

Liquidity premium in CDS markets

Kuate Kanga and Christian Wilde*

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Abstract

We develop a state-space model to decompose bid and ask quotes of CDS into two components, fair default premium and liquidity premium. This approach gives a better estimate of the default premium than mid quotes, and it allows to disentangle and compare the liquidity premium earned by the protection buyer and the protection seller. In contrast to other studies, our model is structurally much simpler, while it also allows for correlation between liquidity and default premia, as supported by empirical evidence. The model is implemented and applied to a large data set of 118 CDS for a period ranging from 2004 to 2010. The model-generated output variables are analyzed in a difference-in-difference framework to determine how the default premium as well as the liquidity premium of protection buyers and sellers evolved during different periods of the financial crisis and to which extent they differ for financial institutions compared to non-financials.

JEL-Classification: C22, G12

Keywords: CDS, liquidity

*Correspondence: House of Finance, Goethe University Frankfurt, Grüneburgplatz 1 (PF H29), 60323 Frankfurt am Main, Germany, Phone: +49 69 798 33699, E-mails: Christel.KuateKanga@hof.uni-frankfurt.de, wilde@finance.uni-frankfurt.de.

1 Introduction

The recent financial turmoil emphasizes the complexity and significance of credit and liquidity risk, and it also reveals the strong but underestimated interlinkage of both types of risks. It furthermore sheds light on the sketchy credit risk and liquidity risk models within the arena of financial risk management for both practitioners and researchers. The aim of our study is to understand the relation between credit risk and liquidity risk in more detail, to decompose the CDS premium into a fair CDS premium and a liquidity premium.

The illiquidity in the credit market has been examined in numerous papers mostly on pre-crisis data. Longstaff, Mithal, and Neis (2005) find that the non-default component of the corporate bond spread can be mainly explained by illiquidity. They derive the default component by using the CDS mid premium. However, they assume that the CDS mid price is a pure measure of default risk and is perfectly liquid. Subsequent studies by Tang and Yan (2007), Bongaerts, de Jong, and Driessen (2011), Chen, Fabozzi, and Sverdlove (2010), Bühler and Trapp (2010) show that CDS mid quotes are actually not the pure default premium, as they contain some liquidity risk premium. Tang and Yan (2007) provide evidence of significant liquidity effects on CDS spreads. They find that the liquidity premium seems to be captured by the protection seller, and that it represents 11% of the mid quote.

Following Tang and Yan (2007), Bongaerts, de Jong, and Driessen (2011) extend the liquidity-adjusted CAPM of Acharya and Pedersen (2005) in order to investigate the liquidity risk in CDS markets. They use the bid-ask spread as a measure of liquidity and show that the liquidity premium is a statistically and economically significant component of the CDS spreads and that the protection seller gets the liquidity premium. However, both papers cannot estimate accurately the level of the risk premium in the mid price and do not study the evolution of the liquidity premium. Chen, Fabozzi, and Sverdlove (2010) and Bühler and Trapp (2010) use the liquidity intensity in a reduced-form model to price CDS contracts. This framework allows a better description of the liquidity premium dynamics, a better estimation of the pure default CDS premium, and therefore allows to investigate the relation between the credit risk and the liquidity risk in CDS contracts. Chen, Fabozzi, and Sverdlove (2010) and Bühler and Trapp (2010) estimate the liquidity risk premium in CDS markets by using bid and ask quotes. Chen, Fabozzi, and Sverdlove

(2010) assume that the bid-ask spread is the upper bound for the liquidity premium. They set the ask premium to be equal to the fair CDS premium and assume that the illiquidity discount factor will reduce the CDS price to the bid price. They therefore exogenously give the liquidity premium to the protection buyer. Most papers above only assume that the liquidity premium and the credit risk are uncorrelated, which has turned out during the financial crisis to be a very unrealistic assumption. Moreover, they focus on American corporate data. Bühler and Trapp (2010) are the first to include the correlation between both risk factors. They extend the model of Longstaff, Mithal, and Neis (2005) by adding liquidity discount factors into the CDS pricing formula. Unlike Chen, Fabozzi, and Sverdløve (2010), they focus on European corporate data and do not assume *prima facie* that the protection buyer captures the liquidity premium. Actually they even find that the liquidity premium is earned by the protection seller and represents 5% of the mid price. Unfortunately, the parameter estimates are not shown, so that the performance of the model and the stability of the estimates are unclear. These are important issues, since the model relies on many parameters (12 for the CDS and 18 in total). Furthermore, they find that the premium stemming from the interlinkage of illiquidity and default risk was insignificant. However, since the beginning of the subprime crisis, the comovement of liquidity risk and default risk becomes obvious; as the CDS premium rises, the bid-ask spread widens. This comovement indicates a strong dependence of liquidity risk on credit risk.

So far, there is mixed evidence on how the liquidity premium is allocated between buyers and sellers. Our main objective is thus, by considering the correlation of credit risk and liquidity risk, in a simple and robust framework to disentangle the CDS premium into the fair CDS premium and liquidity premium and to investigate the evolution of the estimated components over time, in particular during the financial crisis. The premia are investigated and compared for both financial and non-financial institutions during different periods of the financial crisis.

We use CDS bid and ask quotes as lower and upper bounds for the default CDS premium and the bid-ask spread as sum of liquidity premium of the protection buyer and the protection seller. Elaborated upon the approach of Hasbrouck (1999) and Fulop and Lescourret (2009), our model is a state space model with log ask price and bid-ask spread of the log-prices as observed processes. We furthermore introduce the ask liquidity

proportion defined as the ratio of the ask liquidity premium to the bid-ask spread. Thus instead of modeling two liquidity premia as in the aforementioned literature, we just have to use one variable. The model is applied on a dataset that covers a period from January 2004 to September 2010 and contains 118 names of companies belonging to the iTraxx Europe.

We consider five periods, each corresponding to different market conditions. The first period, labeled pre-crisis phase, ranges from January 2004 to July 2007 and covers a time of booming and liquid CDS market. We find that the model-implied fair CDS premium is on average 34.37, while the average mid quote is 33.57. The protection seller charges a liquidity premium of about 25.7% of the bid-ask spread, and the remainder goes to the protection buyer. The financial sector exhibits a lower proportional ask liquidity premium than the non-financial sector. The second period, named subprime phase, covers the subprime credit crisis and ranges from August 2007 to August 2008. During this period, both CDS prices and bid-ask spreads peak on March 2008 at the collapse of Bear Sterns. The model-implied CDS premium on average doubles to 70.119. The average ask liquidity proportion of protection sellers declines to 23.5% of the bid-ask spread (but with a larger volatility). The following phase, the systemic phase, ranging from September 2008 to March 2009, covers the bankruptcy of Lehman Brothers and the bailout of Fannie Mae, Freddie Mac as well as AIG. The deteriorating credit and liquidity conditions lead to a jump of the model-implied CDS premium to 178.680 and the protection seller now requires 41.7% of the bid-ask spread. The fourth phase, the recovery phase, ranging from April 2009 to December 2009, is characterized by uncommon measures implemented by governments and central banks aimed at lowering credit risk and providing liquidity to the financial system. Consequently, the model-implied CDS premium drops from 185.717 in April 2009 to 72.620 by the end of December 2009 with an average of 103.623 while the protection seller charges 32.10% of the lower bid-ask spread. After this slight recovery of the CDS market, the worsening sovereign risk spills over to corporates and pushes up the model-implied CDS premium again from 72.620 to 113.282 during the fifth period, the sovereign phase, ranging from January 2010 to September 2010. In this phase, the protection seller requires a larger portion of the bid-ask spread (43.8%).

We find that as the market becomes less liquid (widening liquidity risk), the protection seller and the protection buyer charge a larger liquidity premium. However, relative to

the bid side, the protection seller acts as liquidity regulator by lowering his liquidity premium in periods of decreasing default risk and augmenting his liquidity requirement in periods of rising default risk. We furthermore show that, during the first three periods, the pre-crisis, subprime, and systemic phases, the fair CDS premia of financial institutions are on average lower than non-financial sector's fair CDS premia. However, afterwards, during the recovery and sovereign phase, the higher sovereign risk spills over to financial institutions and pushes the financial sector's fair CDS premia above the non-financial fair CDS premia. The liquidity premia are on average lower for financial institutions than for non-financial institutions.

We first contribute to the existing literature on liquidity risk in CDS markets by analyzing data before and during the financial turbulence, and we can therefore thoroughly understand the liquidity requirement behavior for both counterparties in a CDS contract. Unlike the recent literature, our model provides a strong and positive interlinkage of the default and the liquidity premia in the CDS market, which strengthens further after the collapse of Lehman Brothers in the systemic phase. Moreover, the major portion of the liquidity premium is allocated to the protection buyer.

The remainder of this paper is structured as follows. Section 2 presents the model and filtration procedure. Section 3 describes the CDS and market data. A preliminary analysis of the comovement of bid-ask spreads and CDS mid quotes is also carried out. In Section 4, we run the filtration and interpret the results. We also investigate the evolution of liquidity risk premia in different periods. In Section 5, we summarize and conclude.

2 State Space Model

2.1 Model

A credit default swap (CDS) is defined as an insurance contract between two counterparties for a protection against default of a name or reference entity. The protection buyer agrees to periodically (quarterly or semi-annually) pay the protection seller a premium up to maturity if the reference entity does not go bankrupt. This side of the bilateral contract is called premium or fixed leg. In case the reference firm defaults, the protection seller compensates the protection buyer for the loss incurred on the underlying asset (mostly bonds). This side of the contract is called protection or default leg. We name

the premium paid by the protection buyer a fair CDS premium and it is computed by matching the expected premium leg to the expected protection leg. Thus, a fair CDS premium mainly takes the following risks into account: the credit default risk of the reference name, the counterparty credit risk of the protection seller (risk that protection seller may not be able to satisfy its obligations) and the recovery rate of the underlying asset. Arora, Gandhi, and Longstaff (2012) empirically show that counterparty credit risk is priced in the CDS market, but its effect on CDS premia is economically insignificant. Concerning recovery risk, some papers such as Bakshi, Madan, and Zhang (2006), Das and Hanouna (2009), Schlaefler and Uhrig-Homburg (2010), Schneider, Soegner, and Veza (2010), Conrad, Dittmar, and Hameed (2011), and others have elaborated methods to estimate the recovery rate implied in CDS premia. Nevertheless, recovery risk in the CDS market is neglected in most studies, since it is difficult to identify the recovery rate from the default rate. Rather, the recovery rate is generally assumed to be constant and fixed at 40, 50 or 60%. The main driver of fair CDS premia and actually the most interest arousing risk in the literature and for practitioners is the credit default risk of reference names. The fair CDS premium can therefore be loosely called the CDS default premium.

A CDS is a bilateral contract traded in the OTC market. Trades in CDS markets face therefore some impediments besides the default risk such as information asymmetries, transaction costs, searching costs, funding costs, etc. We call these obstacles liquidity risk and assume that it affects both the protection seller and the protection buyer. However, we account for the fact that they may be affected in an uneven way by liquidity risk and suppose that the protection seller and the protection buyer of a CDS contract set respectively the ask price S_t^{ask} and the bid price S_t^{bid} such that they can cover all the cost of trading. Therefore, S_t^{ask} and S_t^{bid} are upper and lower bounds for our model-implied CDS premium S_t^{def} . We assume that the model-implied CDS premium is the best approximation of a fair CDS premium. The bid-ask spread, $BA_t = S_t^{ask} - S_t^{bid}$, therefore represents the total liquidity premium that the protection buyer and seller charge. Nevertheless, it does not tell about the liquidity premium required by each size of a trade. For instance, a widening bid-ask spread is possible in three scenarios: either the liquidity premium of the ask side narrows while the bidder requires a larger liquidity premium or the liquidity premium of the ask side enlarges while the bidder requires a lower liquidity premium or the liquidity premium of both sides increases. In order to better model the asymmetry in

the liquidity premium, we define from the perspective of the protection seller a liquidity proportion labeled R_t ,

$$R_t = \frac{S_t^{ask} - S_t^{def}}{S_t^{ask} - S_t^{bid}} = \frac{S_t^{ask} - S_t^{def}}{BA_t}, \quad (1)$$

$0 \leq R_t \leq 1$ represents the liquidity premium (in relation to the bid-ask spread) that the protection seller requires to sell a CDS contract. This measure tells more about the trading behavior in CDS markets and therefore completes the information emanating from the bid-ask spread. R_t is also a better measure of liquidity because it is a relative measure bounded on the interval $[0, 1]$ and can therefore be suitable for analyzing CDS involving different ratings, periods of time, and industry sectors. It is clear that a rise of the bid-ask spread indicates an evaporation of liquidity in CDS markets while an increase of R_t denotes an increasing ask liquidity premium relative to the protection buyer. For $R_t = \frac{1}{2}$, the mid CDS premium is equal to the default premium and the bid liquidity premium is equal to the ask liquidity premium (as assumed in Longstaff, Mithal, and Neis (2005)). We offer a test of this. For $R_t \geq \frac{1}{2}$, the protection seller requires a larger liquidity premium than the buyer. This can be motivated by the fact that he is facing a higher cost for selling protection, or as liquidity provider in CDS markets (see Tang and Yan (2007)), he needs to charge a larger liquidity premium to lower the demand to the supply of CDS. In case of higher demand for CDS contracts, according to Amihud and Mendelson (1988), the market maker wants to keep his book clear and requires therefore a higher liquidity premium for selling protection (sell expensive) and a lower liquidity premium for buying protection (buy expensive). The choice of explicitly modeling R allows to test whether larger default risk is associated with higher R , potentially due to higher insurance demand. In order to explicitly model the liquidity premium of each counterparty in a CDS contract, we use an econometric model similar to Hasbrouck (1999) and Fulop and Lescouret (2009). From Equation 1, we obtain

$$S_t^{ask} = S_t^{def} + R_t BA_t = e^{s_t^{def}} + R_t BA_t, \quad (2)$$

$$\begin{aligned} S_t^{bid} &= S_t^{def} - (1 - R_t) BA_t \\ &= e^{s_t^{def}} - (1 - R_t) BA_t. \end{aligned} \quad (3)$$

