# Knowledge Capital and Aggregate Income Differences: Development Accounting for U.S. States 

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## American Economic Journal: Macroeconomics

Online Appendix

## Appendix A: Construction of Years of Schooling Measures by State

We compile average years of educational attainment for each U.S. state from the Integrated Public Use Microdata Series (IPUMS) data of the Minnesota Population Center (Ruggles et al. (2010)). We concentrate on the working-age population between 20 and 65 years. We also drop all respondents who are still in school at the time of the survey.

For the years 1970 to 2000, we use the 1 percent (1970) and 5 percent (1980, 1990, and 2000) random samples of the American population. The 1 percent sample has about 4 million observations, the 5 percent samples have about 13 to 14 million observations. Beginning in the year 2001, we use census data from the American Community Survey (ACS). The ACS provides annual 1 percent random population samples (with smaller sample sizes between 2001 and 2004). The approximate sample size is 3 million observations each year. Survey weights in the census and the ACS allow us to calculate measures that are representative for the U.S. population.

Until 1980, the Census reported directly the years of schooling or highest grade level completed of each individual. Beginning with the 1990 Census, the Census Bureau has changed the coding of educational categories and reports degrees (Bachelor, Master, etc.) instead. To translate the degree information into years of schooling, we use the estimates of average years of schooling of each degree provided by Jaeger (1997). ${ }^{1}$

Substantial differences in the labor-market performance between GED holders and standard high school graduates (Heckman, Humphries, and Mader (2011)) warrant a special treatment of GED holders. Due to the weak labor-market position of GED holders, we assign them 10 rather than 12 years of schooling.

Only the most recent survey waves identify GED holders in the Census data. We therefore estimate a constant share of GED holders among all high-school graduates from the pooled ACS 2008-2010 samples. The pooled sample is restricted for each year to get approximately the same age cohort of people aged 20-65. For example, for the year 2007, we use all people aged 21-66 in ACS 2008, 22-67 in ACS 2009, and 23-68 in 2010; for the year 1990, we use all people aged 3883 in ACS 2008, 39-84 in ACS 2009, and 40-85 in ACS 2010. Note that 1940 is not adjusted because the GED was introduced in 1942.

[^0]Overall, the GED adjustment affects the average years of schooling only very little, though. In 2007, for example, 15 percent of those who would have received 12 years of schooling otherwise are now assigned 10 years of schooling, reducing the mean of the average years of schooling from 12.33 to 12.27 years. Put differently, accounting for GED holders raises the mean share of those with less than 12 years of schooling from 22.6 percent to 26.7 percent.

Having computed the years of schooling of each individual $i$, the average years of schooling $S$ in state $s$ at time $t$ is then given by combining individual years of schooling by the weighted share of individuals $i$ with education level $e$ in the state at the time:

$$
\begin{equation*}
S_{s t}=\sum_{e} \frac{\sum_{i} \text { person weights }_{\text {iest }}}{\sum_{i} \text { person weights }_{\text {ist }}} * \text { years of schooling }{ }_{e} \tag{A1}
\end{equation*}
$$

This yields the average years of schooling by state over time as shown in Figure 2.

## Appendix B: Construction of Test Score Measures by State

As indicated in section 2.3 of the main text, our construction of cognitive skill measures for each U.S. state proceeds in four steps. This appendix provides methodological details on each step. First, we construct a constant measure of the mean test scores of students of each state (Appendix B.1). Second, we adjust the test scores of the working-age population of each state for interstate migration, thereby placing particular emphasis on the fact that interstate migration is selective (Appendix B.2). Third, test scores are adjusted for immigration from other countries, again with a special focus on selectivity (Appendix B.3). Fourth, we project test scores backward in time to allow for age-varying test scores in each state (Appendix B.4).

## B. 1 Construction of Mean State Test Scores

The National Assessment of Educational Progress (NAEP) studies the educational achievement of American students in grades four and eight in different subjects (National Center for Education Statistics (2014)). In our main analysis, we focus on the mathematics score in grade eight, on which we focus the following description. But as far as possible, we also computed test scores based on reading and grade four, as well as on a combination of subjects and grades.

Since 1990, NAEP math tests have been administered on a representative scale at the state level every two to four years for most states. By 2003, test scores are available for all states.

## Adjustment of Pre-1996 Tests for Accommodation

Since 1996, NAEP allows students with disabilities and English language learners specific accommodations to facilitate test participation. The NAEP test scores before 1996 (in 1990 and 1992) did not permit such accommodation, so that they have to be adjusted in order to be on a common scale with the subsequent tests. Therefore, we rescale the pre-1996 tests as follows: For 1996, NAEP test scores and standard deviations are available for tests with and without accommodation at the national level. By subtracting the 1996 U.S. mean without accommodation from the state score and dividing by the 1996 U.S. standard deviation without accommodation, we standardize test scores to mean 0 and standard deviation of 1 . By multiplying the 1996 U.S. standard deviation with accommodation and adding the 1996 U.S. mean with accommodation, we bring each test score before 1996 to the same scale as the tests that permitted accommodation.

That is, the pre-1996 waves are aligned to the post-1996 scale in the following way:

$$
\begin{equation*}
\operatorname{score}_{s t}^{\text {adj }}=\left(\frac{\text { score }_{s t}-\text { mean }_{U S, t=1996}^{\text {same scale }}}{s_{U S, t=1996}^{\text {samescale }}}\right) * s d_{U S, t=1996}^{\text {new scale }}+\operatorname{mean}_{U S, t=1996}^{\text {new scale }} \tag{B1}
\end{equation*}
$$

where score $_{s t}$ is the raw score (without accommodation) of state $s$ at time $t$, mean refers to the U.S. national mean, sd refers to the U.S. standard deviation, same scale refers to scores without accommodation, and new scale refers to scores with accommodation.

## Normalization of Scales to Base Year 2011

Next, we normalize each scale - eight-grade math, etc. - to have a mean of 500 and a standard deviation of 100 in the common base year 2011. This is done by subtracting from each test score the 2011 U.S. mean and dividing by the 2011 U.S. standard deviation and then multiplying by 100 and adding 500:

$$
\begin{equation*}
\text { score }_{s t}^{s t a n d a r d}=\left(\frac{\text { score }_{s t}^{a d j}-\text { mean }_{U S, t=2011}}{s d_{U S, t=2011}}\right) * 100+500 \tag{B2}
\end{equation*}
$$

## Regression-based Estimation of Mean State Scores by State Fixed Effects

Using the normalized scores, we estimate the average test score of each state over all test scores that are available until 2011. This is done by estimating state fixed effects in a regression
with year fixed effects that take into account systematic differences over time, as well as - in estimations that combine tests across subjects and grades - grade-by-subject fixed effects that takes into account systematic differences between grades and subjects:

$$
\begin{equation*}
\text { scores }_{\text {sgut }}^{\text {standard }}=\sum_{s=1}^{50} \alpha_{s} I_{s}+I_{g} * I_{u}+I_{t}+\epsilon_{\text {sgut }} \tag{B3}
\end{equation*}
$$

$I_{s}$ is the fixed effect of state $s$ that we are interested in. $I_{t}$ are time fixed effects and $I_{g} * I_{u}$ are grade-by-subject fixed effects. By leaving out the indicators that represent math, grade eight, and the year 2011, all state fixed effects refer to this subject, grade, and year. The same adjustments and estimations can also be performed for different subsamples of the population, e.g., by education category of the parents. In further analysis, we estimate average standard deviations by employing the same fixed effects regression framework. ${ }^{2}$

## B. 2 Adjustment for Interstate Migration

## Adjusting for State of Birth

To be able to adjust the state skill measure for interstate migration, we start by computing the birthplace composition of each state from the Census data. In particular, we compute the population shares of people currently living in state $s$ who were born in state $s$ ("state locals"), those born in in another state $k$ ("interstate migrants"), and those born in another country ("international immigrants"). Thus, the population share of individuals $i$ from origin state/country o living in state $s$ at time $t$ is given by

$$
\begin{equation*}
\text { population share }_{\text {ost }}=\frac{\sum_{i} \text { person weights }_{\text {iost }}}{\sum_{i} \text { person weights }_{\text {ist }}} \tag{B4}
\end{equation*}
$$

Each state is composed of individuals educated in other states. To adjust, at least partially, for the differences in schooling that these individuals brought with them to their current state of residence, we construct a series of composite test scores. The idea is that each person who is living in a state receives the test score of his home state. The baseline composite test score of state $s$ at time $t$ is then the weighted sum of test scores from all origin states $o$ which are weighted by the fraction of people born in a particular origin $o$ living in state $s$ at time $t$ :

[^1]\[

$$
\begin{equation*}
\text { score }_{s t}^{\text {adj }}=\sum_{o} \text { population share }{ }_{\text {ost }} \times \text { score }_{o} \tag{B5}
\end{equation*}
$$

\]

Thus, each person currently living in a state is assigned the test score from the respective state of birth.

