Diagnostic Expectations and Credit Cycles

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Introduction

- Renewal of academic interest on the link between credit expansion and subsequent bust
 - Schularick, Taylor (2012): credit growth and financial crisis
 - Mian, Sufi, Verner (2015): household debt and low growth
 - Baron, Xiong (2014): bank credit and crash risk in stocks
 - ▶ Fahlenbrach et al. (2016): loan growth and bank performance
- Complementary findings for corporate debt:
 - Greenwood, Hanson (2013): in credit booms, quality of debt issuers falls. Larger high yield share in bond issuance predicts low (negative) excess returns
 - Gilchrist, Zakrajsek (2012), Krishnamurthy, Muir (2015): credit tightening anticipates the coming recession
 - Lopez-Salido, Stein, Zakrajsek (LSZ 2015): low spreads today predicts rise in credit spreads and low growth afterwards

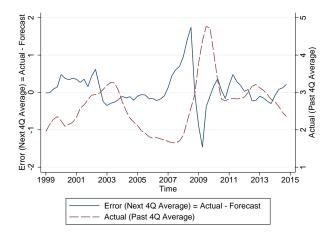
Our Approach

- Build a behavioral model of credit cycles
- Micro-found expectation formation based on Representativeness Heuristic
- Consistent with available evidence but also with predictable expectations errors
- Forward looking; immune to the Lucas critique

Related Literature

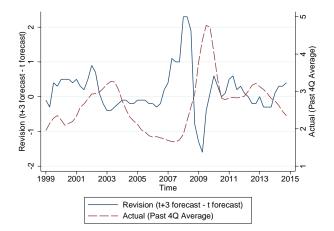
- Financial frictions
 - Those models fail to account for predictable returns and errors
 - Also do not explain where shocks come from
- Extrapolation
 - Greenwood and Shleifer 2014, Barberis et al 2015a,b
 - Our theory micro-founds extrapolation and neglect of risk
- Limited attention
 - Sims 2003, Coibion and Gorodnichenko 2012, 2015, Gabaix 2015
 - These are models of under-reaction, not over-reaction
- Behavioral models of credit cycles
 - Gennaioli, Shleifer, and Vishny 2012, 2015, Greenwood, Hanson, and Jin 2016
 - Our model provides a portable foundation of belief formation

Predictable Expectation Errors of Credit Spreads



Data Source: Blue Chip Financial Forecasts

Predictable Reversals in Expectations of Credit Spreads



Data Source: Blue Chip Financial Forecasts

This Paper

- Model of expectation formation based on Gennaioli and Shleifer's (2010) formalization of Kahneman and Tversky's "representativeness" heuristic
- Inserted into a simple macroeconomic model (no financial frictions), yields many of the previous facts
- What is representativeness?
 - How to model it
 - Implications for Macro-Finance

What is Representativeness?

- KT (1974): we judge the frequency of an attribute by its similarity to, or representativeness for, the parent population
- KT (1983): "an attribute is representative of a class if it is very diagnostic; that is, the relative frequency of this attribute is much higher in that class than in a relevant reference class."
- KT argue that representativeness lies behind systematic, extensively documented biases in probability judgments:
 - ► Base rate neglect, Conjunction Fallacy, Disjunction Fallacy
 - Example (Linda): an intelligent, single woman in her 30's who was an activist in college is deemed more likely to be a feminist bank teller than a bank teller

How to Model Representativeness?

Assess distribution of attribute T in class G

$$h(T = t | G)$$

Following KT, define representativeness of T = t for G as:

$$\frac{h(T=t|G)}{h(T=t|-G)}$$

- ► Distort h(T = t|G) by inflating the probability of values t that score high, neglect / under-weight values that score low
- This model yields the KT biases (GS 2010) and accounts for "kernel of truth" in social stereotypes (Bordalo Coffman Gennaioli Shleifer 2015)

Example: Stereotypes

Hair color distribution among the Irish

h (hair colour | Irish)

• $T \equiv \{red, light, dark\}, G = lrish, -G = World$

hair colour	red	light	dark
Irish	10%	40%	50%
World	1%	14%	85%

The stereotype of Irish overweights red hair:

 $\frac{h(\text{red hair}|\text{Irish})}{h(\text{red hair}|\text{World})} = 10$

Kernel of Truth (Judd and Park 1993, BCGS 2015) Confirmed in data on political, gender, ethnic groups.