We denote by s_t^{ask} , s_t^{bid} the logarithmized ask and bid quotes and define the correspond-

ing bid-ask spread by $ba_t = s_t^{ask} - s_t^{bid}$. The corresponding liquidity proportion of the protection seller is given as $0 \leq r_t = \frac{s_t^{ask} - s_t^{def}}{s_t^{ask} - s_t^{bid}} \leq 1$. We further assume that there exists some liquidity discount factors L_t^{ask} and L_t^{bid} satisfying the relations

$$S_t^{ask} = S_t^{def} L_t^{ask} \quad (4)$$

$$S_t^{bid} = S_t^{def} L_t^{bid}, \quad (5)$$

with $1 \leq L_t^{ask}$ and $0 \leq L_t^{bid} \leq 1$, such that the bid price is always lower than the ask price. It follows from Equation 4 that

$$\log L_t^{ask} = \log S_t^{ask} - \log S_t^{def} \quad (6)$$

$$\begin{aligned} &= \left(\log S_t^{ask} - \log S_t^{def} \right) \frac{\log S_t^{ask} - \log S_t^{bid}}{\log S_t^{ask} - \log S_t^{bid}} \\ &= \frac{\log S_t^{ask} - \log S_t^{def}}{\log S_t^{ask} - \log S_t^{bid}} \left(\log S_t^{ask} - \log S_t^{bid} \right) \\ &= r_t \left(\log S_t^{ask} - \log S_t^{bid} \right). \end{aligned} \quad (7)$$

And analogously we get from Equation 5

$$\log L_t^{bid} = (r_t - 1) \left(\log S_t^{ask} - \log S_t^{bid} \right). \quad (8)$$

We can write from Equations 4 and 7

$$\begin{aligned} \log S_t^{ask} &= \log S_t^{def} + \log L_t^{ask} \\ &= \log S_t^{def} + \frac{\log S_t^{ask} - \log S_t^{def}}{\log S_t^{ask} - \log S_t^{bid}} \left(\log S_t^{ask} - \log S_t^{bid} \right), \\ s_t^{ask} &= s_t^{def} + r_t ba_t. \end{aligned} \quad (9)$$

From Equations 5 and 8 follows the relation

$$s_t^{bid} = s_t^{def} - (1 - r_t) ba_t. \quad (10)$$

In contrast to the approach of Hasbrouck (1999) and Fulop and Lescouret (2009), two state variables, s_t^{def} and r_t , are sufficient to describe the fair default process and the liquidity processes of the ask and bid prices. The log-default premium satisfies the relation

$s_{t+1}^{def} = s_t^{def} + \eta_{t+1}$, where η_{t+1} is a normal random noise with variance σ_η^2 . Longstaff, Mithal, and Neis (2005) show that the non-default component of corporate bond yields is strongly mean reverting and Chen, Fabozzi, and Sverdllove (2010) find a mean reverting behavior of the liquidity rate in corporate CDS. Based on these results, we assume that the liquidity process r_t is a mean-reverting process bounded on $[0, 1]$. Following van Emmerich (2006), the liquidity process r_t has the dynamic $dr_t = \kappa(\theta - r_t)dt + \sigma\sqrt{(1-r_t)r_t}dW_t$. The bounds of the stochastic process are unattractive and unattainable if $2\kappa \geq \left(\frac{\sigma^2}{1-\theta}, \frac{1}{\theta}\right)$. To obtain a model that is linear in the error term, we use the first-order approximation. However, different discrete time processes are obtained depending on the discretization method applied. For example, the explicit Euler discretization gives

$$r_t = \kappa\theta + (1 - \kappa)r_{t-1} + \sqrt{(1 - r_{t-1})r_{t-1}}\epsilon_t, \quad \epsilon_t \sim N(0, \sigma_\epsilon), \quad (11)$$

whereas the implicit Euler discretization yields

$$r_t = \frac{\kappa\theta}{1 + \kappa} + \frac{r_{t-1}}{1 + \kappa} + \frac{1}{1 + \kappa}\sqrt{(1 - r_{t-1})r_{t-1}}\epsilon_t. \quad (12)$$

For our calibration, we choose a general mean reverting process bounded between 0 and 1 given by the dynamic

$$r_t = \alpha + \beta r_{t-1} + \sqrt{(1 - r_{t-1})r_{t-1}}\epsilon_t, \quad \text{with } 0 \leq \alpha \leq 1, -1 \leq \beta \leq 1.$$

Since the log fair CDS premium is a random walk, by means of the first difference of the log ask price

$$\begin{aligned} s_t^{ask} - s_{t-1}^{ask} &= s_t^{def} - s_{t-1}^{def} + r_t b a_t - r_{t-1} b a_{t-1}, \\ y_t &= \eta_t + r_t b a_t - r_{t-1} b a_{t-1}. \end{aligned} \quad (13)$$

The change in ask premium y_t given in Equation 13 represents the observation equation, and the state-variable process is driven by r_t . Since r_t is non-linear in r_{t-1} and ϵ , by means of the Taylor approximation around \hat{r} and 0, we linearize the evolution of the

relative log-liquidity premium

$$r_t = f(r_{t-1}, \epsilon_t) \quad (14)$$

$$\begin{aligned} &\approx f(\hat{r}, 0) + \left. \frac{\partial f}{\partial r} \right|_{\hat{r}, 0} (r_t - \hat{r}) + \left. \frac{\partial f}{\partial \epsilon} \right|_{\hat{r}, 0} (\epsilon_t - 0) \\ &= \alpha + \beta r_{t-1} + \sqrt{(1 - \hat{r})} \hat{r} \epsilon_t \\ &= \alpha + \beta r_{t-1} + \sqrt{(1 - r_{t-1|t-1})} r_{t-1|t-1} \epsilon_t \\ &= \alpha + \beta r_{t-1} + \sqrt{(1 - \hat{r})} \hat{r} \epsilon_t, \quad \text{with} \quad \hat{r} = r_{t-1|t-1} \\ &= \alpha + \beta r_{t-1} + \hat{L}_{t-1} \epsilon_t, \quad \text{with} \quad \hat{L}_{t-1} = \sqrt{(1 - \hat{r})} \hat{r}. \end{aligned} \quad (15)$$

2.2 Kalman Filtering

Given that we avoid a strongly non-linear model as in Fulop and Lescourret (2009), the Kalman filter is used to estimate the unobserved liquidity process r_t . The Kalman filter is a robust and recursive approach to estimate non-observable state variables in linear dynamics systems. It provides a minimum mean squared error of r_t (see Simon (2006) and Kim and Nelson (1999)). We rewrite the observation equation 13 as

$$\begin{aligned} y_t &= r_t b a_t - r_{t-1} b a_{t-1} + \eta_t \\ &= \underbrace{\begin{bmatrix} b a_t & -b a_{t-1} \end{bmatrix}}_{H_t} \underbrace{\begin{bmatrix} r_t \\ r_{t-1} \end{bmatrix}}_{x_t} + \eta_t \\ &= H_t x_t + \eta_t. \end{aligned} \quad (16)$$

The observation or measurement equation gives the relation between the observed process y_t and the state variables x_t . The dynamics of x_t is derived from Equation 15 and given by

$$\begin{aligned} r_t &= \alpha + \beta r_{t-1} + \hat{L}_{t-1} \epsilon_t \\ \underbrace{\begin{bmatrix} r_t \\ r_{t-1} \end{bmatrix}}_{x_t} &= \underbrace{\begin{bmatrix} \alpha \\ 0 \end{bmatrix}}_A + \underbrace{\begin{bmatrix} \beta & 0 \\ 1 & 0 \end{bmatrix}}_F \underbrace{\begin{bmatrix} r_{t-1} \\ r_{t-2} \end{bmatrix}}_{x_{t-1}} + \underbrace{\begin{bmatrix} \hat{L}_{t-1} & 0 \\ 0 & 0 \end{bmatrix}}_{L_{t-1}} \underbrace{\begin{bmatrix} \epsilon_t \\ 0 \end{bmatrix}}_{w_t} \end{aligned}$$

Hence, the transition equation is written as

$$x_t = A + Fx_{t-1} + L_{t-1}w_t. \quad (17)$$

Our state space model can then be summarized as follows

$$\begin{aligned} y_t &= H_t x_t + \eta_t, \\ x_t &= A + Fx_{t-1} + L_{t-1}w_t. \end{aligned}$$

The measurement noise process η_t and the state variable noise process w_t are white, have zero-mean and are correlated with covariance matrices respectively given by σ_η^2 and Q . In the model, we allow for correlation between the default premium and liquidity. In particular, we denote by M the correlation matrix between η_t and w_t . It holds that

$$\begin{aligned} \eta_t &\sim N(0, \sigma_\eta^2), \\ w_t &\sim N(0, Q) \quad Q = \begin{bmatrix} \sigma_\epsilon^2 & 0 \\ 0 & 0 \end{bmatrix}, \\ M = Cov[w_t, \eta_t] &= \begin{bmatrix} \rho_{\epsilon, \eta} \sigma_\epsilon \sigma_\eta \\ 0 \end{bmatrix}. \end{aligned}$$

The Kalman filter is a recursive procedure of prediction or forecasting and updating steps. At time $t \in \mathbb{N}$, $x_{t+1|t}$ denotes the predicted state variable and represents the estimate of x_{t+1} given that information or measurements are available up to time t . Thus, it holds that $x_{t+1|t} = E[x_{t+1}|y_1, \dots, y_t]$. At time $t + 1$, the measurement y_{t+1} is available and we update the estimate of the state variable x_{t+1} by taking the new observation into account. We have $x_{t+1|t+1} = E[x_{t+1}|y_1, \dots, y_{t+1}]$

Forecasting

The predicted or a priori estimate of the state and observed variables are given by

$$\begin{aligned} x_{t+1|t} &= A + Fx_{t|t}, \\ y_{t+1|t} &= H_{t+1}x_{t+1|t}. \end{aligned}$$

We denote by $P_{t+1|t}$ the prediction error covariance matrix and it is defined by

$$P_{t+1|t} = E \left[(x_{t+1} - x_{t+1|t}) (x_{t+1} - x_{t+1|t})' \right]. \quad (18)$$

The corresponding updating error covariance matrix is defined by

$$P_{t|t} = E \left[(x_t - x_{t|t}) (x_t - x_{t|t})' \right]. \quad (19)$$

It follows from Equations 18 and 19, the following relation

$$\begin{aligned} P_{t+1|t} &= E \left[(x_{t+1} - x_{t+1|t}) (x_{t+1} - x_{t+1|t})' \right] \\ &= E \left[(Fx_t + L_t w_{t+1} - Fx_{t|t}) (Fx_t + L_t w_{t+1} - Fx_{t|t})' \right] \\ &= FE \left[(x_t - x_{t|t}) (x_t - x_{t|t})' \right] F' + L_t E \left[(w_{t+1}) (w_{t+1})' \right] L_t' \\ &= FP_{t|t}F' + L_tQL_t'. \end{aligned}$$

Using again Equations 18 and 19, the observation prediction error covariance V_{t+1} of the observed process y_t is derived:

$$\begin{aligned} V_{t+1} &= E \left[(y_{t+1} - y_{t+1|t})^2 \right] \\ &= E \left[(H_{t+1}x_{t+1} + \eta_{t+1} - H_{t+1}x_{t+1|t})^2 \right] \\ &= E \left[(H_{t+1}(x_{t+1} - x_{t+1|t}) + \eta_{t+1})^2 \right] \\ &= H_{t+1}E \left[(x_{t+1} - x_{t+1|t}) (x_{t+1} - x_{t+1|t})' \right] H_{t+1}' + E [\eta_{t+1}^2] + E [2H_{t+1}(x_{t+1} - x_{t+1|t}) \eta_{t+1}] \\ &= H_{t+1}E \left[(x_{t+1} - x_{t+1|t}) (x_{t+1} - x_{t+1|t})' \right] H_{t+1}' + E [\eta_{t+1}^2] + E [2H_{t+1}L_t w_{t+1} \eta_{t+1}] \\ &= H_{t+1}P_{t+1|t}H_{t+1}' + R + 2H_{t+1}L_tM, \quad R = \sigma_\eta^2. \end{aligned}$$

After predicting the state variables, the second step in Kalman procedure is to update the estimates given new observation.

Updating

$$\begin{aligned}
K_{t+1} &= (P_{t+1|t}H'_{t+1} + L_tM) (H_{t+1}P_{t+1|t}H'_{t+1} + R + H_{t+1}L_tM + L'_tM'H'_{t+1})^{-1} \\
&= (P_{t+1|t}H'_{t+1} + L_tM) (H_{t+1}P_{t+1|t}H'_{t+1} + R + 2H_{t+1}L_tM)^{-1}, \\
x_{t+1|t+1} &= x_{t+1|t} + K_{t+1} (y_{t+1} - y_{t+1|t}), \\
P_{t+1|t+1} &= P_{t+1|t} - K_{t+1} (H_{t+1}P_{t+1|t} + (L_tM)').
\end{aligned}$$

For a thorough reading on filtering methods refer to Simon (2006). Before filtering out the state variable $x_{t|t}$, the parameters A , F , Q , R , M and the start values $x_{0|0}$, $P_{0|0}$ should be known. We therefore estimate the parameters by means of the maximum likelihood method. The multivariate distribution of the observation can be written as

$$\begin{aligned}
f(y_1, \dots, y_T) &= f(y_1) f(y_2|y_1) f(y_3|y_1, y_2) \dots \\
&f(y_{T-1}|y_1, y_2, \dots, y_{T-2}) f(y_T|y_1, y_2, \dots, y_{T-2}, y_{T-1}).
\end{aligned}$$

Given that η_t and w_t are normally distributed, y_{t+1} conditional on y_1, y_2, \dots, y_t is normally distributed with mean $y_{t+1|t}$ and variance V_{t+1} . Hence,

$$f(y_{t+1}|y_1, y_2, \dots, y_{t-1}, y_t) = \frac{1}{\sqrt{(2\pi)}\sqrt{V_{t+1}}} \exp\left[-\frac{1}{2} \frac{(y_{t+1} - y_{t+1|t})^2}{V_{t+1}}\right].$$

The log-likelihood function $L = \log(f(y_1, \dots, y_T))$ is written as

$$L = -T \frac{\log(2\pi)}{2} - \frac{1}{2} \sum_{t=1}^T \log V_{t+1} + \frac{(y_{t+1} - y_{t+1|t})^2}{V_{t+1}}, \quad (20)$$

and is maximized according to the parameters $\Theta = [\sigma_\eta \quad \alpha \quad \beta \quad \sigma_\epsilon \quad \rho \quad r_{0|0} \quad P_{0|0}]$. In order to better approximate the global maximum, we start the filtering by choosing $N = 200$ initial parameter vectors $\Theta_0^1, \dots, \Theta_0^N$. The parameters are estimated with a three-step procedure:

1. For each initial parameter set Θ_0^k , $k = 1, \dots, N$, we use the Kalman filter to generate a time series of both $y_{t+1|t}$ and $x_{t|t}$, $t = 1, \dots, T$, and find optimal parameters $\hat{\Theta}^k$, $k = 1, \dots, N$, that maximize L for the observed y_t .
2. Thereafter, the parameter set $\hat{\Theta}$ is chosen from $\{\hat{\Theta}^1, \dots, \hat{\Theta}^N\}$ as the parameter set

that provides the maximum value for L .

3. We reinitialize the procedure by setting $\Theta_0^1 = \hat{\Theta}$ and generate $N - 1$ parameter sets $\Theta_0^2, \dots, \Theta_0^N$ and return to the first step. As soon as the parameters converge, we stop the procedure.