The baseline composite test score thus assigns all locals the mean test score of the state of residence which is also their state of birth, assuming that the locals have not moved during their school career to another state. Assuming that internal migrants have not left their state of birth before finishing grade eight, all internal migrants receive the mean test score of their state of birth. In this variant, the international immigrants receive the mean score of their current state of residence.

## Adjusting for Selective Interstate Migration based on Educational Background

To address selective interstate migration, we compute all population shares separately by educational background. We distinguish two educational categories: Persons with (at least some) university education and persons without university education. For each state, we also construct separate test scores by the education category of the parents (some university education or not).

We then assign separate test scores by educational background $e$ :

$$
\begin{equation*}
\text { score }_{\text {st }}^{\text {sel }}=\sum_{o e} \text { population share }_{o e s t} \times \text { score }_{o e} \tag{B6}
\end{equation*}
$$

For state locals, this adjusted score replaces the average test score of the state of residence with the average test score of the state of residence by education category (university / no university). Likewise, for in-migrants it adjusts the average test scores of by education category. The assumption is that we can assign the population with a university education the test score of children with parents who have a university degree, and equivalently for those without a university education.

## B. 3 Adjustment for International Migration

Our adjustment for international migration combines data from international achievement tests with population shares of immigrants from different countries of origin.

## International Test Score Data

We use international test score data from PISA, TIMSS, and PIRLS for international immigrants residing in one of the U.S. states. ${ }^{3}$ As a first step, the international test data have to be rescaled onto a common scale with the national NAEP data (Hanushek, Peterson, and Woessmann (2012)). To do so, we first standardize all international test scores by subtracting from each mean score on the international scale the U.S. mean value on the international scale by subject, grade, and year and divide this difference by the U.S. standard deviation on the international scale, also by subject, grade, and year. Next, we multiply the standardized value by the U.S. standard deviation of the NAEP score by subject, grade, and year and add the U.S. mean of the NAEP score by subject, grade, and year:
where score $_{\text {sgut }}$ is the raw international test score of country $s$ at grade $g$ in subject $u$ in year $t$.

To compute average test scores for each country, we proceed in the same way as for the national test data. The regression design takes into account systematic differences between grades, subjects, and years. The final estimate of the country average test score is then a country fixed effect:

$$
\begin{equation*}
\text { scores }_{\text {sgut }}^{\text {standard }}=\sum_{s} \alpha_{s} I_{s}+I_{g} * I_{u} * I_{\text {test }}+I_{t}+\epsilon_{\text {sgut }} \tag{B8}
\end{equation*}
$$

where $I_{s}$ is the fixed effect of country $s$ that we are interested in. $I_{t}$ are time fixed effects and $I_{g} * I_{u} * I_{\text {test }}$ are grade times subject times survey fixed effects. The survey fixed effects indicate whether we identify grade 4 in PIRLS or grade 4 in TIMSS. Thus, they are dummy variables for TIMSS, PIRLS, and PISA. Again, the same regression can be estimated for different subsamples of the population. ${ }^{4}$

[^2]Apart from the mean test score, we also estimate the performance of the $75^{\text {th }}$ and the $90^{\text {th }}$ percentile of students in each country for comparison. We also estimate the standard deviation. ${ }^{5}$

In cases where a source country did not participate in the international achievement tests, we impute values from neighboring countries or regions. Table A5 reports the respective imputations for the main source countries of immigrants in the United States.

## Population Shares of Immigrants from Different Countries of Origin

Using Census data, we next calculate the population shares of those born outside U.S. Table A5 shows the main source countries of immigrants who came to the United States over the last 70 years.

In calculating the share of immigrants from different origin countries in the birthplace composition of each state, we take into account the age of immigration. In particular, immigrants arriving in the United States before the age of 6 are assumed to have spent their school career in the U.S. school system, so they are assigned the NAEP score of their state of residence. Those who immigrated after the age of 20 are assigned the test score of their country of origin. And those who immigrated between ages 6 and 20 are assigned a weighted average of the two.

Using the population shares of immigrants from different countries of origin as in equation (B4), we then basically proceed in the same way as with the national test score data. That is, we adjust the composite test score of each state by applying the country-of-origin test scores for international immigrants.

As discussed in the paper, we do not use the average migrant test score because the skills of migrants are not random draws from the home country skill distribution. To estimate the migrant selectivity for each country, we proceed in two steps: First, for each country of origin (country subscripts omitted), we calculate the selectivity parameter for school attainment as the percentile $p$ of the home country distribution from which the average immigrant to the U.S. is drawn:

$$
\begin{equation*}
p=s_{U S}^{p r i} * \frac{1}{2} s_{\text {home }}^{\text {pri }}+s_{U S}^{\text {sec }} *\left(s_{\text {home }}^{\text {pri }}+\frac{1}{2} s_{\text {home }}^{\text {sec }}\right)+s_{U S}^{\text {ter }} *\left(s_{\text {home }}^{\text {pri }}+s_{\text {home }}^{\text {sec }}+\frac{1}{2} s_{\text {home }}^{\text {ter }}\right) \tag{B9}
\end{equation*}
$$

where the respective educational degrees of the population are given by pri= primary, sec = secondary, and ter = tertiary, $s$ refers to the shares of the population with the respective degrees (with $s^{p r i}+s^{s e c}+s^{t e r}=1$ ), home refers to the population in the respective home country, and US

[^3]refers to the immigrants from the specific home country living in the United States. Our baseline estimate uses the attainment selection parameter $p$ to indicate where in the gap between $p$ and perfect selectivity we find the percentile of the cognitive skill distribution for the average immigrant ( $\tilde{p}_{s}$ ). Data are taken from Docquier, Lowell, and Marfouk (2009) and refer to the year 2000. The country of origin test score is than the predicted percentile score indicated by $p$. For each country, we know the mean and standard deviation (as well as the $75^{\text {th }}$ and $90^{\text {th }}$ percentile) of the test score distribution. Assuming a normal distribution, we can calculate the corresponding test score that is adjusted for international migrant selectivity:
\[

$$
\begin{equation*}
\text { score }_{\text {sgut }}^{\text {selectivity }}=\text { invnorm }\left(\tilde{p}_{s}\right) * s d_{\text {sgut }}^{\text {standard }}+\text { score }_{\text {sgut }}^{\text {standard }} \tag{B10}
\end{equation*}
$$

\]

where invnorm ( $\tilde{p}_{s}$ ) are draws of the $p$-th percentile from a normal $(0,1)$ distribution, $\operatorname{score}_{\text {sgut }}^{\text {standard }}$ is the average international test score of country $s$ at grade $g$ in subject $u$ in year $t$, and $s d_{\text {sgut }}^{\text {standard }}$ is the corresponding standard deviation. The comparison of $\operatorname{score} e_{s}^{\text {selectivity }}$ in math, grade 8 , in the year 2007, using $\tilde{p}_{s}=75$ and $\tilde{p}_{s}=90$, respectively, with the countryspecific observed test scores at the $75^{\text {th }}$ and $90^{\text {th }}$ percentile, respectively, show that this prediction works well (correlations almost perfect with $r=99 \%$ in both cases).