Implications for Macro-Finance?

- Given data (Irish), inflate prevalence of hair color (red) whose objective probability goes up the most relative to others
- In a dynamic environment:
 - given news, agents inflate future states of the world whose objective probability goes up the most
 - the context is lagged information
- This yields:
 - extrapolation + neglect of tail risk in a single setup
 - reversals in the absence of news
 - excess volatility
 - immunity to Lucas critique, RE as a special case
 - no learning, rather beliefs distort true process
 - model is portable: unify explanation of lab experiments, social stereotypes, macroeconomic predictions

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Model Ingredients

State of the economy Ω_t at t follows AR(1)

$$\omega_t = b \cdot \omega_{t-1} + \epsilon_t$$

- Diagnostic Expectations about Ω_{t+s}
- Measure 1 of firms of varying risk (different exposure to Ω_t)
- Long lived, risk neutral, representative household that supplies capital to these firms (buys risky debt from them)

Diagnostic Expectations

• After seeing the state ω_t , the agent must represent:

$$h\left(\Omega_{t+1}=\omega_{t+1}|\Omega_t=\omega_t\right)$$

• Here
$$G \equiv \{\Omega_t = \omega_t\}$$

- News assessed relative to -G containing past information
- Main case: reference is information available at t-1

$$-G \equiv \{\Omega_t = b \cdot \omega_{t-1}\}$$

Then representativeness is: $\frac{h(\Omega_{t+1} = \omega_{t+1} | \Omega_t = \omega_t)}{h(\Omega_{t+1} = \omega_{t+1} | \Omega_t = b \cdot \omega_{t-1})}$

Overweighing

• We assume the distorted distribution $h_t^{\theta}(\omega_{t+1})$ to be:

$$h(\Omega_{t+1} = \omega_{t+1} | \Omega_t = \omega_t) \cdot \left[\frac{h(\Omega_{t+1} = \omega_{t+1} | \Omega_t = \omega_t)}{h(\Omega_{t+1} = \omega_{t+1} | \Omega_t = b \cdot \omega_{t-1})} \right]^{\theta} \frac{1}{Z_t}$$

- $\theta \ge 0$ measures the importance of representativeness
- ▶ Rational expectations: special case for $\theta = 0$ or no news $\epsilon_t = 0$
- Inflate density of future states that have become more likely
- Denote Diagnostic Expectations by

$$\mathbb{E}_{t}^{\theta}\left(\omega_{t+1}\right) = \int_{\mathbb{R}} \omega \cdot h_{t}^{\theta}\left(\omega\right) d\omega$$

Specifying -G

Alternative: reference is recent diagnostic expectation

$$\frac{h(\Omega_{t+1} = \omega_{t+1} | \Omega_t = \omega_t)}{h(\Omega_{t+1} = \omega_{t+1} | \Omega_t = \mathbb{E}_{t-1}^{\theta}(\omega_t))}$$

 \blacktriangleright -G influences reaction to news

- different specifications imply different lag structures of expectations
- We proceed with our main case:

$$-G = \{\Omega_t = b \cdot \omega_{t-1}\}$$

and then consider the other cases

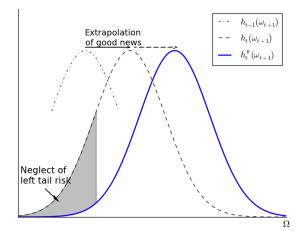
Representation

Proposition 1. When the process for ω_t is AR(1) with normal $(0, \sigma^2)$ shocks, the distribution $h^{\theta}(\omega_{t+1})$ is also normal, with variance σ^2 and mean:

$$\mathbb{E}_{t}^{\theta}\left(\omega_{t+1}\right) = \mathbb{E}_{t}\left(\omega_{t+1}\right) + \theta\left[\mathbb{E}_{t}\left(\omega_{t+1}\right) - \mathbb{E}_{t-1}(\omega_{t+1})\right]$$