3 Data

3.1 CDS Data

We use a dataset of weekly corporate CDS bid and ask prices from Datastream. The dataset covers the period January 2004 - September 2010 and contains 118 5-years CDS names from the iTraxx Europe. Our sample therefore consists of frequently traded CDS contracts. It covers several groups: automobile and industry (29 names), consumer and service (26), financials (23), energy (20), telecommunications (20). Summary statistics of the data are given in Figure 1 and Table 1. Time-series of CDS mid quotes, bid-ask spreads, and relative bid-ask spreads are depicted in Figure 1, and the corresponding numbers are reported in Table 1. We divide the sample in five periods, representing different phases of the financial crisis and different market conditions. The five phases are the pre-crisis phase, subprime phase, systemic phase, recovery phase, and sovereign phase.

The pre-crisis phase begins in January 2004 and ends in July 2007. As Figure 1 illustrates, this period is characterized by a steady decline in CDS prices and a lower volatility (see also Table 1, second column). The International Swap and Derivative Association (ISDA) market survey¹ reports a strong and sustained growth in CDS markets during this first phase. The outstanding amount of CDS rises from 5.44 trillion in the first half of 2004 to \$45.46 trillion in first half of 2007. The higher demand and supply of protection lead to very liquid CDS markets and lower bid-ask spreads with an average of 3.556. Table 1 also suggests that the absolute bid-ask spread is positively correlated with the CDS mid premium (0.590), whereas the relative bid-ask spread moves negatively with the CDS mid price (-0.347). This highlights that there is a relation between the CDS premium and the liquidity premium, and it demonstrates the importance of explicitly

¹www.isda.org/statistics/recent.html

modeling this relation.

This period of flourishing and quiet CDS markets is followed by the financial crisis which covers four periods. The subprime phase ranges from August 2007 to August 2008 and covers the subprime credit crisis. Given that financial markets are highly connected and European financial institutions have large exposures in American mortgage backed securities and derivatives during that period, the subprime crisis spills over from US to European markets. The average CDS price doubles from 33.570 to 68.824 basis points, and its volatility also rises from 31.189 to 41.592. Simultaneously, the bid-ask spread climbs by 46% from 3.556 to 5.207. During this period, the CDS mid premium and the bid-ask spread peak on March 2008 while Bear Stearns collapses. From this period on, the CDS mid premium and the absolute bid-ask spread strongly comove with a correlation larger than 0.69 and which peaks during the systemic phase. The outstanding notional of CDS contracts reaches its highest level in the second half of 2007 (\$62.17 trillion) and declines to \$54 trillion in the first half of 2008. If we assume that the CDS contracts are almost standardized at \$10 million notional values, a larger outstanding amount therefore implies more CDS contracts compared to the previous phase. Although the credit market is going through a turbulent period, enough confident investors or speculators are still capable of providing protection to the increasing demand of insurance against the rising default risk.

The third period, the systemic phase, ranging from September 2008 to March 2009, is a period of severe systemic risk catalyzed by the bankruptcy of Lehman Brothers on September 15, 2008. Stock prices all over the world plummet, the index of global risk aversion measured by Caceres, Guzzo, and Segoviano (2010) jumps by 21%, and the Libor-OIS spread rises significantly. The average CDS price sharply climbs to more than 178, with a dispersion that grows from 41.592 during the subprime phase to 182.795. The bid-ask spread also triples to 14.827. Thus, this implies that both credit quality and liquidity worsen at the same time. The outstanding notional amount sharply declines by about 29% by the end of 2008. The fragile credit market, the high default probability combined with the large risk aversion, and the low funding liquidity available push the protection seller to shy away from risky assets and thus damp down the liquidity of the CDS market. In order to appease the jumpy market, the USA and many European countries intervene by introducing new policy rules such as stimulus packages, by providing guarantees for

risky assets, and by bailing out or nationalizing financial institutions. In September 2008, for example, Fannie Mae and Freddie Mac are placed by the U.S. Treasury into conservatorship of the U.S. government. The Federal Reserve provides AIG with \$85 billion bailout and announces the Troubled Asset Relief Program (TARP) which allows the purchase of illiquid assets. In October 2008, Germany provides 50 billion Euro bailout to Hypo Real Estate and Belgium, France, and Luxemburg provide a 150 billion Euro guarantee of Dexia's debt. These governments' responses to the financial crisis aim at supplying liquidity and at diminishing the systemic risk in the market and thus achieve to lower the CDS prices from April 2009 to December 2009. This latter period is labeled recovery phase. During this period, the price of default protection and its dispersion fall respectively to 102.472 and 77.933 on average. Similarly, the bid-ask spread and its dispersion also slightly decrease. The outstanding notional amount remains almost constant above \$30 trillion. However as Acharya, Drechsler, and Schnabl (2010) point out, these government interventions on the one hand stabilize financial markets, but on the other hand, they allow the transmission of risk from the financial sector to sovereigns.

It follows in several cases the degradation of sovereign creditworthiness and a successive rating downgrade of many European countries such as Greece, Ireland, Italy, Portugal and Spain. Given that a downgrade of a country generally leads to a downgrade of most home companies and given interdependencies and the risk of contagion, the prices of CDS contracts, especially for financial institutions, climb again during the last period (sovereign phase) from January to September 2010. The outstanding notional slightly decreases to about \$27 trillion.

By observing the strong comovement of the bid-ask spread with the mid CDS premium during the financial downturn, we conjecture that market liquidity can be partly explained by default risk. The bid-ask spread is a measure of an asset's liquidity and indicates how easy or at which cost this asset can be traded. As aforementioned, it is one measure of market liquidity. During the financial turmoil, the bid-ask spread increases with the CDS mid premium and peaks in the period following Lehman Brothers' bankruptcy. We therefore investigate the relation of the CDS liquidity premium and the default premium during the five periods proposed. Focusing on the bond market, Dick-Nielsen, Feldhütter, and Lando (2011) note a rise in liquidity premia in investment grade corporate bonds during the crisis. Furthermore, the substantial positive correlation between CDS mid

quotes and the absolute bid-ask spread and the negative correlation between CDS mid quotes and the relative bid-ask spread suggests that the CDS market faces larger cost of trading as the default risk broadens and supports our model choice to include the correlation between the pure default and the liquidity variables. In the next section, we calculate the unobserved CDS premium reflecting default risk, and also the liquidity premium of the protection seller and buyer.

4 Calibration Results

4.1 Parameter Estimates

The calibration is run on the log CDS premia, and the parameter estimates for the entire sample are reported in Table 2. The average volatility of the log CDS premium (state variable s_t), σ_η , exhibits during the first phase the lowest level of 0.08, it rises to 0.13 in the subprime phase and reaches its maximum 0.14 in the systemic phase. During the fourth phase, the corporates recover and the volatility of the log default premium decreases by about 50% to 0.08 as in the first period. During the sovereign phase, because of a high correlation between country risk and corporate default risk as expounded in Section 3, the market becomes more risky and the volatility of the default premium slightly increases to 0.09. The estimated volatilities match quite well the CDS premia and bid-ask-spread evolution of Table 1. Table 2 also shows that the volatility σ_ϵ of the relative log liquidity premium process (liquidity state variable) r_t follows a similar path to the log fair premium volatility. However, in contrast to the default process volatility which slightly increases from 0.13 to 0.14, the liquidity process volatility jumps by more than 60% from 0.66 in the subprime phase to 1.10 in the systemic phase. Afterwards, the volatility of the liquidity process does not recover and even reaches its maximum level 1.29 during the sovereign period. The large volatility of the liquidity process during the systemic, recovery and sovereign phases draws out the general uncertainty in the market since Lehman Brothers' bankruptcy and underlines how important liquidity issues have become after the beginning of the crisis in 2007. It additionally supports the idea that the bankruptcy of Lehman Brothers changes systematically the trading behaviors in CDS markets.

Table 2 furthermore reports a high negative and significant correlation between the default process and liquidity process. According to these results, it seems more realistic

to include a correlation between the default and the liquidity premium when modeling CDS prices. To better interpret the interconnection between default risk and liquidity, we compute the sample correlation between the proportion of the liquidity premium attributable to the ask side (R_t) and the model-implied CDS premium S_t^{def} (see Table 3). The average correlation of the ask liquidity proportion R_t of the protection seller and the model-implied CDS premium S_t^{def} is low during the five phases but changes its sign from slightly negative to slightly positive values $-0.09, -0.07, -0.02, 0.04$ and 0.1 . These low correlations suggest that the ask liquidity proportion weakly depends on the default risk. We notice in Section 3 that the bid-ask spread positively comoves with CDS mid premia, which implies that liquidity decreases as default risk increases. Similarly to bid-ask spread, the correlations between the default premium S_t^{def} and the ask liquidity premium $SL_t^{ask} = S_t^{ask} - S_t^{def}$ and between the default premium and the bid liquidity premium $SL_t^{bid} = S_t^{def} - S_t^{bid}$ are positive during the five phases for most corporates. Thus, the larger the credit risk, the larger is the liquidity premium required by the protection seller and the protection buyer. During the ongoing financial crisis, the default risk goes in line with a worsening liquidity in the CDS market, i.e. a widening bid-ask spread. However, default risk seems to be unrelated to the inequality of the liquidity premium required by the seller and buyer of protection.

Tables 2 displays the estimated parameters separately for financials and non-financials. The numbers show that the default state variable's volatilities σ_η are on average in all phases higher for financial names than for non-financial names, whereas the liquidity variable's volatilities σ_ϵ are on average larger for non-financial companies except during the subprime phase. Moreover, the long run averages of the liquidity variables α are lower for the financial sector and thus signal a lower liquidity premium for the protection sellers on financial entities. It is also noteworthy that during the second phase, both sectors exhibit an almost equal α . The second period corresponds to the beginning of the financial crisis and covers the subprime credit crisis that lays some ambiguity on financial institutions' default probabilities and thus increases the investors' disagreement on banks' real credit worthiness. This could damp down the willingness of financiers to sell protection on financial institutions. There is a very low α for financials during the recovery phase, indicating a low ask liquidity proportion.

4.2 Analysis of Premia

4.2.1 Default premia and absolute liquidity premia

The applied model further yields time-series of default and liquidity premia. The properties of these premia, their comovement in each phase, and their behavior for different sectors (financial and non-financial) can be analyzed. Figure 2 plots the time-series of the model-implied CDS premium S_t^{def} , ask liquidity premium SL_t^{ask} and bid liquidity premium SL_t^{bid} , averaged over names. As the model-implied CDS premium rises, the protection seller and buyer also claim a larger liquidity premium which leads to an enlargement of the bid ask spread. Before August 2007, during the pre-crisis period, the model-implied CDS premium, S_t^{def} , has an average of 34.370 (see Table 4) and the ask and bid liquidity premia are around 0.994 (3.3% of S_t^{def}) and 2.600. The deterioration of credit risk in the market due to the subprime crisis pushes up the average model-implied CDS premium to 70.119 during the second period. The ask and bid liquidity premia also increase respectively to 1.308 and 3.898. During the following period, the systemic phase, the average model-implied CDS premium peaks at 178.680. Given the large turbulence in the market after the bankruptcy of Lehman Brothers, the protection seller more than triples its liquidity premium to 6.951 whereas the protection buyer only doubles its liquidity premium at 7.875. The market recovery in the following phase reduces the average model-implied CDS premium, the average ask and bid liquidity premia respectively to 103.623, 2.772 and 5.074. The last period, the sovereign phase, is characterized by the fact that the model-implied CDS premium on average shrinks to 101.652. However, Table 4 also shows that only the non-financial sector has a lower model-implied CDS premium in the sovereign phase (95.614), whereas the model-implied CDS premium of financial institutions jumps on average to 126.593. This is certainly related to the increasing sovereign risk in several European countries.

Figure 5 and the estimated positive values of skewness in Table 4 reveal that the distributions of the average model CDS premia are right skewed and vary systematically according to the phases. Besides, the model-implied CDS premium has its lowest standard deviation (31.425) before August 2007, during the pre-crisis period. Thereafter, it gradually climbs and peaks in the systemic phase of the crisis (181.198), where it henceforth gradually descends (58.824). Apart from the sovereign period, the standard deviation of the CDS premium is always lower for financials compared to non-financials.

By observing Panel A and B in Figure 2, we see that the bid liquidity premium is on average larger than the ask liquidity premium. Moreover, the comovement of the liquidity premia and the model-implied CDS premium clearly infers that liquidity and credit risks are highly positively correlated, especially since the beginning of the subprime crisis in mid 2007. This contrasts to Chen, Fabozzi, and Sverdlow (2010) who find that the model-implied CDS premium is mostly not related to the liquidity premium. The bid liquidity premium is mostly larger than the ask liquidity premium.

Figure 2, Panel A, and the first column of Table 4 illustrate that non-financial institutions have on average a larger default premium than financial institutions during the first three periods, i.e. pre-crisis phase, subprime phase, and systemic phase. In the recovery phase and especially during the sovereign phase, the model-implied CDS premium of financial institutions rises above the model-implied CDS premium of non-financial institutions (see Table 4 and Panel A in Figure 2). The worsening credit quality in the financial sector is related to increasing sovereign risk which spills over to financial institutions. A BIS study examines the impact of sovereign risk on financial institutions (see Panetta et al. (2011)). This study lists various channels through which sovereign risk may spill over to the financial sector. For instance, banks hold sovereign debt in their portfolio and any downgrade of the countries increases the riskiness of the banks. Moreover, in markets of repurchase agreements, investors commonly use government bonds as collateral in order to get funding. An increase of sovereign risk automatically leads to an augmentation of the repo rates and haircuts. Thus, liquidity dries out in the market. In addition, given the illiquidity of sovereign indices, market participants seem to use financial indices to hedge sovereign risks and therefore strengthen the comovement of sovereign and financial risks. Thus, the increase of the interlinkage of financial and sovereign risks, the growth of funding costs and the deterioration of sovereign portfolios could motivate the increase of the financial model-implied CDS premium over the non-financial CDS fair value during the last phase. This suggests that investors believe the industrial sector to be stabilized but are less confident with financial companies and thus require a larger model-implied CDS premium.

In Panel B of Figure 2, we observe that the ask liquidity premium is on average lower for financial institutions than for non-financial institutions and moreover, the gap between the ask liquidity premium of financials and non-financials rises during the systemic and

recovery phases of the economic downturn. Figure 2, Panel C, indicates that the bid liquidity premium for the financial and non-financial sector seems to hover around the same levels.

