

# Identifying the Heterogeneous Impact of Highly Anticipated Events: Evidence from the Tax Cuts and Jobs Act

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American Economic Association  
January, 2022

- ① Introduction
- ② Evidence from Simulations
- ③ Empirical Model
- ④ Results
- ⑤ Conclusion

- The market's reaction to events can be misleading, especially with high anticipation
  - Huberman and Schwert (1985), Bhattacharya et al. (2000)
- We use simulation to quantify anticipation bias
- We propose a new method to pool data across firms to capture the counterfactual outcome
- The difference is the true event impact

Three main approaches to dealing with anticipation

- Firm characteristics (Malatesta and Thompson, 1985; Brennan, 1990)
- Firm-level counterfactual estimation models using stock and options data
  - Subramanian (2004)
  - Barraclough, Robinson, Smith, Whaley (2013)
  - Borochin (2014)
  - Borochin, Golec (2016)

- Predictive markets
  - Snowberg et al. (2007)
  - Wolfers and Zitzewitz (2009)
  - Snowberg et al. (2011)
- Characteristics need customization, firm-level estimations require identifying assumptions, betting markets are illiquid
- We can recover *heterogeneous* impacts of market-wide events
  - No need to specify an ex-ante ordering of event payoffs
- Eg, TCJA effect on a particular firm depends on its R&D, overseas operations, deferred tax assets

# Main Findings

- Anticipation of an event can break traditional econometric inference in simulations
  - at 80% anticipation, correlation between firm price reaction and true event impact is only 0.44
  - for 96% anticipation (eg. Tax Cuts and Jobs Act) the correlation drops to 0.20
  - the estimated aggregate impact across all firms is near zero at 96% anticipation, despite a meaningful true effect
- We derive a new method to let the data tell us which firm has positive/negative impact from a heterogeneous event, robust to high anticipation
- Applying this method to the TCJA yields an impact estimate of 12.36%, strongest for growth firms with high patent counts

- No satisfactory way to address highly-anticipated economy-wide effects with both winners and losers
- These are usually very important events, and simulation shows that naive analyses are very misleading
- We provide a way to address these events

# Traditional Econometric Setup

The share price of a firm  $i$  exposed to an event is

$$P_{i,T+1} = \begin{cases} s_i x_i \epsilon_i & \text{if } \mathbb{I} = 1 \\ x_i \epsilon_i, & \text{if } \mathbb{I} = 0 \end{cases}$$

Therefore, the stock price of firm  $i$  at time  $T$  is calculated as:

$$\begin{aligned} P_{i,T} &= \mathbb{E}_T[\mathbb{I}(s_i x_i \epsilon_i) + (1 - \mathbb{I})(x_i \epsilon_i) | s_i, x_i] \\ &= q s_i x_i \mathbb{E}[\epsilon_i] + (1 - q) x_i \mathbb{E}[\epsilon_i] \\ &= (q s_i + (1 - q)) x_i \end{aligned}$$

How do we estimate the impact of the event  $s_i$ ?



# Traditional Econometric Setup

Given  $\mathbb{I} = 1$ , this is calculated as:

$$\begin{aligned}\frac{P_{i,T+1} - P_{i,T}}{P_{i,T}} &= \frac{s_i x_i \epsilon_i - (q s_i + (1 - q)) x_i}{(q s_i + (1 - q)) x_i} \\ &= \frac{s_i \epsilon_i - (q s_i + (1 - q))}{(q s_i + (1 - q))}\end{aligned}$$

If the event was completely unanticipated, plugging in  $q = 0$  yields:

$$\frac{P_{i,T+1} - P_{i,T}}{P_{i,T}} = s_i \epsilon_i - 1$$

However, the traditional approach falters when  $q > 0$ , capturing only the unanticipated impact. But is it at least positively correlated?

# Traditional Econometric Setup

To estimate the net aggregate impact of the event  $\mu_S - 1$ , the traditional event study methodology would suggest taking the average of the individual estimates:

$$\frac{1}{N} \sum_{i=1}^N \frac{P_{i,T+1} - P_{i,T}}{P_{i,T}} = \frac{1}{N} \sum_{i=1}^N \frac{s_i \epsilon_i - (q s_i + (1 - q))}{(q s_i + (1 - q))}$$

