

Politicians' Asset Allocation and Economic Bill Proposals*

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Abstract

We study the impact of politicians' asset allocation on their economic bill proposals. Using a novel database of comprehensive financial disclosures of Congress members in South Korea, we find that the Congress members with more real estate assets in their portfolios are less likely to propose economic bills tightening the real estate market. To address endogeneity concerns, we use an instrumental variable uniquely available in our empirical setting; an unexpected earthquake in a local city in South Korea. Controlling for other confounding factors, we argue that the result is mainly driven by politicians' financial interests. Overall, our findings suggest that politicians' personal financial positions have a material impact on their choice of economic bill proposals.

JEL classification: D72, G38, K25, P16

Keywords: real estate, political economy, economic bills, politician

*We thank Viral Acharya, Sumit Agarwal, Murillo Campello, Seungmoon Choi(discussant), Joseph Grimes(discussant), John Hackney, Jin Q Jeon(discussant), Karam Kang, Sungkwan Lee(discussant), David Schoenherr, Johan Sulaeman(discussant), Sheridan Titman, Hye Young You, Alex van de Minne, and Zhengyi Zhou(discussant) for their comments. We also thank the conference participants at 2022 ABFER, 2022 CICF, 2022 APAD, 2022 AsRES-AREUEA Japan, and 2022 FMA for their helpful comments and suggestions. This draft: November 10, 2022.

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1. Introduction

Economic bill proposals are the nascent form of economic policies in democratic countries. In most democratic countries, congress members have an extensive right to propose economic bills and an exclusive privilege to vote for them, making the bills into law. And the laws shaping economic policies have a tremendous impact on financial markets and asset prices. Therefore, it is imperative to understand how congress members propose economic bills in the very first stage of the legislative process. However, the extant literature focuses on politicians' behaviors after some bills are already proposed. A widely studied topic is the politicians' voting behavior in the last stage of the legislative process (e.g., Mian et al. (2010); Tahoun and van Lent (2019)). Prior studies do not offer a rigorous analysis of what determines politicians' economic bill proposals in the early stage of politicians' legislative activities.

In this paper, we study how politicians' personal financial interests affect their economic bill proposals. Specifically, we study how politicians' real estate asset holdings influence their likelihood to propose bills tightening the real estate market. Despite the soaring housing prices inciting social unrest worldwide in the past decade (Crawford (2021)), there was little success in addressing these issues through drawing effective economic policies. Among various reasons for the unsuccessful policies, the misaligned incentive of politicians with their extensive holdings of real estate has been an unexplored but potentially important reason for the failures.

Understanding the impact of politicians' asset allocations on their economic bill proposals is important for various reasons. First, politicians' influence on asset prices is widely conjectured but poorly understood. In this paper, we offer an analysis showing their influence on the asset market through the very first stage of their legislative activities—bill proposals. Real estate prices may not reflect the fundamentals of the housing market because the price movement is driven by politicians' personal interests apart from the fundamentals. A distorted asset price implies that a country's resources are inefficiently allocated throughout the economy.

Second, Congress members' exclusive right to make law is based on the belief that the politicians will represent their constituents' best interests. Our research question sheds addi-

tional light on the misaligned interests between politicians and their constituents in the very early stage of their legislative activities. To the extent that politicians’ personal interest dictates what economic bills are proposed, people should be vigilant of the intention of proposed economic bills or why some preferred bills are never proposed in the first place. Voting members of a community would need to consider implementing an additional monitoring system of politicians’ financial exposures.

To answer our research question, we use a comprehensive financial disclosure data of all South Korean Congress members from 2011 to 2020. The data provides a unique opportunity to investigate the impact of politicians’ personal portfolio holdings on economic bill proposals. We focus on the value of real estate assets in the South Korean Congress members’ personal portfolios as a proxy for their private interests in the real estate market. Although real estate takes a significant portion of household wealth, politicians’ disclosure requirement of real estate assets is less stringent than stock holdings in most countries. For example, in the U.S., compared to the complete disclosures of stock holdings and tradings, there are several exemptions given to the disclosure of real estate assets.¹ The database of South Korean Congress members’ financial positions provides a complete and detailed description of all asset holdings, enabling us to accurately calculate all Congress members’ portfolio composition in a panel dataset. The data includes detailed information on all assets and liabilities owned by Congress members with exact market value or fairly assessed value. The data is also free from biases of selective disclosures because it contains the asset holdings of all family members, including spouses, lineal ascendants, and descendants.

Additionally, our empirical setting has the advantage of measuring the policy directions of all proposed bills by Congress members. Using a textual analysis technique, we identify economic bill proposals related to real estate and classify their policy directions of tightening

¹The regulation requires mandatory disclosure only on the income-generating real estate, providing rooms to exclude primary residences, vacation homes, and vacant second homes (U.S. Senate Select Committee on Ethics (2021); U.S. House of Representatives Committee on Ethics (2021)). Moreover, real estate values are subject to optional disclosure, letting a large number of real estate be reported with 0 dollar value (Baldauf et al. (2021)), which results in a lax screening of real estate assets than other asset classes.

vs. loosening the real estate market. We can also afford to include a comprehensive list of controls capturing various demographic and political characteristics of Congress members.

To preview our result, we find that the reluctance of proposing economic bills tightening the real estate market (*Reluctance of Tightening Real Estate*) increases with the fraction of real estate in a Congress member’s asset portfolio (*Ratio of Real Estate*). In other words, Congress members are less likely to propose a bill that suppresses the real estate market when their interest in real estate is large. We control for various factors that can potentially affect the relation, such as the total asset size, other asset classes in the portfolios, demographic and legislative characteristics of Congress members. The result is also robust to controlling for macro-level variables or year fixed effects. The economic significance is considerable; a one-standard-deviation increase in the *Ratio of Real Estate* increases 16% of the average value of *Reluctance of Tightening Real Estate*.

To substantiate the impact of Congress members’ private interest, it is important to control for other determinants of their legislative behaviors such as constituents’ interests and ideology. In our baseline model, we include the electoral district fixed effect to control for voters’ interests and preferences at the constituent level. In addition, we disaggregate the real estate holdings by the real estate assets location to investigate whether the holdings outside of Congress members’ electoral district also reduce their likelihood to propose economic bills tightening the real estate market. We find it does, indicating that constituents’ interest does not solely drive our results. Moreover, we include affiliated party fixed effects to control for politicians’ ideology. This approach well serves our purpose because we empirically find that their affiliated party distinctively defines the politicians’ ideology in our empirical setting.

If the positive association between the reluctance of proposing tightening real estate bills (*Reluctance of Tightening Real Estate*) and the ratio of real estate assets (*Ratio of Real Estate*) is indeed due to the private interest of Congress members in the real estate market, we expect to find a more significant effect when the private interest is at a larger stake. We consider a case when Congress members have more items of real estate assets in their portfolios, controlling

for the value of real estate assets in their total wealth. The Congress members with more items of real estate assets are likely to hold real estate assets for investment purposes and would be more exposed to the real estate market conditions. We find that the positive association between the *Ratio of Real Estate* and *Reluctance of Tightening Real Estate* gets stronger for the Congress members in the top quintile group in terms of the number of real estate assets.

Politicians are more likely to pursue their private interests when their political positions are more secure and entrenched. We define a Congress member as entrenched 1) if the Congress member won more than 50% of votes in the latest election, or 2) if the votes in the electoral district in the latest election are highly concentrated to the Congress member, or 3) if the Congress member's affiliated party took the majority seats in the metropolitan area in the latest election. We find that the entrenched Congress members are less likely to propose tightening real estate bills when the *Ratio of Real Estate* increases.

We conduct various robustness tests. Our results hold across different types of real estate assets and their actual owners within a family. And the result does not change when using alternative specifications to define the tightening real estate bills. In a placebo test, we estimate the impact of *Ratio of Real Estate* on the likelihood of proposing tightening economic bills unrelated to real estate. We find only insignificant effects, suggesting that our baseline result is not driven by a Congress member's unobservable stance about economic policies. To address the concern that the result is driven by time-varying local interests and economic conditions, we augment the baseline regression model with metropolitan area \times year fixed effects.² The baseline result still holds.

Despite the extensive control variables in our regression analysis, the ratio of real estate assets in Congress members' portfolios is not exogenously determined and prone to potential omitted variable biases. We address the endogeneity concern with an instrumental variable that exogenously affects the ratio of Congress members' real estate assets in their portfolios. The instrumental variable is an earthquake that happened at an industrialized city in South

²Because South Korea has a single constituency legislative system, our empirical setting does not allow including electoral district \times year fixed effects in regression specifications.

Korea, called Pohang, in 2017. South Korea is known as a safe place from seismic activity and the unexpected earthquake—the second-largest one in the country’s modern history—lead to a significant shock to the people resulting in substantial price drop of nearby real estate assets. We first confirm the relevance condition that *Ratio of Real Estate* significantly drops for the Congress members who are exposed to the shock by holding real estate assets near the earthquake epicenter. The exclusion restriction condition reasonably holds because the localized shock would not directly affect the national level economic bill proposals at the Congress. Using the shock as an instrumental variable, the second stage regression supports the positive impact of *Ratio of Real Estate* on the reluctance of proposing bills tightening the real estate market (*Reluctance of Tightening Real Estate*).

When individual Congress members show the legislative behavior favoring their own private interests, a more important question is whether their individual behaviors can have an aggregate impact. To this end, we measure the Congress-level exposure to the real estate market by aggregating all Congress members’ real estate assets values to their total assets. We find that the total number of proposed bills that tighten the real estate market decreases with the Congress-level exposure to real estate. This finding offers an important insight into how aggregate legislative decisions reflect each member’s personal interests which subsequently affect the economic policies at the national level. As corporate board diversity improves firm value by lowering volatility and improving firm performance (e.g., Bernile et al. (2018)), the diversity among Congress members in terms of their financial positions may balance the legislative decisions in the Congress. Our empirical result indicates that Congress without the diversity in the private interests is less likely to propose and approve bills against their aggregate financial positions, suggesting the importance of ongoing monitoring of politicians’ private interests.

Our paper contributes to the literature on the determinants of politicians’ legislative decisions. The literature has debated about the relative importance of constituents’ interest (Peltzman (1984)) and a politician’s ideological preferences (Lee et al. (2004); Bischof et al.

(2020)) in the legislative decisions. Mian et al. (2010) argue that constituents, special interests, and ideology all matter but through different channels. Only a few studies have focused on the effect of the private interest of politicians in their legislative decisions. Benmelech and Moskowitz (2010) examine the private interest of entrenched politicians to set the usury limits. Cohen et al. (2013) report that legislators tend to vote for a bill favorably affecting the companies in which they have private interests. Tahoun and van Lent (2019) find that the politicians owning stocks of financial firms voted for the government support of financial institutions during the Global Financial Crisis. Our paper complements the prior studies by investigating the effect of politicians' asset allocation on their economic bill proposals, the first stage of the legislative process.³ We focus on the politicians' bill proposals rather than their voting on proposed bills because it is well known that not only the private interest but also the decision of party leadership matters when it comes to voting on the legislative proposals (Tahoun and van Lent (2019)).⁴ Therefore, studying the bill proposals would better serve our research purpose of understanding the effect of politicians' private interests on their legislative behaviors related to the real estate market. Moreover, we also show the aggregate effect of the politician's private interest on a country's economic policy, suggesting the danger of politicians' homogeneous exposure to the private interests.

Our paper also contributes to the literature analyzing politicians' private investment using insider information. Many studies in the literature focus on the stock tradings of the members of the U.S. Senate and the House of Representatives and report that politicians outperform in their stock investments by possibly exploiting their insider information (Ziobrowski et al. (2004); Ziobrowski et al. (2011); Karadas (2018); Karadas et al. (2021)). Similarly, Baldauf et al. (2021) find that the U.S. Congress members can time the real estate market to make an abnormal return from real estate investment. Sometimes, firms actively seek building political

³Several recent studies investigate how private real estate assets affect professional decision-makings for CEOs (Cronqvist et al. (2012)), corporate directors (Bahaj et al. (2020)), mutual fund managers (Pool et al. (2019)) and financial advisors (Dimmock et al. (2021)).

⁴Tahoun and van Lent (2019) argue that the vote in a later stage is likely to be affected by the party leadership. Therefore, they claim that the earlier legislative stage, such as the first vote, better reveals Congress members' personal preferences and interests.

networks (Faccio (2006); Goldman et al. (2009); Akey (2015)), through which corporate insider information can potentially flow to politicians. In this study, we raise an unexplored question of whether politicians may actively create insider information themselves by proposing and approving the bills that potentially affect their private interests.

The rest of the paper is organized as follows. Section 2 describes the data and reports summary statistics, Section 3 presents the main empirical results, Section 4 reports the instrumental variable analysis, Section 5 reports potential aggregate effects on legislation, and Section 6 concludes.

2. Data and Summary Statistics

2.1. Financial Disclosure of Congress members in South Korea

South Korea has a unicameral legislative system and there are 300 members in the National Assembly (hereafter “Congress”). General elections are held every four years and term limits are not imposed. Because South Korea has a single constituency system, each electoral district has only one Congress member representing the district.

To measure the financial positions of each Congress member, we use granular data of public officials’ assets and liabilities from the Public Ethics and Transparency Initiative System (PETI System) in South Korea. Since 1993, “The Public Service Ethics Act” requires all public officials in South Korea, who are grades 4 or higher, to disclose their own and immediate family member’s assets and liabilities in detail to the government system annually. To ensure the truthfulness of the disclosure, public officials are subject to disciplinary actions, such as fines or dismissal from public services, with any false disclosure. Notably, the information on senior-level public officials is publicly disclosed. The senior-level public officials include all high-ranked government officials, all Congress members, and all judiciary members.

There are several advantages of using this dataset. First, the data is free from any selection biases because all eligible public officials, including elected officials, must disclose their assets

and liabilities every year. Moreover, the disclosure includes all assets and liabilities owned by public officials, spouses, and lineal ascendants and descendants.⁵ We aggregate all assets and liabilities of family members to construct the assets and liabilities of a public official.