4.2.2 Relative bid-ask spread, relative ask liquidity and bid liquidity premia

Figure 3, Panel A, shows the evolution of the relative bid-ask spread $BA_{t,rel}^{def}$ over time, calculated as average over the names. Panel B and C of Figure 3 shows the decomposition of the relative bid-ask spread into ask ($SL_{t,rel}^{ask}$) and bid ($SL_{t,rel}^{bid}$) liquidity components. Overall, the relative bid-ask spread decreases during the crisis. While financials on average exhibit a larger relative bid-ask spread than non-financials during the pre-crisis period, their spreads falls below that of non-financials in later periods. The decomposition shows that during the crisis, the gap between financials and non-financials is larger for the relative ask premium than for the relative bid-premium.

The histograms in Figure 6 and the numbers in Table 4 show that dispersion of the relative bid-ask spread steadily decreases starting from the pre-crisis period (0.107) and reaches a low in the sovereign period (0.023). A steadily decreasing relative bid-ask spread can be observed for both financials and non-financials. The drop is strongest for financials (from 0.116 to 0.014) compared to non-financials (from 0.102 to 0.024).

The histogram in Figure 7 shows that for both financials and non-financials, the distributional properties of the relative bid liquidity premium are very similar to that of the relative bid-ask spread. In all cases, both mean and dispersion decrease over time.

The relative ask liquidity premium, however, differs substantially. The histograms in Figure 8 show that the shape of the distributions varies in each period. The dispersion of financials' relative ask liquidity premium shrinks from the pre-crisis phase to the sovereign phase. In contrast, for non-financials, there is a peak of mean and dispersion in the systemic phase. While both mean and dispersion are larger for financials during the first periods (pre-crisis phase and subprime phase), they are larger for non-financials during the following periods.

4.2.3 Ask liquidity proportion

Table 4, Figure 4 panel A and Figures 9 describe the distributional properties of the ask liquidity proportion R_t defined as the ratio of the ask liquidity premium to the bid-ask

spread. We suppose R_t better expresses the asymmetry between ask and bid liquidity premia. During the pre-crisis period, the average ask liquidity proportion R_t hovers around its mean of 25.7% and ranges between 0.6% and 57.3% (see Table 4). Although the default risk obviously increases during the subprime phase as indicated in the previous Subsection, the mean of R_t diminishes to 23.5%, however the maximal average ask liquidity proportion reaches 76.4%.

Figure 4 panel A depicts in addition an increase in the volatility. The average proportional liquidity has a mean that jumps to 41.7% in the systemic phase. Thus selling protection becomes more expensive and the ask liquidity premium reaches 97.3% of the bid-ask spread for some names. Given the market recovery during the next period, protection sellers react by easing the trade and consequently requires a lower liquidity risk premium relative to the bid-ask spread for most names. The mean of R_t consequently drops to 32.1%; however, selling protection on some names remains expensive and could reach 82.9% on average. During the last period, the sovereign phase, the average ask liquidity premium ranges now between 2.9% and 94.20% of the bid-ask spread with a mean that climbs to 43.8%. During periods of high credit risk (the systemic and the sovereign phases), the protection seller requires on average a larger percentage of the bid ask spread than in the period of declining or low model-implied CDS premium like the pre-crisis and recovery phase. Besides, Figure 4, Panel A, seems to depict a difference between financial and non-financial institutions. The protection seller requires a larger relative liquidity premium for non-financial names than financial names in all periods except the subprime period where both industrial sectors exhibit an almost equal ask liquidity premium proportion. The difference is particularly large during the last three periods, i.e. systemic, recovery, and sovereign phase.

Figures 9 illustrates the distributions of R_t for financial and non-financial institutions and for different periods. We first notice that the distributions are similar for both industrial sectors. During the pre-crisis and subprime phase, financial and non-financial institutions have a right skewed R_t values. Thus selling protection are generally cheap. However during the systemic phase of the financial crisis, the ask liquidity premia proportion are scattered on the unit interval and thus indicates a considerable more expensive protection selling, specially for non-financial institutions. This could be related to the unorthodox measures and bailout policies applied by several European governments to

save financial institutions and provide liquidity in the financial markets. Thereafter we observe that most values of the ask liquidity proportion are concentrated near the lower bound during the recovery phase, however the non-financial institutions have a larger dispersion. The distribution in the next period, the sovereign period, looks similar to the systemic phase with large dispersion. The ask liquidity proportion R_t covers all the unit interval indicating that selling protection is more expensive. During different phases of the financial downturn, protection sellers seem to require lower liquidity premium for financial than non-financial institutions.

4.2.4 Deviation of CDS mid quotes from model-implied CDS premia

Figure 4, Panel B, shows the difference between mid and model CDS premia (ΔS_t). For both financials and non-financials, the fluctuations of ΔS_t are rather low during the pre-crisis phase and increase substantially during the following crisis periods. While in the pre-crisis phase, the difference is larger for non-financials, it is larger for financials during the crisis periods. For financials, the difference is always, negative, i.e. the model default premium is larger than the mid quote. For non-financials, the difference is also mostly negative, with the exception of certain occasions in the systemic and the sovereign phase.

Figure 10 and Table 4 show that the dispersion of ΔS_t rises from 0.893 (non-financials) and 0.407 (financials) in the pre-crisis phase to a peak of 6.328 (non-financials) and 3.419 (financials) in the systemic phase and subsequently drops again. Overall, the difference between mid quotes and model spreads is negative on average in all periods for both financials and non-financials. The median deviation for financials is largest in the systemic period (-2.613), and for non-financials in the subprime phase (-1.331).

4.2.5 Evolution of liquidity premia in different phases of the financial crisis

We now turn to an investigation of the evolution of liquidity premia over time, in particular during different phases of the financial crisis. We perform a formal statistical analysis of the main variables obtained as output of our model (model-implied CDS premium S_t^{def} , proportion of liquidity premium R_t attributable to the ask side, absolute ask liquidity premium SL_t^{ask} , absolute bid liquidity premium SL_t^{bid} , and relative bid-ask spread). A difference-in-difference framework is applied, with financial institutions as treatment group, non-financial institutions as control group, the pre-crisis period as reference pe-

riod, and several periods during the crisis. For this, we apply several Random-effect Panel regressions, one for each variable from our model. The results are given in Table 5.

The results for the model-implied CDS premium are in the first section of Table 5. The numbers show that overall, the model-implied default premium increased during the crisis and reached a maximum in the systemic period. Financial institutions have a significantly lower default premium than non-financials. Moreover, the interaction terms for financial institutions show that financials have comparatively low default premium during the systemic crisis period, and comparatively large default premium during the sovereign crisis period. This can be explained by government aid and implicit bailout guarantees granted during the crisis that loose their effect with sovereigns becoming more affected by the crisis themselves. Essentially all coefficients in this regression are significant.

The next section of Table 5 shows the regression results for the proportion of the liquidity premium R_t attributable to the ask side. The numbers show that R_t increased significantly after the collapse of Lehman Brothers and remained larger than before. The interaction terms show that there is a significant relative decrease of R_t for financials in these periods.

The next section in Table 5 shows the results for the relative bid-ask spread $BA_{t,rel}^{def}$. There is a strongly negative relation between $BA_{t,rel}^{def}$ and the default premium. Overall, financials have a significantly higher relative bid-ask spread than non-financials. While the relative bid-ask spread decreased substantially after the start of the crisis, it decreased even more for financials, eventually leading to a lower spread than non-financials. All coefficients are highly significant.

The fourth and fifth regressions in Table 5 show the results for two variables, obtained by using our model to decompose the relative bid-ask spread $BA_{t,rel}^{def}$ into its two components, relative ask liquidity premium and relative bid liquidity premium. While the relative bid liquidity premium can be explained very well by the regression, the goodness of fit is very low for the relative ask liquidity premium. It can be seen, that the decrease of the relative bid-ask spread as discussed before can be attributed to the bid premium rather than the ask premium.

The last regression relates the difference between mid quotes and model-implied CDS values to the explanatory variables. The highly significant constant of -0.89676 shows that on average, the model-implied default spread is higher than the mid quote. The

coefficient for financials has a value of 0.25662 and is significantly positive, implying that the difference between mid quotes and model spreads is lower for financials than non-financials during the pre-crisis period, but it is still negative. The dummy variables for the periods indicate that for non-financials, the error increases in the subprime and recovery periods, while it decreases in the systemic and sovereign periods. In all periods, mid-quotes remain lower than model spreads. The interactions terms all have negative signs and are highly significant, implying that the error for financials is even higher than for non-financials. Overall, we find that mid-quotes very consistently underestimate the true default risk for all periods and for both financials and non-financials. The highest errors can be observed for financials during the crisis.

5 Conclusion

Our main objective in this paper is to estimate the CDS default risk and the liquidity premium from CDS bid and ask quotes. In order to better analyze the asymmetry in the liquidity premium of the protection buyer and seller, we introduce the liquidity proportion of the protection seller defined as the ratio of the ask liquidity premium to the bid-ask spread. Moreover, we split our data in five periods that relate to distinct market conditions: the pre-crisis phase, January 2004 - July 2007, characterized by a steady decline in CDS prices with a lower volatility; the subprime phase, August 2007 - August 2008, determined by the subprime credit crisis; the systemic phase, September 2008 - March 2009, catalyzed by the bankruptcy of Lehman Brothers and defined by a severe worsening of credit and liquidity conditions as well as larger market volatility; the recovery phase from April 2009 - December 2009, where CDS mid quotes fall and the market is less volatile; and the last period, the sovereign phase from January - September 2010, determined by widening CDS spreads and growing sovereign risk. We investigate the properties and the determinants of the liquidity premia under different market conditions.

Using a state-space model on the log-prices of 118 European names, we find that the default premium and the liquidity premium are strongly interconnected specially since the beginning of the financial turbulence. During periods of large default premia such as the systemic and sovereign phase, the liquidity proportion rises. Thus, the correlation between the default and liquidity variables should not be ignored when pricing CDS

contracts. Moreover, CDS mid quotes cannot be a pure measure of default risk since we find that for most names, the bid liquidity premium is larger than the ask liquidity premium, i.e. the protection buyer receives a larger liquidity premium than the protection seller. We find that different phases of the financial crisis affect both the default premium, and the liquidity premia earned by protection buyers and sellers. These effects are not symmetrical, and they are different for financial names compared to non-financial names. Overall, the obtained results demonstrate the necessity of applying a fully specified model to obtain a robust estimates for the true default premium and the liquidity premia.

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6 Figures

Figure 1: **Time-series of CDS mid quotes, bid-ask spreads, and relative bid-ask spreads**

This figure depicts time-series of averaged CDS mid quotes S_t^{mid} , bid-ask spreads BA_t , and relative bid-ask spreads $BA_{t,rel}^{mid}$ for the financial and non-financial sector. $BA_{t,rel}^{mid}$ is calculated as the ratio of bid-ask spreads to the mid quotes. The average is taken over individual names. The vertical lines indicate different periods.

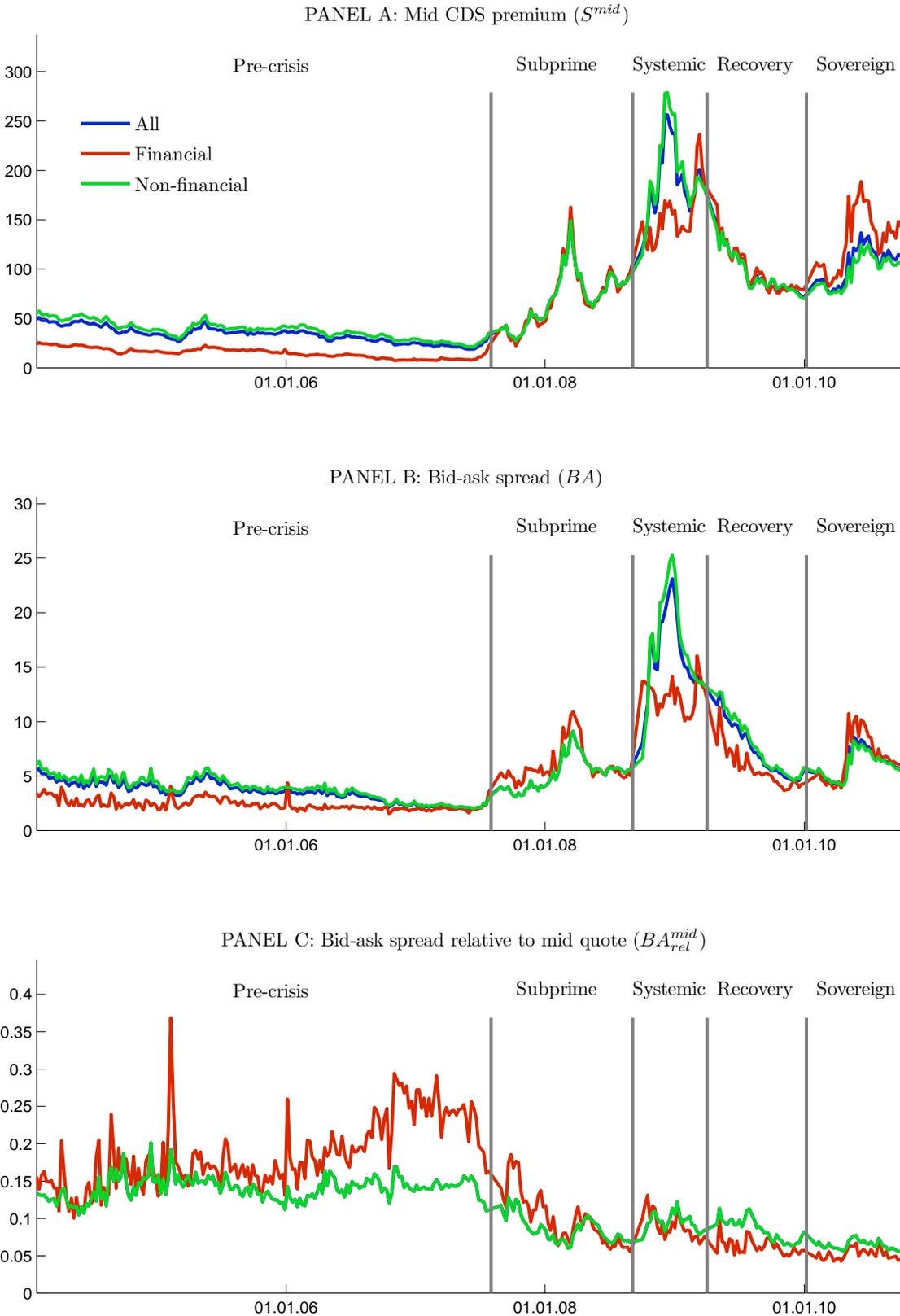


Figure 2: **Time-series of the model-implied CDS premium, ask and bid liquidity premia**

This figure depicts time-series of averaged model-implied CDS premia S_t^{def} , ask SL_t^{ask} and bid SL_t^{bid} liquidity premia for the financial and non-financial sector. The average is taken over individual names. The vertical lines indicate different periods.