- \mathbb{E}_t^{θ} is function of (lagged) rational expectations
- Kernel of truth: overweight incoming news
- Context dependence: the path is important

Neglect of Tail Risk (GSV 2012)



Extrapolation

• Plugging AR(1) in $\mathbb{E}_t^{\theta}(\omega_{t+1})$ we obtain

$$\mathbb{E}_{t}^{\theta}\left(\omega_{t+1}\right) - \omega_{t} = \left[\mathbb{E}_{t}\left(\omega_{t+1}\right) - \omega_{t}\right] + b\theta\left[\omega_{t} - \mathbb{E}_{t-1}(\omega_{t})\right]$$

- Slant toward current objective news $\omega_t \mathbb{E}_{t-1}(\omega_t)$
- Neglect of risk and extrapolation follow from the same psychology of context effects

Diagnostic vs. Adaptive Expectations

Adaptive Expectations

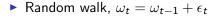
$$\mathbb{E}_t^{\mathsf{a}}(\omega_{t+1}) = \lambda \omega_t + (1-\lambda)\mathbb{E}_{t-1}^{\mathsf{a}}(\omega_t), \quad 0 < \lambda < 1$$

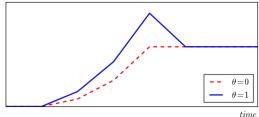
Diagnostic Expectations

$$\mathbb{E}^{ heta}_t\left(\omega_{t+1}
ight) = b(1+ heta)\omega_t - b heta\mathbb{E}_{t-1}(\omega_t)$$

- Overreaction + reversal (rather than momentum)
- Forward looking:
 - Extrapolate only if process is stochastic and persistent, b > 0
 - No mistakes for i.i.d. case, b = 0
 - Immune to Lucas Critique

Sequences of News





- accelerating good news cause sustained optimism
- when good news stop, boom is followed by bust

Non-Fundamental Reversals

• Suppose fundamentals follow a random walk (b = 1). Then:

$$\mathbb{E}_{t}^{\theta}(\omega_{t+1}) = \omega_{t} + \theta\left(\omega_{t} - \omega_{t-1}\right)$$

As a consequence:

$$\mathbb{E}_{t}[\mathbb{E}_{t+1}^{\theta}(\omega_{t+2}) - \mathbb{E}_{t}^{\theta}(\omega_{t+1})] \\ = \mathbb{E}_{t}\left[\left(\omega_{t+1} - \omega_{t}\right)\left(1 + \theta\right) - \theta\left(\omega_{t} - \omega_{t-1}\right)\right] \\ = -\theta\left(\omega_{t} - \omega_{t-1}\right)$$

Excess optimism at t systematically wanes at t + 1 even in the absence of news. A boom is followed by a bust.

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Firms

► Each firm is identified by its risk \(\rho \in \mathbb{R}\) (which is common knowledge), and produces output:

$$y(k|\omega_t, \rho) = \begin{cases} k^{\alpha} & \text{if } \omega_t \ge \rho \\ 0 & \text{if } \omega_t < \rho \end{cases}$$

At t,firm ρ borrows at the interest rate r_{t+1}(ρ) to install capital k_{t+1}(ρ). Maximize expected profit:

$$\max_{k_{t+1}(\rho)} \left[k_{t+1}(\rho)^{\alpha} - k_{t+1}(\rho) \cdot r_{t+1}(\rho) \right] \cdot \mu_t^{\theta}(\rho)$$

where "perceived creditworthiness" is:

$$\mu_t^{\theta}(\rho) = \int_{\rho}^{+\infty} h_t^{\theta}(\omega) \, d\omega$$

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Households

The representative household solves:

$$\max_{D_{s+1}(\rho)} \mathbb{E}_t^{\theta} \left[\sum_{s=t}^{+\infty} \beta^{s-t} c_s \right]$$

with budget constraint

$$c_{s} + \int_{\mathbb{R}} D_{s+1}(\rho) f(\rho) d\rho = w + \int_{\mathbb{R}} I(\rho, \omega_{s}) \cdot [r_{s}(\rho) D_{s}(\rho) + \pi_{s}(\rho)] f(\rho) d\rho$$

