

# The Share of Systematic Variation in Bilateral Exchange Rates \*

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## Abstract

Two factors account for 20% to 90% of the daily, monthly, quarterly, and annual exchange rate movements. These two factors — carry and dollar — are *risk* factors: the former accounts for the cross-section of interest rate-sorted currency returns, while the latter accounts for a novel cross-section of dollar beta-sorted currency returns. The different shares of systematic risk across currencies are related to financial and macroeconomic measures of international comovement. They point to large shares of global shocks in the dynamics of exchange rates, as well as large differences across countries. The results offer new challenges for international macroeconomics and finance models.

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Changes in exchange rates appear random to most investors, central bankers, and researchers alike, except perhaps at very high or very low frequencies. Regressions of changes in individual exchange rates on lagged or contemporaneous interest rate differences or other contemporaneous changes in macroeconomic variables at monthly, quarterly, and annual frequencies deliver very low  $R^2$ s. Contemporaneous industrial production growth and inflation rates, for example, lead to essentially zero adjusted  $R^2$ s on monthly series over the 1983–2010 period for developed countries. As a result, each individual exchange rate movement seems mostly idiosyncratic.

In this paper, to the contrary, I report that two variables — the carry and the dollar factors — account for a substantial share of individual exchange rate time-series in developed countries, as well as in emerging and developing countries with floating exchange rates. All exchange rates are defined with respect to the U.S. dollar, and the carry and the dollar factors are constructed from portfolios of currencies. The carry factor corresponds to the change in exchange rates between baskets of high and low interest rate currencies, while the dollar factor corresponds to the average change in the exchange rate between the U.S. dollar and all other currencies. I regress changes in exchange rates on the carry factor, the same carry factor multiplied by the country-specific interest rate difference (the latter is referred to as “conditional carry”), and the dollar factor. The change in bilateral exchange rate on the left-hand side of these regressions is measured between  $t$  and  $t+1$ ; on the right-hand side, the carry and dollar factors correspond to changes between  $t$  and  $t+1$  too, while the domestic and foreign interest rates are known at date  $t$ . Importantly, the carry and dollar factors do not include the bilateral exchange rate that is the dependent variable.

The factor regressions offer a novel picture of bilateral exchange rate movements. Each factor raises the adjusted  $R^2$ s of the usual macroeconomic regressions by an order of magnitude. With the carry factors, the adjusted  $R^2$ s range from 0% to 23%. With the addition of the dollar factor,  $R^2$ s increase further: as an example, the factor regression for the U.S. dollar / U.K. pound exchange rate has an  $R^2$  of 51%. Crucially, the factor regressions uncover large differences in the shares of systematic variation:  $R^2$ s range from 19% to 91% in developed countries and from 10% to 75% among developing countries with floating currencies.

The substantial  $R^2$ s of the factor regressions do not imply that bilateral exchange rates are easy to forecast: the corresponding regressions use contemporaneous variables, not predictive ones. But if exchange rate movements were independent, the  $R^2$ s on those regressions would be zero. In the data, they are significantly different from zero. Moreover, the distribution of  $R^2$ s on the factor regressions is quite stable across frequencies; similar distributions appear at daily, monthly, quarterly, and annual frequencies.

The carry and dollar factors with their significant slope coefficients thus offer a powerful description of individual currencies. But their most important feature is their risk-based interpretation: both the carry and the dollar factors are *risk* factors in the asset pricing sense.

The risk-based interpretation of the carry factor is well known. Previous research on currency portfolios shows that the carry factor accounts for the cross-section of currency excess returns sorted by interest rates: covariances of the carry factor with currency returns align with the cross-section of average excess returns (cf. Lustig, Roussanov, and Verdelhan, 2011). A consistent result appears here on individual currencies: the higher the interest rate, the larger the loading on the carry risk factor. This is the risk-based explanation of the classic currency carry trade.

This paper shows that, similarly, the dollar factor has a risk-based interpretation. I build portfolios of countries sorted by their time-varying exposures to the dollar factor (i.e., dollar betas). The low dollar-beta portfolio offers an average log excess return of just 0.4% per year for investors who go long foreign currencies when the average forward discount (average foreign minus U.S. interest rates) is positive and short otherwise. The high dollar-beta portfolio offers an average log excess return of 7.6% for similar investments. After transaction costs, the high dollar-beta portfolio still returns 6.3% on average, implying a large Sharpe ratio of almost 0.6 over the last 30 years. Conditioning on the average forward discount, covariances of the dollar factor with portfolio returns account for this new cross-section of average excess returns, while covariances with the carry factor do not. As a result, the carry and dollar factors are two, largely independent, *risk* factors.