In further analysis, we use $\tilde{p}_{s}=p_{s}$ (according to equation (B9)) and $\tilde{p}_{s}=90$, respectively.

## B. 4 Backward Projection of Time-Varying Scores

Finally, we employ two methods of age projections of historical achievement patterns, one based on extrapolation from the available NAEP data and one based on projection from state SAT scores.

## Extrapolation of NAEP Trends

The skill measures developed so far assume that an average test score applies to the whole working-age population. We now aim to project developments of cognitive skills over time by state. Because test score data are not available before 1990 at the state level, we project test scores back in time, incorporating the long-term national trend which dates back to 1978 for eighth-grade math. For the projections, we do not use the 1990 value but rather start in 1992, as the very first test scores seem to differ somewhat from the subsequent trends. The basic idea of our backward projection is to use an average of the linear trend in the state test score and the
observed national trend to predict the test score of the state in a given year until 1978, i.e. from 1978 to 1992.

The national NAEP series that goes back until 1978, called long-term trend NAEP, is on a slightly different scale than the state NAEP series used in the state analysis. First, as scores reported prior to 2004 are reported in a different testing format and both formats are reported for 2004, we align the prior scores by standardization equivalent to the adjustment for scores without accommodation above. Then, to make the scales comparable, we subtract from each long-term trend test score the long-term trend score in 1992 and divide by the U.S. standard deviation in 1992 from the long-term trend. We then multiply this term by the U.S. standard deviation in 1992 from the state NAEP series and add the national mean from the from the state NAEP series.

We start the projection by interpolating the available test scores linearly for each state from 1992 to 2011. ${ }^{6}$ The projection then follows an iterative process: We assume that each test score of state $s$ in $t-1, \widetilde{T}_{s, t-1}$, is equal to the test score in $t, T_{s t}$, minus a simple average of the change in the state-specific linear time trend, i.e. the slope of the time trend, and the change in the national time trend:

$$
\begin{equation*}
\tilde{T}_{s, t-1}=T_{s t}-\frac{1}{2}\left(x_{t} \Delta \text { Linear State Trend }_{s t}+\Delta \overline{\text { Natlonal }_{t}}\right) \tag{B11}
\end{equation*}
$$

where
$\Delta$ Linear State Trend $_{s t}=$ Linear State Trend $_{s t}-$ Linear State Trend $_{s, t-1}$

$$
\Delta \overline{\text { Natıonal }_{t}}=\overline{\text { Natıonal }_{t}}-\overline{\text { Natıonal }_{t-1}}
$$

The Linear State Trend ${ }_{s t}$ is obtained from state-specific regressions of the test score on years. $\overline{\text { Natıonal }_{t}}$ is the long-term trend national average and available backwards until 1978.

To ensure that the (weighted) average of all state test scores is equal to the national average, we adjust the linear state trend with a time-varying constant, $x_{t}$. This adjustment factor is computed by taking the weighted sum of the test score projection on both sides and solving for $x_{t}$ :

[^4]\[

$$
\begin{gather*}
\overline{\text { Natlonal }_{t-1}}=\sum_{s=1}^{51} w_{s} \tilde{T}_{s, t-1}, \text { for } t \leq 1991 \text { and } \sum_{s=1}^{51} w_{s}=1 \\
\Leftrightarrow x_{t}=\frac{2 * \sum_{s=1}^{51} w_{s} T_{s t}-\overline{\text { Natıonal }} \overline{\text { Natıonal }_{t-1}}}{\sum_{s=1}^{51} w_{s} \Delta \text { Linear State Trend }_{s t}} \tag{B12}
\end{gather*}
$$
\]

The weights, $w_{s}$, are based on average daily attendance in public elementary and secondary schools by state from the Digest of Education Statistics (U.S. Department of Education (2013)). To obtain a weight for each state, we divide the average daily attendance in the state by the total national daily attendance. This measure is averaged over the time period 1978 to 1992 as the fractions are rather stable. The cross-sectional correlation between the fractions in 1978 and in 1992 is 98 percent.

This part of the extrapolation is exemplified by Figure A2, which shows both the observed data and the extrapolated state trends for two states: Massachusetts and Mississippi. Massachusetts was above the national average in 2011, but also had a steeper growth trend than the nation as a whole. As such, we shrink the extrapolated trend toward the national trend. Mississippi is different: while it also had a steeper growth trend than the nation as a whole, its scores were below the national average. Again, we shrink the extrapolation to the nationally observed trend.

The projected test score series then uses the available test score information for each state from 1992 to 2011 and the projected scores from the above iterative procedure from 1978 to 1992. Before 1978, we either assume a constant test score or a linear state trend.

The adjusted skill measure is then constructed by taking five-year averages of the projected test score series. These five year averages are then matched to the population shares of the appropriate age. To match the projected test score data, the share of people from origin o living in state $s$ in equation (B4) is computed in five-year age intervals from the Census data, both for the state average and for the education-category subsamples. The adjusted skill measure is then derived as

$$
\begin{equation*}
\text { score }_{s t}^{\text {projected }}=\sum_{\text {oea }} \text { population share } \text { oeast } \times \text { score }_{\text {oea }} \tag{B13}
\end{equation*}
$$

where the population shares and scores now do not only vary by state of origin $o$ and educational category $e$, but also by age category $a$.

## Projection from State SAT Scores

We obtained state-specific SAT scores (in math, writing, and reading) from 1972 to 2013 from the College Board. SAT scores are not representative for the total student population. But College Board also provided information on total participation (number of test takers). We calculate SAT participation rates by dividing the number of SAT participants by the total number of public high school graduates in each state. The latter is collected from various years of the Digest of Education Statistics (filling gaps by linear interpolation between available years).

Regressing the SAT score on the participation rate shows a significant negative relationship, indicating that a higher participation rate is related to a less selective sample and lower test scores. We therefore construct a series of participation-adjusted SAT scores:

$$
\begin{equation*}
\text { SAT score }_{s t}=\alpha_{0}+\alpha_{1} \text { participation rate }{ }_{s t}+\lambda_{s}+\lambda_{t}+\epsilon_{s t} \tag{B14}
\end{equation*}
$$

We use the estimated coefficients to predict SAT test scores with constant participation rates, where we assume that all states have the mean U.S. participation rate over the period 1972 to 2013 of 46.9 percent.

The participation-adjusted SAT scores allow us to predict state NAEP scores before 1992. To do so, we first regress the eighth-grade math test scores in NAEP on the participationadjusted SAT scores by state for the years since 1992 where both test scores series are available:

$$
\begin{equation*}
N A E P \text { score }_{t}=\beta_{0}+\beta_{1} S A T \text { score } e_{t}^{\text {adjusted }}+\epsilon_{t} \tag{B15}
\end{equation*}
$$

Because the SAT is taken around high school graduation, in these regressions we lag the SAT test scores by four years to capture almost the same cohorts as in NAEP. The regressions show that the participation-adjusted SAT score and the NAEP score move together over time in almost all states. ${ }^{7}$

With the estimated coefficients, we can then construct predicted NAEP test scores for each state for the years 1968 to 1991. Applying the same algorithm for the projection of test scores by age as before, we construct new aggregate test scores for each state and year by using the predicted NAEP test scores based on the SAT data.

