# Online Appendix for Severe Air Pollution and Labor Productivity: Evidence from Industrial Towns in China 

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## A. Further discussion of institutions and data

The work day choice. In each 30-day period of department activity (that is, excluding department holidays with zero aggregate output), we observe workers typically producing on 21 to 25 12-hour shifts in Jiangsu, and 17 to 21 12-hour shifts in Henan (Figure A.1). ${ }^{1}$ Recalling Henan's standard shift design-work 12 hours, rest 24 hours - and that mathematically there are 20 such 36 -hour cycles in a 30 -day period, the evidence indicates that worker absenteeism is low. In Jiangsu-work 12 hours, rest 12 hours, and take leave one day every two to four consecutive work days-hours worked are higher than in Henan. We do not observe which of a given worker's days with zero output were scheduled (predetermined) leave days, and which days were unplanned absences. However, to judge by the higher work hours than in Henan, worker absenteeism in Jiangsu is similarly low.
Our empirical analysis considers the possibility of selection on pollution in our worker-day output samples (see the next section). A worker's absence on a given day could be due to pollution-related sickness, or to weather (a correlate of pollution) shifting the value of the outside option, such as clear skies inducing the worker to spend a leisure day outdoors (Shi and Skuterud, 2015). We do not find a significant association between pollution and the probability of positive versus zero output on a given day (Table A. 5 and Figure A.4). This is not surprising given the low worker absenteeism. Our interpretation is that these are relatively healthy Chinese workers who live close to their employer, an employer that has much information about the employee, say to verify a shirking employee's false claim of sickness on justifying an unplanned absence. Moreover, scheduled leaves, like department holidays, are predetermined, so they are unlikely to depend on short-run variation in environmental quality. In practice, we find that attempting to control for selection in the output equation, using exogeneity restrictions based on concurrent weather, makes little difference to our estimates of the relationship between pollution and productivity (Table 4 versus Table A.6).

[^0]Jiangsu's monthly work shift transition. To preserve the continuity of operations, the monthly shift transition (day shift versus night shift) happens at the end of each month over a few days, rather than abruptly on day 1 of the subsequent month. For example, October 2014 records indicate that Team 1 worked the day shift and Team 2 worked the night shift. However, on October 30, 2014, we observe a total of 12 Team 1 workers and 5 Team 2 workers with non-zero output. Given the required staff of 8 to 9 workers per shift ( 17 machines at 2 machines per worker), the asymmetry in the number of workers from each team who worked on October 30 is due to the fact that, while some Team 2 workers took leave, some Team 1 workers already transitioned to working the night shift alongside the other Team 2 workers. Since output records do not mention who the Team 1 workers newly transitioned to the night shift were, nor when the transition started (e.g., October 28 or 29), our regression models include indicators for the last three days of the month interacted with shift, to account for measurement error in the time of day for these few days and for some workers (three Team 1 workers in the example). It is during the last three days of each month that the numbers of Team 1 and Team 2 workers with non-zero output are most unbalanced. Estimates are robust to alternatively dropping these month-end observations.

Worker tasks, products and piece rates. A Jiangsu worker operates machines (most often two) as they wind yarn threads from long narrow ring bobbins into large cones. Over time and following a production schedule, the team leader assigns workers to products and machines, but the worker keeps to her workstation, with the assigned products and machines, for at least an entire shift. As days, weeks, or months pass, a worker may operate different machines, but these machines are similar. Each machine has 16 magazines and each magazine can hold 5 ring bobbins at a time. It takes 10-15 minutes for a machine to completely wind the yarn, depending on the thickness of the yarn, at which point the worker has to replace the bobbins. To do so, one end of thread in a new bobbin needs to be pulled out and attached to the magazine, which requires skill and effort from the worker. In addition, threads naturally break apart during machine operation. The Jiangsu machines - which are more sophisticated than at the Henan department - can connect broken threads automatically, taking 5-10 seconds each time. However, this process may fail, in which case a red light alerts the worker, who needs to intervene and manually reconnect the thread as quickly as possible, for the machine to resume operation. The Jiangsu department produces products with some differentiation, with 80s accounting for the largest share of output in our sample (a mean of $32 \%$ of the within-shift department output across dateshifts), followed by $50 \mathrm{~s}(22 \%)$, 60 s ( $18 \%$ ), $45 \mathrm{~s}(14 \%)$, and $100 \mathrm{~s}(12 \%)$. These five products thus account for $98 \%$ of combined compensation-adjusted output. Another four products (34s, 40s, 70s, and 120s) together account for only $2 \%$ of output.
Henan workers perform similar tasks. They walk up and down the aisles in
their workstation, each worker attending to about five machines, and reconnecting the threads that break during machine operation. Again, a team leader assigns workers to products and machines, and a worker keeps to the assigned products and machines through the shift. One product, 32 s , accounts for about half the department's output (a mean of $48 \%$ of the aggregate output across date-shifts), followed by $40 \mathrm{~s}(18 \%), 21 \mathrm{~s}(13 \%)$, and $24 \mathrm{~s}(10 \%)$. Jointly, these four products account for $89 \%$ of output, and six other products account for the remainder ( $7 \mathrm{~s}, 16 \mathrm{~s}, 18 \mathrm{~s}, 19.5 \mathrm{~s}, 26 \mathrm{~s}$, and 32.5 s ). As in the Jiangsu department, a worker's product-machine assignment is persistent over days and weeks.
Piece rates in Henan, expressed in CNY/kg, are: 0.41 for output of product $40 \mathrm{~s}, 0.37$ for 32 s on an extended machine, 0.29 for 32 s on a regular machine, 0.25 for $24 \mathrm{~s}, 0.235$ for $26 \mathrm{~s}, 0.205$ for $21 \mathrm{~s}, 0.19$ for 19.5 s and $18 \mathrm{~s}, 0.165$ for 16 s , and 0.10 for 7 s . In addition to these products, a "preparatory task" whose output is recorded is compensated at piece rates of $0.22 \mathrm{CNY} / \mathrm{kg}$ on an extended machine and $0.205 \mathrm{CNY} / \mathrm{kg}$ on a regular machine.