Additional empirical evidence in favor of a risk-based approach to exchange rates comes from the

link between the share of systematic risk in currency markets and other measures of cross-country systematic risk. This link is intuitive: when markets are complete, log changes in exchange rates correspond to the differences between domestic and foreign log pricing kernels (also known as stochastic discount factors or inter-temporal marginal rates of substitution). As a result, cross-country differences in currency  $R^2$ s come ultimately from differences in foreign pricing kernels, and thus should appear in other markets. This is indeed the case: countries with a high share of systematic equity risk (measured by their loadings on world aggregate and value risk factors) tend also to have a high share of systematic currency risk. Similarly, higher shares of systematic currency risk correspond to higher shares of fixed income risk.  $R^2$ s on currencies also appear related to measures of comovement in output and consumption growth rates. These findings highlight a novel link across asset and good markets. They also point towards currency markets as potential measures of comovement in inter-temporal marginal rates of substitution: unlike macroeconomic series, they are precisely estimated and available at high frequencies. Unlike equity returns, used in the context of world market integration studies,  $R^2$ s on currencies do not depend on the properties of the domestic and foreign dividend cash flows.

The cross-country differences in currency systematic risk revealed in this paper have key implications for the class of no-arbitrage models, without ruling out more behavioral explanations of the carry and dollar factors. I focus here on the rational, preference-free interpretation and implications, while also providing an interpretation of the dollar and carry factors in a reduced-form term structure model, as well as in a macro-finance general equilibrium model.

Without loss of generality, each pricing kernel can be decomposed into country-specific and world shocks. When markets are complete, bilateral exchange rates thus depend on (home minus foreign) differences in country-specific and world shocks. In large baskets of currencies, foreign country-specific shocks average out. The carry factor, defined as a difference in baskets of exchange rates, is dollar-neutral and depends on world shocks. The dollar factor, defined as the average of all the domestic minus foreign pricing kernels, depends on both U.S.-specific and world shocks, but not on foreign-specific shocks. The cross-country differences in dollar betas have two implications:

foreign pricing kernels must differ in their loadings on global shocks, and the dollar factor must have a global component, implying that the U.S. pricing kernel loads differently than the average pricing kernel on world shocks. If all shocks to the foreign pricing kernels were global, then the carry and dollar factors would perfectly account for all the changes in exchange rates. If all shocks to the pricing kernels were local, then the carry factors would be insignificant, the dollar factor would only account for U.S.-specific shocks, and all the  $R^2$ s would be the same across countries. The data suggest otherwise. Changes in exchange rates respond to local and global risk shocks.

Building on this intuition, the relative importance of the local and global shocks in exchange rate movements can be measured precisely, without any assumption on preferences. To focus on the global component of the dollar factor, I substitute the dollar factor with the change in exchange rates of the high dollar-beta portfolio minus the change in exchange rates of the low dollar-beta portfolio, thus eliminating the U.S.-specific shocks. Regressions of changes in bilateral exchange rates on the global component of the dollar factor and the carry factors deliver  $R^2$ s between 18% and 87%, pointing to large shares of global shocks in the dynamics of exchange rates, as well as large differences across countries.

The share of global shocks in exchange rates and pricing kernels is a key moment to consider in macroeconomic models. As an example, in the model of Colacito and Croce (2011), global long-run risk shocks drive most of the variation in the pricing kernels, but they do not affect exchange rates (i.e., the differences in pricing kernels). As a result, the model solves the Backus and Smith (1993) and Brandt, Cochrane, and Santa-Clara (2006) puzzle: pricing kernels are volatile and equity risk premia are high, yet exchange rates are as volatile as in the data. The findings in this paper raise the bar: global shocks cannot cancel out from domestic and foreign pricing kernels because they actually account for a large part of exchange rate variation. More generally, the factor regressions offer a set of precisely estimated, useful, and challenging new moments to match.

The paper is organized as follows. Section 1 reviews the related literature. Section 2 shows that the dollar and carry factors explain a large share of bilateral exchange rates. Section 3 uncovers a new cross-section of currency excess returns that is explained by the dollar risk factor. Section 4

compares the shares of currency systematic risk to measures of world market integration. Section 5 spells out the preference-free implications of the previous findings, relating them to recent models in international finance. Section 6 concludes. A data Appendix at the end of this document describes the data set. The changes in exchange rates, the interest rates, the carry and dollar factors, as well as the time-varying loadings on those factors and the dollar beta portfolios, are available on my website and thus the results in this paper can be easily replicated. A separate Appendix, also available on my website, reports several robustness checks and extensions, as well as model proofs and simulation details.

## 1 Related Literature

Numerous studies in the 1970s and early 1980s report large  $R^2$ s in regressions of *levels* of exchange rates on various macroeconomic variables [see, for example, Frankel (1979) and Hooper and Morton (1982)]. But both sides of those regressions feature highly persistent variables, and in-sample fits do not lead to out-of-sample accurate predictions. Meese and Rogoff (1983) show that a large class of models fails to outperform the random walk in forecasting changes in exchange rates for individual currency pairs out-of-sample. Since Meese and Rogoff (1983), the standard view in international economics is that individual exchange rates follow random walks, with perhaps small departures from random walks at very high frequencies (Evans and Lyons, 2005). Engel and West (2005) show that exchange rates are very close to random walks when fundamentals are not stationary and risk premia are constant. The findings of this paper, which pertain to *changes* in exchange rates, are not inconsistent with the random walk view of exchange rates: the dollar factor, and conditional and unconditional carry factors are not persistent variables and the common shocks that account for each currency pair could be random walks.