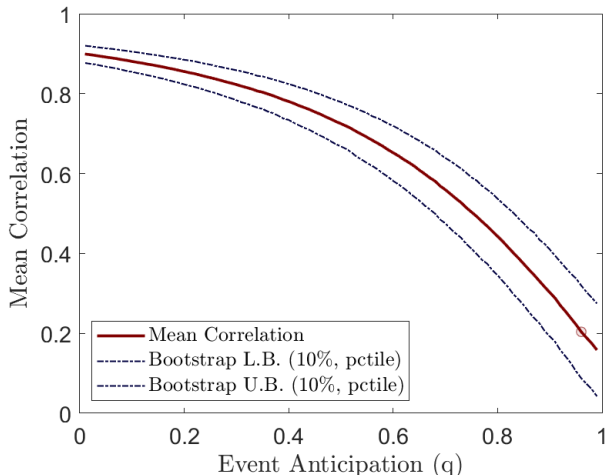
If  $q = 0$ , this would be a reasonable estimator:

$$\mathbb{E} \left[ \frac{1}{N} \sum_{i=1}^N \frac{P_{i,T+1} - P_{i,T}}{P_{i,T}} \right] = \frac{1}{N} \sum_{i=1}^N (\mathbb{E}[s_i] - 1) = \mu_S - 1$$

However, for positive values of  $q$ , this estimator would only capture the unanticipated fraction of the aggregate impact of the event. Given that  $q$  is not clear, a low estimate may be due to a low true impact (low  $\mu_S$ ) or high anticipation (high  $q$ ).

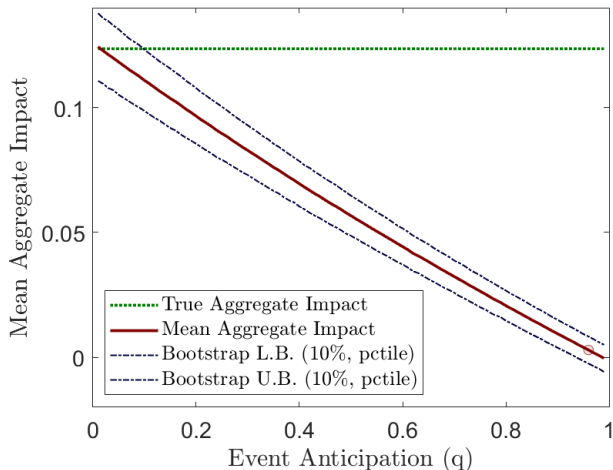
# Simulation of Performance of Estimator of Firm Impact

The Correlation Between  $(P_{i,T+1} - P_{i,T})/P_{i,T}$  and  $s_i - 1$



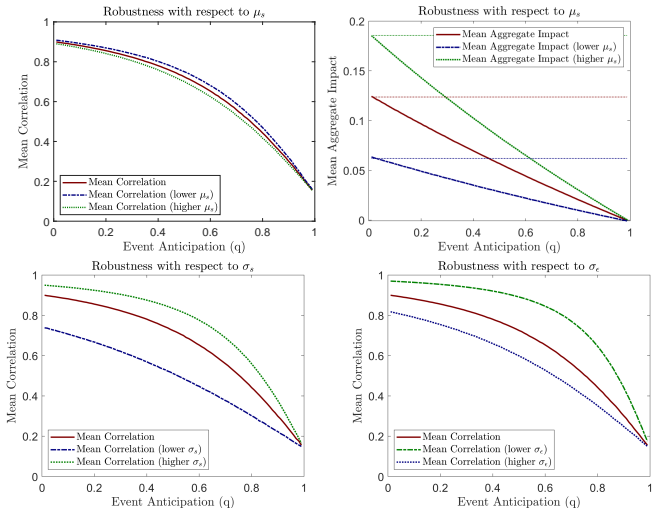
# Simulation of Performance of Estimator of Aggregate Impact

The True and Estimated Net Aggregate Impact  $\mu_s - 1$



# Robustness of Simulation

Do magnitudes of true impact and noise matter?



- The 2017 Tax Cuts and Jobs Act (TCJA) was a highly anticipated event
- Moreover, the TCJA had winners and losers
- Our simulation results show that the traditional approach can fail to properly estimate firm impacts of such an event, even after it occurs
- Empirical analysis presents a roadmap for other applications of our general methodology

# Estimating the Counterfactual

For a firm  $i$  given the stock price  $S_{i,t}$  and  $N$  call options with unique strikes  $K_j$  and a common  $\tau$  that ends after the event, the prices of the  $N + 1$  assets are:

$$\begin{aligned} S_{i,t} &= E_t(q) \cdot E_t(S_{i,u}) + (1 - E_t(q)) \cdot E_t(S_{i,d}) \\ c_{i,1,t} &= E_t(q) \cdot C(E_t(S_{i,u}), E_t(\sigma_{i,u}), K_1, \tau) + (1 - E_t(q)) \cdot C(E_t(S_{i,d}), E_t(\sigma_{i,d}), K_1, \tau) \\ &\dots \\ c_{i,N,t} &= E_t(q) \cdot C(E_t(S_{i,u}), E_t(\sigma_{i,u}), K_N, \tau) + (1 - E_t(q)) \cdot C(E_t(S_{i,d}), E_t(\sigma_{i,d}), K_N, \tau) \end{aligned}$$

Why not puts?