To determine compensation, the Jiangsu department aggregates output across individual products using the following weights, expressed in adjusted cases/physical case for each product: 1.50 for output of 120s, 1.40 for $100 \mathrm{~s}, 1.30$ for $80 \mathrm{~s}, 1.15$ for $70 \mathrm{~s}, 1.06$ for $60 \mathrm{~s}, 1.00$ for $50 \mathrm{~s}, 0.98$ for 45 s and 40 s , and 0.96 for 34 s . A worker's aggregate number of adjusted cases is converted into "counts" at a rate of 105 counts/adjusted case. For output produced in the 12-hour shift, expressed in counts, a worker enjoys a bonus of 0.01 CNY per count in excess of 11,000 counts but no greater than 12,000 counts (i.e., 10 CNY per 1,000 excess counts), a bonus of 0.02 CNY per count in excess of 12,000 counts but no greater than 13,000 counts, and a bonus of 0.03 CNY per count for output in excess of 13,000 counts. The penalty scheme for 12 -hour output that falls short of 11,000 counts is symmetric to the bonus scheme around the reference output of 11,000 counts. This variable compensation scheme can be visualized in Figure A.2.

Differences in piece rates across products reflect differences in the standard rates of output from variation in product thread. Worker tasks are very similar across products. Figure A. 3 shows that the assignment of product varieties to workers exhibits strong day-to-day persistence at both work sites. To prepare the plots by work site, we take the four or five varieties in a work site with doubledigit output shares, and aggregate the remaining varieties into a residual category labeled "Other." We consider pairs of consecutive work shifts by a worker, with 12 hours and up to 24 hours of rest in between work in Jiangsu and Henan, respectively, given the standard shift designs. We then compute the empirical probability that, conditional on a worker producing a given variety on a shift, she also produced that variety in her preceding shift (likely among others-Table 2). These product-level transition probabilities within worker are of the order of 60 to $80 \%$.

Cleaning the worker output samples. For each department, we observe worker output by date-shift pair. Table A. 1 describes the minimal data cleaning
we apply to the raw records. We drop the one shift that precedes (resp., follows) a department holiday, during which machines are turned off early (resp., turned on late) for cleaning (resp., setting up). We drop a few date-shifts for which exceptional conditions were annotated, such as the occurrence of a power outage that affects all workers who worked on that shift. For the analysis of individual output, we additionally drop a few observations for which exceptional conditions were annotated, such as the worker being unusually assigned multiple tasks during the shift, beyond working at her single workstation, or was not working in the main ground-level production facility (Henan).

Air pollution data availability. Hourly PM2.5 and $\mathrm{SO}_{2}$ mass concentration measurements at official air monitoring sites in both towns that host the workplaces we study are quite complete during the sample periods. ${ }^{2}$ The exception is the Henan workplace during 2014, for which ambient air data are available only at three neighboring cities at most 60 km away. As stated below, the few missing observations are fairly evenly distributed throughout the sample periods and do not cluster on specific days, suggesting that the sites are well maintained.

For the PM2.5 monitor 3.7 km from the Jiangsu workplace, only $3 \%$ of all 7272 hours during the period August 2, 2014 ( 30 days prior to the start of the output sample) to May 31, 2015 exhibit missing concentrations in the raw data, say due to equipment failure. We impute the relatively few hourly values that are missing for the closest monitor using the mean value recorded contemporaneously at three other PM2.5 monitors located in the same town. Pairwise correlation coefficients for PM2.5 measured hour by hour at the closest monitor and at each of the other three same-town monitors range between 0.93 and 0.96 , indicating high levels of spatial correlation within the Jiangsu town.

Similarly, for the $\mathrm{SO}_{2}$ monitor 3.7 km from the Jiangsu workplace, $4 \%$ of all hourly observations are missing during the period, and we impute the few missing values using the mean value recorded at three other $\mathrm{SO}_{2}$ monitors located in the same town. We noticed that the $\mathrm{SO}_{2}$ series for a given monitor can exhibit shortlived and isolated spikes, fluctuations that are often not observed around the same hours at the other $\mathrm{SO}_{2}$ monitors. For example, the five hourly readings starting at 8 am on August 15, 2014 at the closest monitor are 13, 15, 219, 30 and 39 $\mu \mathrm{g} / \mathrm{m}^{3}$. For this reason, when using $\mathrm{SO}_{2}$ in our specifications, we take mean $\mathrm{SO}_{2}$ levels across monitors in the town. The correlation coefficient between mean $\mathrm{SO}_{2}$ across monitors in the town and PM2.5 at the monitor closest to the work site is 0.61 (again, hourly series).

For the PM2.5 monitor 3.4 km from the Henan workplace, only $2 \%$ of all 3624 hourly records during the period January 1, 2015 to May 31, 2015 are missing. We recode to missing 24 successive hourly records starting at noon on January 18, 2015 that are fixed at $60 \mu \mathrm{~g} / \mathrm{m}^{3}$, when observations at three other monitors in

[^1]the same town vary from hour to hour and in step. As in Jiangsu, we then impute the few values that are missing for the workplace's closest monitor using the mean contemporaneous value at the same town's three other PM2.5 monitors. Again, hourly PM2.5 readings exhibit high levels of spatial correlation; pairwise correlation coefficients for measurements at the closest monitor and measurements at the other same-town monitors are at least 0.90. Starting in March 2014 (one month prior to the start of the output sample), hourly data availability is similarly high at PM2.5 monitors in three neighboring cities at most 60 km away, 17 monitors in all. For the overlapping 2015 period, the correlation coefficient for PM2.5 measured 3.4 km from the Henan workplace and mean PM2.5 across the neighboring cities is 0.80 .
Similar comments to those above apply to hourly $\mathrm{SO}_{2}$ concentrations measured in the town that hosts the Henan workplace ( 4 monitors, the closest being 3.4 km away) and at the neighboring cities ( 17 monitors). For the overlapping 2015 period, the correlation coefficient for mean $\mathrm{SO}_{2}$ measured at the host town and mean $\mathrm{SO}_{2}$ measured across the neighboring cities is 0.76 .

Surface and atmospheric meteorological data availability. As stated in the text, we obtain local surface-level temperature, humidity and precipitation readings, at three-hour intervals, from NASA. For each workplace, we select the geographic cell ( 0.25 degree by 0.25 degree) in the NASA data that is centered on its host town. We obtain surface-level readings for wind speed and direction, and for atmospheric temperature differences, at 12 -hour intervals, from NOAA. For each workplace, we select the reference location in the NOAA data that is closed to the host town, namely station identification codes CHM00058362 for the Jiangsu site and CHM00057083 for the Henan site. ${ }^{3}$

Both datasets are quite complete during the sample periods. For example, surface-level data in the NOAA data for stations CHM00058362 and CHM00057083 are missing for one and six 12 -hour readings, respectively. Whenever the wind direction or wind speed readings on the surface are missing, we take the corresponding parameter reading above the surface at an atmospheric pressure point of 1000 mb , if the latter is available; failing that, we take the wind direction or speed reading further above the surface at an atmospheric pressure point of 925 mb .