This paper is related to principal component analyses of exchange rates: the dollar factor is close to the first principal component, although the carry factors are different from the other principal components. Early examples of principal component analyses include Diebold and Nerlove (1989) who propose a multivariate latent-variable model of seven currencies in which the common factor

displays ARCH. Bollerslev (1990) estimates a GARCH model with constant conditional correlation on a set of five weekly exchange rates. More recently, Engel, Mark, and West (2009) propose a principal component decomposition of exchange rates and use the components to predict bilateral exchange rates. None of these papers reports the share of common variation of each currency pair. More importantly, they do not offer any risk-based interpretation of their principal components. To the contrary, the current paper only uses risk factors to account for exchange rates in U.S. dollars. The paper shows that the dollar factor is a *risk* factor in the asset pricing sense by building a novel cross-section of currency excess returns and linking systematic currency to equity and fixed income risk. The paper thus focuses on two economically motivated factors, which account for the cross-section of currency excess returns and have a natural interpretation in any no-arbitrage model.

Although this paper builds on Lustig, Roussanov and Verdelhan (2011), it is clearly distinct from it: Lustig et al. (2011) do not report  $R^2$ s on any time-series regressions of bilateral exchange rates. They focus on the dynamics of *portfolios* of currencies. When they check their asset pricing results on bilateral exchange rates, they report only measures of cross-sectional, not time-series, fit. More generally, the current paper is part of a growing literature that focuses on currency portfolios to study currency risk.<sup>1</sup> Portfolios are a very useful tool to extract and study risk premia: they are built in order to average out idiosyncratic components and to focus only on systematic risk. Thus, by construction, they are silent on the share of systematic versus idiosyncratic variation in *each* currency pair, which is the objective of this paper.

Finally, the findings point to similar results obtained on equity and bond markets. Roll (1988) studies contemporaneous regressions of large individual U.S. stock returns on systematic risk factors

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<sup>1</sup>Following Lustig and Verdelhan (2005, 2007), DeSantis and Fornari (2008), Jurek (2008), Farhi, Fraiberger, Gabaix, Ranciere, and Verdelhan (2009), Galsband and Nitschka (2010), Christiansen, Rinaldo, and Soderlind (2011), Gilmore and Hayashi (2011), Hassan and Mano (2012), Menkhoff, Sarno, Schmeling, and Schrimpf (2012), and Mueller, Stathopoulos and Vedolin (2012) study the properties of one-month interest rate-sorted portfolios of currency excess returns. Ang and Chen (2010), Hu, Pan, and Wang (2010), Kozak (2011) consider new sorts, focusing on properties of the foreign yield curves at longer horizons or on liquidity risk. Lustig, Roussanov, and Verdelhan (2012) study the predictability of the dollar risk factor (thus focusing on one single currency portfolio), while Maggiori (2012b) uses a conditional-Capital Asset Pricing Model to price the dollar excess return. Rinaldo and Soderlind (2008) and Hoffmann and Suter (2010) study the risk characteristics of the Swiss franc and other safe-haven currencies.

and on the returns of other stocks in the same industry; he reports an average  $R^2$  of about 35% on monthly data and 20% on daily data. Steeley (1990) and Litterman and Scheinkman (1991) uncover a clear factor structure in bond returns, where three factors account for more than 95% of the total return variance. Currency markets do not appear much different.

## 2 Measuring Currency Systematic Variations

This section shows that the dollar and carry factors explain a large part of each currency pair variations, uncovering cross-country differences in the shares of systematic currency risk. I start from the usual UIP and macroeconomic regressions and then turn to the carry and dollar factors.

### 2.1 Interest Rates and Other Macroeconomic Variables

**Notation and Data** A lower case  $s$  denotes the log of the nominal spot exchange rate in units of foreign currency per U.S. dollar, and  $f$  the log of the one-month forward exchange rate, also in units of foreign currency per U.S. dollar. An increase in  $s$  means an appreciation of the home currency. Interest rate differences are derived from forward rates. In normal times, forward rates satisfy the covered interest rate parity condition; the forward discount is equal to the interest rate differential:  $f_t - s_t \approx i_t^* - i_t$ , where  $i^*$  and  $i$  denote the foreign and domestic nominal risk-free rates over the maturity of the contract.<sup>2</sup>

End-of-month series are built from daily spot and forward exchange rates in U.S. dollars and the sample period runs from November 1983 to December 2010. These data are collected by Barclays and Reuters and available on Datastream. Spot and forward exchange rates correspond to midpoint quotes. The Data Appendix lists all of the 13 developed and 18 emerging countries in the data set.

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<sup>2</sup>Akram, Rime, and Sarno (2008) study high-frequency deviations from covered interest rate parity (CIP). They conclude that CIP holds at daily and lower frequencies. While this relation was violated during the extreme episodes of the financial crisis in the fall of 2008, including or excluding those observations does not have a major effect on the results.