# Identification

- Five parameters:  $q_i$  the probability of TCJA passage,  $S_{i,u}$  and  $\sigma_{i,u}$  the value and volatility of firm  $i$  if TCJA passes,  $S_{i,d}$  and  $\sigma_{i,d}$  if not
- Label switching is a problem, prior settings (eg M&A, Obamacare) allow for identifying assumption such as  $S_{i,u} > S_{i,d}$  that does not apply here
- Our innovation is to consider both possibilities simultaneously, and pick the more applicable one

$$V_{i,t,winner}(q, \theta) = \min_{\theta} |P_{i,t} - \hat{P}_{i,t}(q, \theta)| \quad s.t. \quad S_{i,u} \geq S_{i,d}$$

$$V_{i,t,loser}(q, \theta) = \min_{\theta} |P_{i,t} - \hat{P}_{i,t}(q, \theta)| \quad s.t. \quad S_{i,u} < S_{i,d}$$



- We classify the firm as a “winner” if the  $S_{i,u} \geq S_{i,d}$  restriction results in a better fit more than half of the time over the 30-day period from November 10 to December 2, 2017
- It is a “loser” otherwise. In other words, we let the data tell us which identifying restriction is more appropriate for each firm.
- Once this has been established, we repeat the optimization with the appropriate restriction in place:

$$V_{i,t}(q, \theta) = \begin{cases} \min_{\theta} |P_{i,t} - \hat{P}_{i,t}(q, \theta)| & \text{s.t. } S_{i,u} \geq S_{i,d} & \text{if firm } i \text{ is a “winner”} \\ \min_{\theta} |P_{i,t} - \hat{P}_{i,t}(q, \theta)| & \text{s.t. } S_{i,u} < S_{i,d} & \text{otherwise} \end{cases}$$

- Getting  $q$  right is key, so we develop a specialized version of the estimator for a common  $q$  across all  $M$  firms

$$W_{i,t}(q, \theta_i) = \begin{cases} |1 - \hat{P}_{i,t}(q, \theta_i)/P_{i,t}| & \text{s.t. } S_{i,u} \geq S_{i,d} \quad \text{if firm } i \text{ is a "winner"} \\ |1 - \hat{P}_{i,t}(q, \theta_i)/P_{i,t}| & \text{s.t. } S_{i,u} < S_{i,d} \quad \text{otherwise} \end{cases}$$

$$\left( q_t, \{\theta_{i,t}\}_{i=1}^M \right) = \arg \min_{q, \{\theta_i\}_{i=1}^M} \left\{ \sum_{i=1}^M W_{i,t}(q, \theta_i) \right\}$$

- Benefit is clear, we get  $M \times N$  call option restrictions (600 in our application) on  $q_t$
- Cost is daunting,  $4M + 1$  parameters must be jointly estimated

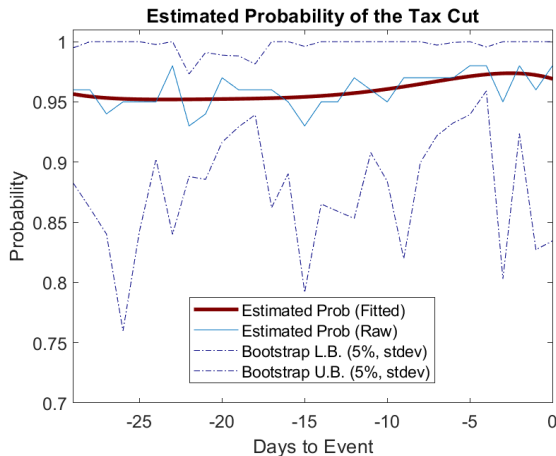
- However, note that the firm-specific parameters  $\{\theta_{i,t}\}_{i=1}^M$  depends on the results from other firms only through the common event probability  $q$
- Thus, we first divide the range  $q \in [0, 1]$  into a discrete grid  $Q$
- Then, for each firm  $i$ , on each date  $t$ , and for every  $q \in Q$ , we estimate  $\theta_{q,i,t}$  which minimizes  $W_{i,t}(q, \theta_i)$  given  $q$
- The common probability on day  $t$  can be calculated as:

$$q_t = \arg \min_q \left\{ \sum_{i=1}^M W_{i,t}(q, \theta_{q,i,t}) \right\}$$

- OptionMetrics and Compustat data over debate window, November 10 to December 2, 2017
- Top 100 firms by option volume, cumulatively between 29% and 50% of daily call volume in 2017Q4, 22% of Compustat market cap
- Adding more firms increases reliance on zero-volume contracts, ie uninformative and stale prices
- We also add innovation measures and tax assets and liabilities