## B. A selection model.

The empirical design is based on a more general model of the worker's problem, where the worker's productive but costly effort choice $e$, conditional on attending work, is embedded in the choice of attending work over an unscheduled absence

[^2]with reservation utility $\phi$ :
\[

$$
\begin{equation*}
\max \left\{\max _{e} w q(e, a)-c(e, Z), \phi(Z, W)\right\} . \tag{B.1}
\end{equation*}
$$

\]

Conditional on selecting into work on a given day, the worker is paid a piece rate $w$ per unit of individual output quantity $q$. Besides its dependence on effort, output can shift with the worker's ability $a$. The worker's exposure to ambient air pollution $Z$, both concurrent to and preceding the day of work, shifts her cost of effort and, thus, the payoff from work.
Pollution exposure $Z$ may also affect selection into work on a given day. ${ }^{4}$ The weather-sheltered workplace in our settings-temperature and humidity control, indoor protection from rain-provides a natural exclusion restriction for identifying the selection equation, since concurrent weather $W$ can shift the value of outdoor leisure $\phi$ but not the payoff from work. The epidemiological studies cited in the text do not model a distributed lag for weather.
The empirical counterpart to framework (B.1) is then:

$$
\begin{align*}
d_{i t}^{*} & =\gamma_{0}+Z_{t} \delta+W_{t} \gamma_{1}+\gamma_{t}+\gamma_{i}+\zeta_{i t}, \quad d_{i t}=1\left[d_{i t}^{*}>0\right],  \tag{B.2}\\
q_{i t} & =\alpha_{0}+Z_{t} \beta+X_{i t} \alpha_{1}+\alpha_{t}+\alpha_{i}+\epsilon_{i t} \quad \text { if } d_{i t}=1, \tag{B.3}
\end{align*}
$$

where $d_{i t}=1\left[d_{i t}^{*}>0\right]$ indicates worker $i$ 's choice of coming to work on day $t$, and $q_{i t}$ is her conditional output. Vector $Z_{t}$ consists of concurrent pollution and, more generally, lagged-day pollution levels. $W_{t}$ are concurrent weather covariates, namely, temperature, humidity and rain that may shift the worker's reservation utility. As explained in the text, $X_{i t}$ is a vector of observed worker-day output controls. To account for systematic seasonality and worker heterogeneity, both selection and output equations include time fixed effects, $\gamma_{t}$ and $\alpha_{t}$, and individual worker fixed effects, $\gamma_{i}$ and $\alpha_{i}$. $\zeta_{i t}$ and $\epsilon_{i t}$ are idiosyncratic error terms.

In a regression of (B.3) without the worker fixed effects $\alpha_{i}$, a sufficient condition for consistently estimating $\beta$, pollution's effect on output, using pooled OLS is $E\left[\alpha_{i}+\epsilon_{i t} \mid Z_{t}, X_{i t}, \alpha_{t}, d_{i t}=1\right]=0$ (e.g., Dustmann and Rochina-Barrachina, 2007). The OLS estimator will be biased downward, for example, if workers with higher ability $\alpha_{i}$ are more likely to call in absent when faced with high pollution exposure $Z_{t}$. Or say that more vulnerable workers are more likely to call in sick when $Z_{t}$ rises, and these workers are less productive than other workers, in which case the estimate for $\beta$ will biased upward.
The inclusion of worker fixed effects in (B.3) can remove the bias caused by selection on $\alpha_{i}$. To obtain a consistent estimate of $\beta$, a sufficient condition is (Wooldridge, 1995):

$$
E\left[\epsilon_{i t}-\epsilon_{i s} \mid Z_{t}, X_{i t}, \alpha_{t}, Z_{s}, X_{i s}, \alpha_{s}, d_{i t}=d_{i s}=1\right]=0 \text { for periods } s \neq t
$$

This condition will only be violated if there is selection on $\epsilon_{i t}$. For example, a

[^3]worker who experiences a positive output shock on day $t$ (high $\epsilon_{i t}$ ) is more likely to call in absent when exposed to high rather than low pollution. Further, the worker would need to observe $\epsilon_{i t}$ prior to making the work versus non-work choice. We can then use (B.2) to estimate the probability that a worker chooses work on each day, and correct for this probability in (B.3).

## C. Further results (more prolonged exposure)

Worker attendance. Panels in the third column of Figure A. 4 illustrate the relationship between attendance and cumulative pollution exposure (a $P$-day lag structure). The outcome variable is one if the worker produced output on a department work day, and zero otherwise. We consider a sample of worker by day observations in which the worker either produced output or her first day of zero output immediately following a day she worked. As in the concurrent pollution analysis of Table A.5, concurrent weather $W$ is allowed to shift attendance. We plot 2SLS estimates, ${ }^{5}$ using the full sample period for Henan. Point estimates of the effect on attendance from cumulative pollution exposure, as we raise $P$, are negative in Jiangsu and mostly positive in Henan, but in both cases they are at best marginally statistically significant and economically modest, peaking at $-1 \%$ and $+1 \%$, respectively, per $+10 \mu \mathrm{~g} / \mathrm{m}^{3}$ sustained increase in PM2.5. Effects for $\mathrm{SO}_{2}$ are higher than for PM2.5 as $\mathrm{SO}_{2}$ levels in $\mu \mathrm{g} / \mathrm{m}^{3}$ are lower.
Individual heterogeneity. We estimate output equation (1) allowing the response to prolonged pollution exposure to vary by worker $i$, i.e., write $\beta_{i}$ instead of $\beta$. The estimation sample, by site, includes the 25 workers with the most output observations. We interact each exposure variable $Z_{p, t}$ (24-hour PM2.5 or $\mathrm{SO}_{2}$ at daily lag $p$ of output day $t$ ) with a set of worker fixed effects for the 25 individuals. Within worker, exposure coefficients $\beta_{i, p}$ are constrained over $p$ according to her own $\operatorname{PDL}(P, 4)$. The Henan sample is the full April 2014 to May 2015 period. Panels in the second column of Figure A. 4 plot the 25th, 50th and 75th percentiles in the cross-worker distributions of 2SLS estimates for $\sum_{p=0}^{P} \hat{\beta}_{i, p}$ as we vary $P$. The median effect across workers is similar to the mean effect in panels (a) to (d) of Figure 6(I)-6(II); at the peak, the 75 th percentile can reach twice the median. Results for specific workers, of varying output sensitivity to pollution, are summarized in panels (e) and (f) of Figure 6(I)-6(II) (sensitivity is ranked based on the $P=20$ model).
Figure A. 5 further examines the individual worker response to sustained pollution exposure. Panels (a) and (b) show that a worker's rank in terms of output sensitivity to pollution exposure (" $q$ sens.") is similar whether we use PM2.5 or $\mathrm{SO}_{2}$. We also rank workers according to their mean output in the sample and, in panels (c) to (f), find that a worker's output rank tends to correlate positively with her output sensitivity to pollution. We further rank workers according to their attendance sensitivity to pollution (" $d$ sens."), similarly based on a PDL( 20,4 ), but