**UIP Redux** According to the UIP condition, the expected change in exchange rates should be equal to the interest rate differential between foreign and domestic risk-free bonds. The UIP condition is equivalent to an Euler equation for risk-neutral investors. It implies that a regression of exchange rate changes on interest rate differentials should produce a slope coefficient of one. Instead, empirical work following Tryon (1979), Bilson (1981) and Fama (1984) consistently reveals a slope coefficient that is smaller than one and very often negative. The international economics literature refers to these negative UIP slope coefficients as the UIP puzzle or forward premium anomaly.<sup>3</sup>

Negative slope coefficients mean that currencies with higher than average interest rates tend to appreciate, not to depreciate as UIP would predict. Investors in foreign one-period discount bonds thus earn the interest rate spread, which is known at the time of their investment, plus the bonus from the currency appreciation during the holding period. As a result, the forward premium anomaly implies positive predictable excess returns for investments in high interest rate currencies and negative predictable excess returns for investments in low interest rate currencies.

Panel I of Table 1 reports country-level results from the usual UIP tests:

$$\Delta s_{t+1} = \alpha + \beta(i_t^* - i_t) + \varepsilon_{t+1}.$$

As in the rest of the literature, betas are always below one, and most of them are negative. All but one are statistically insignificant. Out of 13 countries, only three lead to positive betas, none of which is statistically significant. As in the rest of the paper, all  $R^2$ s are adjusted for the degrees of freedom. The adjusted  $R^2$ s on these regressions are tiny, often negative, with a maximum of 1.7%, and an average of -0.2%. UIP tests are predictability tests because the returns on the nominal notes between  $t$  and  $t+1$  are known at date  $t$ . Yet  $R^2$ s are not much higher when using contemporaneous

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<sup>3</sup>The UIP condition appears to be a reasonable description of the data only in four cases. Bansal and Dahlquist (2000) show that UIP is not rejected at high inflation levels, and likewise Huisman, Koedijk, Kool and Nissen (1998) find that UIP holds for very large forward premia. Chaboud and Wright (2005) show that UIP is valid at very short horizons but is rejected for horizons above a few hours. Meredith and Chinn (2005) find that UIP cannot be rejected at horizons above 5 years. Lothian and Wu (2005) find positive UIP slope coefficients for France/U.K. and U.S./U.K. on annual data over 1800-1999, because of the 1914-1949 subsample.

macroeconomic variables.

[Table 1 about here.]

**Industrial Production and Inflation** To focus on the monthly frequency, I use U.S. industrial production growth rates and inflation differentials to account for changes in exchange rates. Panel II of Table 1 reports country-level results from the following regression:

$$\Delta s_{t+1} = \alpha + \beta(i_t^* - i_t) + \gamma \Delta ip_{t+1} + \delta(\Delta cpi_{t+1}^* - \Delta cpi_{t+1}) + \varepsilon_{t+1},$$

where  $\Delta ip_{t+1}$  denotes the log difference in the industrial production index and  $\Delta cpi_{t+1}^* - \Delta cpi_{t+1}$  denotes the foreign minus domestic log difference in the consumer price indices. Many adjusted  $R^2$ s are negative, and the highest value is 2.5%. Using foreign minus domestic (instead of U.S.) industrial production growth rates limits the sample in terms of countries and time windows but delivers similar results. The difference in industrial production appears significant only for Canada over the last 16 years. The adjusted  $R^2$ s are all below 3% and none of them appears significantly different from zero.

This experiment does not mean that exchange rates are unrelated to macroeconomic variables. It simply highlights how difficult it is to uncover such links at the monthly frequency and for individual currency pairs (instead of, for example, an annual frequency and portfolios of currencies). At the monthly frequency, I do not know of any macroeconomic variable that would account for changes in exchange rates as well as the carry and dollar factors.

## 2.2 Carry and Dollar Factors

**Developed Countries** Following Lustig and Verdelhan (2005, 2007), and like Lustig et al. (2011), I build six *portfolios* of currencies by sorting all developed and emerging countries each month according to their interest rates. By averaging out idiosyncratic risk and conditioning on interest rates, these portfolios deliver a cross-section of exchange rates and currency risk premia. The reader is referred to Lustig et al. (2011) for a detailed description of the portfolio statistics.

The carry factor, denoted  $Carry_{t+1}$ , is the average change in exchange rate between countries in the last portfolio (high interest rate countries) and those in the first portfolio (low interest rate countries).

The dollar factor is the average change in the dollar versus all the other currencies; it corresponds to the average change in exchange rate across all six portfolios at each point in time. Table 2 reports country-level results from the following regression:

$$\Delta s_{t+1} = \alpha + \beta(i_t^* - i_t) + \gamma(i_t^* - i_t)Carry_{t+1} + \delta Carry_{t+1} + \tau Dollar_{t+1} + \varepsilon_{t+1},$$

where  $Dollar_{t+1}$  corresponds to the average change in exchange rates against the U.S. dollar. Each variable is expressed in percentage points; as a result, the slope coefficient on the conditional carry factor is 100 times smaller than if all series were not in percentage points. As already noted, for each currency put on the left-hand side of a regression, that currency is excluded from any portfolio that appears on the right-hand side.