# Probability of TCJA Passage

Confidence interval from bootstrap standard deviation, 1000 simulations each day of 100 firms drawn with replacement



# Comparison of Estimates

$S_u/S_d - 1$	RET	CAR[-3,0]	CAR[-5,0]	CAR[-7,0]	CAR[-10,0]
12.36%	0.68%	0.96%	0.89%	1.09%	2.29%
$S_u/S_d - 1$	RET	CAR[-3,0]	CAR[-5,0]	CAR[-7,0]	CAR[-10,0]
1.000	0.135	0.183	0.133	0.197	0.233

# What Makes Winners and Losers?

We split the sample using the  $S_u/S_d$  ratio, and check whether characteristics are significantly different across the two groups

Firm Attributes	Differences	t-statistics
R&D Intensity	0.021***	(4.05)
Patent Count	0.498***	(5.37)
Total Citations	0.489***	(4.86)
Total Originality	0.345***	(4.53)
Total Generality	0.201***	(4.09)
Average Citations	-0.014	(-0.86)
Average Originality	-0.018***	(-3.41)
Average Generality	-0.009**	(-2.16)
Tangibility	-0.029***	(-2.98)
Sales Growth	0.040***	(6.44)
Asset Growth	0.022***	(3.39)
Employment Growth	0.011**	(2.51)
Cash Effective Tax Rate	-1.616***	(-3.59)
Indefinitely Reinvested Foreign Earnings/Assets	-0.048***	(-5.14)
Net Tax Assets/Assets	-0.023***	(-4.50)
Cash/Asset	-0.006*	(-1.76)
Market to Book Ratio	-0.088	(-1.64)
Size (log(assets))	0.232***	(3.56)
Leverage	-0.012*	(-1.95)
Asset Maturity	-0.727***	(-4.13)
Property Plant and Equipment/Assets	-0.029***	(-2.98)
Profitability	-0.014***	(-3.95)
Return on Assets (ROA)	-0.015***	(-4.97)
Return on Equity (ROE)	-0.057***	(-3.48)
Whited-Wu Index	-0.011***	(-3.01)
Advertising Expenses	-0.007***	(-5.54)

# Economic Significance

We split the sample into two using the median of firm characteristics, and test the  $S_u/S_d$  ratio difference

$S_u/S_d$ ratio	Differences	t-statistics
R&D Intensity	0.026***	(3.74)
Patent Count	0.036***	(6.52)
Total Citations	0.036***	(6.52)
Total Originality	0.027***	(4.83)
Total Generality	0.020***	(3.55)
Average Citations	0.005	(0.86)
Average Originality	-0.034***	(-6.09)
Average Generality	-0.017***	(-3.00)
Tangibility	-0.006	(-1.07)
Sales Growth	0.020***	(3.67)
Asset Growth	0.012**	(2.19)
Employment Growth	0.002	(0.37)
Cash Effective Tax Rate	-0.021***	(-3.71)
Indefinitely Reinvested Foreign Earnings/Assets	-0.021***	(-3.42)
Net Tax Assets/Assets	-0.006	(-1.11)
Cash/Asset	-0.022***	(-4.03)
Market to Book Ratio	-0.011*	(-1.94)
Size (log(assets))	0.019***	(3.46)
Leverage	-0.004	(-0.79)
Asset Maturity	-0.004	(-0.66)
Property Plant and Equipment/Assets	-0.006	(-1.07)
Profitability	-0.008	(-1.37)
Return on Assets (ROA)	-0.031***	(-5.58)
Return on Equity (ROE)	-0.017***	(-3.10)
Whited-Wu Index	-0.017***	(-3.10)
Advertising Expenses	-0.015**	(-2.24)



- We develop a method to estimate ex-ante event probabilities for highly anticipated events, which is also robust to firm-level heterogeneity in the impact of the event
  - TCJA passage more than 90% likely
  - Impact is estimated at 12.36% across a sample of the 100 large firms, compared to an average of 0.68% when ignoring anticipation
- We demonstrate the existence of a downward bias for the aggregate impact of an event across multiple firms from anticipation

- TCJA application shows that large firms with high patent counts and growth prospects are the greatest relative winners
- Two innovation strategies: the production of a few high-quality patents versus a large number of mediocre patents
- TCJA appears to encourage the latter

Thank You!