[^5]
## Additional Appendix Tables

Table A2: Selectivity of Migrant Sending Countries

| Country | School-attainment selectivity |  | Country <br> Adjusted | School-attainment selectivity |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Adjusted | Unadjusted |  | Adjusted | Unadjusted |
| Niger | 1.000 | 0.990 | Egypt | 0.982 | 0.867 |
| Congo, Dem. Rep. | 0.998 | 0.959 | Nepal | 0.982 | 0.865 |
| Gambia, The | 0.998 | 0.96 | Pakistan | 0.981 | 0.861 |
| Kenya | 0.998 | 0.956 | Suriname | 0.981 | 0.863 |
| Lesotho | 0.998 | 0.956 | Tunisia | 0.981 | 0.861 |
| Tanzania | 0.998 | 0.952 | Afghanistan | 0.980 | 0.857 |
| Cameroon | 0.997 | 0.947 | Kiribati | 0.980 | 0.858 |
| Cote d'Ivoire | 0.997 | 0.946 | Maldives | 0.980 | 0.858 |
| Madagascar | 0.997 | 0.945 | Mauritius | 0.980 | 0.858 |
| Malawi | 0.997 | 0.949 | Tuvalu | 0.980 | 0.858 |
| Mongolia | 0.997 | 0.948 | Sri Lanka | 0.979 | 0.855 |
| Uganda | 0.997 | 0.946 | Bahrain | 0.978 | 0.852 |
| Zambia | 0.997 | 0.943 | Iran | 0.978 | 0.853 |
| Nigeria | 0.996 | 0.936 | Qatar | 0.978 | 0.853 |
| Zimbabwe | 0.995 | 0.930 | Saudi Arabia | 0.978 | 0.852 |
| Mauritania | 0.994 | 0.921 | United Arab Emirates | 0.978 | 0.853 |
| Palau | 0.994 | 0.921 | Singapore | 0.977 | 0.850 |
| Burkina Faso | 0.993 | 0.917 | Bangladesh | 0.976 | 0.845 |
| Sao Tome and Principe | 0.992 | 0.908 | Senegal | 0.976 | 0.845 |
| Chad | 0.991 | 0.903 | Eritrea | 0.974 | 0.840 |
| Mozambique | 0.991 | 0.904 | Kuwait | 0.974 | 0.839 |
| Guinea | 0.990 | 0.898 | Liechtenstein | 0.974 | 0.838 |
| Guinea-Bissau | 0.990 | 0.902 | Switzerland | 0.972 | 0.833 |
| Liberia | 0.990 | 0.898 | Burma (Myanmar) | 0.971 | 0.828 |
| Mali | 0.990 | 0.901 | Bhutan | 0.970 | 0.828 |
| Rwanda | 0.990 | 0.902 | Taiwan | 0.970 | 0.827 |
| Ethiopia | 0.989 | 0.895 | Angola | 0.969 | 0.823 |
| India | 0.989 | 0.897 | Congo, Rep. of the | 0.969 | 0.823 |
| Indonesia | 0.989 | 0.894 | Venezuela | 0.969 | 0.825 |
| Namibia | 0.989 | 0.893 | Equatorial Guinea | 0.968 | 0.822 |
| Burundi | 0.988 | 0.891 | Benin | 0.967 | 0.819 |
| Macedonia | 0.988 | 0.893 | Libya | 0.967 | 0.817 |
| Papua New Guinea | 0.988 | 0.891 | Brazil | 0.966 | 0.816 |
| Sierra Leone | 0.988 | 0.891 | Turkey | 0.966 | 0.817 |
| Botswana | 0.987 | 0.887 | Seychelles | 0.965 | 0.812 |
| Central African Rep. | 0.987 | 0.885 | Djibouti | 0.964 | 0.809 |
| Brunei | 0.986 | 0.883 | Monaco | 0.963 | 0.808 |
| Comoros | 0.986 | 0.880 | Palestinian Territory | 0.963 | 0.809 |
| Sudan | 0.986 | 0.883 | Thailand | 0.962 | 0.806 |
| Ghana | 0.985 | 0.877 | Bolivia | 0.956 | 0.791 |
| Algeria | 0.984 | 0.872 | Malaysia | 0.956 | 0.790 |
| Togo | 0.984 | 0.875 | France | 0.954 | 0.785 |
| Andorra | 0.983 | 0.870 | Swaziland | 0.954 | 0.785 |
| Gabon | 0.983 | 0.871 | Georgia | 0.953 | 0.784 |
| Morocco | 0.983 | 0.869 | San Marino | 0.952 | 0.780 |
| South Africa | 0.983 | 0.870 | Vatican | 0.951 | 0.778 |

(continued on next page)

Table A2 (continued)

| Country | School-attainment selectivity |  | Country <br> Adjusted | School-attainment selectivity |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Adjusted | Unadjusted |  | Adjusted | Unadjusted |
| Paraguay | 0.951 | 0.779 | Austria | 0.912 | 0.704 |
| Lebanon | 0.949 | 0.775 | Serbia and Montenegro | 0.912 | 0.704 |
| Belize | 0.948 | 0.771 | Slovakia | 0.912 | 0.703 |
| China, Hong Kong SAR | 0.947 | 0.769 | Saint Vincent | 0.911 | 0.702 |
| China, Macao SAR | 0.947 | 0.769 | Bahamas, The | 0.910 | 0.700 |
| Somalia | 0.946 | 0.767 | Nicaragua | 0.910 | 0.699 |
| Haiti | 0.945 | 0.765 | Romania | 0.909 | 0.699 |
| Azerbaijan | 0.944 | 0.763 | Costa Rica | 0.908 | 0.697 |
| Spain | 0.944 | 0.764 | Vanuatu | 0.907 | 0.694 |
| Philippines | 0.943 | 0.760 | Australia | 0.906 | 0.693 |
| Latvia | 0.941 | 0.758 | Czech Republic | 0.906 | 0.693 |
| Uzbekistan | 0.940 | 0.755 | Cyprus | 0.905 | 0.692 |
| Grenada | 0.939 | 0.754 | Israel | 0.905 | 0.693 |
| Bulgaria | 0.938 | 0.750 | Oman | 0.905 | 0.692 |
| United Kingdom | 0.938 | 0.752 | Peru | 0.905 | 0.692 |
| Vietnam | 0.938 | 0.750 | Antigua and Barbuda | 0.904 | 0.690 |
| Japan | 0.936 | 0.747 | Armenia | 0.903 | 0.688 |
| Micronesia | 0.936 | 0.746 | Korea | 0.902 | 0.688 |
| Marshall Islands | 0.935 | 0.746 | Albania | 0.901 | 0.685 |
| Kazakhstan | 0.933 | 0.742 | Saint Kitts and Nevis | 0.899 | 0.682 |
| Panama | 0.933 | 0.741 | Dominica | 0.898 | 0.68 |
| Colombia | 0.932 | 0.738 | Uruguay | 0.898 | 0.681 |
| Estonia | 0.932 | 0.739 | Luxembourg | 0.894 | 0.674 |
| Denmark | 0.931 | 0.737 | Norway | 0.894 | 0.674 |
| New Zealand | 0.930 | 0.736 | Saint Lucia | 0.893 | 0.673 |
| Trinidad and Tobago | 0.930 | 0.736 | Cambodia | 0.892 | 0.671 |
| Sweden | 0.929 | 0.733 | Fiji | 0.892 | 0.671 |
| Belgium | 0.928 | 0.731 | Cape Verde | 0.890 | 0.668 |
| China | 0.928 | 0.732 | Bosnia and Herzegovina | 0.889 | 0.666 |
| Belarus | 0.927 | 0.729 | Malta | 0.885 | 0.660 |
| Chile | 0.927 | 0.730 | Poland | 0.885 | 0.661 |
| Kyrgyzstan | 0.927 | 0.729 | Croatia | 0.884 | 0.660 |
| Hungary | 0.926 | 0.728 | Barbados | 0.88 | 0.653 |
| Tajikistan | 0.926 | 0.728 | Ireland | 0.879 | 0.652 |
| Turkmenistan | 0.926 | 0.729 | Tonga | 0.879 | 0.652 |
| Finland | 0.925 | 0.725 | Honduras | 0.876 | 0.647 |
| Jamaica | 0.925 | 0.725 | Germany | 0.872 | 0.643 |
| Netherlands | 0.923 | 0.723 | Laos | 0.867 | 0.636 |
| Argentina | 0.922 | 0.721 | Ecuador | 0.865 | 0.632 |
| Lithuania | 0.921 | 0.719 | Portugal | 0.865 | 0.633 |
| Ukraine | 0.919 | 0.716 | Italy | 0.863 | 0.629 |
| Moldova | 0.918 | 0.714 | Greece | 0.850 | 0.613 |
| Syria | 0.918 | 0.713 | Slovenia | 0.850 | 0.613 |
| Russia | 0.917 | 0.712 | Guatemala | 0.847 | 0.609 |
| Guyana | 0.916 | 0.709 | Dominican Republic | 0.843 | 0.604 |
| Samoa | 0.916 | 0.710 | Cuba | 0.836 | 0.595 |
| Iceland | 0.914 | 0.708 | El Salvador | 0.827 | 0.584 |
| Iraq | 0.914 | 0.707 | Canada | 0.774 | 0.525 |
| Jordan | 0.914 | 0.707 | Mexico | 0.710 | 0.461 |
| Yemen | 0.913 | 0.704 | Puerto Rico | 0.500 | 0.500 |