The loadings on the dollar factor are positive and statistically significant in all 13 developed countries, with values ranging from 0.3 to 1.6. The loadings on the dollar factor reflect the existence of a clear principal component in the dollar exchange rates. When the dollar appreciates, it does so against all developed currencies, but in different proportions. This common component explains a large share of the variation in bilateral exchange rates. The adjusted  $R^2$ s are now all between 19% and 91%. The average  $R^2$  among the 13 developed countries is 61%. Without the carry factors, the average  $R^2$  is 57%; unsurprisingly, the difference is particularly large for Australia and Japan (26% vs. 20% and 30% vs 24%), two textbook examples of carry traders' favorites. Without the dollar factor, adjusted  $R^2$ s range from 0% to 23%, with an average of 7%.

The conditional carry loadings ( $\gamma$ ) in Table 2 are positive in 11 out of 13 countries (the only exceptions are Canada and Japan, where the coefficients are negative but insignificant). They are positive and statistically significant in 9 out of 13 countries. These findings are consistent with those of Lustig et al. (2011): as already noted, in their portfolios of currencies sorted by interest rates, the higher the interest rate (i.e., going from the first to the last portfolio), the larger

the loading on the carry factor. The findings in the current paper are, however, different from those of Lustig et al. (2011), which pertain to cross-sectional differences in interest rates (i.e., whether one currency has a higher interest rate than another). Here, the conditional carry loading indicates that, for a given country, times of larger interest rate differences are also times of higher comovement with the carry factor. By focusing on one bilateral exchange rate at a time, each test explores time-series, not cross-sectional, variations (see Hassan and Mano (2012) for more on this difference).

The total sensitivity of each bilateral exchange rate to the carry factor depends on the conditional and unconditional carry components. Table 2 shows that the corresponding slope coefficients are jointly statistically significant at the 1% (10%) confidence level in 11 (12) out of 13 countries (the only exception is the U.K.).<sup>4</sup> A total sensitivity that is positive means that the foreign currency depreciates when the carry does too. In the data, the Japanese yen tends to appreciate when the carry factor tanks, while the Australian dollar tends to depreciate. Such currency movements correspond to the funding and investment roles of these currencies that are commonly reported in articles on carry trades. The Japanese yen appreciates in bad times, while the Australian dollar depreciates: this difference is at the heart of any risk-based explanation of carry trades. But for many countries, the total sensitivity to the carry factor switches sign along the sample. For example, the Swiss franc depreciated in the 1980s (a time of relatively high Swiss interest rates) when the carry factor paid badly; recently, it has appreciated. This result is particularly clear for estimates based on rolling windows. The “safe-haven” characteristic of the Swiss franc thus appears sample-dependent, and linked to the interest rate level.

[Table 2 about here.]

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<sup>4</sup>The interest rate difference, the conditional carry, and the unconditional carry factors are correlated, and thus standard errors offer only a partial view of the significance of each coefficient. Table 2 thus reports the result of a Wald test where the null hypothesis is that the loadings  $\gamma$  and  $\delta$  on the conditional and unconditional carry factors are jointly zero. The null hypothesis is rejected at the 10% confidence interval for all developed countries except the U.K. At a daily frequency, the null hypothesis is rejected at the 1% level for all countries. Results are reported in the separate Appendix.

**Emerging Markets** To explore exchange rate dynamics further, and as an initial robustness check, Table 3 reports similar tests on a set of 18 developing countries (using the same factors as for developed countries). Adjusted  $R^2$ s tend to be high for floating currencies. A simple finding emerges: for floating currencies, the results tend to be similar to those of developed countries, whereas pegs, reassuringly, appear different. On the one hand, loadings on the dollar factor are positive and significant for 15 out of 18 countries; loadings on the carry factors are jointly significant for nine countries; and  $R^2$ s in Table 3 range from 10% to 75% for floating currencies. On the other hand, Hong Kong, Saudi Arabia, and the United Arab Emirates, which have pegged their currencies to the U.S. dollar at some point in the sample, do not exhibit significant loadings on the dollar factor. This broad dichotomy hides more subtle nuances: for example, Thailand and Malaysia, although they also experienced currency pegs, do not appear much different from the other developing countries. The carry and dollar factors thus highlight the uncertainty behind exchange rate regime classifications, and rolling window estimates could be used to refine such classifications.

[Table 3 about here.]

I consider additional robustness checks (other factors, like momentum; daily, quarterly, and annual frequencies; bid-ask spreads; and different base currencies) and also study the stability of the factor structure, comparing the implied pseudo-predictability to those of random walks and principal components, as well as the time-variation in the share of systematic currency risk. The results are presented and commented in the separate Appendix. They all reinforce the findings presented in the main text.