Second, our data provide a complete description of the assets and liabilities of public officials. The data includes cash and deposits; all types of securities, such as public equities, private equities, government bonds, municipal bonds, corporate bonds; all types of real estate assets, such as the ownership of land and buildings, superficies, and the lumpsum deposit on rented residential properties (known as “Jeon-Se” in Korean); and other assets, such as vehicles, precious metals (gold, silver, platinum), jewelry, antiques, work of arts, intellectual property, and golf club memberships. The data also includes all types of debts associated with the public officials.

Third, the data includes detailed information on the characteristics of assets and liabilities. Our data has the exact market value (or fairly assessed value) of all types of assets and liabilities, which differs from other countries’ datasets. For example, the members in the U.S. Congress are required to file their financial disclosure by indicating the range of their asset value but not reporting exact values. Our data also provide the property type, location, and market value (or appraisal value) for real estate assets and liabilities.

In this paper, we focus on Congress members to examine the role of personal portfolio composition on their professional law-making decisions. We particularly focus on the Congress members affiliated with the committees that make laws related to the real estate market. There are seven committees that are highly relevant to making laws affecting the government’s real estate policies. These committees include “Land Infrastructure and Transport Committee,” “Public Administration and Security Committee,” “Trade, Industry, Energy, SMEs, and Startups Committee,” “Strategy and Finance Committee,” “Legislation and Judiciary Committee,” “National Policy Committee,” and “Agriculture, Food, Rural Affairs, Oceans, and

⁵There is an exemption rule that the lineal ascendants and descendants who are independent of the public officials can refuse to disclose. But the exemption should be approved by the Public Service Ethics Committee to prevent any intentional reduction or concealment of their assets. The exemption should be renewed every 3 years.

Fisheries Committee.” The committee membership can change during a legislative session and there are 181 Congress members on average in these committees.⁶

2.2. Tightening Real Estate Bills Proposed by Congress Members

In this paper, the first reason for focusing on the politicians’ bill proposals (rather than their voting behaviors) is that a bill proposal is the first legislative step reflecting a politician’s personal or ideological interests. Any Congress member can freely initiate bill proposals when more than ten Congress members support it. While some major bill proposals with a specific political agenda can be proposed by the senior representative members of a political party, most other bill proposals are independently initiated by Congress members.

And the data offers another reason; avoiding a possible selection bias. Not all proposed bills are approved eventually. For example, when multiple congress members propose similar bills, the committee’s chair sometimes proposes a new combined bill, which will be referred to the full Congress for a vote. Sometimes, only the proposed bills that are likely to pass the vote at the full Congress would be referred. Figure 1 reports the likelihood of a bill being approved by the full Congress when the bill is referred. There is a clear pattern that the votes for approving the bills are well above 90%, indicating that referred bills are very likely to be approved. For such a reason, in the following sections, we focus on the bill proposals instead of the voting on those proposed bills.

We obtain a complete list of bill proposals from the database provided by the Congress of South Korea. There are a total of 46,569 bills from 2011 to 2020 proposed by all Congress members. The database offers detailed information on all proposed bills, such as the title of the proposed bill, a summary of the bill, the proposal date of the bill, assigned committee, related ministry in government, and the detailed outcomes in the legislative progress of the

⁶As the Committee assignments are based on qualifications such as seniority and areas of expertise, the Congress members in the seven Committees are likely to be an expert on real estate and likely to interact more with the real estate market. Therefore, we focus on the cross-sectional comparison within this group as they would face a stronger conflict of interest between their private interest in real estate and their professional duty to propose laws regarding the real estate market.

proposed bill. We also identify the name of the Congress member who proposed the bill, including information on whether the Congress member is a primary sponsor or not.

We are particularly interested in bills related to real estate as a possible vehicle for the Congress members to convey their personal incentives in drafting their proposals. To identify real estate bills, we use two-step screening. We first narrow down all proposed bills to those associated with government ministries implementing real estate policy. The ministries are “Ministry of Land, Infrastructure, and Transport,” “Ministry of the Interior and Safety,” “Ministry of Economy and Finance,” “Ministry of Justice,” “Ministry of Trade, Industry and Energy,” and “Financial Services Commission.” Among the total 46,569 proposed bills, 19,869 bills are associated with those six ministries of government.

We then apply a textual analysis technique of keyword searching for the real estate bills. We first extract the general keywords for real estate bills from the titles of pre-existing real estate laws classified by the Korean Law Information Center.⁷ The general keywords include “Real Estate,” “Housing,” “Land,” “Development Gain,” “Real Estate Agent,” “Appraisal Value,” “Renter,” “Residence,” “Reconstruction,” and “New Home Sales.”

Since the bills associated with a particular ministry are likely to have some specific keywords by the ministry’s objective, we also extract the ministry-specific keywords that enhance the identification of real estate-related bills of the particular ministry. For example, the real estate bills associated with the “Ministry of Economy and Finance” are mostly tax-related, so we use “Tax” for the ministry-specific keyword. On the other hand, there are bills associated with “Ministry of Land, infrastructure, and Transport” that contain our general keywords but are related to infrastructure, which is not our primary interest. To remove those bills, we drop bills with a list of ministry-specific keywords such as “Harbor,” “Airport,” or “Ground Water.” Applying the general and ministry-specific keywords to the title and summary of proposed bills, we find 2,560 proposed bills that are associated with real estate.

To measure the policy direction of the real estate bills, we use the keyword search to dif-

⁷We observe that most of the proposed bills are revising pre-existing laws (94%) rather than legislating entirely new laws.

ferentiate tightening real estate bills from other real estate bills. The tightening keywords are “Enforcement,” “Enhancement,” “Elimination,” “Obligation,” “Prohibition,” “Permission,” “Restriction,” “Sanction,” “Speculation,” “Unfair,” “Violation.” The loosening keywords are “Reduction,” “Tax Exemption,” “Tax Credit,” “Tax Benefit,” “Allowance,” “Unnecessary,” “Incentive,” and “Abolition.” We then count the number of tightening keywords and loosening keywords from the summary of a real estate bill. We define a real estate bill as tightening if the number of tightening keywords is more than that of loosening keywords. Among 2,560 real estate bills, we find 849 tightening real estate bills. Figure 2 reports the number of real estate bills and the number of tightening real estate bills by year. On average, 33% of the real estate bills are tightening bills.

To ensure the validity of our measure of tightening bills, we cross-check our measure with an alternative measure using the official list of tightening bills categorized by the South Korean government from 2015 to 2020.⁸ While the sample period of this alternative measure is limited from 2015 to 2020, we find that our measure is positively correlated with the alternative measure with statistical significance at a 1% level, indicating the consistency of our measure. Due to the shortened sample period, we use this measure for checking the robustness of our main results in a later section.

2.3. Description of Variables and Summary Statistics

Using the financial disclosure data of Congress members, we construct variables describing their portfolio compositions and report the summary statistics in Panel B of Table 1.

The Congress members have total assets of 2.4 billion Korean Won (KRW) on average, which is about 2 million USD as of 2021.⁹ We aggregate the detailed components of assets into five categories: Real Estate, Cash and Deposits, Residential Deposits, Stocks, and Other Assets. Real estate takes the largest part of total assets (*Ratio of Real Estate*) with about

⁸See the South Korean government website for explaining government regulations at <https://www.better.go.kr/rz.law.AssemblyLawListSIPL.laf>

⁹All monetary amounts are inflation-adjusted with GDP deflator.

48% on average, with a standard deviation of 24%. Cash and deposits (*Ratio of Cash*) also take a considerable portion in the Congress members' portfolios with an average ratio of 29% and a standard deviation of 17%. Residential deposits (*Ratio of Residential Deposits*), which take about 12% of the total assets, are the lumpsum cash deposit on rented residential properties.¹⁰ Stock holdings (*Ratio of Stocks*) take about 1% of the total assets and Other Assets (*Ratio of Other Assets*) with 7%. On average, Congress members also maintain 22% leverage (*Leverage*), defined as total debts to total assets.

The data suggests that real estate asset is a significant part of the Congress members' portfolios while the fraction of stock ownership is limited. Therefore, it is more plausible that the real estate would carry more weight in determining Congress members' personal interests among all asset classes and we will use the ratio of real estate assets in Congress members' portfolios (*Ratio of Real Estate*) as our main independent variable. Figure 3 reports the average portfolio ratio for each asset class and leverage ratio by year. The figure clearly shows that real estate asset takes the largest portion of Congress members' financial assets over the sample period, suggesting that real estate has been serving as a crucial asset class affecting Congress members' financial interests over the years.

To measure the propensity of Congress members being reluctant to propose tightening real estate bills, we construct a dummy variable, *Reluctance of Tightening Real Estate*, that equals to 1 if a Congress member i does not propose any tightening real estate bill in year t , and 0 if otherwise. A dummy variable capturing full abstinence from any tightening real estate bills will better serve to answer our research question because it closely reflects a Congress member's reluctance to proposing tightening real estate bills and it is less prone to potential biases from misclassifying the type of individual bills. Panel B reports the summary statistics of the variable. On average, 25% of Congress members do not propose any tightening real estate bill during the sample period.

¹⁰In Korea, renters usually post a large deposit and do not pay monthly rents until the expiration of the lease. At the end of the lease contract, the deposit is returned to the renter. This unique contract, called "Jeon-se," is the dominant form of rental contracts in Korea. For non-homeowners, the residential deposits usually take a large portion of their total asset.

We also collect the detailed characteristics of Congress members¹¹ and report the summary statistics in Panel C. *Conservative Party* is an indicator variable that equals to 1 if the Congress member i is associated with the conservative party in year t , which is less likely to regulate the housing market, and 0 otherwise. *Terms Served* is a Congress member i 's number of serving terms as a Congress member as of year t . On average, Congress members have two terms of experience. *Primary Sponsor* is an indicator variable that equals to 1 if the Congress member i proposes at least one real estate bill as the primary sponsor in year t , and 0 otherwise.

Age is the age of the Congress member i in year t . The average age of Congress members is 58, with six years of standard deviation. *Female* is a dummy variable that equals to 1 if the Congress member i is female, and 0 otherwise. 12% of the Congress members are female. To measure the educational level of Congress members, we divide Congress members into three groups by their final degrees and define three indicator variables: *Education (high school or below)*, *Education (college)*, and *Education (postgraduate)*. 62% of Congress members have postgraduate degrees, 36% have a college degree, and only 2% have final education lower than high school.

In Panel D, we report the variables that capture various aspects of real estate holdings. For example, *Number of Real Estate* is the number of real estate properties owned by a Congress member. The average (median) number of real estate properties is 7(4), but this number is highly skewed with a 90th-percentile of 15. On average, 34% of total assets are residential real estate properties (*Ratio of Residential Real Estate*), and 13% of total assets are non-residential real estate properties (*Ratio of Non-residential Real Estate*). By the type of ownership, 27% of total assets are directly owned by Congress members (*Ratio of Owned by Congressmen*), and 20% of total assets are owned by their family members (*Ratio of Owned by Family*). While 36% of total assets are the real estate assets in the metropolitan area (*Ratio of Metropolitan Area*), 12% of total assets are the real estate assets in the non-metropolitan area (*Ratio of*

¹¹The data is available at Open Assembly Information Website (<https://open.assembly.go.kr/>).

Non-metropolitan Area). Interestingly, Congress members hold 13% of total assets in the real estate in their electoral districts (*Ratio of Own Electoral District*) but have 34% of total assets in the real estate outside of their electoral districts (*Ratio of Other Electoral Districts*).

3. Empirical Results

3.1. The Effect of Congress Members' Real Estate Holdings on Proposing Tightening Real Estate Bills

To test whether Congress members' financial positions affect their legislative decisions, we estimate a linear probability model for the effect of their real estate assets on their economic bill proposal behavior. Our regression specification is as follows:

$$\begin{aligned} & \textit{Reluctance of Tightening Real Estate}_{i,t} \\ &= \alpha + \beta \cdot \textit{Ratio of Real Estate}_{i,t-1} + \gamma \cdot X_{i,t-1} + \delta \cdot M_{t-1} + \eta_i + \epsilon_{i,t}, \end{aligned} \quad (1)$$

where the dependent variable, *Reluctance of Tightening Real Estate*_{*i,t*}, is an indicator variable that equals to 1 if the Congress member *i* does not propose any bill that tightens the real estate market in year *t*, and the main independent variable, *Ratio of Real Estate*_{*i,t-1*}, is the ratio of real estate assets to total assets owned by a Congress member *i* in year *t* − 1.

*X*_{*i,t-1*} includes various control variables related to a Congress member's individual characteristics such as 1) other components of asset portfolio such as *Log (Total Assets)*, *Leverage*, *Ratio of Cash*, and *Ratio of Residential Deposits*, 2) demographic variables such as *Age* and *Female*, and 3) party affiliation, *Conservative Party*. η_i include additional individual characteristics such as education level, *Terms Served*, *Primary Sponsor* and *Electoral District*, all of which are included as fixed effects.

Figure 4 reports the univariate relationship between the dependent variable, *Reluctance of Tightening Real Estate*, and our primary independent variable, *Ratio of Real Estate*. We

clearly observe that the propensity to be reluctant to propose tightening real estate bills increases as *Ratio of Real Estate* increases. For example, while the Congress members with an average fraction of real estate wealth (48%) are likely not to propose tightening bills with a probability of 26%, the Congress members at one standard deviation (SD) above the mean (72%) do not propose tightening bills with a probability of 27.3% and the Congress members at one SD below the mean (24%) do not propose with a probability of 23.9%.

Panel A of Table 2 reports the estimated coefficients from regressions of *Reluctance of Tightening Real Estate* on *Ratio of Real Estate*. In Column (1), controlling for the Congress members' other portfolio positions, we find that the estimated coefficient on *Ratio of Real Estate* is 0.162, with *t*-statistic of 2.28, suggesting that an increase in the real estate holdings increases the propensity of reluctance of proposing bills that tighten real estate market, and this relation is statistically significant. In terms of economic significance, a one SD increase in the *Ratio of Real Estate* in the previous year increases 16% of the average value of *Reluctance of Tightening Real Estate* ($=1 \text{ SD of the } Ratio \text{ of Real Estate } (0.24) \times 0.162 / \text{mean of the } Reluctance \text{ of Tightening Real Estate } (0.25)$).