[^4]do not find an association with the worker's output sensitivity rank-see panels (g) to (j).

## REFERENCES

Dustmann, Christian, and María Engracia Rochina-Barrachina. 2007. "Selection correction in panel data models: An application to the estimation of females' wage equations." Econometrics Journal, 10(2): 263-293.

Shi, J., and M. Skuterud. 2015. "Gone fishing! Reported sickness absenteeism and the weather." Economic Enquiry, 53(1): 388-405.

Wooldridge, Jeffrey M. 1995. "Selection corrections for panel data models under conditional mean independence assumptions." Journal of Econometrics, 68(1): 115-132.

Table A.1-: Procedure to prepare the raw worker output data

| Procedure | Observations |  | No. of workers |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Dropped | Remaining | Dropped | Remaining |
| (I) Jiangsu |  |  |  |  |
| Initial number of observations (worker by day pairs) |  | 4579 |  | 47 |
| (-) One shift prior to/after dept. holiday (to clean/setup) | 32 | 4547 | 0 | 47 |
| (=) Number of observations for aggregate analysis |  | 4547 |  | 47 |
| (-) Observations with exceptional duties recorded | 47 | 4500 | 1 | 46 |
| (=) Number of observations for individual analysis |  | 4500 |  | 46 |
| Proportion of raw data in the final sample |  | 98\% |  | 98\% |
| (II) Henan |  |  |  |  |
| Initial number of observations (worker by day pairs) |  | 5614 |  | 83 |
| (-) One shift prior to/after dept. holiday (to clean/setup) | 124 | 5490 | 0 | 83 |
| (-) Shifts with power outage recorded | 52 | 5438 | 0 | 83 |
| (=) Number of observations for aggregate analysis |  | 5438 |  | 83 |
| (-) Observations with exceptional duties recorded | 131 | 5307 | 3 | 80 |
| (=) Number of observations for individual analysis |  | 5307 |  | 80 |
| Proportion of raw data in the final sample |  | 95\% |  | 96\% |

Table A. 2 -: Summary statistics for atmospheric ventilation condiTIONS, BY WORK SITE

| Variables | N | Mean | Std.Dev. | Min. | Max. |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Jiangsu |  |  |  |  |  |
| Wind speed at the surface (m/s) | 525 | 2.66 | 1.46 | 0.00 | 8.00 |
| Wind direction at the surface (all hours from a given direction=1) |  |  |  |  |  |
| ... from North | 525 | 0.20 | 0.36 | 0.00 | 1.00 |
| ... from East | 525 | 0.42 | 0.46 | 0.00 | 1.00 |
| ... from South | 525 | 0.12 | 0.30 | 0.00 | 1.00 |
| ... from West | 525 | 0.26 | 0.43 | 0.00 | 1.00 |
| Temperature difference ( ${ }^{\circ} \mathrm{C}$ ) for increasing altitudes at standard atmospheric pressure levels |  |  |  |  |  |
| ... from surface to 1000 mb | 525 | -0.43 | 1.75 | -3.20 | 9.00 |
| ... from 1000 to 925 mb | 525 | -3.50 | 2.15 | -6.60 | 8.00 |
| ... from 925 to 850 mb | 525 | -2.58 | 2.19 | -6.60 | 5.40 |
| ... from 850 to 700 mb | 525 | -5.50 | 3.66 | -12.90 | 5.60 |
| ... from 700 to 500 mb | 525 | -14.12 | 2.79 | -20.90 | -3.80 |
| Henan |  |  |  |  |  |
| Wind speed at the surface ( $\mathrm{m} / \mathrm{s}$ ) | 778 | 2.00 | 1.18 | 0.00 | 8.00 |
| Wind direction at the surface (all hours from a given direction=1) |  |  |  |  |  |
| ... from North | 784 | 0.12 | 0.28 | 0.00 | 1.00 |
| ... from East | 784 | 0.41 | 0.45 | 0.00 | 1.00 |
| ... from South | 784 | 0.25 | 0.38 | 0.00 | 1.00 |
| ... from West | 784 | 0.21 | 0.37 | 0.00 | 1.00 |
| Temperature difference ( ${ }^{\circ} \mathrm{C}$ ) for increasing altitudes at standard atmospheric pressure levels |  |  |  |  |  |
| ...from surface to 1000 mb | 778 | 0.35 | 1.30 | -2.00 | 8.90 |
| ...from 1000 to 925 mb | 784 | -3.18 | 2.25 | -6.40 | 7.80 |
| ...from 925 to 850 mb | 784 | -3.73 | 1.82 | -7.20 | 4.30 |
| ...from 850 to 700 mb | 784 | -8.29 | 3.48 | -15.60 | 3.00 |
| ...from 700 to 500 mb | 784 | -14.72 | 3.03 | -22.60 | -1.40 |

Note: An observation is a date-shift in the corresponding output sample. Sample periods are September 1, 2014 to May 31, 2015 (Jiangsu) and April 1, 2014 to May 31, 2015 (Henan).
A. 12