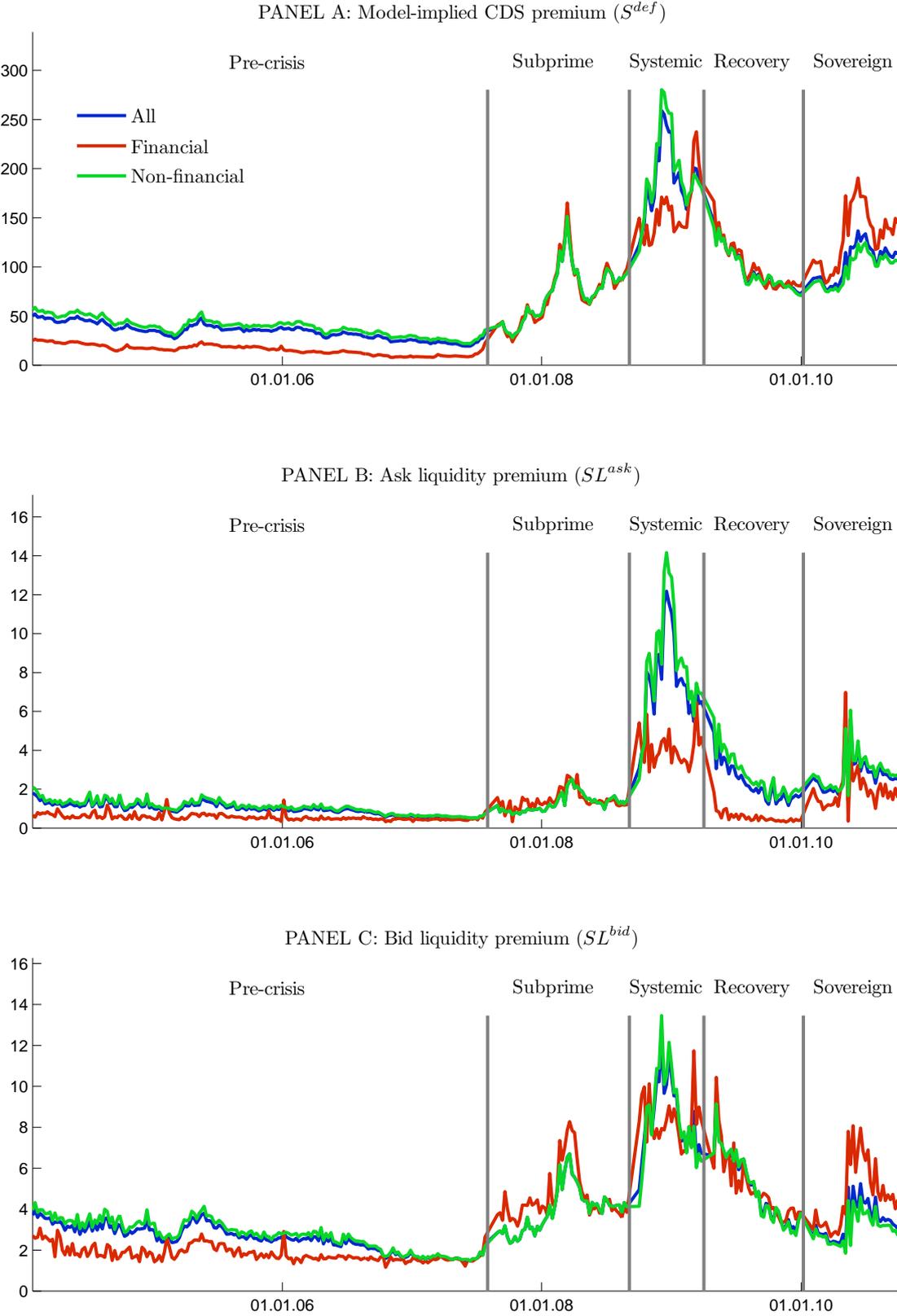


Figure 3: Time-series of the relative bid-ask spread and, ask and bid liquidity premia

This figure depicts time-series of the relative bid-ask spread $BA_{t,rel}^{def}$, the relative ask $SL_{t,rel}^{ask}$ and bid $SL_{t,rel}^{bid}$ liquidity premia for the financial and non-financial sector. They are computed relatively to the model default premia. The average is taken over individual names. The vertical lines indicate different periods.

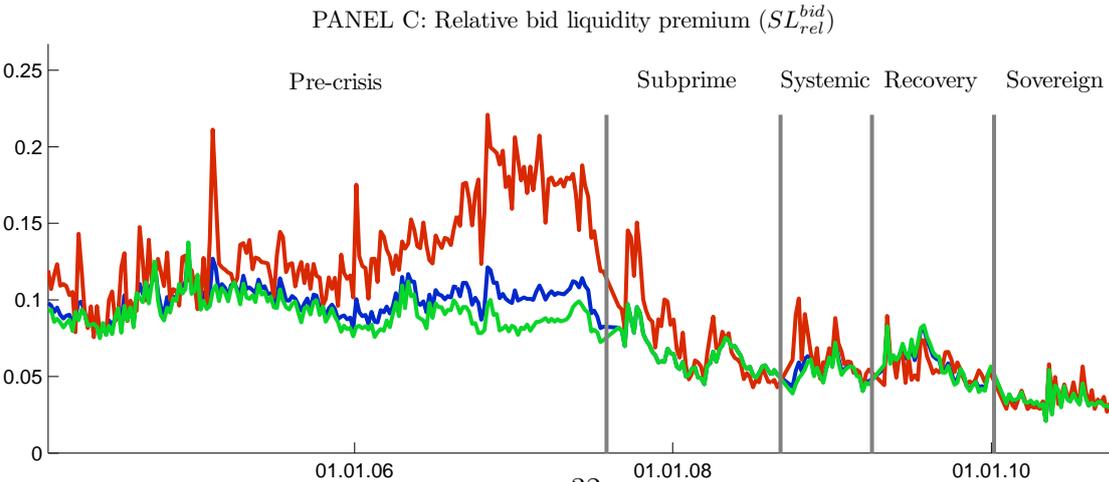
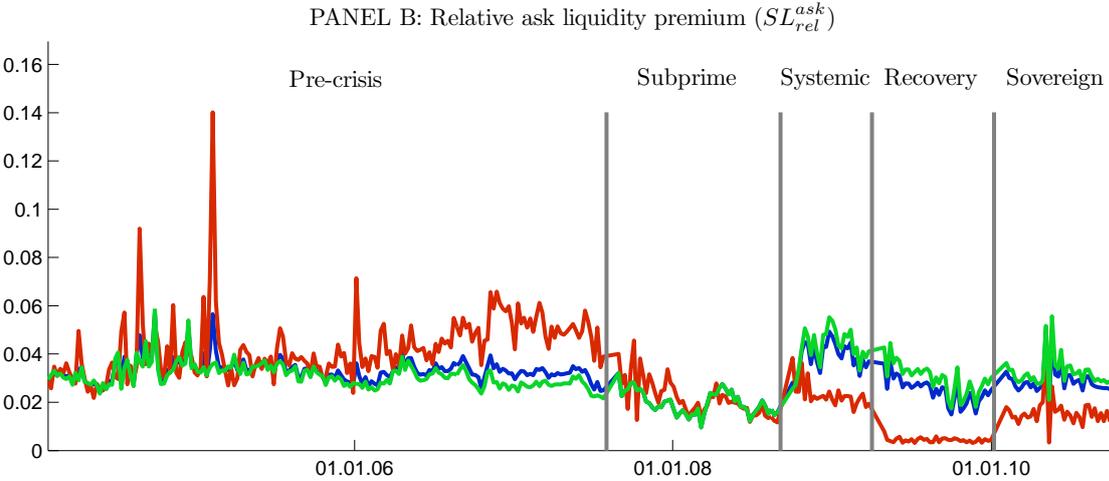
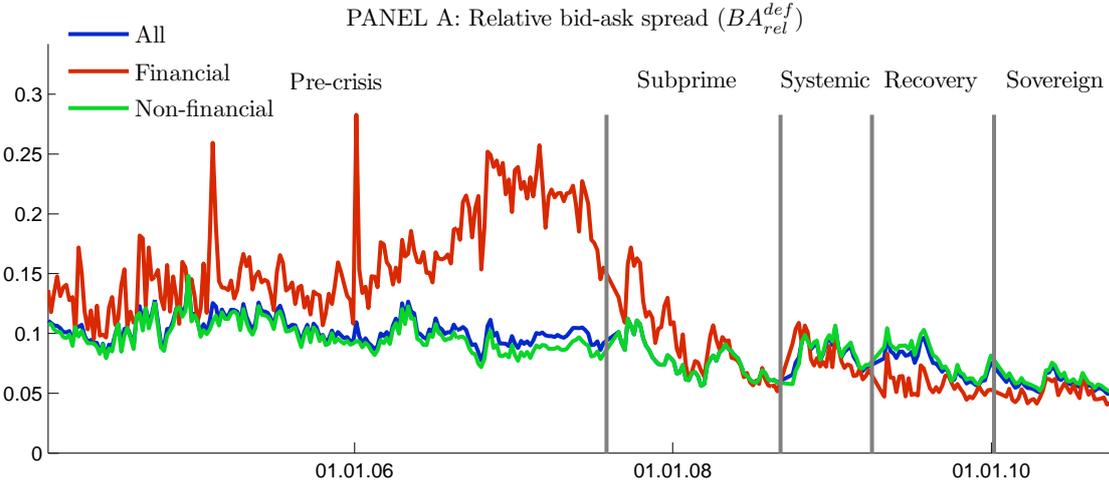


Figure 4: **Time-series of the proportional ask liquidity premia and the deviation of the mid quotes from the model-implied fair CDS premia**

This figure depicts time-series of the averaged proportional ask liquidity premia R_t and of the difference between the mid quotes and the model-implied CDS premia $\Delta S_t = S_t^{mid} - S_t^{def}$. The average is taken over individual names. The vertical lines indicate different periods.

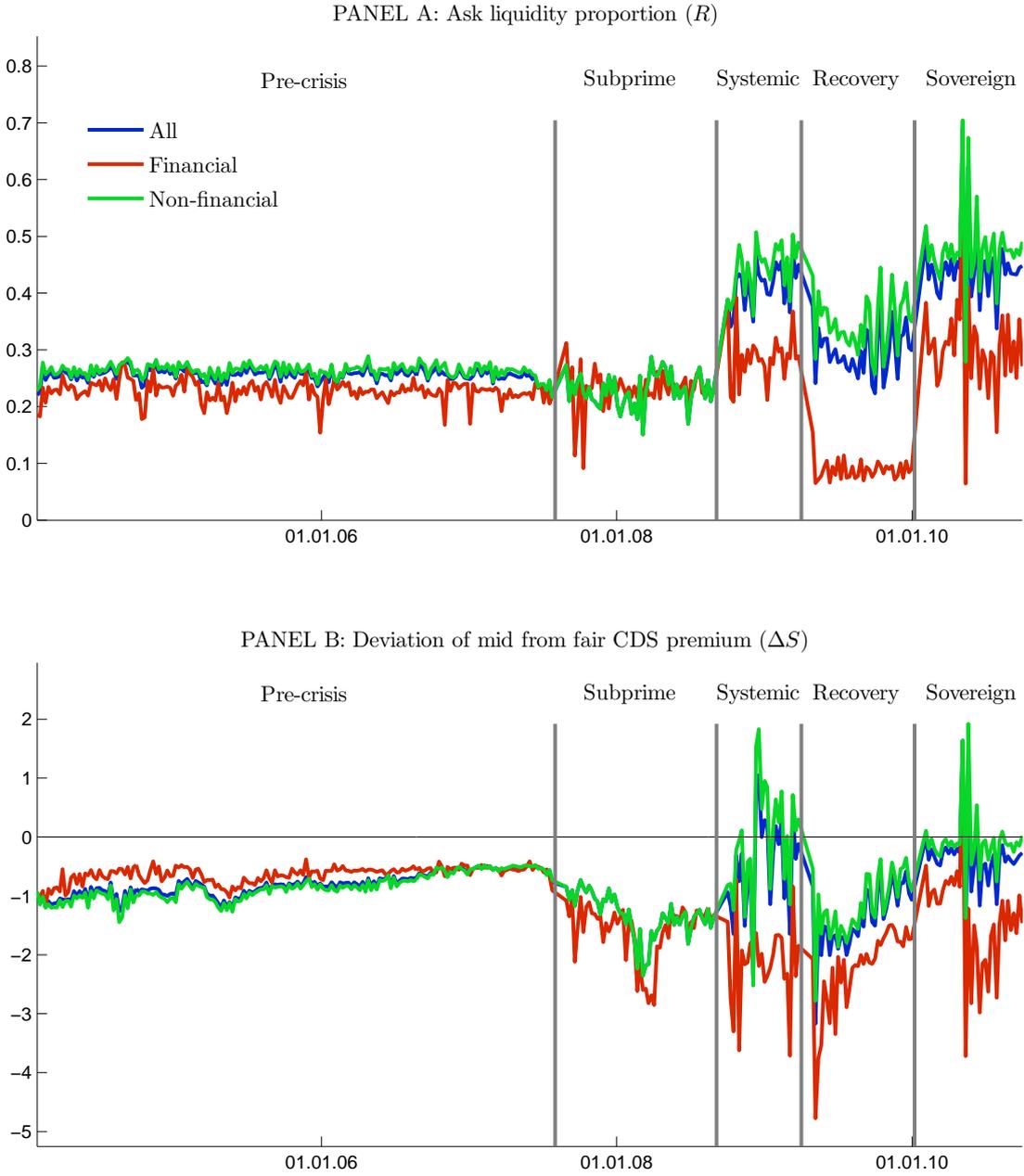


Figure 5: **Histogram of the model-implied CDS premium S_t^{def}**

This figure depicts the distribution of the model-implied CDS premium S_t^{def} for the financial and non-financial sector and also exhibits its evolution over different periods, from the pre-crisis period over different phases of the financial crisis.