## 2.3 Comparison to the Literature

The large cross-country differences in  $R^2$ s and the significant loadings on the dollar and carry factors are the first benchmark results of this paper. I now contrast these results to the previous literature on UIP tests and to pairwise currency correlations.

**Omitted Variables** The literature on exchange rates contains many estimations of the UIP condition. Froot and Thaler (1990) reported that, in a survey of 75 published estimates, the slope coefficient of the regression of changes in exchange rates on interest rate differences was always below unity (positive in a very few cases, and  $-0.88$  on average). Many more papers have run similar tests and offered potential explanations. A simple search in *Scopus* returns 310 articles published since 1990 that mention “exchange rates” and either “uncovered interest rate parity” or “forward premium” or “carry trade” in their title, abstract or keywords. Engel (1996) and Chinn (2006) provide recent surveys.

This paper shows that all these UIP tests miss key explanatory variables. The conditional carry and dollar factors are jointly statistically significant for almost all currencies and they are correlated to interest rate differences. These findings imply that all classic UIP tests suffer from an omitted variable bias. Slope coefficients on interest rate differences vary when the carry, conditional carry, and dollar factors are added as explanatory variables. They tend to increase from an average of  $-0.72$  in UIP regressions to an average of  $-0.19$  with the carry and dollar factors. Neither the former nor the latter coefficients are significant. The latter still does not measure the interest rate sensitivities of exchange rates because the conditional carry and the dollar factor also respond to interest rate differences.

**How Large Are the  $R^2$ s?** The factor regressions deliver larger  $R^2$ s than interest rates and macroeconomic variables. The large  $R^2$ s can also be favorably compared to those of simple pairwise currency regressions. For each currency  $i$ , let us regress it on a constant and each currency  $j$  available in the sample ( $i \neq j$ ):

$$\Delta s_{t+1}^i = \alpha + \beta \Delta s_{t+1}^j + \varepsilon_{t+1}^{i,j}.$$

Each regression delivers a pair-specific  $R^2$ , denoted  $R^{2^{i,j}}$ . This is the  $R^2$  equivalent to the Option-Metrics correlation coefficient that many practitioners use. Some currencies are highly correlated, and thus lead to high  $R^2$ s: the Australian and New Zealand dollars offer one such example. But

without knowing ex ante the intrinsic correlation matrix, one would expect, for each currency  $i$ , an  $R^2$  that corresponds to the mean of all the estimates that involve currency  $i$ . For each currency  $i$ , the standard deviation of all the estimates  $R^{2^{i,j}}$  gives a measure of the uncertainty around its mean  $R^2$ .

Figure 1 summarizes the findings, comparing  $R^2$ s obtained with the carry and dollar factors (vertical axis) to those obtained from random univariate regressions (horizontal axis). If the carry and dollar factors are of any help in capturing exchange rate variation, all points in Figure 1 should be above the 45-degree line, which is clearly the case. It turns out that the factors'  $R^2$ s are, for all developed countries, at least one-standard deviation above the mean  $R^2$  estimated from the pairwise univariate regressions above. The same is true for all developing countries, except three: Saudi Arabia and the United Arab Emirates, which are unsurprising outliers since they have pegged their currencies to the U.S dollar, and Indonesia, which has few observations. For developed currencies, one can always handpick another developed currency that is highly correlated with the currency under study, and thus will lead to a high  $R^2$ . But there is no currency that delivers this feature for all exchange rates. Conversely, the dollar and the carry factor have the same economic interpretation for all currencies.<sup>5</sup> They deliver  $R^2$ s that are large, not only compared to those of UIP tests (arguably, a low bar) and macroeconomic variables, but also to those of bivariate exchange rate regressions.

[Figure 1 about here.]

### 3 Dollar Risk

The carry and dollar factors offer more than a simple statistical description of individual exchange rates: they are *risk* factors, in the asset pricing sense. Thus, they capture the share of systematic risk in bilateral exchange rates. The literature review above presents evidence in favor of a risk-based interpretation of the carry factor. This section presents new evidence on the dollar risk

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<sup>5</sup>The carry and dollar factors only differ across currencies because of the exclusion of the currency under study. But simple carry and dollar factors that use the whole sample produce similar results.

factor.

### 3.1 Portfolios of Countries Sorted by Dollar Exposures

Lustig et al. (2011) show that the average forward discount rate of developed countries (i.e., the average interest rate difference between foreign and U.S. short-term interest rates) predicts the returns on the aggregate currency portfolio and its exchange rate component (i.e., the dollar factor). They report a high average excess return, along with a high Sharpe ratio on a simple investment strategy that exploits this predictability by going long the aggregate currency market when the average forward discount rate is positive and short otherwise. The resulting excess return is the aggregate conditional dollar excess return.