Column (2) in Panel A of Table 2 reports the regression coefficients with additional controls of demographic and legislative variables. As Mian et al. (2010) find the effects of various factors such as constituents' interest and ideology on Congress members' legislative behaviors, it is crucial to control for the other possible channels to identify the role of personal interests. Therefore, we include *Electoral District* fixed effects to control for the particular policy demand from the voters in the electoral district of the Congress member.¹² In addition, to control for Congress members' ideology, we include *Conservative Party* fixed effects. This is because the ideology of Congress members in South Korea is starkly divided by their party affiliations. Figure 5 reports the distribution of Congress members' ideology scores¹³ in the 20th Congress

¹²There are two types of Congress members in the Congress of South Korea: electorate members and list members. Electorate members are those the voters elect by an electoral district, and more than 80% of the Congress members belong to this type. List members for a party are elected by the proportion of votes the political party gets. Since there is no electoral district assigned to this group, we assign a same indicator variable for all Congress members in this group.

¹³The score is W-NOMINATE score (Poole and Rosenthal (1985), Poole (2005), Poole and Rosenthal (2007))

by their affiliated party, showing that the *Conservative Party* fixed effects are sufficient to control for the ideology of politicians in South Korea. Again, we find a similar result as in Column (1), indicating that the fraction of real estate assets in Congress members' portfolios decreases the likelihood of proposing tightening real estate bills.

Column (3) in Panel A of Table 2 reports the estimated regression coefficients with controls of macroeconomic factors: *GDP Growth* (lagged real GDP Growth) and *HPI Growth* (lagged House Price Index Growth minus Consumer Price Index Growth). We find a similar result that Congress members with more fraction of real estate assets in their portfolios are less likely to propose tightening real estate bills. Furthermore, Column (4) includes year fixed effects to control for any time-specific macro factors affecting the relation. We find a robust effect of the fraction of real estate assets in Congress members' portfolios on their likelihood of proposing tightening bills.

The real estate ownership in Congress members' electoral district is closely related to the constituent's economic interest and it may drive our results. To rule out this possibility, in Panel B of Table 2, we re-estimate the baseline regression model (Equation 1) but replace *Ratio of Real Estate* with *Ratio of Own Electoral District* and *Ratio of Other Electoral Districts*. We find that the likelihood of reluctance of tightening real estate bills increases with both *Ratio of Own Electoral District* and *Ratio of Other Electoral Districts*, indicating that the effect is not entirely driven by economic interests of local constituents. Instead, the effect of real estate holdings outside of members' electoral districts shows a more significant effect on their legislative behavior.

calculated by The Joongang, Pollab, and PolMetriX for the 20th Congress using Congress members' legislative voting behavior from June 2016 to end of 2019 (<https://www.joongang.co.kr/article/23724222>). If a legislator's ideology score is closer to -1, the legislator's vote is more liberal, while the vote is more conservative if the legislator's score is closer to 1.

3.2. Heterogeneity Analysis by the Scope of Real Estate Assets and Political Entrenchment

In the previous section, we find that the Congress members with a higher fraction of real estate assets in their portfolios are less likely to propose bills that tighten real estate markets. If the private interest is one of the main factors driving Congress members' legislative behavior, we expect to find a more substantial effect when the private interest is at a larger stake. To this end, we evaluate the effect of the real estate asset on their legislative actions in two scenarios with varying degrees of personal interests.

First, holding the amounts of total wealth and the fraction of real estate in the portfolio, we expect private incentives to increase with the number of real estate in their portfolios because the scope of their legislative decisions' potential impacts is greater. To this end, we define *Large Number*, an indicator variable that equals to 1 if the number of real estate assets is above the top quintile of the distribution in each year t and 0 otherwise. The *Large Number* variable has a mean of 0.22 with a standard deviation of 0.42, indicating that about 22% of Congress members own more than 7 real estate assets. We augment the baseline regression with the interaction term between *Ratio of Real Estate* and the *Large Number* dummy variable. In Table 3, we report the estimated coefficient on the interaction term between *Ratio of Real Estate* and *Large Number*. The result shows that the effect of *Ratio of Real Estate* gets stronger with a large number of real estate assets. In terms of economic significance, a one-standard-deviation increase in the *Ratio of Real Estate* in the previous year increases 32% of the average value of *Reluctance of Tightening Real Estate* for the group with a large number of real estate assets ($=1 \text{ SD of the } Ratio \text{ of Real Estate } (0.24) \times 0.332 / \text{mean of the } Reluctance \text{ of Tightening Real Estate } (0.25)$).

Second, entrenched Congress members are more likely to pursue their private incentives with less concerns about their reelections. To classify entrenched Congress members, we use the latest election results. We first define the *High Votes*, an indicator variable that equals to 1 if a Congress member i wins 50% of votes from the latest election and 0 otherwise. Column

(1) of Table 4 reports the result interacting *Ratio of Real Estate* with *High Votes*. We employ the same regression specification as in the Column (4) of Table 2, which has year fixed effects. We find that, when a Congress member won an election by significant votes, the positive effect of *Ratio of Real Estate* on the likelihood to be reluctant to propose tightening real estate bills becomes stronger.

We also classify entrenched Congress members using the concentration of votes in his/her latest election proxied by the Herfindahl index of all candidates' votes in an electoral district. High Herfindahl index of all candidates' votes implies the winner (i.e., the current Congress member) received disproportionately more votes compared to their losing rivals. We define the *Vote Concentration* indicator variable that equals to 1 if the Herfindahl index is above the median in the distribution of the Herfindahl index of the election and 0 otherwise. We add the interaction term between the *Ratio of Real Estate* and *Vote Concentration* to the same stringent regression specification as in the Column (4) of Table 2, which has year fixed effects. Column (2) of Table 4 reports the estimated coefficient on the interaction term between *Ratio of Real Estate* and *Vote Concentration*. We again find that the more the Congress members won with a high Herfindahl index, the more the positive effect of *Ratio of Real Estate* on the reluctance of the proposal of tightening real estate bills.

Sometimes, the Congress members' affiliated party is more influential to the voters than the Congress member himself. In this case, we expect the Congress members to be more entrenched if they are representing the electoral districts in the province where the party takes a majority of seats. Therefore, we construct the *Party Shares*, a binary variable equals to 1 if the fraction of the party's seats in the total number of seats in the province is above the top tercile of the distribution and 0 otherwise. Column (3) of Table 4 reports the regression result interacting *Ratio of Real Estate* with *Party Shares*. We again find that the more the party of Congress members won in the province, the stronger the positive effect of *Ratio of Real Estate* on the reluctance of proposing tightening real estate bills becomes.

3.3. Robustness Checks

We find that *Ratio of Real Estate* negatively affects the propensity to propose tightening real estate bills. For checking robustness of the results, we disaggregate our main independent variable (*Ratio of Real Estate*) into several components and examine whether the results are driven by a particular component of *Ratio of Real Estate*.

First, we disaggregate *Ratio of Real Estate* by the owner of the real estate asset. As explained in the data section, we aggregate all assets owned by the Congress member himself and his family members to define the portfolio of a Congress member. In Panel D of Table 1, we report the summary statistics of *Ratio of Real Estate* by ownership. Among the 48% of real estate assets in Congress members' portfolios, 27% of total assets are the real estate assets directly owned by Congress members (*Ratio of Owned by Congressmen*), but 20% of total assets are the real estate assets owned by their family members (*Ratio of Owned by Family*).

In Panel A of Table 5, we re-estimate the baseline regression model (Equation 1) but replace *Ratio of Real Estate* with *Ratio of Owned by Congressmen* and *Ratio of Owned by Family*. For brevity, we do not report the estimated coefficients on other control variables. Again, we find that the reluctance of tightening real estate bills increases with both *Ratio of Owned by Congressmen* and *Ratio of Owned by Family*. Moreover, the magnitude of effects is similar, suggesting that the family members' interests are as important as Congress members' direct interests.

Second, we disaggregate *Ratio of Real Estate* by the property type of the real estate assets. For example, in Panel D of Table 1, we report that 34% of total assets are residential real estate (*Ratio of Residential Real Estate*) and 13% are non-residential real estate (*Ratio of Non-residential Real Estate*). In Panel B of Table 5, we re-estimate the baseline regression model (Equation 1) but replace *Ratio of Real Estate* with *Ratio of Residential Real Estate* and *Ratio of Non-residential Real Estate*. We find that the reluctance of tightening real estate bills increases with both *Ratio of Residential Real Estate* and *Ratio of Non-residential Real*

Estate.

Third, we disaggregate *Ratio of Real Estate* by the source of the portfolio changes. For example, *Ratio of Real Estate* can change because of the changes in real estate value in the portfolio or the new purchase or sales of the real estate assets. We define a dummy variable, *Composition Changes*, that equals to 1 if a Congress member in the year has any new acquisition, sales, and real estate inheritance and 0 otherwise. While 27% of member-year observations in our data have such an event, 73% of member-year observations do not have any composition changes of real estate assets in their portfolios. In Panel C of Table 5, we rerun the regression specification in Table 2, Column (4) but include the interaction term between *Composition Changes* dummy and *Ratio of Real Estate*. Again, we find that the effect of *Ratio of Real Estate* is not different across the samples with and without compositional changes, indicating that our results are not driven by any particular source of changes in *Ratio of Real Estate*.

We also test the robustness of our results regarding the specification of the dependent variable. First, we use alternative categorization of bills to construct our dependent variable, *Reluctance of Tightening Real Estate*. For the subset of years from 2015 to 2020, the Korean government officially categorizes all proposed bills on their policy direction (i.e., tightening vs. loosening). Therefore, we can alternatively define the dependent variable using the categorization imputed by the Korean government, *Reluctance of Tightening Real Estate (Government sorted)*. We use *Reluctance of Tightening Real Estate (Government sorted)* as a dependent variable. Table 6 reports the estimated coefficients from the baseline regression model (Equation 1) with this dependent variable. We still find that the reluctance of a Congress member to propose a tightening real estate bill increases as the ratio of real estate assets increases.

Second, we construct our dependent variable excluding all bill proposals that are potentially not genuine but a copy of other politicians' bills with a minor tweak. We calculate the similarity of 849 tightening real estate bills using TF-IDF (Term Frequency-Inverse Document Frequency) and drop those bills above a certain level of similarity from our construction of

the dependent variable. As shown in Appendix Table 2, our baseline results in Table 2 remain the same after removing potentially duplicative bill proposals from our dependent variable.

Third, we use an alternative measure of Congress members’ bill proposal behaviors. Instead of using the indicator variable of any tightening bill proposal by Congress members, we calculate the ratio of tightening real estate bills to total real estate bills proposed by Congress members in each year and define a dummy that equals to 1 if the ratio is below the median in the distribution of the ratio in the Congress in the year (*Reluctance of Tightening Real Estate (Alt)*). Again, we find similar results as in Table 2, reported in Appendix Table 3.

In a placebo test (Table 7), we re-estimate the baseline regression model (Equation 1) using the Congress members’ bill proposals related to economic policies other than the real estate market. If our results are solely driven by politicians’ unobservable stance to economic policies, we expect to find similar legislative behaviors on other economic bills. Using all other economic bill proposals that are not related to the real estate market, we construct *Reluctance of Tightening Economic Policy*.¹⁴ In Table 7, we find that *Ratio of Real Estate* does not change Congress members’ behavior on the bills related to other economic policies, suggesting that our results are not solely driven by politicians’ distinct stance to economic policies.

In Appendix Table 5, we further rule out the possibility that our results are driven by time-varying local interests or economic conditions of a Congress member’s constituents. In Panel A, following Mian et al. (2010), we augment the baseline regression model with time-varying variables representing local socioeconomic conditions at the constituent level such as industry composition, education attainment, and poverty level. We still find the baseline result holds. In Panel B, we include metropolitan area \times year fixed effects in Equation (1) to further control for unobservable time-varying local conditions.¹⁵ Although the metropolitan area is broader than a single constituent, this specification would still capture the time-varying local interests

¹⁴We consider all economic bills related to the “Ministry of Economy and Finance” and the “Financial Services Commission” and classify them as non-real estate bills based on the textual analysis as in Section 2.

¹⁵Because South Korea has a single constituency legislative system, our empirical setting does not allow for including electoral district \times year fixed effects in regression specifications.

in the real estate market. The analysis result remains consistent with the baseline result.

4. Instrumental Variable Analysis

In the above section, using the rich panel structure of our data to control for various endogenous factors, we find that the fraction of real estate assets in Congress members' portfolios positively affects the Congress member's reluctance of proposing a bill tightening the real estate market. However, there still remains a concern on biases arising from potential omitted variables that we were not able to control since *Ratio of Real Estate* is not exogenously determined. In this section, we exploit a uniquely available instrumental variable in Korea that exogenously change the *Ratio of Real Estate* of Congress members and link them to their bill proposal behaviors.

We use an unpredicted earthquake in South Korea as an instrumental variable for *Ratio of Real Estate*. South Korea had been known as a safe place from seismic activity compared to neighboring countries like Japan or China. However, on November 15, 2017, a 5.4-magnitude earthquake occurred at Pohang, an industrialized city located in the south-eastern part of South Korea. The earthquake was the second-largest one in the country's modern history. The earthquake caused a significant shock to the people around Pohang since the quake was not expected, particularly at Pohang where a number of nuclear power plants are present.¹⁶ In addition, the earthquake caused significant damage to the densely populated industrialized city, injuring 135 people and damaging 57,000 structures that cost around \$123 million. Afterward, Pohang suffered from reduced real estate prices, the number of tourists decreased, the population outflowed, and local businesses shrank. (Lee (2019)).