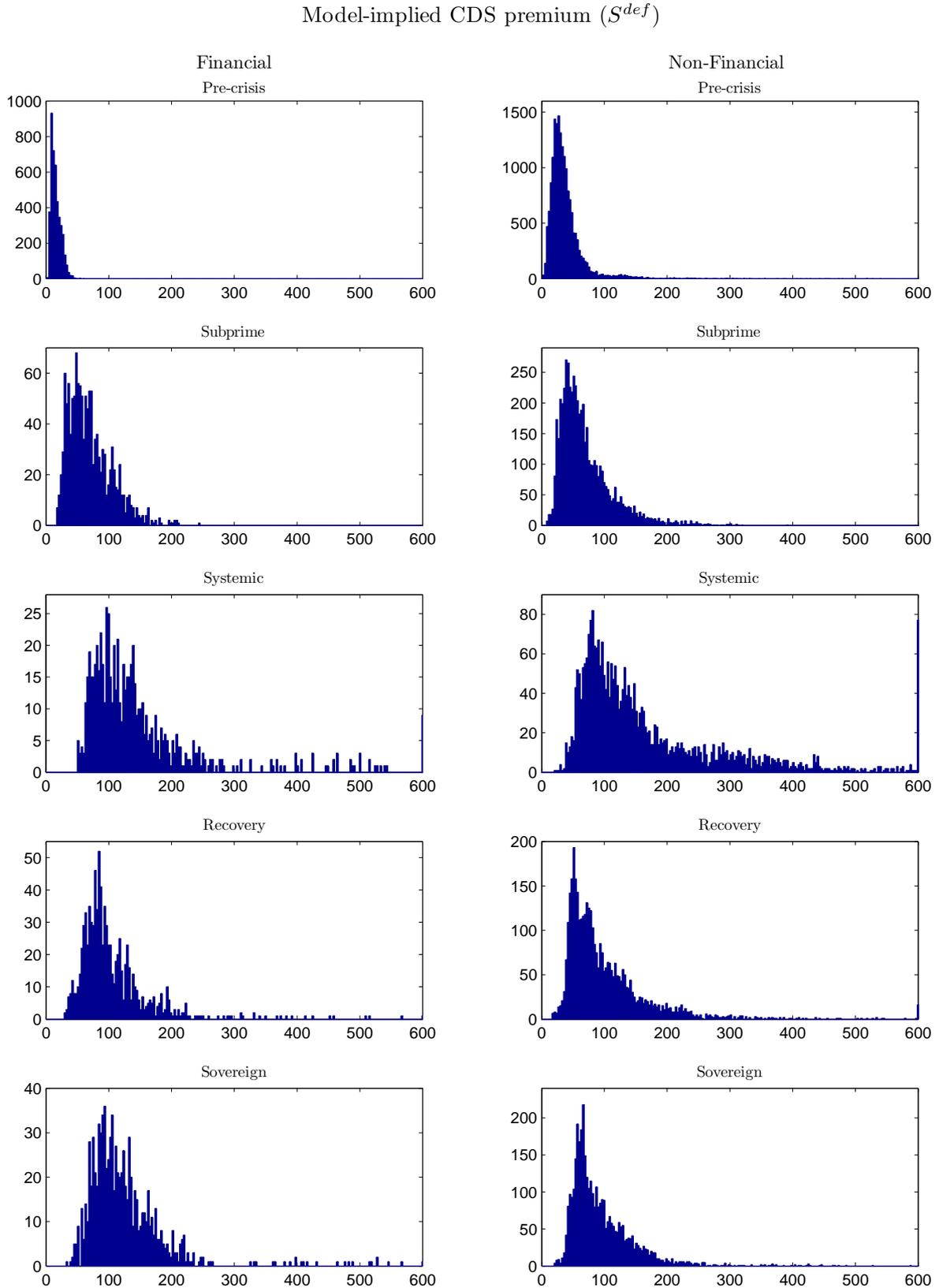


Figure 6: **Histogram of the relative bid-ask spread**

This figure depicts the distribution of the relative bid ask spread $BA_{t,rel}^{def} = \frac{BA_t}{S_t^{def}}$ for the financial and non-financial sector and also exhibits its evolution over different periods, from the pre-crisis period over different phases of the financial crisis.

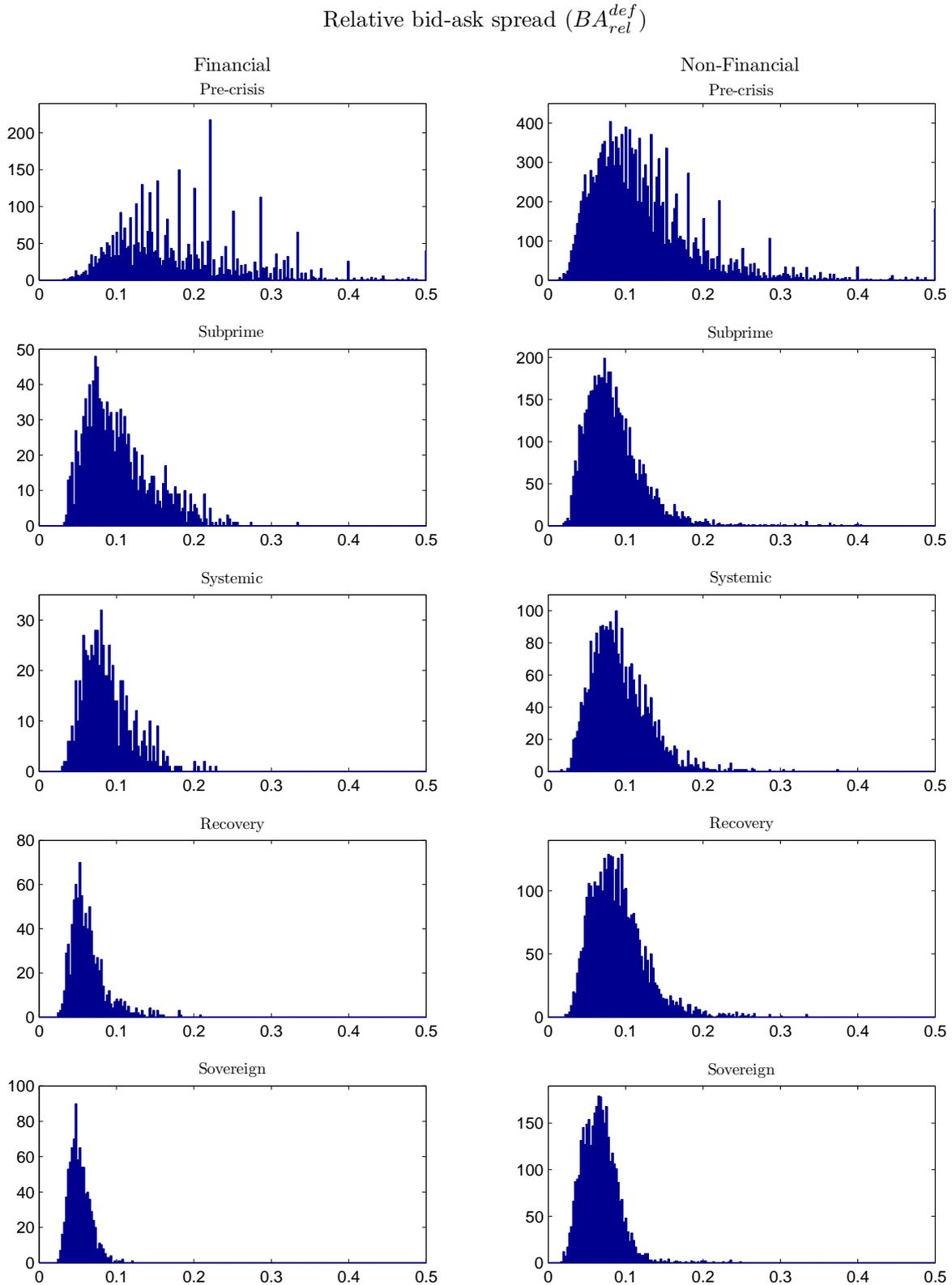


Figure 7: **Histogram of the relative liquidity premium of the protection buyer**

This figure depicts the distribution of the relative liquidity premium of the protection buyer $SL_{t,rel}^{bid} = \frac{SL_t^{bid}}{S_t^{def}}$ for the financial and non-financial sector and also exhibits its evolution over different periods, from the pre-crisis period over different phases of the financial crisis.

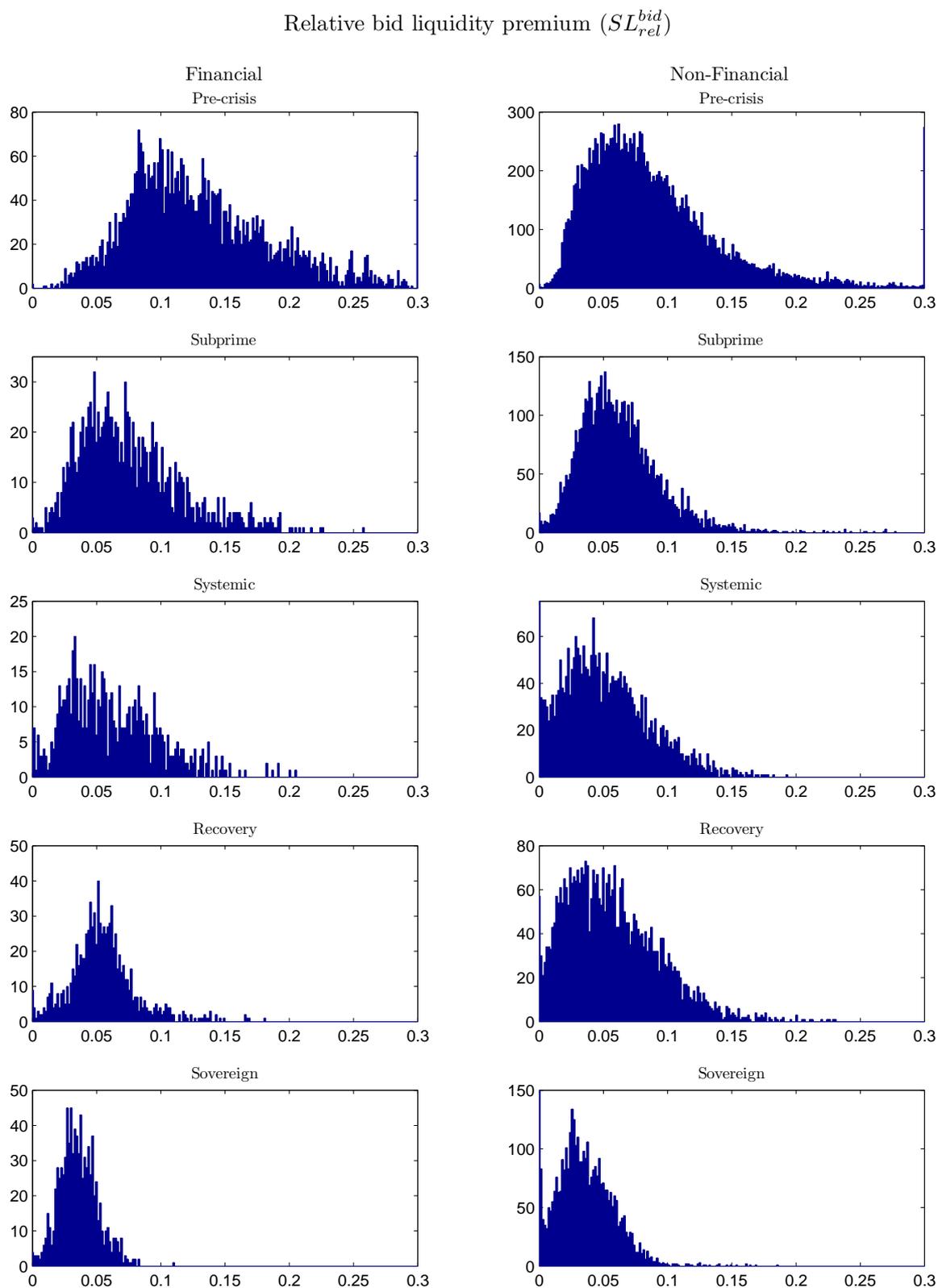


Figure 8: **Histogram of the relative liquidity premium of the protection seller**

This figure depicts the distribution of the relative liquidity premium of the protection seller $SL_{t,rel}^{ask} = \frac{SL_t^{ask}}{S_t^{def}}$ for the financial and non-financial sector and also exhibits its evolution over different periods, from the pre-crisis period over different phases of the financial crisis.

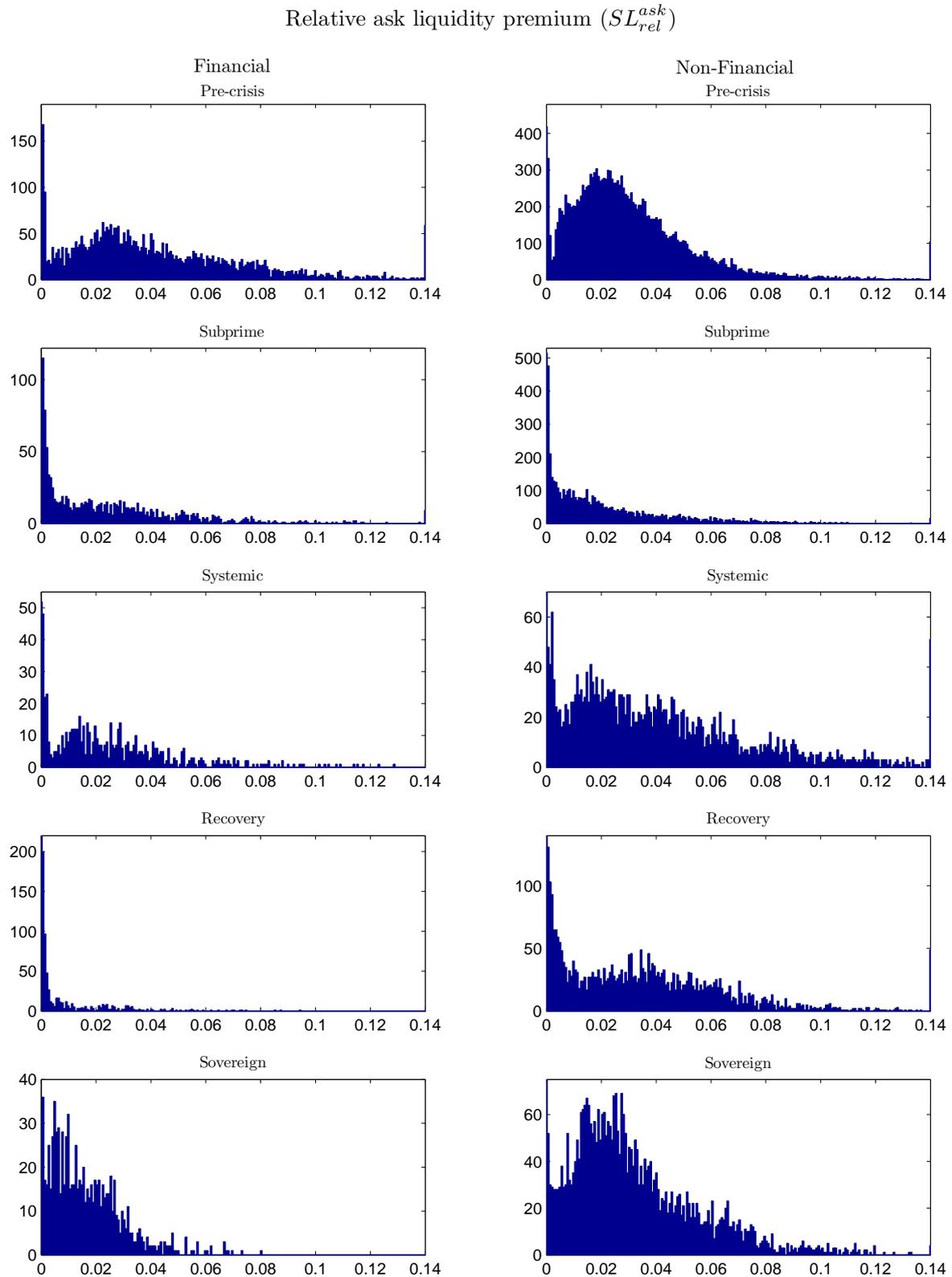


Figure 9: **Histogram of the ask liquidity premium proportional to the bid-ask spread**

This figure depicts the distribution of the liquidity premium of the protection seller proportional to the bid-ask spread $R_t = \frac{SL_t^{ask}}{BA_t}$ for the financial and non-financial sector and also exhibits its evolution over different periods, from the pre-crisis period over different phases of the financial crisis.