I combine the dollar predictability shown in Lustig et al. (2011) with the heterogeneity in the loadings on the dollar shown in the previous section in order to build a new large cross-section of portfolio excess returns. The new portfolios are based on each currency's time-varying exposures to the dollar factor. At each date  $t$ , each currency  $i$ 's change in exchange rate is regressed on a constant and the dollar and carry factors, as in the previous section, using a 60-month rolling window that ends in period  $t - 1$ . Currency  $i$ 's exposure to the dollar factor is denoted  $\tau_t^i$ ; it only uses information available at date  $t$ . Currencies are then sorted into six groups at time  $t$  based on the slope coefficients  $\tau_t^i$ . Portfolio 1 contains currencies with the lowest exposures ( $\tau$ ), while portfolio 6 contains currencies with the highest exposures. I refer to these portfolios as dollar beta-sorted. At each date  $t$  and for each portfolio, the investor goes long if the average forward discount rate among all developed countries is positive and goes short otherwise.

Panel I of Table 4 reports summary statistics on the portfolios of countries sorted by dollar exposures. Average log excess returns range from 1.3% to 7.1% on an annual basis from portfolios 1 to 6. Mean excess returns on portfolios 2 to 6 are statistically different from zero (the standard errors are obtained by bootstrapping and thus take into account the sample size). Taking bid-ask spreads into account reduces the average excess returns (from 0.6% to 5.8%) but does not change the cross-sectional pattern. Unsurprisingly, the volatility of excess returns increases from portfolio



1 to 6: portfolio 6 contains more systematic variation than portfolio 1 by construction.

[Table 4 about here.]

The cross-section of dollar beta-sorted currencies is novel. How does it relate to previous work? The comparison of unconditional and conditional currency excess returns links this result to exchange rate predictability: without conditioning on the average forward discount rate, the dollar beta-sorted portfolios deliver a cross-section of gross average excess returns ranging from 0.3% to 2.4%. Because the average forward discount rate predicts future dollar returns, the conditional average excess returns are much larger, particularly for large loadings on the dollar factor: the higher the loading on the dollar factor, the more predictable the future currency excess returns, and thus the higher the average conditional excess returns. The set of average currency excess returns is thus consistent with the aggregate predictability results in Lustig, Roussanov, and Verdelhan (2012).

The cross-section of dollar beta-sorted currencies is key to estimating the price of dollar risk. This estimation does not appear in previous work on currency carry trades, because all portfolios of currencies sorted by interest rates load in the same way on the dollar factor. They do not offer the different dollar exposures needed to estimate the price of dollar risk. As a result, the dollar factor plays the role of a constant in the second stage of a Fama-McBeth regression on currency carry trades. The novel cross-section of dollar beta-sorted currencies, on the contrary, implies a significant market price of dollar risk.

### 3.2 The Price of Dollar Risk

Panel II of Table 4 reports Generalized Method of Moments (GMM) and Fama-McBeth (FMB) asset pricing results. The market price of risk  $\lambda$  is positive, significant, and close to the mean of the risk factor, as implied by a no-arbitrage condition.<sup>6</sup> Average excess returns of the dollar beta-sorted portfolios correspond to the covariances between excess returns and a single risk factor,

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<sup>6</sup>The Euler equation implies a beta-pricing model:  $E[R^e] = \beta\lambda$ , where  $\beta$  measures the quantity of risk and  $\lambda$  the price of risk. Since the risk factor is a return, the Euler equation applies to the risk factor itself, which as a beta of one, and thus implies that  $E[R_{Dollar}^e] = \lambda$ .

the aggregate conditional dollar excess return. The pricing errors are not statistically significant. Panel III of Table 4 reports Ordinary Least Squares (OLS) estimates of the factor betas. Betas increase monotonically from 0.11 to 1.52; they are precisely estimated. The alphas (which measure the returns after correction for their risk exposure) are not statistically different from zero. The dollar risk differs from both the carry risk and the equity risk. The carry trade risk factor (which is dollar neutral) and the aggregate U.S. stock market excess returns *cannot* account for the excess returns of portfolios sorted on dollar exposures: for the CAPM, loadings (not reported) tend to increase from portfolios 1 to 6, but they are too small, implying a large market price of risk that is not in line with the mean U.S. stock market excess return. Likewise, the loadings on the carry trade risk factor are small and imply large and statistically significant pricing errors. Sorts on dollar exposures thus reveal a novel cross-section of currency risk premia.

Figure 2 reports the realized and predicted average excess returns. Each portfolio  $j$ 's actual excess return is regressed on a constant and the conditional dollar excess return to obtain the slope coefficient  $\beta^j$ . Each predicted excess return then corresponds to the OLS estimate  $\beta^j$  multiplied by the mean of the conditional dollar excess return. Figure 2 clearly shows that predicted excess returns are aligned with their realized counterparts. An investor who takes on more dollar risk is rewarded by higher excess returns on average. The dollar risk is intuitive. When the U.S. economy approaches a recession, U.S. short-term interest rates tend to be low relative to other developed economies: the average forward discount is thus positive. If U.S. investors then buy a basket of currency forward contracts, they are long foreign currencies and short the U.S. dollar. They thus run the risk of a dollar appreciation during difficult times for them. The dollar appreciation is not unlikely: if markets are complete, the U.S. dollar should actually appreciate when pricing kernels are higher in the U.S. than abroad, i.e., when the U.S. experiences relatively bad times. Shorting the dollar is thus a risky strategy and risk-averse investors expect to be compensated for bearing that risk.