We use this event as an exogenous shock to our main independent variable, *Ratio of Real Estate*. We restrict our sample to the Congress members with financial disclosures information

¹⁶Later, it turns out that the earthquake was a man-made disaster from the introduction of a geothermal plant, which harvests energy by injecting high-pressure water deep into Earth. See <https://www.science.org/content/article/second-largest-earthquake-modern-south-korean-history-tied-geothermal-plant>

around the time of the Pohang earthquake during 2017 and 2020, which contain 457 member-year observations. We first measure the distance of all real estate assets owned by the Congress members to the epicenter of the Pohang earthquake based on their longitudes and latitudes.¹⁷ Then, for each Congress member i , we calculate the fraction of the real estate value within 40 miles from the epicenter of the earthquake to the total real estate value of the Congress member at the beginning of 2017. This distance includes all nearby areas to Pohang but does not include neighboring major metropolitan regions such as Daegu and Busan, where real estate prices are likely to be driven by other local shocks.

In Figure 6, we report the 40-mile radius from the earthquake’s epicenter. About 7% of Congress members have some real estate assets within 40 miles from the epicenter before the earthquake. We define *Ratio Within 40 Miles* as the fraction of real estate value in the radius to the total real estate assets owned by Congress members, with the mean of 3% and a standard deviation of 0.14, and use it as an instrumental variable for our main independent variable, *Ratio of Real Estate*.

Columns (1)-(2) of Table 8 report the instrumental variable regression results using the Pohang earthquake. Column (1) reports the first stage regression of *Ratio of Real Estate* on *Ratio Within 40 Miles*. The regression specification is similar to Column (4) of Table 2 with time-fixed effects, but we replace the *Electoral District* fixed effects with the *Metropolitan* fixed effects due to the much narrower area coverage in the data than the main sample. We find that *Ratio Within 40 Miles* decreases *Ratio of Real Estate* by -0.146 with t -statistics of -3.29. The reduction in *Ratio of Real Estate* of Congress members indicates that the real estate price within 40 miles from the earthquake’s epicenter significantly fell after the earthquake or that Congress members sold real estate assets nearby the epicenter after the earthquake. In any case, the instrumental variable satisfies the relevance condition.

For exclusion restriction, we argue that the fraction of Congress members’ real estate within 40 miles of the earthquake’s epicenter would not directly change the likelihood of

¹⁷More specifically, we use the Haversine formula, a way of measuring distances between two locations on the surface of a sphere.

proposing tightening bills except through *Ratio of Real Estate*. The Congress usually propose laws affecting the overall country rather than a specific city or town. In Korea, proposing a bill for a specific region is often harshly criticized as favoritism and Congress members actively avoid such scandals.

Column (2) of Table 8 reports the second stage regression results of instrumenting *Ratio of Real Estate* with *Ratio Within 40 Miles*. We find that the Congress members with a higher instrumented *Ratio of Real Estate* are more likely to be reluctant to propose tightening real estate bills. Overall, the instrumental variable analysis based on the unexpected earthquake supports our main finding that Congress members with a more financial interest in the real estate market are less likely to propose tightening real estate bills.

5. Aggregate Effect on Legislation Decisions

The Congress makes laws influencing people’s daily lives through legislative debate and compromise. When the fraction of real estate assets in Congress members’ portfolios decreases the likelihood of proposing tightening real estate bills, an important question will be whether these individual behaviors have aggregate effects.

Similar to the portfolio diversification benefit, we expect that the politicians’ self-interested behaviors will be diversified away when there exists a reasonable heterogeneity in the *Ratio of Real Estate* among Congress members. If so, we would not see any adverse aggregate impact. However, when real estate is a shared private interest among Congress members, the politicians’ individual effects would not be canceled away but summed to an aggregate impact. To evaluate the aggregate impact, we report below a scatter plot showing the relation between the aggregated ratio of Congress members’ real estate assets and their collective legislative actions related to real estate bills.

To measure the Congress-level exposure to real estate assets, we define *Aggregated Ratio of Real Estate*, the ratio of aggregate real estate asset values in all Congress members’ portfolios

to the aggregate total asset values of all Congress members in the year t . *Aggregated Ratio of Real Estate* has the mean of 0.43 with a standard deviation of 0.11, indicating that 43% of aggregated assets of all Congress members are real estate assets over the sample period. The aggregated number of proposed tightening real estate bills at the annual level has a mean of 84.09 with a standard deviation of 30.2.

Figure 7 reports the scatter plot using *Aggregated Ratio of Real Estate* on the x-axis and the aggregated number of proposed tightening real estate bills on the y-axis. Since our sample period is from 2011 to 2020, we have ten yearly observations on the figure. We find that the aggregated number of proposed tightening real estate bills are negatively associated with *Aggregated Ratio of Real Estate*, and the coefficient of the linear fit is -207.29 with the t-statistics of -3.39. The economic significance is large. A one-standard deviation increase in *Aggregated Ratio of Real Estate* reduces 75% of the standard deviation of the aggregated number of proposed tightening real estate bills.

Overall, the aggregate effect analysis suggests that the personal interests on bill proposals are not canceled out among the Congress members, implying the significant impact of the total personal interests on their legislative decisions. This finding offers an important insight into how aggregate legislative decisions reflect each member's personal interests which subsequently affect the economic policies at the national level. As corporate board diversity improves firm value by lowering volatility and improving firm performance (e.g., Bernile et al. (2018)), the diversity among Congress members in terms of their financial positions may balance the legislative decisions in the Congress. Our empirical result indicates that, without the diversity in the private interests, Congress is less likely to propose and approve bills against their aggregate financial positions, suggesting the importance of ongoing monitoring of politicians' private interests.

6. Conclusion

This paper finds that the fraction of real estate assets in Congress members' portfolios decreases the likelihood of proposing tightening real estate bills. That is, the private interest of Congress members in real estate can affect their behavior of proposing bills on the real estate market. We use an instrumental variable—an unexpected earthquake—to find a causal effect of Congress members' personal interests on their legislative behaviors. Finally, by extending our analysis to the Congress-level, we also find that the Congress-level exposure to real estate assets is negatively associated with the total number of tightening bills proposed and approved by Congress.

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Figure 1: Voting Behavior on Tightening Real Estate Bills

The stacked bar graph reports the aggregate vote counts by year on the passage of tightening real estate bills. The votes are divided into Approval, Objection, and Abstention. Since there are multiple bills voted in a year, we aggregate vote counts for all referred tightening real estate bills in a year. The orange solid line graph plots the proportion of the approval votes to total vote counts.

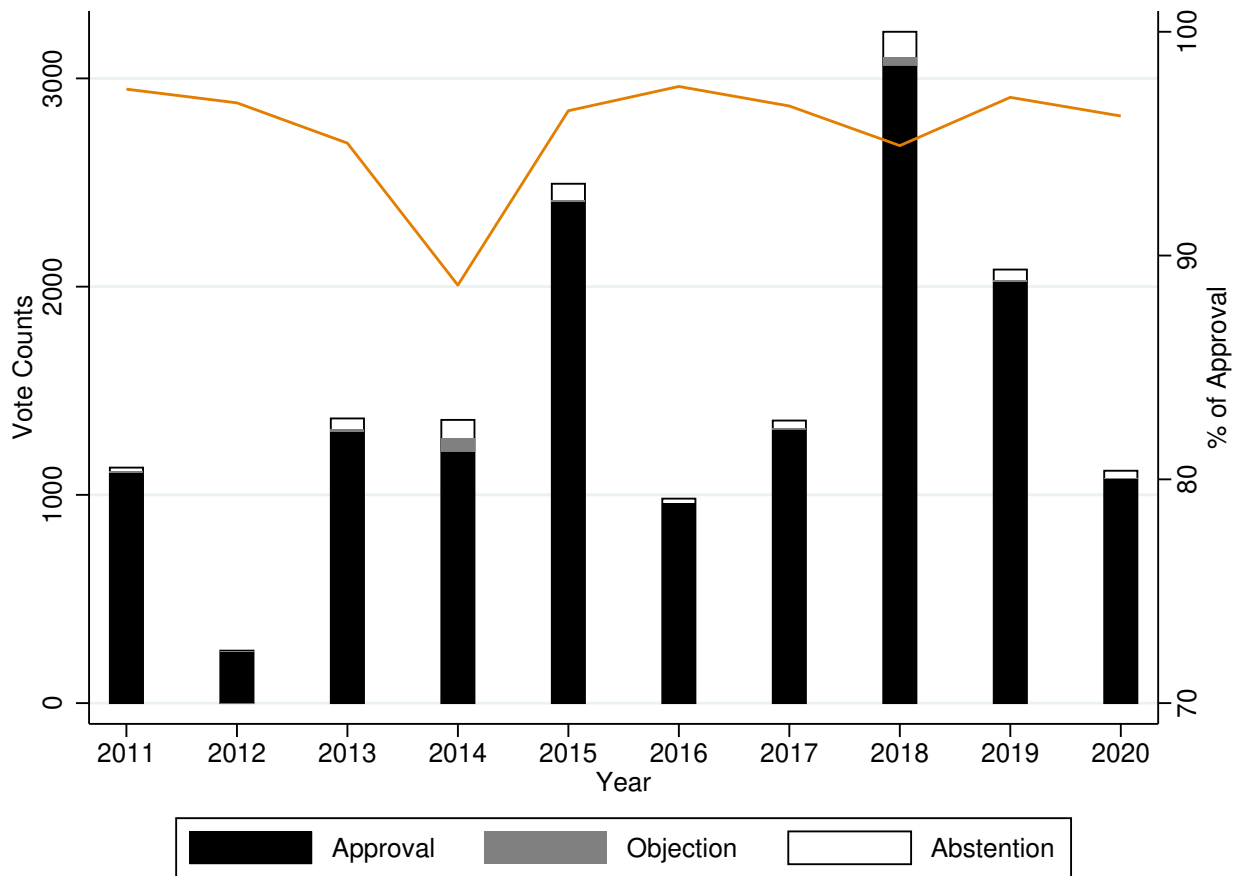


Figure 2: The Number of Proposed Bills by Year

This figure reports the number of proposed bills related to the real estate market and the number of tightening real estate bills among the proposed bills based on the textual analysis from 2011 to 2020. The proportion of tightening real estate bill proposals ranges from 27.4% (in 2016) to 38.7% (in 2012) over the sample period.

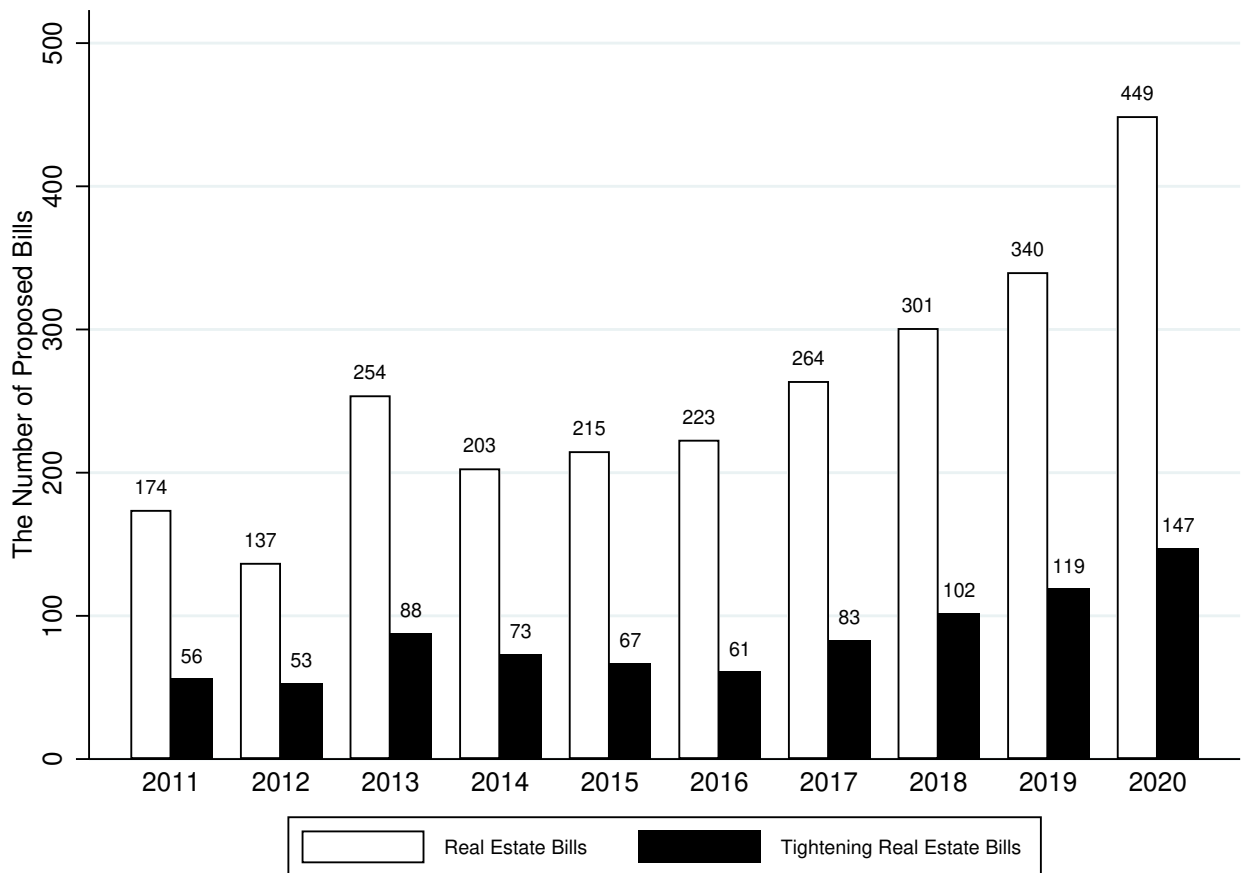


Figure 3: Asset Composition of Congress Members' Portfolio

The figure reports annual asset compositions in the portfolios of the Congress members of the National Assembly of South Korea from 2011 to 2020. Total assets are divided into *Ratio of Real Estate*, *Ratio of Cash*, *Ratio of Residential Deposits*, *Ratio of Stocks*, and *Ratio of Other Assets* and reported in the bar graph. *Leverage* is also computed as the ratio to the total assets and is reported with the negative sign.