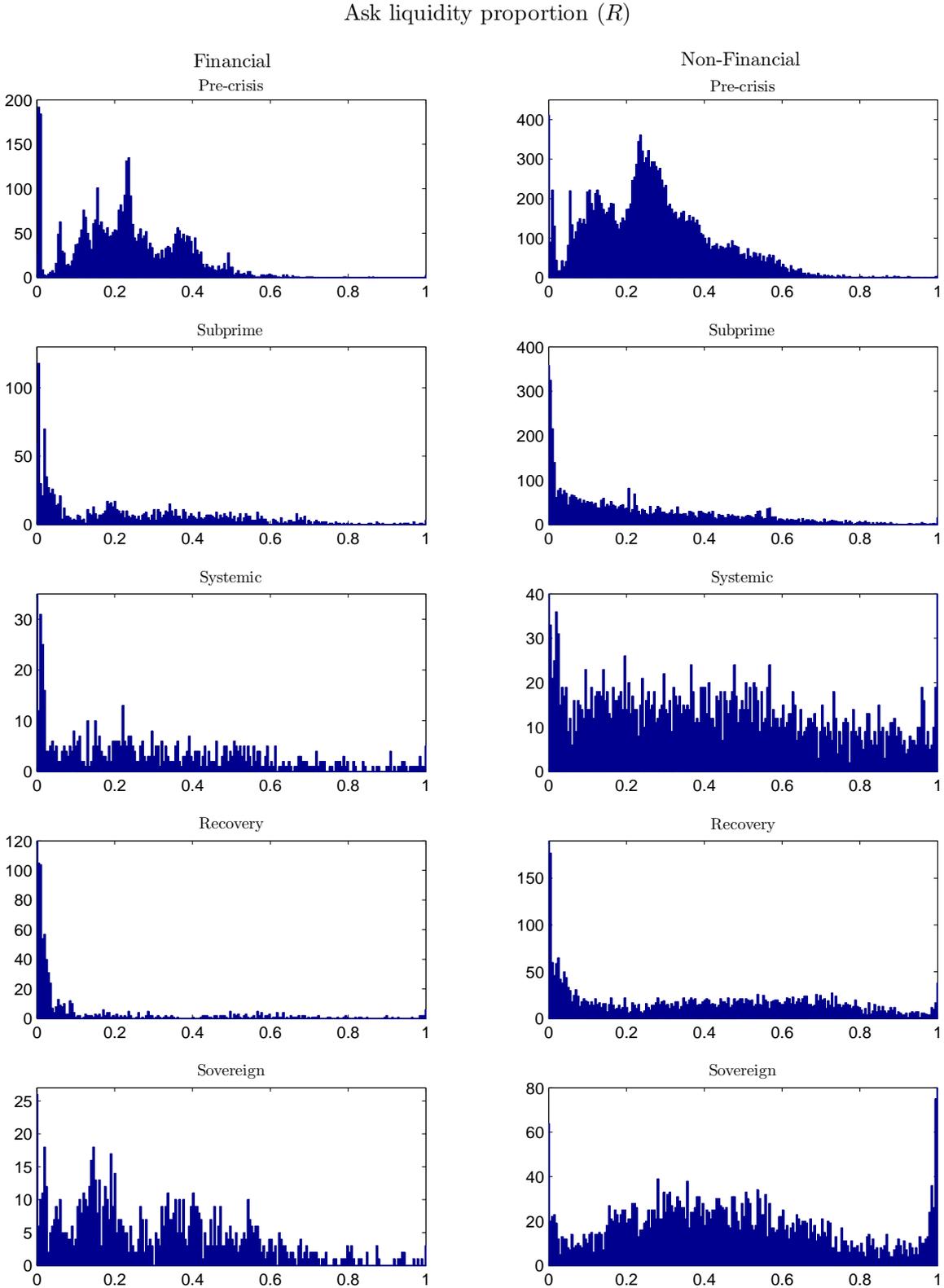
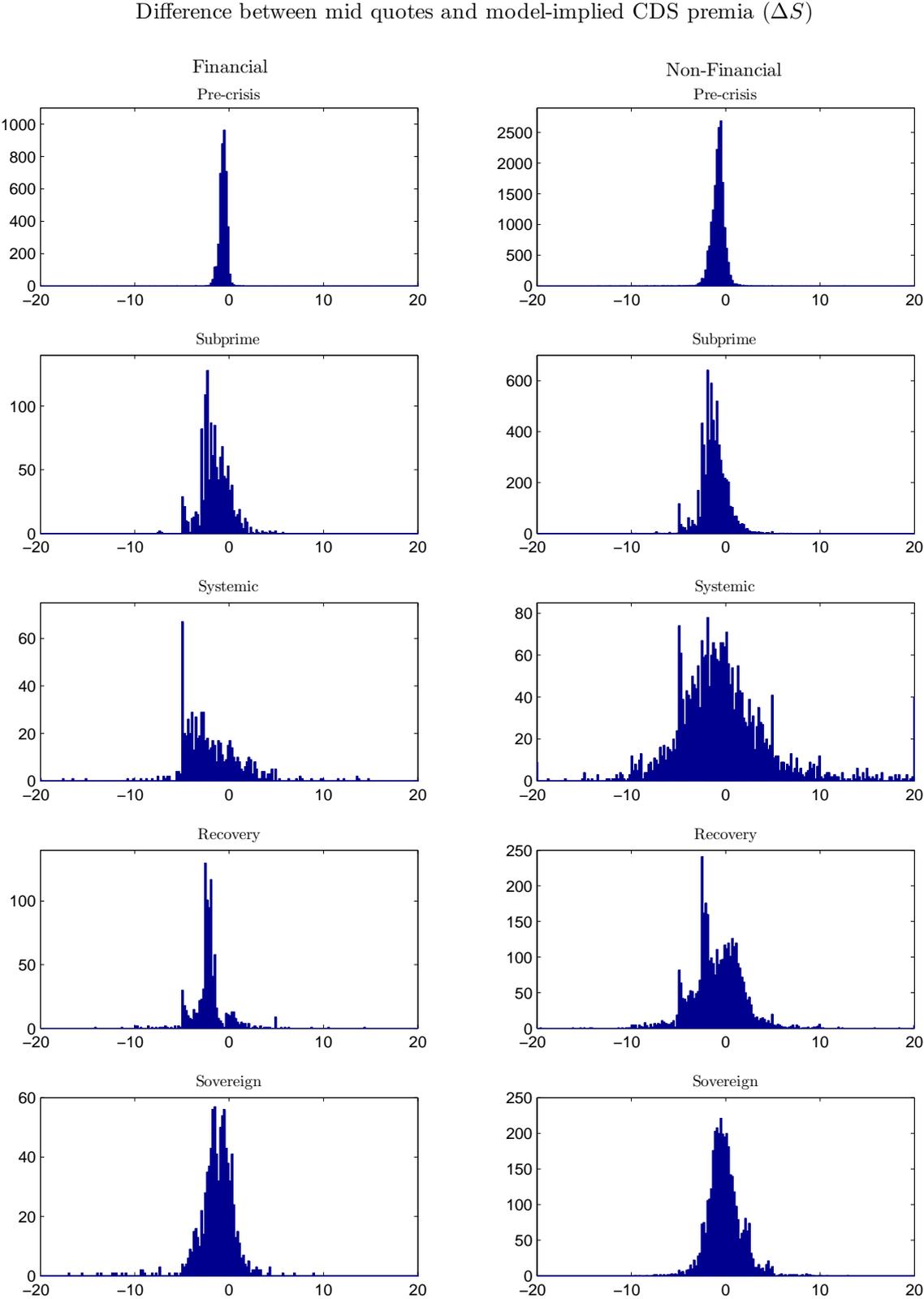


Figure 10: **Histogram of the deviation of the mid quotes from the model-implied CDS premia**

This figure depicts the distribution of the difference between the mid quotes and the model-implied CDS premia $\Delta S_t = S_t^{mid} - S_t^{def}$ for the financial and non-financial sector and also exhibits its evolution over different periods, from the pre-crisis period over different phases of the financial crisis.



7 Tables

Table 1: **Summary Statistics**

This table reports the means, the minimums, the maximums, the medians and standard deviations (std) of the collected data. S_t^{mid} is the CDS mid premium. BA_t is the absolute bid-ask spread. $BA_{t,rel}^{mid}$ is the relative bid-ask spread computed as the ratio of the bid-ask spread to the mid premium. ρ_{mid} denotes the empirical correlation between the CDS mid premium and a corresponding variable (absolute bid-ask spread or relative bid-ask spread).

	ALL			NON-FINANCIALS			FINANCIALS		
Pre-crisis Phase January 2004 - July 2007									
	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$
mean	33.570	3.556	0.140	38.075	3.840	0.129	15.211	2.398	0.186
min	3.000	1.000	0.012	3.000	1.000	0.012	3.500	1.000	0.030
max	528.300	82.000	1.985	528.300	82.000	1.985	58.500	48.000	1.983
median	26.800	3.000	0.118	30.700	3.000	0.109	13.500	2.000	0.166
std	31.189	2.997	0.107	33.086	3.209	0.102	7.491	1.395	0.116
ρ_{mid}	1.000	0.590	-0.347	1.000	0.579	-0.321	1.000	0.271	-0.492
Subprime Phase August 2007 - August 2008									
	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$
mean	68.824	5.207	0.089	68.912	5.010	0.085	68.307	6.004	0.103
min	8.500	1.000	0.020	8.500	2.000	0.020	0.000	0.000	0.034
max	390.000	25.000	0.404	390.000	25.000	0.404	244.200	20.000	0.333
median	57.700	5.000	0.080	57.100	5.000	0.078	61.450	5.000	0.093
std	41.592	2.426	0.041	43.134	2.443	0.039	34.629	2.192	0.045
ρ_{mid}	1.000	0.690	-0.524	1.000	0.717	-0.506	1.000	0.606	-0.672
Systemic Phase September 2008 - March 2009									
	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$
mean	178.218	14.827	0.092	184.947	15.547	0.093	150.424	11.849	0.088
min	23.200	3.000	0.017	23.200	3.000	0.017	45.000	4.000	0.029
max	2915.600	310.000	0.374	2915.600	310.000	0.374	775.000	58.000	0.229
median	125.000	10.000	0.086	127.500	10.000	0.087	120.000	10.000	0.082
std	182.795	16.031	0.036	196.081	17.384	0.037	108.109	7.705	0.032
ρ_{mid}	1.000	0.891	-0.227	1.000	0.895	-0.215	1.000	0.839	-0.399
Recovery Phase April 2009 - December 2009									
	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$
mean	102.472	7.845	0.085	101.722	8.236	0.090	105.565	6.232	0.064
min	19.510	2.340	0.022	19.510	2.650	0.022	27.490	2.340	0.026
max	1320.900	67.000	0.333	1320.900	67.000	0.333	635.380	40.000	0.208
median	81.343	5.570	0.079	78.975	6.000	0.085	88.283	5.000	0.058
std	77.933	5.272	0.036	81.130	5.530	0.036	62.979	3.613	0.024
ρ_{mid}	1.000	0.807	-0.310	1.000	0.820	-0.320	1.000	0.824	-0.310
Sovereign Phase January 2010 - September 2010									
	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$	S_t^{mid}	BA_t	$BA_{t,rel}^{mid}$
mean	101.301	5.949	0.064	95.499	5.830	0.067	125.268	6.439	0.052
min	20.110	1.970	0.018	20.110	1.970	0.018	32.480	2.400	0.026
max	592.190	40.770	0.249	580.975	26.730	0.249	592.190	40.770	0.119
median	86.788	4.990	0.062	79.355	4.970	0.065	110.530	5.200	0.050
std	58.544	3.163	0.023	54.577	2.867	0.024	67.618	4.135	0.014
ρ_{mid}	1.000	0.774	-0.418	1.000	0.731	-0.448	1.000	0.890	-0.139

Table 2: **Parameters estimates**

The estimates of the parameters of the log-fair CDS premium $s_t = s_{t-1} + \eta_t$ and the liquidity process $r_t = \alpha + \beta r_{t-1} + \sqrt{(1 - r_r)} r_t \epsilon$ are reported in this table.