We assume the endowment is large enough to obtain interior solutions despite risk neutrality, w ≥ (αβ)¹/_{1-α}

Equilibrium

 Firms invest until MPK in case of success equals contract interest rate

$$k_{t+1}(\rho) = \left[\frac{\alpha}{r_{t+1}(\rho)}\right]^{\frac{1}{1-\rho}}$$

 Households buy debt until expected return equals inverse discount factor

$$r_{t+1}(\rho) \cdot \mu_t^{\theta}(\rho) = \frac{1}{\beta} \quad \leftrightarrow \quad r_{t+1}(\rho) = \frac{1}{\beta \mu_t^{\theta}(\rho)}$$

 \blacktriangleright Equilibrium spread between risky firm ρ and safe firm $\rho \rightarrow -\infty$

$$S\left(\rho, \mathbb{E}_{t}^{\theta}(\omega_{t+1})\right) \equiv r_{t+1}(\rho) - \frac{1}{\beta} = \frac{1}{\beta} \left(\frac{1}{\mu_{t}^{\theta}(\rho)} - 1\right)$$

Equilibrium

• Debt issuance/installed capital of firm ρ :

$$k_{t+1}(\rho) = \left[\alpha\beta\mu_t^{\theta}(\rho)\right]^{\frac{1}{1-\alpha}} = \left[\frac{\alpha}{1/\beta + S\left(\rho, \mathbb{E}_t^{\theta}(\omega_{t+1})\right)}\right]^{\frac{1}{1-\alpha}}$$

Total debt issued and investment (full depreciation):

$$\mathcal{K}_{t+1} = \int_{\mathbb{R}} \left[lpha eta \mu_t^{ heta}(
ho)
ight]^{rac{1}{1-lpha}} f(
ho) d
ho$$

• Future output in state ω_{t+1} :

$$Y_{t+1}(\omega_{t+1}) = \int_{-\infty}^{\omega_{t+1}} \left[\alpha \beta \mu_t^{\theta}(\rho) \right]^{\frac{1}{1-\alpha}} f(\rho) d\rho$$

Spreads and Issuance

Define average spread at t (inverse measure of optimism)

$$S_t = \int_{\mathbb{R}} S\left(
ho, \mathbb{E}^{ heta}_t(\omega_{t+1})
ight) f\left(
ho
ight) d
ho$$

• **Proposition.** Higher S_t (lower optimism at t) causes:

- disproportionate rise in spread of riskier firms: $\frac{\partial^2 S}{\partial S_t \partial \rho} > 0$
- disproportionate decline in debt issuance and investment by riskier firms: $\frac{\partial}{\partial S_t} \frac{k_{t+1}(\rho_1)}{k_{t+1}(\rho_2)} < 0$ for $\rho_1 > \rho_2$
- holds for $\theta \ge 0$

Accordingly, GH (2013) show junk share rises as spreads fall

Dynamics of Credit Spreads

• Linearise model for $\mathbb{E}^{\theta}_{t}(\omega_{t+1})$ near long-term mean $\overline{\omega} = 0$

$$S_t = \sigma_0 - \sigma_1 \mathbb{E}_t^{\theta} \left(\omega_{t+1} \right)$$

• **Proposition.** Average spread *S*_t follows process:

$$S_t = (1-b)\sigma_0 + b \cdot S_{t-1} - (1+ heta)b\sigma_1\epsilon_t + heta b^2\sigma_1\epsilon_{t-1}$$

- for θ = 0 (rational expectations), spreads follow AR(1), just like fundamentals
- for $\theta > 0$, spreads instead follow ARMA(1,1)

Dynamics of Credit Spreads

Spreads follow ARMA(1,1):

$$S_t = (1-b)\sigma_0 + b \cdot S_{t-1} - (1+ heta)b\sigma_1\epsilon_t + heta b^2\sigma_1\epsilon_{t-1}$$