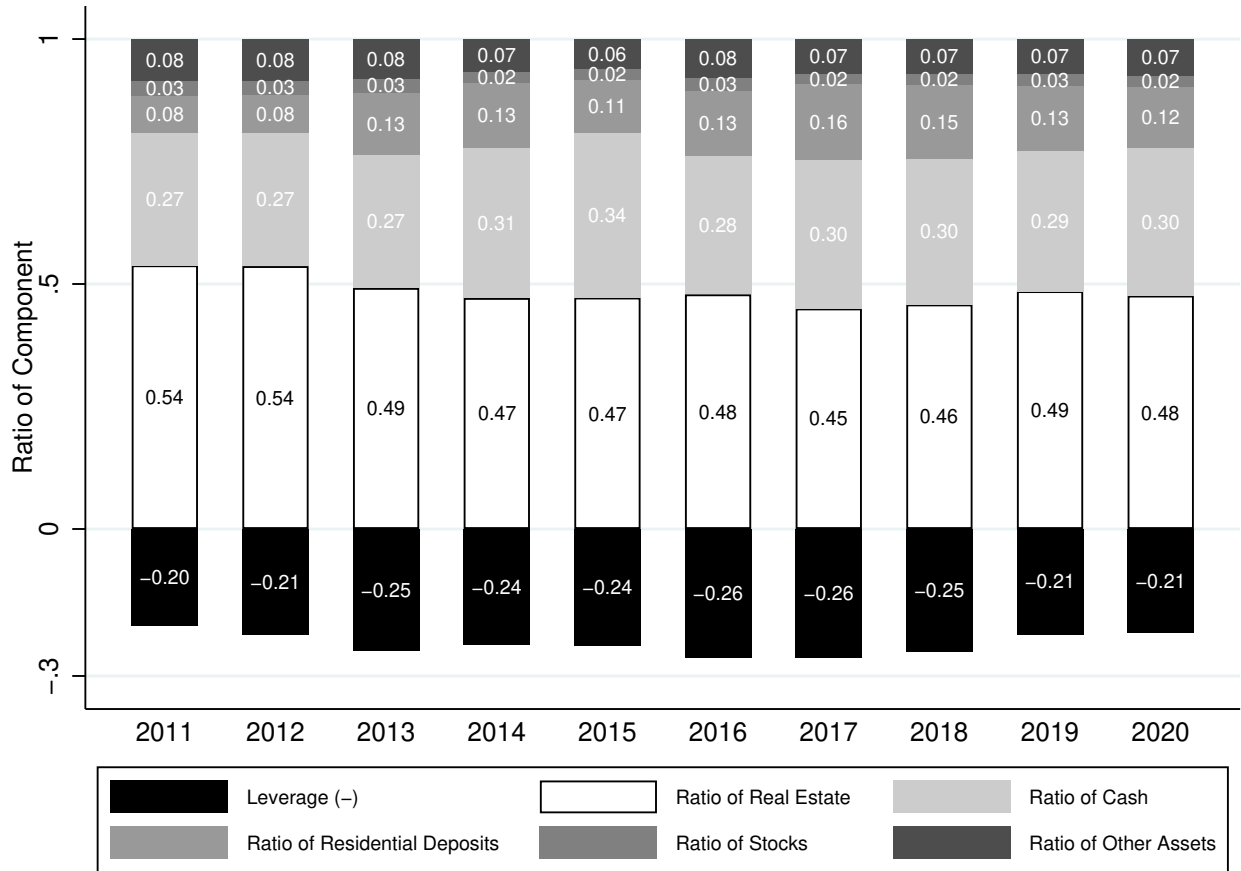


Figure 4: Ratio of Real Estate and Propensity for Reluctance of Tightening Real Estate Bills

The figure reports the propensity for the *Reluctance of Tightening Real Estate* by the *Ratio of Real Estate*, using univariate local polynomial regression. The solid line indicates the propensity for reluctance of tightening real estate bills for a given level of *Ratio of Real Estate*, and the 95% confidential intervals are reported in the dotted line.

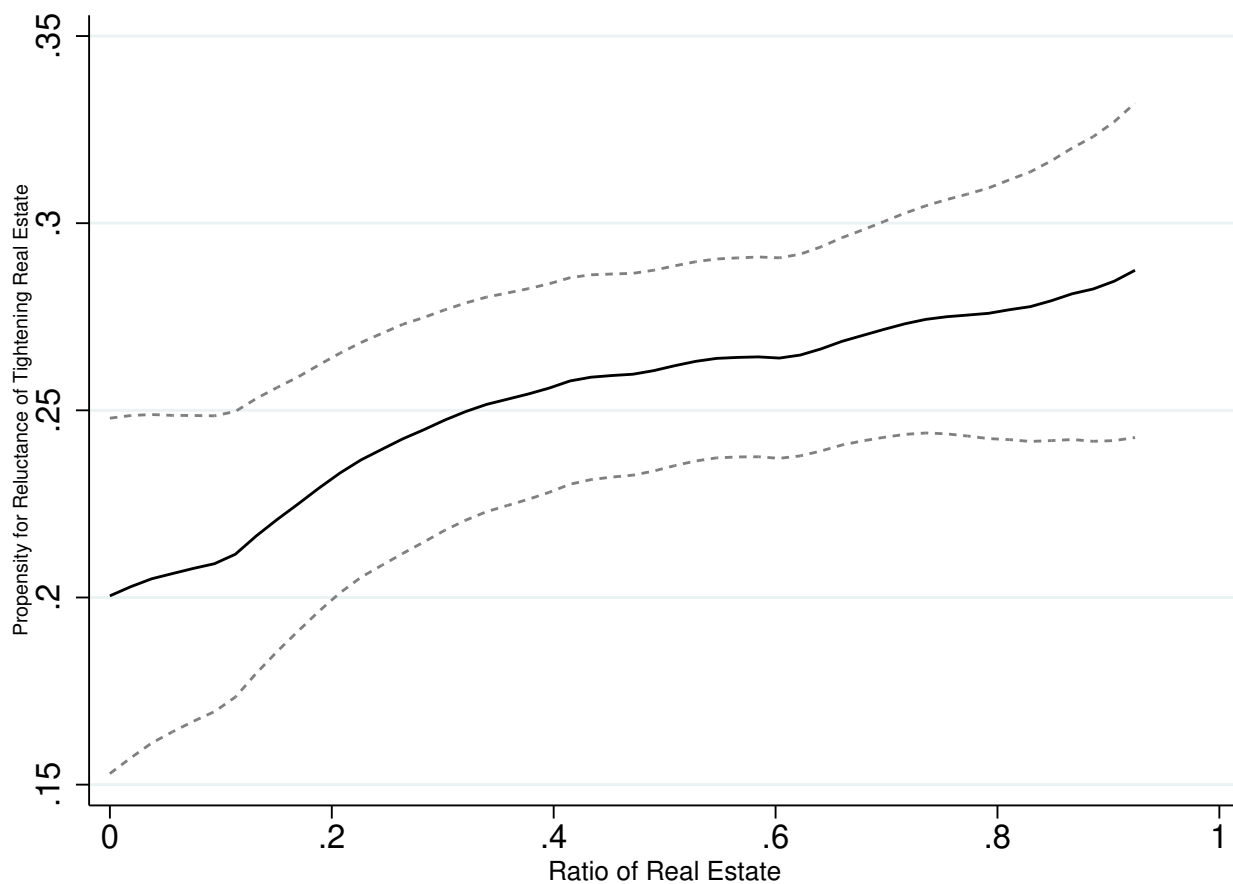


Figure 5: Ideology Score and Affiliated Party

The figure illustrates the distribution of Congress members' ideology score by their affiliated party. Ideology score is based on each Congress member's legislative voting behavior (Poole and Rosenthal (1985), Poole (2005), Poole and Rosenthal (2007)). We present the distribution of the ideology scores for the 20th National Assembly based on Congress members' legislative voting behavior from June 2016 to December 2019. If the score is closer to -1, the legislator's vote is more liberal, while the vote is more conservative if the legislator's score is closer to 1. The colored and white bar graphs represent the number of members in each interval of ideology scores for the conservative and democratic parties, respectively. The line and dash graphs are kernel density estimations of the bar graph.

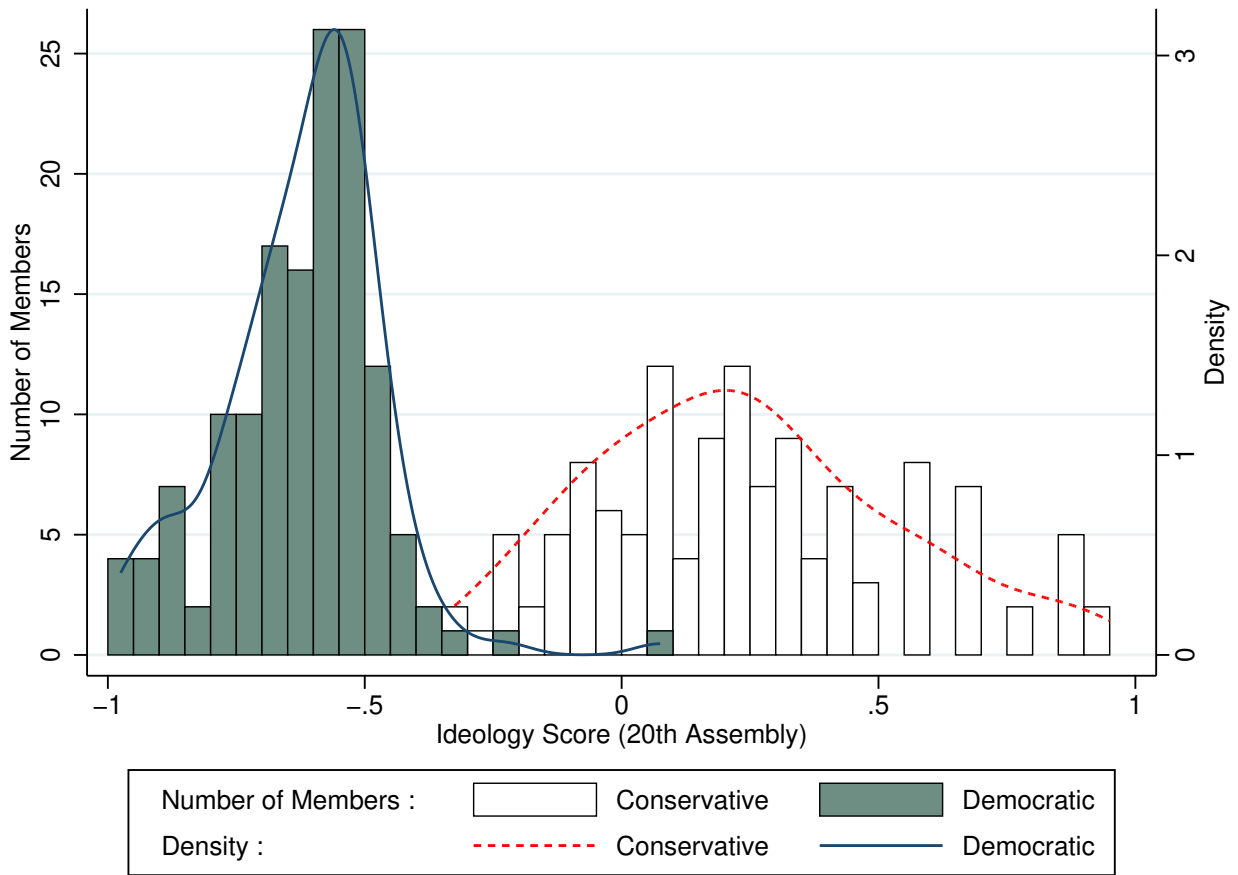


Figure 6: Regions within 40 Miles from the Pohang Earthquake Epicenter

The figure reports the epicenter of the Pohang earthquake in 2017 and the surrounding regions in the south-eastern part of South Korea. The red circle represents the 40-mile radius from the epicenter. Note that the radius does not include neighboring major metropolitan areas such as Daegu and Busan.

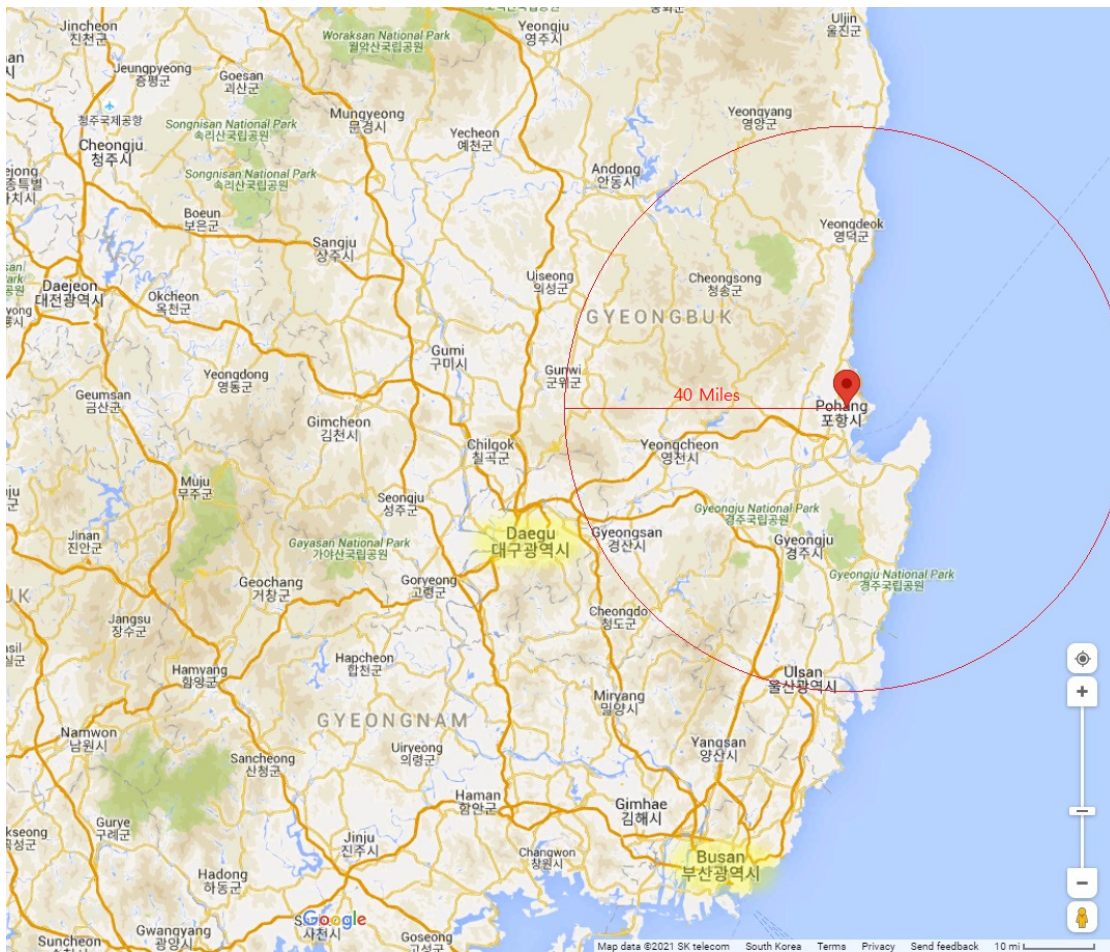


Figure 7: Aggregated Effect of Ratio of Real Estate Assets

The figure reports the scatter plot between the aggregated real estate holdings of the National Assembly of South Korea members (*Aggregated Ratio of Real Estate*) and their annual activities regarding tightening real estate bills (*Aggregated # of Proposed Tightening Real Estate Bills*). The solid line is the linear prediction of *Aggregated # of Proposed Tightening Real Estate Bills* on *Aggregated Ratio of Real Estate*, which is defined as the ratio of the total aggregate amount of all Congress members' real estate assets to the total aggregate amount of all Congress members' assets in year t , and 95% confidential intervals are reported in the dashed line.