## Additional Appendix Tables

Table A2: Selectivity of Migrant Sending Countries

| Country | School-attainment selectivity |  | Country | School-attainment selectivity |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Adjusted | Unadjusted |  | Adjusted | Unadjusted |
| Mongolia | 0.997 | 0.948 | England | 0.938 | 0.752 |
| Indonesia | 0.989 | 0.894 | Scotland | 0.938 | 0.752 |
| Macedonia | 0.988 | 0.893 | United Kingdom | 0.938 | 0.752 |
| Botswana | 0.987 | 0.887 | American Samoa | 0.936 | 0.746 |
| Ghana | 0.985 | 0.877 | Guam | 0.936 | 0.746 |
| Southern Africa | 0.985 | 0.878 | Japan | 0.936 | 0.747 |
| Africa | 0.984 | 0.872 | Overseas Territories | 0.936 | 0.746 |
| Algeria | 0.984 | 0.872 | U.S. Virgin Islands | 0.936 | 0.746 |
| Morocco | 0.983 | 0.869 | Israel/Palestine | 0.934 | 0.751 |
| South Africa | 0.983 | 0.870 | Kazakhstan | 0.933 | 0.742 |
| Egypt | 0.982 | 0.867 | Panama | 0.933 | 0.741 |
| Northern Africa | 0.982 | 0.867 | Colombia | 0.932 | 0.738 |
| Tunisia | 0.981 | 0.861 | Estonia | 0.932 | 0.739 |
| Bahrain | 0.978 | 0.852 | Baltic States | 0.931 | 0.738 |
| Iran | 0.978 | 0.853 | Denmark | 0.931 | 0.737 |
| Qatar | 0.978 | 0.853 | New Zealand | 0.930 | 0.736 |
| Saudi Arabia | 0.978 | 0.852 | Trinidad and Tobago | 0.930 | 0.736 |
| United Arab Emirates | 0.978 | 0.853 | Sweden | 0.929 | 0.733 |
| Singapore | 0.977 | 0.850 | Western Europe | 0.929 | 0.741 |
| Kuwait | 0.974 | 0.839 | Belgium | 0.928 | 0.731 |
| Liechtenstein | 0.974 | 0.838 | Former USSR without Russia | 0.928 | 0.734 |
| Switzerland | 0.972 | 0.833 | Chile | 0.927 | 0.730 |
| Taiwan (Chinese Taipei) | 0.970 | 0.827 | Former USSR | 0.927 | 0.731 |
| Southeast Asia + Iran | 0.968 | 0.825 | Kyrgyzstan | 0.927 | 0.729 |
| Brazil | 0.966 | 0.816 | Hungary | 0.926 | 0.728 |
| Turkey | 0.966 | 0.817 | Finland | 0.925 | 0.725 |
| Southeast Asia | 0.965 | 0.820 | South America | 0.925 | 0.730 |
| Palestinian Nat'l Auth. | 0.963 | 0.809 | Total Average | 0.925 | 0.744 |
| Thailand | 0.962 | 0.806 | Netherlands | 0.923 | 0.723 |
| Malaysia | 0.956 | 0.790 | Argentina | 0.922 | 0.721 |
| Asia | 0.955 | 0.798 | Lithuania | 0.921 | 0.719 |
| Middle East | 0.955 | 0.798 | Northern Europe | 0.921 | 0.720 |
| France | 0.954 | 0.785 | Ukraine | 0.919 | 0.716 |
| Georgia | 0.953 | 0.784 | Moldova | 0.918 | 0.714 |
| East Asia | 0.950 | 0.791 | Oceania | 0.918 | 0.715 |
| Lebanon | 0.949 | 0.775 | Syrian Arab Republic | 0.918 | 0.713 |
| Hong Kong | 0.947 | 0.769 | Europe | 0.914 | 0.714 |
| Macao-China | 0.947 | 0.769 | Iceland | 0.914 | 0.708 |
| Azerbaijan | 0.944 | 0.763 | Jordan | 0.914 | 0.707 |
| Spain | 0.944 | 0.764 | Antarctica | 0.913 | 0.706 |
| Philippines | 0.943 | 0.760 | Austria | 0.912 | 0.704 |
| Latvia | 0.941 | 0.758 | Montenegro | 0.912 | 0.704 |
| Bulgaria | 0.938 | 0.750 | Serbia | 0.912 | 0.704 |

(continued on next page)

Table A2 (continued)

| Country | School-attainment selectivity |  | Country | School-attainment selectivity |  |
| :--- | :---: | :---: | :--- | :---: | :---: |
|  | Adjusted | Unadjusted |  | Adjusted | Unadjusted |
| Slovak Rep. | 0.912 | 0.703 | Central America | 0.891 | 0.677 |
| Czechoslovakia | 0.909 | 0.698 | Bosnia and Herzegovina | 0.889 | 0.666 |
| Romania | 0.909 | 0.699 | Malta | 0.885 | 0.660 |
| Eastern Europe | 0.907 | 0.698 | Poland | 0.885 | 0.661 |
| Australia | 0.906 | 0.693 | Croatia | 0.884 | 0.66 |
| Czech Rep. | 0.906 | 0.693 | Ireland | 0.879 | 0.652 |
| Yugoslavia | 0.906 | 0.707 | Honduras | 0.876 | 0.647 |
| Cyprus | 0.905 | 0.692 | Germany | 0.872 | 0.643 |
| Oman | 0.905 | 0.692 | Portugal | 0.865 | 0.633 |
| Peru | 0.905 | 0.692 | Italy | 0.863 | 0.629 |
| Armenia | 0.903 | 0.688 | Greece | 0.850 | 0.613 |
| Korea, Rep. | 0.902 | 0.688 | El Salvador | 0.827 | 0.584 |
| Albania | 0.901 | 0.685 | Canada | 0.774 | 0.525 |
| Southern Europe | 0.899 | 0.697 | North America | 0.774 | 0.525 |
| Uruguay | 0.898 | 0.681 | Mexico | 0.710 | 0.461 |
| Luxembourg | 0.894 | 0.674 | Puerto Rico | 0.500 | 0.500 |
| Norway | 0.894 | 0.674 |  |  |  |