	ALL					NON-FINANCIALS					FINANCIALS				
Pre-crisis Phase: January 2004 - July 2007															
	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$
mean	0.0766	0.2066	0.1502	0.3482	-0.4223	0.0749	0.2108	0.1709	0.3524	-0.4204	0.0837	0.1891	0.0648	0.3308	-0.4301
std	0.0001	0.0024	0.0049	0.0024	0.0702	0.0001	0.0027	0.0052	0.0026	0.0719	0.0001	0.0009	0.0032	0.0006	0.0622
min	0.0456	0.0003	-0.8626	0.0001	-1.0000	0.0456	0.0003	-0.8626	0.0001	-1.0000	0.0684	0.0024	-0.7150	0.0045	-0.9996
max	0.1609	0.8764	0.9625	1.5916	1.0000	0.1609	0.8764	0.8813	1.5916	1.0000	0.1069	0.4599	0.9625	0.7557	0.9992
Subprime Phase: August 2007 - August 2008															
	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$
mean	0.1323	0.2556	0.0729	0.6645	-0.3832	0.1264	0.2558	0.0622	0.6314	-0.4212	0.1566	0.2547	0.1168	0.8016	-0.2263
std	0.0003	0.0038	0.02	0.0026	0.0194	0.0003	0.0042	0.0172	0.0029	0.0210	0.0003	0.0004	0.0287	0.0006	0.0124
min	0.0932	0.0001	-0.9329	0.0000	-1.0000	0.0932	0.0001	-0.9329	0.0000	-1.0000	0.1272	0.0016	-0.7329	0.1339	-1.0000
max	0.1735	0.9742	0.8438	2.6176	0.6873	0.1615	0.9742	0.8438	2.6176	0.6026	0.1735	0.8663	0.6647	1.7718	0.6873
Systemic Phase: September 2008 - March 2009															
	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$
mean	0.1417	0.4251	0.0120	1.1036	-0.4663	0.1332	0.4474	0.0128	1.1110	-0.5186	0.1767	0.3331	0.0087	1.0730	-0.2503
std	0.0008	0.0117	0.0301	0.0108	0.0609	0.0008	0.0081	0.0336	0.0120	0.0669	0.0009	0.0210	0.0016	0.0010	0.0002
min	0.0704	0.0007	-0.7421	0.0000	-1.0000	0.0704	0.0019	-0.7421	0.0000	-1.0000	0.1432	0.0007	-0.6698	0.1916	-1.0000
max	0.2317	1.0000	0.8143	3.0000	0.9630	0.2116	1.0000	0.8143	3.0000	0.9630	0.2317	0.8576	0.7257	2.7136	0.7180
Recovery Phase: April 2009 - December 2009															
	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$
mean	0.0814	0.2757	0.2022	0.9441	-0.4683	0.0798	0.3212	0.1770	1.0334	-0.5644	0.0878	0.0874	0.3066	0.5751	-0.0712
std	0.0003	0.0011	0.0109	0.0393	0.3498	0.0002	0.0012	0.0113	0.0021	0.3858	0.0003	0.0003	0.0095	0.0898	0.0001
min	0.0395	0.0000	-0.6071	0.0001	-1.0000	0.0395	0.0000	-0.6071	0.0001	-1.0000	0.0584	0.0000	-0.4193	0.0004	-0.9235
max	0.1204	0.9034	0.8363	2.6808	1.0000	0.1204	0.9034	0.8363	2.6808	1.0000	0.1171	0.6647	0.5654	1.8897	0.2541
Sovereign Phase: January 2010 - September 2010															
	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$	σ_{η_t}	α	β	σ_{ϵ_t}	$\rho_{\eta_t, \epsilon_t}$
mean	0.0867	0.4431	-0.0105	1.2909	-0.3859	0.0786	0.4696	-0.0189	1.3044	-0.4796	0.1204	0.3339	0.0244	1.2351	0.0010
std	0.0002	0.0008	0.0011	0.0023	0.0684	0.0003	0.0009	0.0012	0.0025	0.0765	0.0001	0.0003	0.0003	0.0016	0.0012
min	0.0323	0.0075	-0.8546	0.0387	-1.0000	0.0323	0.0075	-0.7715	0.0387	-1.0000	0.0813	0.0098	-0.8546	0.2522	-0.6606
max	0.2539	1.0000	0.7076	2.8260	0.9970	0.2539	1.0000	0.7076	2.8260	0.9859	0.1730	0.9394	0.6083	2.7752	0.9970

Table 3: **Empirical Correlations**

This table presents the mean, the minimum and the maximum of the sample correlation between the model-implied CDS premium and respectively the relative liquidity premium $\rho(S_t^{def}, R_t)$, the ask liquidity premium $\rho(S_t^{def}, SL_t^{ask})$ and the bid liquidity premium $\rho(S_t^{def}, SL_t^{bid})$.

	ALL			NON-FINANCIALS			FINANCIALS		
Pre-crisis Phase: January 2004 - July 2007									
	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$
mean	-0.0880	0.3266	0.4771	-0.1324	0.2429	0.4172	-0.0773	0.3469	0.4916
min	-0.8140	-0.1976	-0.1154	-0.8140	-0.0966	0.1305	-0.5667	-0.1976	-0.1154
max	0.3840	0.7897	0.8987	0.1028	0.5010	0.5799	0.3840	0.7897	0.8987
Subprime Phase: August 2007 - August 2008									
	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$
mean	-0.0663	0.3982	0.6393	0.0309	0.3932	0.5360	-0.0899	0.3994	0.6643
min	-0.5749	-0.1829	-0.0667	-0.3742	-0.0869	-0.0134	-0.5749	-0.1829	-0.0667
max	0.6450	0.8219	0.9212	0.5262	0.6895	0.7360	0.6450	0.8219	0.9212
Systemic Phase: September 2008 - March 2009									
	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$
mean	-0.0169	0.3054	0.3696	0.1693	0.3121	0.1686	-0.0620	0.3038	0.4183
min	-0.6636	-0.5418	-0.4189	-0.4950	-0.2604	-0.3229	-0.6636	-0.5418	-0.4189
max	0.7226	0.9125	0.9107	0.6912	0.7507	0.7734	0.7226	0.9125	0.9107
Recovery Phase: April 2009 - December 2009									
	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$
mean	0.0380	0.4115	0.4511	0.0148	0.3030	0.5085	0.0437	0.4378	0.4372
min	-0.5446	-0.3882	-0.3138	-0.5342	-0.3555	-0.0932	-0.5446	-0.3882	-0.3138
max	0.9471	0.9002	0.9164	0.6966	0.7483	0.8570	0.9471	0.9002	0.9164
Sovereign Phase: January 2010 - September 2010									
	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$	$\rho(S_t^{def}, R_t)$	$\rho(S_t^{def}, SL_t^{ask})$	$\rho(S_t^{def}, SL_t^{bid})$
mean	0.1044	0.3882	0.2810	0.1089	0.4888	0.5177	0.1033	0.3639	0.2237
min	-0.5415	-0.3563	-0.4752	-0.5221	0.2468	-0.0286	-0.5415	-0.3563	-0.4752
max	0.6914	0.8622	0.9388	0.3736	0.7496	0.9335	0.6914	0.8622	0.9388

Table 4: Model-implied fair CDS and liquidity premia distribution properties

This table shows the mean, the minimum, the maximum and the median of the time-series average of the model-implied CDS premium S_t^{def} , the ratio of the protection seller's liquidity premium to the bid-ask spread R_t , the relative ask liquidity premium $SL_{t,rel}^{ask} = \frac{SL_t^{ask}}{S_t^{def}}$, the relative bid liquidity premium $SL_{t,rel}^{bid} = \frac{SL_t^{bid}}{S_t^{def}}$ and the deviation of the mid quotes from the model-implied fair CDS premia $\Delta S_t = S_t^{mid} - S_t^{def}$.

	ALL					NON-FINANCIALS					FINANCIALS				
Pre-crisis Phase															
January 2004 - July 2007															
	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t
mean	34.37	0.257	0.033	0.100	-0.80	38.92	0.264	0.031	0.092	-0.84	15.85	0.228	0.043	0.133	-0.64
min	3.81	0.000	0.000	0.000	-20.00	3.82	0.000	0.000	0.000	-20.00	3.81	0.000	0.000	0.000	-5.62
max	539.21	1.000	1.181	0.993	29.83	539.21	1.000	1.181	0.993	29.83	60.83	0.999	1.108	0.992	1.50
median	27.57	0.244	0.027	0.085	-0.73	31.49	0.252	0.026	0.077	-0.76	13.95	0.224	0.033	0.120	-0.60
std	31.43	0.151	0.038	0.071	0.82	33.33	0.154	0.034	0.068	0.89	7.63	0.137	0.050	0.070	0.41
skew	5.16	0.620	11.497	3.601	-1.16	5.04	0.629	12.647	3.928	-1.01	1.05	0.453	9.035	3.704	-0.89
Subprime Phase															
August 2007 - August 2008															
	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t
mean	70.12	0.235	0.021	0.065	-1.30	70.15	0.235	0.021	0.062	-1.24	70.00	0.236	0.024	0.076	-1.54
min	9.38	0.000	0.000	0.000	-10.00	9.38	0.000	0.000	0.000	-10.00	17.99	0.000	0.000	0.000	-7.50
max	391.67	1.000	0.313	0.277	9.78	391.67	1.000	0.313	0.277	9.78	244.25	1.000	0.258	0.258	5.69
median	58.99	0.174	0.013	0.060	-1.38	58.19	0.169	0.013	0.058	-1.33	62.99	0.187	0.016	0.069	-1.66
std	41.93	0.227	0.025	0.034	1.42	43.51	0.227	0.025	0.032	1.37	34.70	0.227	0.028	0.040	1.59
skew	1.79	0.981	2.503	1.290	0.15	1.86	1.008	2.421	1.341	0.13	1.12	0.869	2.635	0.957	0.32
Systemic Phase															
September 2008 - March 2009															
	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t
mean	178.68	0.417	0.038	0.053	-0.46	185.06	0.444	0.041	0.051	-0.11	152.34	0.305	0.024	0.062	-1.91
min	22.18	0.000	0.000	0.000	-72.05	22.18	0.000	0.000	0.000	-72.01	49.99	0.000	0.000	0.000	-24.99
max	2958.54	1.000	0.287	0.205	76.94	2958.54	1.000	0.287	0.193	76.94	790.17	1.000	0.129	0.205	14.86
median	126.57	0.386	0.029	0.048	-1.14	128.12	0.417	0.033	0.046	-0.79	123.05	0.240	0.019	0.055	-2.61
std	181.20	0.299	0.034	0.035	5.92	194.36	0.300	0.035	0.035	6.33	107.63	0.266	0.023	0.037	3.42
skew	5.48	0.372	1.512	0.737	1.89	5.34	0.282	1.425	0.715	1.79	2.73	0.760	1.480	0.804	0.44
Recovery Phase															
April 2009 - December 2009															
	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t
mean	103.62	0.321	0.027	0.056	-1.15	102.64	0.370	0.032	0.057	-0.92	107.67	0.120	0.007	0.055	-2.11
min	17.83	0.000	0.000	0.000	-19.60	17.83	0.000	0.000	0.000	-19.60	29.98	0.000	0.000	0.000	-14.16
max	1313.28	1.000	0.332	0.230	26.45	1313.28	1.000	0.332	0.230	26.45	629.39	1.000	0.094	0.181	14.46
median	82.82	0.259	0.019	0.052	-1.57	79.67	0.359	0.027	0.051	-0.98	90.47	0.017	0.001	0.053	-2.20
std	77.62	0.305	0.032	0.035	2.80	80.77	0.303	0.033	0.037	2.91	62.82	0.223	0.013	0.026	2.04
skew	4.83	0.517	2.239	0.924	0.74	4.97	0.290	2.120	0.886	0.62	3.25	2.230	2.932	1.006	1.04
Sovereign Phase															
January 2010 - September 2010															
	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t	S_t^{def}	R_t	$SL_{t,rel}^{ask}$	$SL_{t,rel}^{bid}$	ΔS_t
mean	101.65	0.438	0.029	0.035	-0.35	95.61	0.469	0.032	0.035	-0.12	126.59	0.307	0.016	0.036	-1.33
min	20.30	0.000	0.000	0.000	-17.07	20.30	0.000	0.000	0.000	-9.60	33.67	0.000	0.000	0.000	-17.07
max	601.49	1.000	0.168	0.186	12.93	586.92	1.000	0.168	0.186	12.93	601.49	1.000	0.080	0.110	8.92
median	86.92	0.408	0.024	0.033	-0.46	79.57	0.441	0.027	0.033	-0.28	111.38	0.272	0.013	0.035	-1.19
std	58.82	0.272	0.023	0.021	1.99	54.48	0.273	0.023	0.023	1.88	68.76	0.222	0.013	0.015	2.12
skew	2.84	0.409	1.328	0.897	-0.07	2.68	0.318	1.213	0.911	0.70	3.18	0.686	1.333	0.465	-1.95

Table 5: **Panel regressions explaining the effect of the financial crisis on the relative bid-ask spread, the model-implied default and the liquidity premia**

This Table shows the results of random effects panel regressions and explains the impact of the crises periods on the model-implied CDS premium S_t^{def} , the relative bid-ask spread $BA_{t,rel}^{def}$, the ratio of the liquidity premia R_t , the relative liquidity premium of the protection seller $SL_{t,rel}^{ask}$, and the relative liquidity premium of the protection buyer $SL_{t,rel}^{bid}$. The relative bid-ask spread $BA_{t,rel}^{def} = \frac{BA_t}{S_t^{def}}$ is computed as the ratio of the bid-ask spread to the model-implied CDS premium. The ratio $R_t = \frac{SL_t^{ask}}{BA_t}$ represents the proportion of the bid-ask spread attributed to the protection seller and equals the ratio of the ask liquidity premium to the bid-ask spread. $SL_{t,rel}^{ask} = \frac{SL_t^{ask}}{S_t^{def}}$ is the liquidity premium of the protection seller and $SL_{t,rel}^{bid} = \frac{SL_t^{bid}}{S_t^{def}}$ is the liquidity premium of the protection seller buyer relative to the model-implied CDS premium. ΔS_t designates the difference between the mid quote and the model-implied CDS premium. x_t^{def} is defined as $x_t^{def} = -1000e^{-S_t^{def}}$ such that the model default premium is linearly related to the explained variables.

	S_t^{def}	t-stat	R_t	t-stat	$BA_{t,rel}^{def}$	t-stat	$SL_{t,rel}^{ask}$	t-stat	$SL_{t,rel}^{bid}$	t-stat	ΔS_t	t-stat
S_t^{def}			0.0001	1.30								
x_t^{def}					-0.0230	-9.63	-0.0027	-1.96	-0.0204	-12.76	-0.0207	-1.63
financial	-24.21	-7.51	-0.0315	-1.12	0.0494	4.93	0.0113	2.16	0.0381	4.73	0.2566	2.61
subprime	30.09	10.30	-0.0298	-1.33	-0.0385	-10.63	-0.0103	-4.65	-0.0282	-7.75	-0.3387	-2.59
systemic	145.00	9.09	0.1686	5.11	-0.0294	-6.10	0.0104	3.18	-0.0398	-7.95	0.7857	1.70
recovery	62.58	10.71	0.1021	3.40	-0.0326	-6.58	0.0015	0.44	-0.0340	-7.02	-0.0228	-0.12
sovereign	55.56	12.59	0.2022	8.50	-0.0544	-11.35	0.0010	0.43	-0.0554	-11.54	0.7817	5.27
fin_sub	24.07	6.06	0.0327	0.72	-0.0327	-3.52	-0.0082	-1.44	-0.0245	-2.82	-0.5611	-2.09
fin_sys	-8.51	-0.34	-0.1045	-1.84	-0.0557	-5.09	-0.0285	-4.02	-0.0272	-2.92	-2.0564	-3.33
fin_rec	29.24	2.71	-0.2187	-4.43	-0.0766	-6.74	-0.0369	-6.22	-0.0397	-4.26	-1.4448	-4.99
fin_sov	55.19	4.46	-0.1335	-2.78	-0.0646	-6.21	-0.0271	-4.47	-0.0375	-4.55	-1.4665	-4.20
const	40.06	13.10	0.2584	19.36	0.1215	22.93	0.0309	18.52	0.0906	18.81	-0.8968	-11.37
Adj R sqd	0.29		0.1369		0.2267		0.0527		0.2770		0.0342	
N Obs	41654		41654		41654		41654		41654		41654	