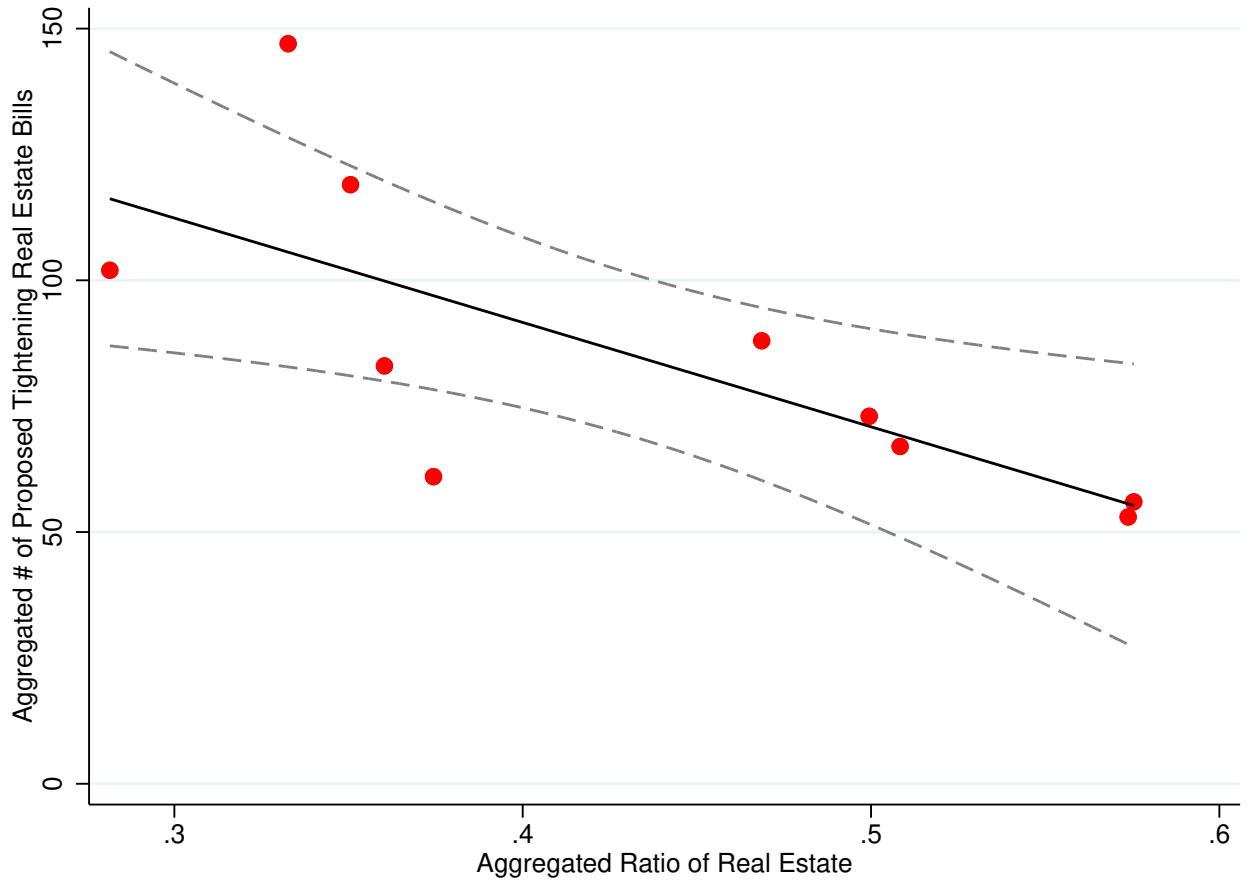


Table 1: Summary Statistics

We report the summary statistics of variables in our analysis. We restrict our sample to Congress members in the National Assembly of South Korea, who are affiliated with the committees that process real estate laws. The sample period is from 2011 to 2020. See Appendix Table 1 for the variable descriptions. All variables are constructed at the member-year level. We winsorize all variables at the 5% and 95% levels.

	Obs	Mean	Std.Dev.	10th pctl.	Median	90th pctl.
Panel A: Tightening Real Estate Bills						
<i>Reluctance of Tightening Real Estate</i>	1,809	0.25	0.44	0	1	1
<i>Reluctance of Tightening Real Estate</i> (Government sorted)	1,076	0.42	0.49	0	0	1
Panel B: Portfolio of Congress Members						
<i>Total Assets (mils)</i>	1,809	2,435	2,163	608	1,638	6,002
<i>Log (Total Assets)</i>	1,809	7.46	0.82	6.41	7.40	8.70
<i>Ratio of Real Estate</i>	1,809	0.48	0.24	0.12	0.49	0.80
<i>Ratio of Cash</i>	1,809	0.29	0.17	0.08	0.27	0.56
<i>Ratio of Residential Deposits</i>	1,809	0.12	0.15	0	0.05	0.36
<i>Ratio of Stocks</i>	1,809	0.01	0.03	0	0	0.06
<i>Ratio of Other Assets</i>	1,809	0.07	0.09	0	0.03	0.21
<i>Leverage</i>	1,809	0.22	0.20	0	0.18	0.53
Panel C: Characteristics of Congress Members						
<i>Conservative Party</i>	1,809	0.49	0.50	0	0	1
<i>Terms Served</i>	1,809	2.01	1.12	1	2	4
<i>Primary Sponsor</i>	1,809	0.39	0.49	0	0	1
<i>Age</i>	1,809	58.3	6.47	49	58	67
<i>Female</i>	1,809	0.12	0.32	0	0	1
<i>Education (high school or below)</i>	1,809	0.02	0.13	0	0	0
<i>Education (college)</i>	1,809	0.36	0.48	0	0	1
<i>Education (postgraduate)</i>	1,809	0.62	0.48	0	1	1
Panel D: Decomposing Ratio of Real estate						
<i>Number of Real Estate</i>	1,809	7.01	10.60	1	4	15
By the type of ownership						
<i>Ratio of Owned by Congressmen</i>	1,809	0.27	0.23	0	0.25	0.64
<i>Ratio of Owned by Family</i>	1,809	0.20	0.19	0	0.15	0.50
By the type of real estate						
<i>Ratio of Residential Real Estate</i>	1,809	0.34	0.22	0.02	0.34	0.64
<i>Ratio of Non-residential Real Estate</i>	1,809	0.13	0.19	0	0.04	0.47
By the type of location						
<i>Ratio of Own Electoral District</i>	1,809	0.13	0.19	0	0.03	0.45
<i>Ratio of Other Electoral Districts</i>	1,809	0.34	0.25	0	0.35	0.71
<i>Composition Changes</i>	1,809	0.27	0.45	0	0	1

Table 2: The Effect of Real Estate Ownership on the Reluctance of Proposing Bills Tightening the Real Estate Market

We report the coefficient estimates from panel regressions of the National Assembly of South Korea members' real estate bill proposal behavior (*Reluctance of Tightening Real Estate*) on their real estate asset ratios in their portfolios (*Ratio of Real Estate*). We use the member-year observations from 2011 to 2020. Panel A reports our baseline regression. The dependent variable is *Reluctance of Tightening Real Estate*, which is a dummy variable that equals to 1 if a Congress member does not propose any tightening real estate bills in a year and 0 otherwise. The main independent variable of interest is *Ratio of Real Estate*, which is the fraction of real estate assets in the Congress members' total assets. In Column (1), we include control variables on other components in the portfolio, such as *Log (Total Assets)*, *Leverage*, *Ratio of Cash*, and *Ratio of Residential Deposits*. Column (2) includes Congress members' characteristics and Electoral District fixed effects. *Other Controls* include *Terms Served*, *Primary Sponsor*, and *Education*. We additionally include macro variables in Column (3), such as *GDP Growth* and *HPI Growth*. We include year fixed effects in Column (4). In Panel B, we disaggregate *Ratio of Real Estate* by the location of the real estate; *Ratio of Own Electoral District*, the fraction of real estate assets located in the Congress members' own electoral district, and *Ratio of Other Electoral Districts*, the fraction of real estate assets located in other electoral districts. See Appendix Table 1 for the variable descriptions. The *t*-statistics reported in parentheses are based on standard errors clustered at the Congress member level. ***, **, * denote 1%, 5%, and 10% statistical significance.

Panel A: Baseline Regression	(1)	(2)	(3)	(4)
	<i>Reluctance of Tightening Real Estate</i>			
<i>Ratio of Real Estate</i>	0.162** (2.28)	0.132* (1.66)	0.150* (1.93)	0.165** (2.13)
<i>Log (Total Assets)</i>	0.048*** (3.26)	0.037* (1.73)	0.041* (1.96)	0.028 (1.35)
<i>Leverage</i>	0.007 (0.10)	0.156** (1.98)	0.142* (1.82)	0.121 (1.59)
<i>Ratio of Cash</i>	0.152* (1.65)	0.236** (2.26)	0.264*** (2.60)	0.270*** (2.69)
<i>Ratio of Residential Deposits</i>	0.072 (0.71)	0.094 (0.79)	0.123 (1.05)	0.137 (1.20)
<i>Conservative Party</i>		0.106*** (3.25)	0.093*** (2.91)	0.113*** (3.31)
<i>Age</i>		0.004 (1.60)	0.004* (1.69)	0.004 (1.54)
<i>Female</i>		0.050 (1.19)	0.051 (1.23)	0.049 (1.15)
<i>GDP Growth</i>			5.844*** (7.38)	
<i>HPI Growth</i>			0.211 (0.39)	
Observations	1,809	1,809	1,809	1,809
Adjusted R-squared	0.010	0.136	0.169	0.244
Other Controls	NO	YES	YES	YES
Electoral District FE	NO	YES	YES	YES
Year FE	NO	NO	NO	YES

Continued on the next page

Table 2 continues

Panel B: By Property Location	(1)	(2)	(3)	(4)
	<i>Reluctance of Tightening Real Estate</i>			
<i>Ratio of Own Electoral District</i>	0.110 (1.33)	0.157 (1.55)	0.172* (1.75)	0.175* (1.78)
<i>Ratio of Other Electoral Districts</i>	0.195*** (2.67)	0.158** (1.97)	0.176** (2.25)	0.188** (2.42)
<i>Log (Total Assets)</i>	0.043*** (2.85)	0.037* (1.68)	0.041* (1.89)	0.028 (1.28)
<i>Leverage</i>	0.009 (0.14)	0.156** (1.99)	0.143* (1.82)	0.121 (1.59)
<i>Ratio of Cash</i>	0.159* (1.76)	0.254** (2.43)	0.280*** (2.76)	0.282*** (2.82)
<i>Ratio of Residential Deposits</i>	0.062 (0.62)	0.113 (0.93)	0.140 (1.19)	0.146 (1.27)
<i>Conservative Party</i>		0.105*** (3.19)	0.092*** (2.84)	0.111*** (3.23)
<i>Age</i>		0.004 (1.56)	0.004 (1.64)	0.003 (1.50)
<i>Female</i>		0.050 (1.19)	0.051 (1.22)	0.048 (1.13)
<i>GDP Growth</i>			5.851*** (7.37)	
<i>HPI Growth</i>			0.203 (0.37)	
Observations	1,809	1,809	1,809	1,809
Adjusted R-squared	0.011	0.136	0.169	0.244
Other Controls	NO	YES	YES	YES
Electoral District FE	NO	YES	YES	YES
Year FE	NO	NO	NO	YES

Table 3 : The Effect of Real Estate Ownership on the Reluctance of Tightening the Real Estate Market: By The Number of Real Estate Assets

We report the coefficient estimates from panel regressions of the National Assembly of the South Korea members' real estate bill proposal behavior (*Reluctance of Tightening Real Estate*) on their real estate asset ratios in their portfolios (*Ratio of Real Estate*) by the number of real estate assets. We use the member-year observations from 2011 to 2020. The dependent variable is *Reluctance of Tightening Real Estate*, which is a dummy variable that equals to 1 if a Congress member does not propose any tightening real estate bills in a year and 0 otherwise. *Large Number* is an indicator variable equals to 1 if the number of real estate assets is in the top quintile of the distribution in each year t and 0 otherwise. Our main variable of interest is the interaction term between *Ratio of Real Estate*, which is the fraction of real estate assets in the Congress members' total assets, and *Large Number*. Regression specifications are the same as in Table 2. See Appendix Table 1 for the variable descriptions. The t -statistics reported in parentheses are based on standard errors clustered at the Congress member level. ***, **, * denote 1%, 5%, and 10% statistical significance.

