal Funds futures rates around FOMC announcement dates
- US monthly data, 1990m1 to 2012m6 (270 observations)

SQLPI application: monetary policy shocks



SQLPI application – effect of 1 p.p. interest rate shock