AMERICAN ECONOMIC JOURNAL



Table A. 3-: Output and concurrent pollution: Robustness to untrimmed individual-Level samples
Table A.4-: Multi-pollutant models (PM2.5 and $\mathrm{SO}_{2}$ ) of concurrent exposure and output

| Dependent variable is average output per worker | Jiangsu work site |  | Henan work site |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | (1) | (2) | (3) | (4) | (5) |
|  | OLS | 2SLS | OLS | 2SLS | 2SLS, neighbor poll. |
| $\overline{\text { PM2.5 (per } 10 \mu \mathrm{~g} / \mathrm{m}^{3} \text { ) }}$ | -0.01 | 0.07 | -0.03 | -0.12 | 0.00 |
|  | (0.04) | (0.08) | (0.13) | (0.17) | (0.13) |
| $\mathrm{SO}_{2}\left(\right.$ per $\left.10 \mu \mathrm{~g} / \mathrm{m}^{3}\right)$ | -0.17 | -0.47 | 0.14 | 0.24 | 0.40* |
|  | (0.13) | (0.39) | (0.16) | (0.18) | (0.24) |
| Combined effect 1: $+10 \mu \mathrm{~g} / \mathrm{m}^{3} \mathrm{PM} 2.5$ and $+10 \mu \mathrm{~g} / \mathrm{m}^{3} \mathrm{SO}_{2}$ | -0.18 | -0.40 | 0.11 | 0.12 | 0.40 |
|  | (0.13) | (0.33) | (0.11) | (0.09) | (0.25) |
| Combined effect 2: $+10 \mu \mathrm{~g} / \mathrm{m}^{3} \mathrm{PM} 2.5$ and $+5 \mu \mathrm{~g} / \mathrm{m}^{3} \mathrm{SO}_{2}$ | -0.09 | -0.17 | 0.04 | 0.00 | 0.20 |
|  | (0.08) | (0.14) | (0.09) | (0.10) | (0.16) |
| Date-shift characteristics | Yes | Yes | Yes | Yes | Yes |
| Observations | 524 | 524 | 258 | 258 | 756 |
| Number of regressors | 24 | 24 | 21 | 21 | 30 |
| First-stage F-statistic (excluded instruments) |  | 53 |  | 134 | 43 |
| R-squared | 0.85 | 0.85 | 0.76 | 0.76 | 0.77 |
| Mean value of dependent variable (cases or CNY worker ${ }^{-1}$ day $^{-1}$ ) | 111.4 | 111.4 | 119.6 | 119.6 | 127.4 |
| Dependent variable is individual worker output | Jiangsu work site |  | Henan work site |  |  |
|  | (6) | (7) | (8) | (9) | (10) |
|  | OLS | 2SLS | OLS | 2SLS | 2SLS, neighbor poll. |
| PM2.5 (per $10 \mu \mathrm{~g} / \mathrm{m}^{3}$ ) | 0.02 | 0.11 | -0.09 | -0.19 | -0.08 |
|  | (0.04) | (0.08) | (0.13) | (0.14) | (0.10) |
| $\mathrm{SO}_{2}\left(\right.$ per $\left.10 \mu \mathrm{~g} / \mathrm{m}^{3}\right)$ | -0.08 | -0.45 | 0.16 | 0.14 | 0.26 |
|  | (0.10) | (0.37) | (0.13) | (0.16) | (0.21) |
| Combined effect 1: $+10 \mu \mathrm{~g} / \mathrm{m}^{3} \mathrm{PM} 2.5$ and $+10 \mu \mathrm{~g} / \mathrm{m}^{3} \mathrm{SO}_{2}$ | -0.06 | -0.34 | 0.07 | -0.05 | 0.18 |
|  | (0.10) | (0.32) | (0.10) | (0.14) | (0.20) |
| Combined effect 2: $+10 \mu \mathrm{~g} / \mathrm{m}^{3}$ PM2.5 and $+5 \mu \mathrm{~g} / \mathrm{m}^{3} \mathrm{SO}_{2}$ | -0.02 | -0.11 | -0.01 | -0.12 | 0.05 |
|  | (0.06) | (0.14) | (0.10) | (0.12) | (0.12) |
| Worker fixed effects | Yes | Yes | Yes | Yes | Yes |
| Worker-date-shift characteristics | Yes | Yes | Yes | Yes | Yes |
| Observations | 4,443 | 4,443 | 2,085 | 2,085 | 5,054 |
| Number of regressors | 70 | 70 | 65 | 65 | 106 |
| First-stage F-statistic (excluded instruments) |  | 55 |  | 186 | 46 |
| R-squared | 0.83 | 0.83 | 0.79 | 0.79 | 0.70 |
| Mean value of dependent variable (cases or CNY worker ${ }^{-1}$ day $^{-1}$ ) | 112.4 | 112.4 | 119.6 | 119.6 | 126.6 |
| Year-month fixed effects | Yes | Yes | Yes | Yes | Yes |
| Day-of-week fixed effects | Yes | Yes | Yes | Yes | Yes |
| Time-of-day fixed effects | Yes | Yes | Yes | Yes | Yes |

$\overline{\text { Note: }}$ The dependent variable is average output per worker in columns (1) to (5), or a worker's output in columns (6) to (10). An observation is a date-shift (top panel, or a worker by date-shift with non-zero output (bottom panel; we trim the sample at three standard deviations of the sample mean). Sample periods are September 1, 2014 to May 31, 2015 (Jiangsu), January 1, 2015 to May 31, 2015 (Henan) or April 1, 2014 to May 31 , 2015 (Henan with Table 3(II) (Henan) for date-shift characteristics and worker-date-shift characteristics. OLS estimates or 2SLS estimates, where measured pollution, $Z$, is instrumented using pollution predicted from atmospheric ventilation conditions, $\hat{Z}$ (see note 20). To reflect common emission sources, we report an
 at $1 \%,{ }^{* *}$ at $5 \%,{ }^{*}$ at $10 \%$.



 the worker either produced output or her first day of leave excluding department holidays. Sample periods are September 1,2014 to May 31,2015 in the