	(1)	(2)	(3)	(4)
	<i>Reluctance of Tightening Real Estate</i>			
<i>Ratio of Real Estate</i> \times <i>Large Number</i>	0.204 (1.62)	0.296** (2.16)	0.283** (2.09)	0.307** (2.39)
<i>Ratio of Real Estate</i>	0.128* (1.68)	0.073 (0.83)	0.092 (1.06)	0.098 (1.12)
<i>Large Number</i>	-0.184** (-2.55)	-0.218*** (-2.77)	-0.199** (-2.54)	-0.198*** (-2.63)
<i>Log (Total Assets)</i>	0.061*** (4.02)	0.042** (1.97)	0.045** (2.12)	0.030 (1.41)
<i>Leverage</i>	0.010 (0.15)	0.163** (2.07)	0.151* (1.92)	0.131* (1.72)
<i>Ratio of Cash</i>	0.140 (1.57)	0.213** (2.05)	0.244** (2.41)	0.251** (2.51)
<i>Ratio of Residential Deposits</i>	0.034 (0.35)	0.044 (0.36)	0.074 (0.61)	0.081 (0.68)
<i>Conservative Party</i>		0.100*** (3.05)	0.088*** (2.74)	0.109*** (3.18)
<i>Age</i>		0.004* (1.72)	0.004* (1.81)	0.004* (1.67)
<i>Female</i>		0.048 (1.16)	0.050 (1.20)	0.048 (1.14)
<i>GDP Growth</i>			5.788*** (7.28)	
<i>HPI Growth</i>			0.239 (0.44)	
Observations	1,809	1,809	1,809	1,809
Adjusted R-squared	0.014	0.138	0.171	0.246
Other Controls	NO	YES	YES	YES
Electoral District FE	NO	YES	YES	YES
Year FE	NO	NO	NO	YES

Table 4 : The Effect of Real Estate Ownership on the Reluctance of Proposing Bills Tightening the Real Estate Market: Entrenched Congress Members

We report the coefficient estimates from panel regressions of the National Assembly of South Korea members' real estate bill proposal behavior on their real estate asset ratios in their portfolios by the degree of political entrenchment for Congress members. The dependent variable is *Reluctance of Tightening Real Estate*, which is a dummy variable that equals to 1 if a Congress member does not propose any tightening real estate bills in a year and 0 otherwise. We restrict our sample to Electorate members who have their own electoral district, by excluding the List members who are elected based on their own party's total votes. We use *High Votes* in Column (1), *Vote Concentration* in Column (2), and *Party Shares* in Column (3) as the measurement of *Entrenched*. Our main variable of interest is the interaction term between *Ratio of Real Estate*, which is the fraction of real estate assets in the Congress members' total assets, and *Entrenched*. Regression specifications are the same as in Table 2, Column (4). See Appendix Table 1 for the variable descriptions. The *t*-statistics reported in parentheses are based on standard errors clustered at the Congress member level. ***, **, * denote 1%, 5%, and 10% statistical significance.

<i>Measure of Entrenchment:</i>	(1)	(2)	(3)
	<i>High Votes</i>	<i>Vote Concentration</i>	<i>Party Shares</i>
	<i>Reluctance of Tightening Real Estate</i>		
<i>Ratio of Real Estate</i> \times <i>Entrenched</i>	0.274** (2.57)	0.193** (1.98)	0.221** (1.99)
<i>Ratio of Real Estate</i>	0.067 (0.67)	0.126 (1.32)	0.134 (1.48)
<i>Entrenched</i>	-0.129** (-2.32)	-0.120** (-2.29)	-0.106** (-2.00)
<i>Log (Total Assets)</i>	0.029 (1.14)	0.030 (1.19)	0.031 (1.26)
<i>Leverage</i>	0.194** (2.28)	0.185** (2.19)	0.194** (2.33)
<i>Ratio of Cash</i>	0.353*** (3.39)	0.354*** (3.34)	0.350*** (3.33)
<i>Ratio of Residential Deposits</i>	0.210 (1.64)	0.220* (1.70)	0.195 (1.50)
<i>Conservative Party</i>	0.136*** (3.61)	0.129*** (3.47)	0.137*** (3.53)
<i>Age</i>	0.007*** (2.75)	0.007*** (2.75)	0.007*** (2.66)
<i>Female</i>	0.047 (0.88)	0.048 (0.89)	0.055 (1.04)
Observations	1,610	1,610	1,610
Adjusted R-squared	0.238	0.237	0.237
Other Controls	YES	YES	YES
Electoral District FE	YES	YES	YES
Year FE	YES	YES	YES

Table 5: Robustness of the Baseline Regression: Property Ownership and Type

We report the robustness of our results in Table 2 by disaggregating *Ratio of Real Estate* into its components. We disaggregate *Ratio of Real Estate* by the owner of the real estate in Panel A, by the property type of the real estate in Panel B, and by the source of the changes in Panel C. The regression specifications are the same as in Table 2, but we do not report the estimated coefficients on other variables for brevity. See Appendix Table 1 for the variable descriptions. The *t*-statistics reported in parentheses are based on standard errors clustered at the Congress member level. ***, **, * denote 1%, 5%, and 10% statistical significance.

Panel A: By Property Ownership	(1)	(2)	(3)	(4)
	<i>Reluctance of Tightening Real Estate</i>			
<i>Ratio of Owned by Congressmen</i>	0.143*	0.099	0.117	0.142*
	(1.92)	(1.13)	(1.37)	(1.66)
<i>Ratio of Owned by Family</i>	0.132	0.140	0.158*	0.157*
	(1.61)	(1.57)	(1.82)	(1.81)
Panel B: By Property Type	(1)	(2)	(3)	(4)
	<i>Reluctance of Tightening Real Estate</i>			
<i>Ratio of Residential Real Estate</i>	0.180**	0.137*	0.151*	0.159**
	(2.38)	(1.73)	(1.95)	(2.04)
<i>Ratio of Non-residential Real Estate</i>	0.132	0.186*	0.209**	0.246**
	(1.53)	(1.90)	(2.16)	(2.57)
Panel C: By Type of Changes	(1)	(2)	(3)	(4)
	<i>Reluctance of Tightening Real Estate</i>			
<i>Ratio of Real Estate \times Composition Changes</i>	-0.115	-0.143	-0.133	-0.131
	(-1.17)	(-1.38)	(-1.32)	(-1.41)
<i>Ratio of Real Estate</i>	0.192***	0.160*	0.176**	0.190**
	(2.60)	(1.86)	(2.09)	(2.25)
<i>Composition Changes</i>	0.073	0.104*	0.098*	0.098*
	(1.29)	(1.83)	(1.75)	(1.92)
Observations	1,809	1,809	1,809	1,809
Other Controls	NO	YES	YES	YES
Electoral District FE	NO	YES	YES	YES
Macro Controls	NO	NO	YES	NO
Year FE	NO	NO	NO	YES

Table 6: Robustness of the Baseline Regression: Alternative Categorization of Proposed Bills

We report the robustness of our results in Table 2 using an alternative dependent variable based on the official list of tightening bills categorized by the South Korean government. We use the member-year observations from 2015 to 2020, which is shorter than the full sample period due to data availability. The dependent variable is *Reluctance of Tightening Real Estate (Government sorted)*, a dummy variable equals to 1 if a Congress member does not propose any real estate bill in the official list of tightening bills categorized by the government in a year, and 0 otherwise. Regression specifications are the same as in Table 2. See Appendix Table 1 for the variable descriptions. The *t*-statistics reported in parentheses are based on standard errors clustered at the Congress member level. ***, **, * denote 1%, 5%, and 10% statistical significance.

	(1)	(2)	(3)	(4)
	<i>Reluctance of Tightening Real Estate (Government sorted)</i>			
<i>Ratio of Real Estate</i>	0.246** (2.08)	0.273* (1.73)	0.266 (1.65)	0.308* (1.85)
<i>Log (Total Assets)</i>	0.073*** (3.35)	0.062** (2.15)	0.054* (1.82)	0.063** (2.07)
<i>Leverage</i>	0.011 (0.11)	0.097 (0.78)	0.104 (0.84)	0.056 (0.45)
<i>Ratio of Cash</i>	0.266* (1.73)	0.245 (1.41)	0.210 (1.19)	0.340* (1.89)
<i>Ratio of Residential Deposits</i>	0.293* (1.81)	0.382* (1.93)	0.363* (1.80)	0.415** (2.04)
<i>Conservative Party</i>		0.061 (1.43)	0.083* (1.88)	0.061 (1.38)
<i>Age</i>		0.001 (0.32)	-0.001 (-0.26)	0.003 (0.67)
<i>Female</i>		-0.166*** (-2.62)	-0.176*** (-2.72)	-0.157** (-2.34)
<i>GDP Growth</i>			-43.413*** (-4.01)	
<i>HPI Growth</i>			-6.190*** (-4.06)	
Observations	1,076	1,076	1,076	1,076
Adjusted R-squared	0.014	0.161	0.182	0.278
Other Controls	NO	YES	YES	YES
Electoral District FE	NO	YES	YES	YES
Year FE	NO	NO	NO	YES

Table 7: The Effect of Real Estate Ownership on Other Tightening Economic Policy Bills

We report the regression results on the effect of real estate ownership on other tightening economic policy bills. The dependent variable is *Reluctance of Tightening Economic Policy*, a dummy variable that equals to 1 if a Congress member does not propose any tightening non-real estate economic policy in a year and 0 otherwise. Regression specifications are the same as in Table 2. See Appendix Table 1 for the variable descriptions. The t -statistics reported in parentheses are based on standard errors clustered at the Congress member level. ***, **, * denote 1%, 5%, and 10% statistical significance.

	(1)	(2)	(3)	(4)
	<i>Reluctance of Tightening Economic Policy</i>			
<i>Ratio of Real Estate</i>	0.046 (0.81)	-0.065 (-0.88)	-0.041 (-0.57)	-0.006 (-0.09)
<i>Log (Total Assets)</i>	0.043*** (4.12)	0.059*** (3.99)	0.063*** (4.40)	0.049*** (3.49)
<i>Leverage</i>	0.043 (0.89)	0.103 (1.64)	0.094 (1.54)	0.108* (1.81)
<i>Ratio of Cash</i>	0.036 (0.54)	0.036 (0.42)	0.060 (0.73)	0.127 (1.56)
<i>Ratio of Residential Deposits</i>	-0.087 (-1.14)	-0.060 (-0.63)	-0.023 (-0.25)	0.022 (0.25)
<i>Conservative Party</i>		0.071*** (2.81)	0.057** (2.33)	0.079*** (3.20)
<i>Age</i>		0.005*** (2.79)	0.006*** (2.93)	0.006*** (3.00)
<i>Female</i>		-0.023 (-0.67)	-0.015 (-0.45)	-0.008 (-0.25)
<i>GDP Growth</i>			5.936*** (8.16)	
<i>HPI Growth</i>			-0.312 (-0.81)	
Observations	1,809	1,809	1,809	1,809
Adjusted R-squared	0.012	0.183	0.230	0.315
Other Controls	NO	YES	YES	YES
Electoral District FE	NO	YES	YES	YES
Year FE	NO	NO	NO	YES

Table 8: Instrumental Variable Analysis

We report the instrumental variables regression results using an exogenous event, the 2017 Pohang Earthquake in Columns (1)-(2). Column (1) reports the first stage regression result of *Ratio of Real Estate* on *Ratio Within 40 Miles*. Column (2) reports the second stage regression result of *Reluctance of Tightening Real Estate* on the instrumented *Ratio of Real Estate* using the earthquake shock. We include control variables, such as *Log (Total Assets)*, *Leverage*, *Ratio of Cash*, and *Ratio of Residential Deposits*. We also include other Congress members' characteristics and Metropolitan fixed effects. See Appendix Table 1 for the variable descriptions. The table reports point estimates with *t*-statistics in parentheses. Standard errors in Column (1) are clustered at the individual level. Standard errors in Column (2) are calculated using the bootstrap method, and clustered at the individual level. ***, **, * denote 1%, 5%, and 10% statistical significance.