Table A3: Summary Statistics

|  | Obs. | Mean | Std. dev. | Min. | Max. |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Real GDP per capita, 2007 | 47 | 41,218 | 6,388 | 29,302 | 59,251 |
| Years of schooling, 2007 | 47 | 13.11 | 0.35 | 12.52 | 13.74 |
| Test scores: |  |  |  |  |  |
| Baseline: local average adjusted for interstate migrants | 47 | 499.9 | 15.98 | 460.4 | 527.7 |
| + Adjustment of locals by education category | 47 | 494.4 | 15.46 | 454.9 | 521.3 |
| + Adjustment of interstate migrants by education category | 47 | 493.9 | 15.80 | 453.1 | 522.0 |
| + Adjustment of international migrants scores by selectivity | 47 | 497.7 | 15.57 | 454.8 | 524.7 |
| Age adjustment with extrapolation of NAEP trends by education category | 47 | 442.4 | 22.04 | 381.9 | 476.5 |
| Age adjustment with projection from SAT scores | 47 | 407.2 | 27.52 | 321.5 | 456.6 |
| Growth rate of real GDP per capita, 1970-2007 | 47 | 2.24 | 0.31 | 1.56 | 2.89 |
| Change in years of schooling, 1970-2007 | 47 | 2.02 | 0.45 | 0.78 | 2.86 |
| Estimated annual change in test scores, 1968-2011 | 47 | 3.17 | 1.21 | 1.17 | 6.77 |

Notes: See sections 2.2, 2.3, and 3.1 for details on the data. Test scores refer to eighth-grade math. Locals are all persons who report a state of birth equal to the current state of residence. Interstate migrants report another state of birth than state of residence. International migrants report another country of birth than the United States. "By education category" indicates that individuals with/without university education are assigned the test scores of children of parents with/without university education.

Table A4: Main Data by State

(continued on next page)

Table A4 (continued)

|  | Real GDP per capita 2007 | Years of schooling 2007 | Test scores |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Average NAEP score | Baseline score | Adjusted for selective migration | Projection by NAEP trends | Projection from SAT scores |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
| New Jersey | 51,337 | 13.48 | 519.0 | 513.9 | 516.5 | 465.5 | 433.2 |
| New Mexico | 35,313 | 12.71 | 468.2 | 480.9 | 480.0 | 428.4 | 403.3 |
| New York | 53,165 | 13.27 | 497.8 | 498.4 | 508.8 | 460.1 | 426.4 |
| North Carolina | 41,123 | 12.98 | 497.3 | 497.1 | 492.3 | 416.1 | 358.3 |
| North Dakota | 41,329 | 13.47 | 531.8 | 527.0 | 520.3 | 472.8 | 456.6 |
| Ohio | 38,389 | 13.13 | 510.2 | 506.8 | 500.8 | 432.5 | 394.4 |
| Oklahoma | 36,504 | 12.84 | 488.5 | 491.4 | 486.2 | 437.8 | 412.4 |
| Oregon | 42,422 | 13.18 | 511.0 | 503.1 | 503.9 | 450.7 | 420.1 |
| Pennsylvania | 39,951 | 13.21 | 509.2 | 507.7 | 501.5 | 444.3 | 406.7 |
| Rhode Island | 42,274 | 13.05 | 489.4 | 495.8 | 495.5 | 445.4 | 411.0 |
| South Carolina | 33,539 | 12.85 | 490.5 | 492.7 | 486.8 | 414.8 | 354.5 |
| South Dakota | 41,649 | 13.12 | 521.6 | 518.6 | 508.7 | 460.5 | 427.1 |
| Tennessee | 37,068 | 12.74 | 475.8 | 482.1 | 477.3 | 415.5 | 374.3 |
| Texas | 45,502 | 12.52 | 502.7 | 499.8 | 496.8 | 438.1 | 400.2 |
| Utah | 39,464 | 13.26 | 506.5 | 502.9 | 497.2 | 454.7 | 434.7 |
| Vermont | 36,445 | 13.63 | 525.2 | 517.1 | 511.7 | 447.5 | 400.1 |
| Virginia | 47,501 | 13.44 | 508.3 | 501.8 | 501.6 | 441.0 | 402.6 |
| Washington | 47,553 | 13.37 | 513.8 | 506.8 | 514.2 | 460.2 | 391.5 |
| West Virginia | 29,302 | 12.53 | 475.7 | 483.0 | 472.8 | 411.9 | 380.2 |
| Wisconsin | 39,841 | 13.28 | 521.1 | 516.5 | 509.5 | 463.1 | 433.3 |
| Wyoming | 59,558 | 13.22 | 514.1 | 509.4 | 504.9 | 452.2 | 423.6 |

Notes: (1) Real GDP per capita in 2005 U.S. dollars. (2) Mean years of completed schooling, 2007. (3) Estimated average eighth-grade math NAEP score from 1992 to 2011, obtained from a regression of NAEP test scores on time and state fixed effects; see Appendix B.1. (4) Baseline: local average adjusted for interstate migrants by average test score of their state of birth. (5) Baseline + adjustment of locals by education category + adjustment of interstate migrants by education category + adjustment of international migrants by selectivity. (6) Age adjustment with extrapolation of NAEP trends by education category; see Appendix B.4. (7) Age adjustment with projection from SAT scores; see Appendix B.4.

## Table A5: Main Source Countries

| Country of Birth | Total Census Observations, 1940-2010 | Share of all immigrants (in percent) | Imputation of international test scores |
| :---: | :---: | :---: | :---: |
| Mexico | 1,054,264 | 24.14 |  |
| Philippines | 192,335 | 4.40 |  |
| Puerto Rico | 184,529 | 4.22 | NAEP |
| Germany | 138,950 | 3.18 |  |
| India | 136,515 | 3.13 | Southeast Asia: Indonesia, Malaysia, Philippines, Singapore, Thailand + Iran |
| Canada | 136,424 | 3.12 |  |
| Cuba | 115,914 | 2.65 | Central America: El Salvador, Panama, Honduras, Trinidad\&Tobago |
| China | 115,670 | 2.65 | East Asia: Shanghai-China, Hong Kong, Macao-China, Mongolia, Taiwan (Chinese Taipei), Japan, Korea, Rep. |
| Vietnam | 111,037 | 2.54 | Southeast Asia: Indonesia, Malaysia, Philippines, Singapore, Thailand |
| Italy | 102,190 | 2.34 |  |
| El Salvador | 93,766 | 2.15 |  |
| Korea | 87,184 | 2.00 | South Korea |
| England | 81,712 | 1.87 |  |
| USA, Unknown State | 72,212 | 1.65 | NAEP |
| Poland | 71,464 | 1.64 |  |
| Dominican Republic | 67,583 | 1.55 | Central America |
| Japan | 62,327 | 1.43 |  |
| Jamaica | 58,633 | 1.34 | Central America |
| Colombia | 57,598 | 1.32 |  |
| Guatemala | 55,451 | 1.27 | Central America |
| Abroad, ns | 52,545 | 1.20 | Total Average |
| Other USSR/Russia | 44,915 | 1.03 | USSR: Russia, Moldova, Ukraine, Armenia, Azerbaijan, Georgia, Kazakhstan, Kyrgyzstan |
| Taiwan | 40,817 | 0.93 |  |
| Haiti | 40,287 | 0.92 | Central America |
| West Germany | 36,231 | 0.83 | Germany |
| Iran | 34,117 | 0.78 |  |
| Ecuador | 32,475 | 0.74 | South America: Argentina, Brazil, Chile, Colombia, Peru, Uruguay |
| Peru | 32,047 | 0.73 |  |
| Portugal | 31,728 | 0.73 |  |
| Honduras | 31,141 | 0.71 |  |
| Ireland | 30,295 | 0.69 |  |
| Greece | 29,979 | 0.69 |  |
| France | 28,703 | 0.66 |  |
| Brazil | 25,754 | 0.59 |  |
| United Kingdom | 25,565 | 0.59 |  |
| Hong Kong | 25,324 | 0.58 |  |
| Nicaragua | 23,920 | 0.55 | Central America |
| Pakistan | 23,123 | 0.53 | Southeast Asia + Iran |
| Guyana/British Guiana | 22,425 | 0.51 | South America |
| Laos | 21,998 | 0.50 | Southeast Asia |
| Trinidad and Tobago | 21,731 | 0.50 |  |

Notes: Main source countries/regions of immigrants living in the United States. Only countries with a share of the total immigrant inflow of at least 0.5 percent. Averages over all available Census years. Imputation: Countries/ region by which test scores are imputed in cases without international test score data. Source: Authors’ calculations based on Ruggles et al. (2010).