As S_t is function of expectations at t:

- ▶ has autoregressive component, $S_t \sim b \cdot S_{t-1}$
- but S_{t-1} overreact to news at t-1
- t-1 overreaction subsides at t, does not contaminate S_t
- add correction term $\theta b^2 \sigma_1 \epsilon_{t-1}$ (moving average component)

Credit Spreads Forecasts

To compute forecasts of spreads, note:

$$\mathbb{E}_{t}^{\theta}\left(\mathbb{E}_{t+s}^{\theta}\left(\omega_{t+\tau}\right)\right) = \mathbb{E}_{t}^{\theta}\left(\omega_{t+\tau}\right)$$

diagnostic expectations satisfy law of iterated expectations

Revisions of expectations are unpredictable to investors.
 Forecasts of credit spreads then follow:

$$\mathbb{E}_{t}^{\theta}\left(S_{t+T}\right) = \sigma_{0}\left(1 - b^{T}\right) + b^{T}S_{t}$$

- Actual spreads follow ARMA(1,1) but forecasts follow AR(1)
 - introduces systematic errors, that can account for our motivating evidence

Credit Spreads Forecasts

- **Proposition.** Conditional on information at *t*:
 - forecast error at t + 1 is predictable:

$$\mathbb{E}_{t}\left[S_{t+1} - \mathbb{E}_{t}^{\theta}\left(S_{t+1}\right)\right] = \theta b^{2} \sigma_{1} \epsilon_{t}$$

revision of forecasts are predictable:

$$\mathbb{E}_{t}\left[\mathbb{E}_{t+s}^{\theta}\left(S_{t+T}\right) - \mathbb{E}_{t}^{\theta}\left(S_{t+T}\right)\right] = \theta b^{T+1} \sigma_{1} \epsilon_{t}$$

- Good news predict that S_t and E^θ_t(S_{t+1}) are too low, and that future spreads are revised upwards
- Consistent with our evidence:
 - ▶ negative correlation between S_t and error $S_{t+1} \mathbb{E}_t^{\theta}(S_{t+1})$
 - ► also between S_t and revision $\mathbb{E}^{\theta}_{t+s}(S_{t+T}) \mathbb{E}^{\theta}_t(S_{t+T})$

Predictable Returns and Excess Volatility

▶ Corollary. Let $S_t^r \equiv S_t|_{\theta=0}$ be rational spread. For $\theta > 0$:

avg returns are predictably low (high) on good (bad) news

$$S_t - S_t^r = -\theta b \sigma_1 \epsilon_t$$

spreads exhibit excess volatility

$$Var_{t-1}[S_t] = (1+\theta)^2 Var_{t-1}[S_t^r]$$

- Consistent with evidence
 - high junk shares predict low (even negative) returns (GH 2013)
 - fundamentals account for small share of volatility (Collin-Dufresne et al 2001)

Non-fundamental boom-bust cycles

Corollary. Let $\theta > 0$ and S_{t-1} be low due to good news $\epsilon_{t-1} > 0$. Then, controlling for fundamentals at t - 1:

- spreads predictably rise at t
- aggregate investment at t, and aggregate production at t + 1, predictably drop
- consistent with Lopez-Salido, Stein, Zakrajsek (2015)

Summary

- We present a psychologically founded, forward looking model of expectation formation.
 - in a simple macro model, it reproduces several features of credit cycles.
 - also reproduces facts about spreads forecasts
- Features in common with RE model:
 - spread compression in good times
 - relatively high issuance by high yield firms
- ► Features arising from diagnostic expectations (θ > 0)
 - extrapolative expectations of fundamentals
 - excess volatility of credit spreads
 - predictably low returns in good times
 - systematic non-fundamental reversals (bad credit shocks)

Future Avenues

Expectations

- consider a variety of time series and their expectations
- understand sources of under- and over-reaction
- Financial Frictions
 - Special role of debt / risk misallocation (GSV 2012, 2013)

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Asymmetric effect of busts