	1st Stage (1) <i>Ratio of Real Estate</i>	2nd Stage (2) <i>Reluctance of Tightening Real Estate</i>
<i>Ratio Within 40 Miles</i>	-0.146*** (-3.29)	
<i>Ratio of $\widehat{Real Estate}$</i>		1.379** (1.98)
<i>Log (Total Assets)</i>	-0.045** (-2.34)	0.103*** (2.85)
<i>Leverage</i>	0.130* (1.69)	0.155 (0.92)
<i>Ratio of Cash</i>	-0.729*** (-9.17)	1.144** (2.08)
<i>Ratio of Residential Deposits</i>	-0.891*** (-14.70)	1.214* (1.74)
<i>Conservative Party</i>	0.024 (0.88)	0.025 (0.58)
<i>Age</i>	-0.001 (-0.43)	0.004 (1.58)
<i>Female</i>	0.006 (0.22)	0.059 (0.74)
Observations	457	457
Other Controls	YES	YES
Metropolitan FE	YES	YES
Year FE	YES	YES

Appendix Table 1: Definitions of Variables

Variable	Description
Tightening Real Estate Bills	
<i>Reluctance of Tightening Real Estate</i>	1 if a Congress member i does not propose any tightening real estate bill in year t and 0 if otherwise
<i>Reluctance of Tightening Real Estate (Government sorted)</i>	1 if a Congress member i does not propose any tightening (categorized by government) real estate bill in year t and 0 if otherwise
Portfolio of Congress Members	
<i>Total Assets (mils)</i>	Total amount of all assets disclosed by a Congress member i in year t in million KRW
<i>Log (Total Assets)</i>	The logarithm of <i>Total Assets</i>
<i>Leverage</i>	The ratio of total debt amount to total asset amount disclosed by a Congress member i in year t
<i>Ratio of Real Estate</i>	The ratio of total amount of land and buildings ownership to total asset amount disclosed by a Congress member i in year t
<i>Ratio of Cash</i>	The ratio of total amount of cash and deposits to total asset amount disclosed by a Congress member i in year t
<i>Ratio of Residential Deposits</i>	The ratio of total amount of residential deposits to total asset amount disclosed by a Congress member i in year t
<i>Ratio of Stocks</i>	The ratio of total amount of stocks and blind trusts to total asset amount disclosed by a Congress member i in year t
<i>Ratio of Other Assets</i>	The ratio of total amount of other assets to total asset amount disclosed by a Congress member i in year t
Characteristics of Congress Members	
<i>Conservative Party</i>	1 if a Congress member i is associated with the conservative party in year t and 0 otherwise
<i>Terms Served</i>	The number of terms of a Congress member i serving as Congress member by year t
<i>Primary Sponsor</i>	1 if a Congress member i proposes at least one real estate bill as the primary sponsor in year t and 0 otherwise
<i>Age</i>	The age of a Congress member i in year t
<i>Female</i>	1 if a Congress member i is a female and 0 otherwise
<i>Education (high school or below)</i>	1 if a Congress member i 's final education is high school or below and 0 otherwise
<i>Education (college)</i>	1 if a Congress member i 's final education is college and 0 otherwise
<i>Education (postgraduate)</i>	1 if a Congress member i 's final education is postgraduate and 0 otherwise
Macroeconomic Factors	
<i>GDP Growth</i>	Two-year lagged real GDP growth
<i>HPI Growth</i>	Two-year lagged House Price Index growth minus Consumer Price Index growth

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Appendix Table 1 continues

Variable	Description
The Number of Real Estate Assets	
<i>Number of Real Estate</i>	The number of real estate assets owned by a Congress member i in year t
<i>Large Number</i>	1 if the number of real estate assets owned by a Congress member i is above the top quintile of the distribution in each year t and 0 otherwise
Decomposing Ratio of Real Estate	
By the type of ownership	
<i>Ratio of Owned by Congressmen</i>	The ratio of total amount of real estate directly owned by a member i to total asset amount disclosed by a Congress member i in year t
<i>Ratio of Owned by Family</i>	The ratio of total amount of real estate held by a member i 's family to total asset amount disclosed by a Congress member i in year t
By the type of real estate	
<i>Ratio of Residential Real Estate</i>	The ratio of total amount of residential real estate to total asset amount disclosed by a Congress member i in year t
<i>Ratio of Non-residential Real Estate</i>	The ratio of total amount of non-residential real estate to total asset amount disclosed by a Congress member i in year t
By the type of location	
<i>Ratio of Own Electoral District</i>	The ratio of total real estate amount located in a Congress member i 's own electoral district to total asset amount disclosed by the member i in year t
<i>Ratio of Other Electoral Districts</i>	The ratio of total real estate amount located in the other electoral district to total asset amount disclosed by a Congress member i in year t
Type of changes	
<i>Composition Changes</i>	1 if a Congress member i in year t has any new acquisition, sales, and real estate inheritance and 0 otherwise
Measure of Entrenchment	
<i>High Votes</i>	1 if a Congress member i wins 50% of votes from the latest election and 0 otherwise
<i>Vote Concentration</i>	1 if the Herfindahl index of candidates' votes in an electoral district is above the median in the distribution of that in the latest election and 0 otherwise
<i>Party Shares</i>	1 if the fraction of the party's seats in the total number of seats in the province is above the top tercile of the distribution and 0 otherwise

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Appendix Table 1 continues

Variable	Description
IV	
<i>Ratio Within 40 Miles</i>	The ratio of total real estate amount within 40 miles from Pohang epicenter to total real estate amount disclosed by a Congress member i at the beginning of 2017
Additional Analyses	
<i>Reluctance of Tightening Real Estate (Alt)</i>	1 if the ratio of tightening real estate bills to total real estate bills proposed by a Congress member i is below the median in the distribution of that in year t
<i>Reluctance of Tightening Economic Policy</i>	1 if a Congress member i does not propose any tightening non-real estate economic policy in year t and 0 otherwise
Local Socioeconomic Variables	
<i>Fraction of Workforce in Financial Industry</i>	Source : Korean Statistical Information Service (kosis.kr) Fraction of the number of workforce in financials (District-level)
<i>Fraction of Workforce in Real Estate Industry</i>	Fraction of the number of workforce in real estate industry (District-level)
<i>Fraction of Under High School</i>	Fraction of people who are educated under high school (District-level)
<i>Fraction of Graduates</i>	Fraction of people who have master degree (District-level)
<i>Fraction of People in Poverty</i>	Fraction of people who are in poverty (Metropolitan-level)

Appendix Table 2: Excluding Duplicative Bills by Similarity Score

We report our baseline regression after excluding tightening real estate bills that are similar to others. The dependent variable is *Reluctance of Tightening Real Estate*. We calculate a similarity score between 0 and 1 for every two bills in our 849 tightening real estate bills using TF-IDF (Term Frequency-Inverse Document Frequency). We exclude bills if there is a tightening real estate bill proposed earlier in the same assembly by other members but is similar at a certain level. In Columns (1)-(4), we repeat our baseline regressions in Table 2 after dropping the duplicative bills with a similarity score above 0.5. Columns (5) and (6) show the results in Column (4) with the higher thresholds of similarity score. The *t*-statistics reported in parentheses are based on standard errors clustered at the Congress member level. ***, **, * denote 1%, 5%, and 10% statistical significance.

<i>Measure of Similar Bills:</i>	(1)	(2)	(3)	(4)	(5)	(6)
		Similarity > 0.5			> 0.75	> 0.9
	<i>Reluctance of Tightening Real Estate</i>					
<i>Ratio of Real Estate</i>	0.163** (2.24)	0.136* (1.72)	0.153** (1.99)	0.166** (2.06)	0.153** (2.00)	0.163** (2.10)
<i>Log (Total Assets)</i>	0.060*** (4.06)	0.036* (1.76)	0.040** (1.98)	0.029 (1.41)	0.031 (1.47)	0.028 (1.32)
<i>Leverage</i>	0.057 (0.88)	0.222*** (2.80)	0.209*** (2.65)	0.183** (2.37)	0.141* (1.83)	0.116 (1.52)
<i>Ratio of Cash</i>	0.168* (1.85)	0.233** (2.24)	0.260** (2.57)	0.252** (2.45)	0.252** (2.54)	0.262*** (2.60)
<i>Ratio of Residential Deposits</i>	0.141 (1.36)	0.225* (1.93)	0.253** (2.21)	0.246** (2.13)	0.132 (1.16)	0.136 (1.19)
<i>Conservative Party</i>		0.140*** (4.55)	0.128*** (4.20)	0.151*** (4.59)	0.114*** (3.35)	0.115*** (3.35)
<i>Age</i>		0.005** (2.02)	0.005** (2.12)	0.004* (1.85)	0.004* (1.76)	0.004 (1.57)
<i>Female</i>		0.029 (0.74)	0.031 (0.77)	0.030 (0.72)	0.055 (1.27)	0.050 (1.19)
<i>GDP Growth</i>			5.444*** (6.71)			
<i>HPI Growth</i>			0.331 (0.59)			
Observations	1,809	1,809	1,809	1,809	1,809	1,809
Adjusted R-squared	0.010	0.142	0.169	0.224	0.246	0.245
Other Controls	NO	YES	YES	YES	YES	YES
Electoral District FE	NO	YES	YES	YES	YES	YES
Year FE	NO	NO	NO	YES	YES	YES

Appendix Table 3: Alternative Measure of Bill Proposals

We report our baseline regression using an alternative measure of Congress members' bill proposal behaviors. We define *Reluctance of Tightening Real Estate (Alt)* as a dummy that equals to 1 if the ratio of tightening real estate bills to total real estate bills proposed by a Congress member is below the median in the distribution of that in a year. Regression specifications are the same as in Table 2. The *t*-statistics reported in parentheses are based on standard errors clustered at the Congress member level. ***, **, * denote 1%, 5%, and 10% statistical significance.

	(1)	(2)	(3)	(4)
	<i>Reluctance of Tightening Real Estate (Alt)</i>			
<i>Ratio of Real Estate</i>	0.142 (1.47)	0.213** (2.03)	0.210** (2.00)	0.211** (2.00)
<i>Log (Total Assets)</i>	0.020 (1.29)	0.032 (1.48)	0.031 (1.42)	0.029 (1.33)
<i>Leverage</i>	0.037 (0.54)	0.100 (1.09)	0.101 (1.09)	0.120 (1.29)
<i>Ratio of Cash</i>	0.066 (0.59)	0.231* (1.83)	0.230* (1.81)	0.251** (1.97)
<i>Ratio of Residential Deposits</i>	0.091 (0.74)	0.305** (2.15)	0.298** (2.09)	0.296** (2.08)
<i>Conservative Party</i>		0.079** (2.40)	0.083** (2.50)	0.096*** (2.80)
<i>Age</i>		0.008*** (3.15)	0.008*** (3.17)	0.008*** (3.03)
<i>Female</i>		0.068 (1.43)	0.068 (1.42)	0.068 (1.40)
<i>GDP Growth</i>			-1.157 (-1.29)	
<i>HPI Growth</i>			0.542 (0.71)	
Observations	1,809	1,809	1,809	1,809
Adjusted R-squared	0.001	0.035	0.035	0.037
Other Controls	NO	YES	YES	YES
Electoral District FE	NO	YES	YES	YES
Year FE	NO	NO	NO	YES

Appendix Table 4: Controlling for Time-Varying Local Interests or Economic Conditions

We report the robustness of our results with additional controls for time-varying local interests or economic conditions. Panel A includes additional controls at the electoral district level in the same manner of Mian et al. (2010). Column (1) reports the same result as in Column (4) of Table 2, but with a smaller sample size due to missing additional controls. Column (2) includes *Fraction of Workforce in Financial Industry* and *Fraction of Workforce in Real Estate Industry*. Column (3) includes variables regarding education: *Fraction of Under High School* and *Fraction of Graduates*. Column (4) includes *Fraction of People in Poverty*. In Panel B, we re-estimate Table 2 by replacing *Electoral District* fixed effects with *Metropolitan* or *Metropolitan × Year* fixed effects. See Appendix Table 1 for the variable descriptions. The *t*-statistics reported in parentheses are based on standard errors clustered at the Congress member level. ***, **, * denote 1%, 5%, and 10% statistical significance.

Panel A: District-level Controls	(1)	(2)	(3)	(4)
	<i>Reluctance of Tightening Real Estate</i>			
<i>Ratio of Real Estate</i>	0.163** (2.09)	0.170** (2.16)	0.164** (2.07)	0.164** (2.07)
<i>Log (Total Assets)</i>	0.028 (1.33)	0.028 (1.34)	0.027 (1.27)	0.027 (1.29)
<i>Leverage</i>	0.121 (1.59)	0.125 (1.65)	0.116 (1.50)	0.112 (1.44)
<i>Ratio of Cash</i>	0.269*** (2.68)	0.276*** (2.75)	0.271*** (2.69)	0.268*** (2.66)
<i>Ratio of Residential Deposits</i>	0.135 (1.18)	0.148 (1.27)	0.152 (1.31)	0.154 (1.33)
<i>Conservative Party</i>	0.113*** (3.31)	0.115*** (3.39)	0.116*** (3.45)	0.113*** (3.31)
<i>Age</i>	0.004 (1.55)	0.004 (1.55)	0.004 (1.64)	0.004* (1.69)
<i>Female</i>	0.049 (1.16)	0.049 (1.15)	0.046 (1.09)	0.048 (1.12)
<i>Fraction of Workforce in Financial Industry</i>		2.090 (1.01)	2.887 (1.43)	3.093 (1.50)
<i>Fraction of Workforce in Real Estate Industry</i>		-4.407 (-1.10)	-4.375 (-1.10)	-4.570 (-1.14)
<i>Fraction of Under High School</i>			4.122* (1.90)	4.026* (1.83)
<i>Fraction of Graduates</i>			9.566 (1.17)	9.710 (1.19)
<i>Fraction of People in Poverty</i>				-4.234 (-0.87)
Observations	1,807	1,807	1,807	1,807
Adjusted R-squared	0.244	0.244	0.246	0.245
Other Controls	YES	YES	YES	YES
Electoral District FE	YES	YES	YES	YES
Year FE	YES	YES	YES	YES

Continued on the next page

Appendix Table 4 continues

Panel B: Metropolitan \times Year FE	(1)	(2)	(3)	(4)
	<i>Reluctance of Tightening Real Estate</i>			
<i>Ratio of Real Estate</i>	0.162** (2.28)	0.144** (2.42)	0.156*** (2.65)	0.168*** (2.70)
<i>Log (Total Assets)</i>	0.048*** (3.25)	0.030** (2.09)	0.033** (2.26)	0.028* (1.84)
<i>Leverage</i>	0.014 (0.22)	0.092 (1.48)	0.084 (1.36)	0.056 (0.87)
<i>Ratio of Cash</i>	0.157* (1.72)	0.161** (2.03)	0.178** (2.28)	0.173** (2.17)
<i>Ratio of Residential Deposits</i>	0.069 (0.69)	0.096 (1.08)	0.123 (1.42)	0.160* (1.80)
<i>Conservative Party</i>		0.076*** (3.15)	0.068*** (2.81)	0.067** (2.55)
<i>Age</i>		0.002 (0.89)	0.002 (1.08)	0.001 (0.51)
<i>Female</i>		0.057 (1.48)	0.060 (1.55)	0.043 (1.10)
<i>GDP Growth</i>			5.921*** (8.11)	
<i>HPI Growth</i>			0.050 (0.10)	
Observations	1,815	1,815	1,815	1,815
Adjusted R-squared	0.009	0.151	0.183	0.271
Other Controls	NO	YES	YES	YES
Metropolitan FE	NO	YES	YES	NO
Metropolitan \times Year FE	NO	NO	NO	YES