 Observations
Number of reg PM2.5 $\geq 150$ or $\mathrm{SO}_{2} \geq 75 \mu \mathrm{~g} / \mathrm{m}^{3}$ (yes=1, previous day)
PM2.5 or $\mathrm{SO}_{2}$ (per $10 \mu \mathrm{~g} / \mathrm{m}^{3}$, previous day) PM2.5 $\in[100,150)$ or $\mathrm{SO}_{2} \in[50,75) \mu \mathrm{g} / \mathrm{m}^{3}$ (yes $=1$, previous day)
PM2.5 $\geq 150$ or $\mathrm{SO}_{2} \geq 75 \mu \mathrm{~g} / \mathrm{m}^{3}$ (yes=1, previous day)
PM2.5 $\in[50,100)$ or $\mathrm{SO}_{2} \in[25,50) \mu \mathrm{g} / \mathrm{m}^{3}$ (yes $=1$, previous day)

| Jiangsu work site | OLS categorical |  | OLS linear |  | 2SLS linear |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Dependent variable is individual works on day (yes=1) | (1) | (2) | (3) | (4) | (5) | (6) |  |  |
|  | PM2.5 | $\mathrm{SO}_{2}$ | PM2.5 | $\mathrm{SO}_{2}$ | PM2.5 | $\mathrm{SO}_{2}$ |  |  |
| $\overline{\text { PM2.5 } \in[50,100) ~ \text { or } \mathrm{SO}_{2} \in[25,50) ~} \mathrm{mg} / \mathrm{m}^{3}$ (yes $=1$, previous day) | -0.02** | -0.01 |  |  |  |  |  |  |
|  | (0.01) | (0.01) |  |  |  |  |  |  |
| PM2.5 $\geq 100$ or $\mathrm{SO}_{2} \geq 50 \mu \mathrm{~g} / \mathrm{m}^{3}$ (yes $=1$, previous day) | $\begin{aligned} & -0.01 \\ & (0.01) \end{aligned}$ | $\begin{gathered} -0.01^{* *} \\ (0.01) \end{gathered}$ |  |  |  |  |  |  |
| PM2.5 or $\mathrm{SO}_{2}$ (per $10 \mu \mathrm{~g} / \mathrm{m}^{3}$, previous day) |  |  | $\begin{gathered} -0.00 \\ (0.00) \end{gathered}$ | $\begin{aligned} & -0.00 \\ & (0.00) \end{aligned}$ | $\begin{aligned} & -0.00 \\ & (0.00) \end{aligned}$ | $\begin{gathered} 0.00 \\ (0.01) \end{gathered}$ |  |  |
| Indicator for an initial two months at the firm | Yes | Yes | Yes | Yes | Yes | Yes |  |  |
| Observations | 5,671 | 5,671 | 5,671 | 5,671 | 5,671 | 5,671 |  |  |
| Number of regressors | 88 | 88 | 87 | 87 | 87 | 87 |  |  |
| First-stage F-statistic (excluded instruments) |  |  |  |  | 159 | 155 |  |  |
| R-squared | 0.03 | 0.03 | 0.03 | 0.03 | 0.03 | 0.03 |  |  |
| Mean value of dependent variable (works on day $=1$ ) | 0.79 | 0.79 | 0.79 | 0.79 | 0.79 | 0.79 |  |  |
| Henan work site | OLS ca | gorical | OLS | near | 2SLS | near | 2SLS | ghbor poll. |
| Dependent variable is individual works on day (yes=1) | (7) | (8) | (9) | (10) | (11) | (12) | (13) | (14) |
|  | PM2.5 | $\mathrm{SO}_{2}$ | PM2.5 | $\mathrm{SO}_{2}$ | PM2.5 | $\mathrm{SO}_{2}$ | PM2.5 | $\mathrm{SO}_{2}$ |

Table A.5-: The relationship Between attendance and immediate exposure to pollution

Table A.6-: Robustness to AdDing A SElECTION CORRECTION IN THE OUTPUT EQUATION

| Lag structure$(P)$ | Jiangsu |  | Henan |  | Henan, neighbor poll. |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | PM2.5 | $\mathrm{SO}_{2}$ | PM2.5 | $\mathrm{SO}_{2}$ | PM2.5 | $\mathrm{SO}_{2}$ |
| 5 | $\begin{aligned} & -0.06 \\ & (0.14) \end{aligned}$ | $\begin{gathered} -0.55^{* *} \\ (0.26) \end{gathered}$ | $\begin{gathered} 0.02 \\ (0.23) \end{gathered}$ | $\begin{gathered} 0.33 \\ (0.33) \end{gathered}$ | $\begin{gathered} -0.09 \\ (0.14) \end{gathered}$ | $\begin{gathered} 0.35 \\ (0.33) \end{gathered}$ |
| 10 | $\begin{aligned} & -0.19 \\ & (0.19) \end{aligned}$ | $\begin{aligned} & -0.19 \\ & (0.29) \end{aligned}$ | $\begin{gathered} -0.66^{* *} \\ (0.33) \end{gathered}$ | $\begin{gathered} 0.46 \\ (0.51) \end{gathered}$ | $\begin{aligned} & -0.17 \\ & (0.18) \end{aligned}$ | $\begin{aligned} & -0.39 \\ & (0.44) \end{aligned}$ |
| 15 | $\begin{aligned} & -0.30 \\ & (0.22) \end{aligned}$ | $\begin{aligned} & -0.11 \\ & (0.27) \end{aligned}$ | $\begin{aligned} & -0.75 \\ & (0.57) \end{aligned}$ | $\begin{aligned} & -0.70 \\ & (0.81) \end{aligned}$ | $\begin{aligned} & -0.05 \\ & (0.20) \end{aligned}$ | $\begin{gathered} -1.05^{* *} \\ (0.52) \end{gathered}$ |
| 20 | $\begin{gathered} -0.47^{* *} \\ (0.22) \end{gathered}$ | $\begin{aligned} & -0.12 \\ & (0.40) \end{aligned}$ | $\begin{aligned} & -1.24 \\ & (0.80) \end{aligned}$ | $\begin{gathered} -0.70 \\ (1.07) \end{gathered}$ | $\begin{aligned} & -0.29 \\ & (0.19) \end{aligned}$ | $\begin{gathered} -2.08^{* * *} \\ (0.65) \end{gathered}$ |
| 25 | $\begin{gathered} -0.62^{* * *} \\ (0.21) \end{gathered}$ | $\begin{gathered} -1.27^{* *} \\ (0.54) \end{gathered}$ | $\begin{gathered} -1.26 \\ (1.03) \end{gathered}$ | $\begin{aligned} & -1.83 \\ & (1.39) \end{aligned}$ | $\begin{gathered} -0.44^{*} \\ (0.27) \end{gathered}$ | $\begin{gathered} -2.26^{* * *} \\ (0.63) \end{gathered}$ |
| 30 | $\begin{gathered} -0.34^{*} \\ (0.20) \end{gathered}$ | $\begin{gathered} -1.23^{* *} \\ (0.50) \end{gathered}$ | $\begin{aligned} & -1.97^{*} \\ & (1.06) \end{aligned}$ | $\begin{gathered} -4.89^{* * *} \\ (1.54) \end{gathered}$ | $\begin{aligned} & -0.67^{*} \\ & (0.35) \end{aligned}$ | $\begin{gathered} -2.53^{* * *} \\ (0.71) \end{gathered}$ |