Table A6: Development Accounting Results for Different Years

| Test score specification | Year | Total knowledge capital |  | Test scores |  | Years of schooling |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Baseline: local average adjusted for interstate migrants | 2007 | $0.150 * * *$ | (0.045) | 0.057** | (0.025) | $0.093 * * *$ | (0.023) |
|  | 2000 | $0.149^{* * *}$ | (0.047) | $0.061 * *$ | (0.026) | $0.088^{* * *}$ | (0.024) |
|  | 1990 | $0.127^{* *}$ | (0.048) | 0.031 | (0.029) | $0.096 * * *$ | (0.023) |
|  | 1980 | $0.155^{* *}$ | (0.078) | 0.024 | (0.038) | $0.131 * * *$ | (0.044) |
|  | 1970 | $0.179^{* * *}$ | (0.060) | 0.028 | (0.033) | $0.151^{* * *}$ | (0.032) |
| + Adjustment of locals by education category | 2007 | $0.159^{* * *}$ | (0.043) | $0.066^{* * *}$ | (0.024) | $0.093 * * *$ | (0.023) |
|  | 2000 | $0.157^{* *}$ | (0.046) | $0.069^{* * *}$ | (0.025) | $0.088^{* * *}$ | (0.024) |
|  | 1990 | $0.138 * * *$ | (0.046) | 0.042 | (0.027) | $0.096{ }^{* * *}$ | (0.023) |
|  | 1980 | 0.181** | (0.076) | 0.050 | (0.035) | $0.131^{* *}$ | (0.044) |
|  | 1970 | $0.198 * * *$ | (0.059) | 0.047 | (0.031) | $0.151^{* * *}$ | (0.032) |
| + Adjustment of interstate migrants by education category | 2007 | $0.169^{* * *}$ | (0.043) | $0.076 * * *$ | (0.024) | $0.093 * * *$ | (0.023) |
|  | 2000 | $0.165^{* *}$ | (0.047) | $0.077^{* * *}$ | (0.025) | $0.088^{* * *}$ | (0.024) |
|  | 1990 | $0.145^{* *}$ | (0.046) | 0.049* | (0.026) | $0.096{ }^{* * *}$ | (0.023) |
|  | 1980 | $0.178 * *$ | (0.075) | 0.047 | (0.034) | $0.131^{* * *}$ | (0.044) |
|  | 1970 | $0.186^{* * *}$ | (0.057) | 0.035 | (0.029) | $0.151^{* * *}$ | (0.032) |
| + Adjustment of international migrants by selectivity | 2007 | $0.190^{* * *}$ | (0.041) | $0.097^{* * *}$ | (0.022) | $0.093 * * *$ | (0.023) |
|  | 2000 | $0.180^{* * *}$ | (0.045) | $0.092^{* * *}$ | (0.024) | $0.088^{* * *}$ | (0.024) |
|  | 1990 | $0.169^{* *}$ | (0.043) | 0.073 *** | (0.023) | $0.096{ }^{* * *}$ | (0.023) |
|  | 1980 | 0.195** | (0.076) | 0.064* | (0.034) | $0.131^{* * *}$ | (0.044) |
|  | 1970 | $0.203{ }^{* * *}$ | (0.056) | 0.052* | (0.028) | $0.151^{* * *}$ | (0.032) |

Notes: Development accounting results (covariance measure) for 47 U.S. states with different test score specifications. Test scores refer to eighth-grade math. Locals are all persons who report a state of birth equal to the current state of residence. Interstate migrants report another state of birth than state of residence. International migrants report another country of birth than the United States. "By education category" indicates that individuals with/without university education are assigned the test scores of children of parents with/without university education. Calculations assume a return of $w=0.17$ per standard deviation in test scores and a return of $r=0.08$ per year of schooling. Bootstrapped standard errors in parentheses with 1,000 replications. Significance levels: ${ }^{*} \mathrm{p}<0.10,{ }^{* *} \mathrm{p}<0.05,{ }^{* * *} \mathrm{p}<0.01$.

Table A7: Five-State Measure: Alternative Numbers of Top and Bottom States

|  | Total knowledge capital | Test scores | Years of schooling |
| :--- | :---: | :---: | :---: |
| Five-state measure | 0.306 | 0.186 | 0.120 |
| Three-state measure | 0.307 | 0.170 | 0.137 |
| Seven-state measure | 0.261 | 0.164 | 0.097 |

Notes: Development accounting results (five-state measure) for 47 U.S. states with different numbers of countries used at the top and bottom of the state distribution. Test score specification adjusts locals and interstate migrants by age-education category based on extrapolation of NAEP trends by education category and international migrants by selectivity. Test scores refer to eighth-grade math. Calculations assume a return of $w=0.17$ per standard deviation in test scores and a return of $r=0.08$ per year of schooling.

Table A8: Growth Accounting by State, 1970-2007

|  | Average annual growth rate of real GDP per capita (percent) | Absolute change in years of schooling | Estimated annual change in test scores | Average annual growth rate accounted for by |  |  | Percent of total growth |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Total knowledge capital | Test scores | Years of schooling | Total knowledge capital | Test scores | Years of schooling |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
| Alabama | 2.35 | 2.65 | 2.77 | 0.70 | 0.31 | 0.38 | 29.6 | 13.4 | 16.3 |
| Arizona | 2.03 | 1.60 | 2.71 | 0.54 | 0.31 | 0.23 | 26.5 | 15.1 | 11.4 |
| Arkansas | 2.39 | 2.50 | 2.85 | 0.68 | 0.32 | 0.36 | 28.5 | 13.5 | 15.0 |
| California | 2.14 | 1.01 | 2.22 | 0.40 | 0.25 | 0.15 | 18.5 | 11.8 | 6.8 |
| Colorado | 2.58 | 1.57 | 3.38 | 0.61 | 0.38 | 0.23 | 23.6 | 14.8 | 8.8 |
| Connecticut | 2.79 | 2.25 | 2.77 | 0.64 | 0.31 | 0.32 | 22.9 | 11.3 | 11.6 |
| Florida | 2.16 | 1.98 | 3.93 | 0.73 | 0.45 | 0.29 | 33.9 | 20.6 | 13.2 |
| Georgia | 2.44 | 2.66 | 2.90 | 0.71 | 0.33 | 0.38 | 29.2 | 13.5 | 15.7 |
| Hawaii | 1.63 | 1.96 | 3.28 | 0.65 | 0.37 | 0.28 | 40.2 | 22.9 | 17.4 |
| Idaho | 2.02 | 1.53 | 2.08 | 0.46 | 0.24 | 0.22 | 22.6 | 11.7 | 10.9 |
| Illinois | 2.03 | 2.03 | 3.36 | 0.67 | 0.38 | 0.29 | 33.1 | 18.7 | 14.4 |
| Indiana | 2.01 | 1.85 | 3.28 | 0.64 | 0.37 | 0.27 | 31.8 | 18.5 | 13.3 |
| Iowa | 2.32 | 1.64 | 1.17 | 0.37 | 0.13 | 0.24 | 15.9 | 5.7 | 10.2 |
| Kansas | 2.43 | 1.63 | 3.10 | 0.59 | 0.35 | 0.23 | 24.1 | 14.4 | 9.6 |
| Kentucky | 1.86 | 2.62 | 3.64 | 0.79 | 0.41 | 0.38 | 42.6 | 22.2 | 20.3 |
| Louisiana | 2.41 | 2.33 | 4.30 | 0.82 | 0.49 | 0.34 | 34.2 | 20.3 | 14.0 |
| Maine | 2.20 | 2.20 | 1.63 | 0.50 | 0.18 | 0.32 | 22.8 | 8.4 | 14.4 |
| Maryland | 2.41 | 2.32 | 3.94 | 0.78 | 0.45 | 0.33 | 32.5 | 18.6 | 13.9 |
| Massachusetts | 2.56 | 2.21 | 5.47 | 0.94 | 0.62 | 0.32 | 36.7 | 24.2 | 12.5 |
| Michigan | 1.56 | 1.97 | 2.74 | 0.59 | 0.31 | 0.28 | 38.1 | 19.9 | 18.2 |
| Minnesota | 2.37 | 1.96 | 2.88 | 0.61 | 0.33 | 0.28 | 25.6 | 13.7 | 11.9 |
| Mississippi | 2.36 | 2.46 | 5.16 | 0.94 | 0.58 | 0.35 | 39.7 | 24.8 | 15.0 |
| Missouri | 1.89 | 2.10 | 2.56 | 0.59 | 0.29 | 0.30 | 31.3 | 15.3 | 16.0 |
| Montana | 2.10 | 1.68 | 1.42 | 0.40 | 0.16 | 0.24 | 19.2 | 7.7 | 11.5 |
| Nebraska | 2.42 | 1.67 | 1.54 | 0.42 | 0.17 | 0.24 | 17.1 | 7.2 | 10.0 |
| Nevada | 1.69 | 0.78 | 3.12 | 0.47 | 0.35 | 0.11 | 27.6 | 21.0 | 6.7 |
| New Hampshire | 2.56 | 2.16 | 2.85 | 0.64 | 0.32 | 0.31 | 24.8 | 12.6 | 12.2 |
| New Jersey | 2.41 | 2.25 | 3.41 | 0.71 | 0.39 | 0.32 | 29.5 | 16.1 | 13.5 |