Note: 2SLS estimates of $\operatorname{PDL}(P, 4)$ output equations that include a selection correction imputed from a UDL $(P)$ probit attendance regression, where $P$ is the number of lags in days. Each row and column reports the coefficient and standard error (in parentheses, clustered on day) on the cumulative effect on worker output from concurrent and lagged pollution exposure, $\sum_{p=0}^{P} \beta_{p}$, estimated from a separate quartic distributed lag model, and pollution variables are daily 24 -hour means. Sample periods to May 31, 2015 start on September 1, 2014 (Jiangsu), January 1, 2015 (Henan with host-town pollution) or April 1, 2014 (Henan with neighboring-city pollution). As in Table 4, we report estimates as a proportion of mean worker output. ${ }^{* * *}$ Significant at $1 \%,{ }^{* *}$ at $5 \%$, *at $10 \%$.

figure A.1. : Number of 12-hour shifts worked per 30-DAy Period
Notes: Frequency chart across workers (46 in Jiangsu, 80 in Henan) of the number of 12-hour shifts worked per 30 -day period of department activity. This is calculated as a worker's number of positive output records (work days) in the sample divided by the number of days with department activity between the worker's first and last work day in the sample, multiplied by 30 days (i.e., per 30-day period).

figure A.2. : Variable compensation scheme in Jiangsu

Notes: A Jiangsu department worker's output during a 12-hour shift is aggregated across individual products into "adjusted cases" using the weights listed in the Appendix text, and then converted into "counts" at a rate of 105 counts/adjusted case. The worker is paid a bonus of: 0.01 CNY per count in excess of 11,000 counts but no greater than 12,000 counts, 0.02 CNY per count in excess of 12,000 counts but no greater than 13,000 counts, and 0.03 CNY per count for output in excess of 13,000 counts. The penalty scheme for output that falls short of 11,000 counts is symmetric to the bonus scheme around the reference output of 11,000 counts.

(b) Henan worker sample
figure A.3. : A worker's shift-To-shift own-Product transitions

Notes: To prepare each plot, we take the five and four varieties in each work site with double-digit output shares, and aggregate the remaining varieties into a residual category labeled "Other." We consider pairs of consecutive work shifts by a worker, with 12 hours and up to 24 hours of rest in between work in Jiangsu and Henan, respectively, given the standard shift designs. We then compute the empirical probability that, conditional on a worker producing a given variety on a shift, she also produced that variety in her preceding shift.

figure A.4. : The cumulative impact of exposure to PM2.5 or $\mathrm{SO}_{2}$
Notes: Robustness tests of the cumulative impact on worker supply of $+10 \mu \mathrm{~g} / \mathrm{m}^{3}$ exposure to pollution (PM2.5 or $\mathrm{SO}_{2}$ ) on each of the concurrent and previous $P$ days. Panels in the first column report the output effect estimated by OLS (expressed as a proportion of mean output in the respective sample). Panels in the second column report individual output effects (25th, 50th and 75 th percentiles). Panels in the third column report the attendance effect (in percentage points). In the two last columns, 2SLS uses $\hat{Z}$ to instrument for $Z$. Each value reports the point estimate and $95 \%$ CI on $\sum_{p=0}^{P} \beta_{p}$ from a different quartic distributed lag model, $\operatorname{PDL}(P, 4)$, as we raise $P$ along the horizontal axis.


## figure A.5. : Individual heterogeneity

Notes: In each panel, an observation is a worker in the sample of 25 workers with most output observations at the given work site. Workers are ranked (i.e., $1,2, \ldots, 25$ ) according to their mean output in the sample ("output rank"), and their output sensitivity (" $q$ sens.") or their attendance sensitivity (" $d$ sens.") to cumulative pollution exposure, according to 20-day quartic distributed lag models, PDL(20, 4), estimated by 2 SLS .

figure A.6. : Forms of autocorrelation
Notes: Robustness to allowing for alternative forms of autocorrelation in the analysis of prolonged pollution exposure. 2SLS output regressions implemented on the pooled sample that combines observations from both Jiangsu and Henan sites, with $\hat{Z}$ instrumenting for $Z$. As in Figure 7, we show the cumulative impact on output of $+10 \mu \mathrm{~g} / \mathrm{m}^{3}$ exposure to a pollutant ( PM 2.5 or $\mathrm{SO}_{2}$ ) on each of the concurrent and previous $P$ days. Panels (a) (PM2.5) and (b) $\left(\mathrm{SO}_{2}\right)$ compute standard errors using two-way clusters on day and on worker. Panels (c) (PM2.5) and (d) ( $\mathrm{SO}_{2}$ ) compute standard errors using day clusters (within-day correlation across workers) and an autocorrelation bandwidth of 30 days (correlation over time). Other notes to Figure 7 apply here.


Figure A.7. : Unconstrained Distributed lag
Notes: Robustness to removing the smoothness constraint on the lagged pollution coefficients, $\beta_{p}$ : Unconstrained distributed lag, UDL $(P)$. 2SLS output regressions implemented separately by site (full output sample with neighboring-city pollution in the case of Henan), using $\hat{Z}$ to instrument for $Z$. As in Figure 6(I)-6(II), we show the cumulative impact on output of $+10 \mu \mathrm{~g} / \mathrm{m}^{3}$ exposure to a pollutant (PM2.5 or $\mathrm{SO}_{2}$ ) on each of the concurrent and previous $P$ days. Each value reports the point estimate and $95 \%$ CI on $\sum_{p=0}^{P} \beta_{p}$ from a different UDL $(P)$ model, as we raise $P$ along the horizontal axis. We report the effect on output as a proportion of mean output in the respective sample.


Figure A.8. : Quadratic distributed lag
Notes: Robustness to constraining the lagged pollution coefficients, $\beta_{p}$, further: Quadratic distributed lag, $\operatorname{PDL}(P, 2)$. 2SLS output regressions implemented separately by site (full output sample with neighboring-city pollution in the case of Henan), using $\hat{Z}$ to instrument for $Z$. As in Figure 6(I)-6(II), we show the cumulative impact on output of $+10 \mu \mathrm{~g} / \mathrm{m}^{3}$ exposure to a pollutant ( PM 2.5 or $\mathrm{SO}_{2}$ ) on each of the concurrent and previous $P$ days. Each value reports the point estimate and $95 \% \mathrm{CI}$ on $\sum_{p=0}^{P} \beta_{p}$ from a different $\operatorname{PDL}(P, 2)$ model, as we raise $P$ along the horizontal axis. We report the effect on output as a proportion of mean output in the respective sample.

figure A.9. : Output sample untrimmed
Notes: Robustness to not trimming the output sample at three standard deviations of the sample mean. 2SLS output regressions implemented separately by site (full output sample with neighboring-city pollution in the case of Henan), using $\hat{Z}$ to instrument for $Z$. As in Figure 6(I)-6(II), we show the cumulative impact on output of $+10 \mu \mathrm{~g} / \mathrm{m}^{3}$ exposure to a pollutant ( PM 2.5 or $\mathrm{SO}_{2}$ ) on each of the concurrent and previous $P$ days. Each value reports the point estimate and $95 \% \mathrm{CI}$ on $\sum_{p=0}^{P} \beta_{p}$ from a different quartic distributed lag model, $\operatorname{PDL}(P, 4)$, as we raise $P$ along the horizontal axis. We report the effect on output as a proportion of mean output in the respective sample.

figure A.10. : SEasonality controls
Notes: Robustness to controls for seasonality: Replacing year-month fixed effects by a quadratic time trend coupled with indicators for winter months, November to February, during which pollution tends to be most severe. 2SLS output regressions implemented separately by site (full output sample with neighboring-city pollution in the case of Henan), using $\hat{Z}$ to instrument for $Z$. As in Figure 6(I)-6(II), we show the cumulative impact on output of $+10 \mu \mathrm{~g} / \mathrm{m}^{3}$ exposure to a pollutant ( PM 2.5 or $\mathrm{SO}_{2}$ ) on each of the concurrent and previous $P$ days. Each value reports the point estimate and $95 \% \mathrm{CI}$ on $\sum_{p=0}^{P} \beta_{p}$ from a different quartic distributed lag model, $\operatorname{PDL}(P, 4)$, as we raise $P$ along the horizontal axis. We report the effect on output as a proportion of mean output in the respective sample.

figure A.11. : Exogenous ventilation
Notes: Robustness to the set of exogenous ventilation conditions, in which we drop wind direction from the ventilation conditions $V$ that are used to form an instrument $\hat{Z}$ for measured pollution $Z$. 2SLS output regressions implemented separately by site (full output sample with neighboring-city pollution in the case of Henan). As in Figure 6(I)-6(II), we show the cumulative impact on output of $+10 \mu \mathrm{~g} / \mathrm{m}^{3}$ exposure to a pollutant (PM2.5 or $\mathrm{SO}_{2}$ ) on each of the concurrent and previous $P$ days. Each value reports the point estimate and $95 \%$ CI on $\sum_{p=0}^{P} \beta_{p}$ from a different quartic distributed lag model, $\operatorname{PDL}(P, 4)$, as we raise $P$ along the horizontal axis. We report the effect on output as a proportion of mean output in the respective sample.

figure A.12. : The logarithm of output

Notes: Robustness to specifying the dependent variable, individual worker output, in logarithms instead of levels. 2SLS output regressions implemented separately by site (full output sample with neighboringcity pollution in the case of Henan), using $\hat{Z}$ to instrument for $Z$. As in Figure 6(I)-6(II), we show the cumulative impact on output of $+10 \mu \mathrm{~g} / \mathrm{m}^{3}$ exposure to a pollutant (PM2.5 or SO 2 ) on each of the concurrent and previous $P$ days. Each value reports the point estimate and $95 \% \mathrm{CI}$ on $\sum_{p=0}^{P} \beta_{p}$ from a different quartic distributed lag model, $\operatorname{PDL}(P, 4)$, as we raise $P$ along the horizontal axis. We report the effect on output in log points $(\times 100)$.


## figure A.13. : Multi-Pollutant models

Notes: Multi-pollutant models (PM2.5 and $\mathrm{SO}_{2}$ ) of more prolonged exposure and output. We implement each $\operatorname{PDL}(P, 4)$ output regression on the pooled sample that combines observations from both Jiangsu and Henan sites, with $\hat{Z}$ (PM2.5 and $\mathrm{SO}_{2}$ ) instrumenting for $Z$ (both PM2.5 and $\mathrm{SO}_{2}$ in the regression equation). Panel (a) shows the cumulative impact on output (point estimate and $95 \%$ CI) of combined doses of $+10 \mu \mathrm{~g} / \mathrm{m}^{3}$ PM2.5 and $+10 \mu \mathrm{~g} / \mathrm{m}^{3} \mathrm{SO}_{2}$ on each of the concurrent and previous $P$ days. To reflect common emission sources and a mean PM2.5 to $\mathrm{SO}_{2}$ concentration ratio of about 2, panel (b) shows the cumulative impact on output of combined doses of $+10 \mu \mathrm{~g} / \mathrm{m}^{3} \mathrm{PM} 2.5$ and $+5 \mu \mathrm{~g} / \mathrm{m}^{3} \mathrm{SO}_{2}$ on each of the concurrent and previous $P$ days. Other notes to Figure 7 (for single-pollutant models) apply here.


[^0]:    ${ }^{1}$ We ignore the density at 30 days for Henan, as this consists of 15 temporary workers who appear briefly in the records, with positive output for up to three consecutive days only (i.e., 3 days with output / 3 days from worker's first day to last day in sample $\times 30$ days in a 30 -day period).

[^1]:    ${ }^{2}$ We acquired local pollutant concentrations from the Chinese Ministry of Environmental Protection, via http://www.pm25.in. PM2.5 concentrations from the US State Department for its regional US embassies, which we describe in the text, are available via www.stateair.net/web/post/1/1.html.

[^2]:    ${ }^{3}$ NASA and NOAA data are available via http://daac.gsfc.nasa.gov/ and http://www1.ncdc. noaa.gov/pub/data/igra/.

[^3]:    ${ }^{4}$ Since the work shift starts at a predetermined time for the worker's team and is of fixed 12-hour duration, we ignore these additional margins of labor supply.

[^4]:    ${ }^{5}$ Atmospheric ventilation $V$ influences pollution but is excluded from the attendance equation (B.2).