[^6]Table A8 (continued)

|  | Average annual growth rate of real GDP per capita (percent) | Absolute change in years of schooling | Estimated annual change in test scores | Average annual growth rate accounted for by |  |  | Percent of total growth |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Total knowledge capital | Test scores | Years of schooling | Total knowledge capital | Test scores | Years of schooling |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
| New Mexico | 2.01 | 1.71 | 1.86 | 0.46 | 0.21 | 0.25 | 22.7 | 10.5 | 12.3 |
| New York | 2.12 | 2.05 | 3.49 | 0.69 | 0.40 | 0.30 | 32.6 | 18.7 | 13.9 |
| North Carolina | 2.30 | 2.76 | 6.06 | 1.08 | 0.69 | 0.40 | 47.2 | 29.9 | 17.3 |
| North Dakota | 2.86 | 2.38 | 1.46 | 0.51 | 0.17 | 0.34 | 17.7 | 5.8 | 12.0 |
| Ohio | 1.80 | 1.92 | 3.86 | 0.71 | 0.44 | 0.28 | 39.7 | 24.3 | 15.4 |
| Oklahoma | 2.26 | 1.71 | 1.93 | 0.47 | 0.22 | 0.25 | 20.6 | 9.7 | 10.9 |
| Oregon | 2.31 | 1.58 | 2.13 | 0.47 | 0.24 | 0.23 | 20.3 | 10.4 | 9.8 |
| Pennsylvania | 2.04 | 2.20 | 3.24 | 0.68 | 0.37 | 0.32 | 33.5 | 18.0 | 15.5 |
| Rhode Island | 2.32 | 2.19 | 2.67 | 0.62 | 0.30 | 0.32 | 26.6 | 13.1 | 13.6 |
| South Carolina | 2.30 | 2.86 | 5.35 | 1.02 | 0.61 | 0.41 | 44.2 | 26.4 | 17.9 |
| South Dakota | 2.89 | 1.89 | 2.94 | 0.61 | 0.33 | 0.27 | 20.9 | 11.5 | 9.4 |
| Tennessee | 2.29 | 2.52 | 3.59 | 0.77 | 0.41 | 0.36 | 33.7 | 17.8 | 15.9 |
| Texas | 2.48 | 1.85 | 4.43 | 0.77 | 0.50 | 0.27 | 30.9 | 20.2 | 10.7 |
| Utah | 2.41 | 1.22 | 1.93 | 0.39 | 0.22 | 0.18 | 16.4 | 9.1 | 7.3 |
| Vermont | 2.00 | 2.19 | 4.02 | 0.77 | 0.46 | 0.32 | 38.5 | 22.8 | 15.8 |
| Virginia | 2.69 | 2.66 | 3.74 | 0.81 | 0.42 | 0.38 | 30.0 | 15.8 | 14.3 |
| Washington | 2.24 | 1.48 | 6.77 | 0.98 | 0.77 | 0.21 | 43.8 | 34.3 | 9.5 |
| West Virginia | 1.67 | 2.33 | 2.88 | 0.66 | 0.33 | 0.34 | 39.6 | 19.5 | 20.1 |
| Wisconsin | 2.17 | 1.94 | 2.26 | 0.54 | 0.26 | 0.28 | 24.7 | 11.8 | 12.9 |

Notes: Estimated annual change in test scores: in percent of a standard deviation, obtained from a regression of test scores (NAEP scores projected based on participation-corrected SAT scores as derived in section 2.3.4) on years for each state, 1968-2011.

Figure A1: Cognitive Skills and Years of Schooling across U.S. States, 2007


Notes: Scatterplot of cognitive skill measure (adjusted for selective interstate and for international migration by selectivity) and average years of schooling of the working-age population across U.S. states, 2007. Source: Authors’ calculations based on data from Ruggles et al. (2010) and National Center for Education Statistics (2014).

Figure A2: Projection of Test Scores for Massachusetts and Mississippi


Notes: NAEP test score in eighth-grade math. Source: Authors' calculations based on data from National Center for Education Statistics (2014).


[^0]:    ${ }^{1}$ Some Census years only report educational categories that cover several years of schooling. For these years, we assume the same fraction for this educational category as in the closest survey with full information.

[^1]:    ${ }^{2}$ Standard deviations are also adjusted to be on the same scale by $s d_{s t}^{s t a n d a r d}=\left(\frac{s d_{s t}^{a d j}{ }_{s d} d_{U S, t=2011}}{s d_{U S, t=2011}}\right) * 100+$ 100.

[^2]:    ${ }^{3}$ We draw the data from the International Data Explorer (IDE) of the National Center of Education Statistics (http://nces.ed.gov/surveys/international/ide/).
    ${ }^{4}$ When estimating separate scores by the education category of the father, in PISA we use a simple average of the test scores in ISCED categories 0-4 for non-university education and ISCED categories 5a and 6 for university education. In TIMSS 1995 and 1999, we use the average of the categories until "finished secondary" for nonuniversity education and "finished university" for university education. In the subsequent TIMSS waves, we use ISCED categories 0-4 for non-university education and ISCED categories 5a and more than 5a for university education. The IDE does not report educational background variables for PIRLS and TIMSS grade 4.

[^3]:    ${ }^{5}$ Standard deviations are again adjusted to be on the same scale with NAEP.

[^4]:    ${ }^{6}$ A few states started representative NAEP testing later than 1992. These are Alaska, Montana, Oregon, Vermont, and Washington in 1996, Illinois, Kansas, and Nevada in 2000, and South Dakota in 2003. We project their scores back to 1992 with a simple backward projection method: $\tilde{T}_{s, t-1}=T_{s t}-\frac{1}{2}\left(\Delta\right.$ Linear State Trend $_{s t}+$ $\Delta \overline{\text { Natıonal }_{t}}$ ).

[^5]:    ${ }^{7}$ Exceptions are Kansas, Nevada, and South Dakota, which are also the states that start relatively late in NAEP, thereby impeding the prediction of a reliable connection between NAEP and SAT. For these states, we use the U.S. average coefficient.

[^6]:    (continued on next page)