[Figure 2 about here.]

As a robustness check, I run country-level Fama and MacBeth (1973) tests, using country-level

excess returns as test assets. The country-level results confirm the previous findings: the dollar risk is priced in currency markets, and the price of risk is not statistically different from the mean of the risk factor's excess return, as no-arbitrage implies. The pricing errors are unsurprisingly larger than those obtained on currency portfolios, but the null hypothesis that all pricing errors are jointly zero cannot be rejected.

Portfolios of countries sorted by dollar exposures and conditional excess returns at the country level thus offer clear evidence in favor of a risk-based explanation of exchange rates. Providing additional evidence, the next section explores the link between shares of systematic risk in different markets.

## 4 World Comovement and Systematic Currency Variation

If financial markets are complete, the change in the nominal exchange rate  $\Delta s$  between the home country and foreign country  $i$  is equal to:

$$\Delta s_{t+1}^i = m_{t+1} - m_{t+1}^i,$$

where  $m$  and  $m^i$  denote the nominal pricing kernels of respectively the domestic and foreign investors.<sup>7</sup> Pricing kernels, also known as stochastic discount factors (SDFs), correspond to the intertemporal marginal rates of substitution.

Since all exchange rates are defined with respect to the U.S. dollar,  $m$  is a common component of all exchange rates. Shares of systematic currency risk vary across countries because the properties of foreign stochastic discount factors vary. Such variation should affect other asset prices (e.g., equity and fixed income markets). This section thus compares the cross-sectional variation in the

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<sup>7</sup>This result derives from the Euler equations of the domestic and foreign investors buying any asset  $R^i$  that pays off in foreign currency:  $E_t[M_{t+1}R^i S_t/S_{t+1}] = 1$  and  $E_t[M_{t+1}^i R^i] = 1$ . When markets are complete, the pricing kernel is unique and thus exchange rates are defined as  $S_{t+1}/S_t = M_{t+1}/M_{t+1}^i$ , or in logs  $\Delta s_{t+1} = m_{t+1} - m_{t+1}^i$ . This definition of exchange rates holds even if markets are incomplete. In that case, exchange rates correspond to the projections of  $M$  and  $M^i$  on the space of traded assets (which includes exchange rates). If the law of one price applies and investors can form portfolios freely, then there is a unique pricing kernel in the space of traded assets. The projections of  $M$  and  $M^i$  on that space are thus unique and exchange rates are well defined. See Chapter 4 in Cochrane's (2005) textbook.

shares of systematic currency risk to the cross-sectional variation in the shares of systematic equity and fixed income risk, as well as other macroeconomic measures of comovement, often considered in the context of world integration studies. I first define these different measures of systematic risk, and then study their links.

## 4.1 Measures of Systematic Risk in Asset and Good Markets

**Systematic Risk in Asset Markets** Comovements between stock returns across countries are the object of a large literature that is too wide to survey here. The reader is referred to a recent survey article by Lewis (2011) and recent evidence at the industry level by Bekaert, Hodrick, and Zhang (2009) and Bekaert, Harvey, Lundblad and Siegel (2011).

I adopt a simple approach, focusing on the world Capital Asset Pricing Model (CAPM) and Fama and French (1993) factors. The share of systematic equity risk at the country-level corresponds to the share of each foreign country's equity returns explained by world equity factors, i.e. the  $R^2$  of the following regression:

$$r_{t+1}^{m,\$} = \alpha + \beta r_{t+1}^{m,world,\$} + \gamma r_{t+1}^{hml,world,\$} + \varepsilon_{t+1},$$

where  $r_{t+1}^{m,\$}$  denotes the returns on a given foreign country's MSCI stock market index,  $r_{t+1}^{m,world,\$}$  is the MSCI world return equity index, and  $r_{t+1}^{hml,world,\$}$  is the difference between the returns on the world MSCI value equity index and the world MSCI growth equity index (i.e., high minus low book-to-market equity returns). Equity indices correspond to country-level indices, and thus no additional local risk factor is used. Pukthuanthong and Roll (2009) argue in favor of similar  $R^2$ s to measure world market integration.<sup>8</sup> As in Bekaert and Harvey (1995), Bekaert, Hodrick, and Zhang (2009), and Pukthuanthong and Roll (2009), all returns are expressed in U.S. dollars.

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<sup>8</sup>Note that such measures of comovement are not necessarily informative about asset market integration or risk-sharing in general; their interpretation in those terms depends on the model of the economy. For example, models with segmented markets may exhibit a large correlation of the pricing kernels of domestic and foreign market participants in an overall poorly diversified economy [see e.g. Alvarez, Atkeson, and Kehoe (2002, 2009)]. Likewise, asset market integration may be assessed through the significance of alphas in the regression above, but the assessment then depends on the validity of the risk factors.

Systematic equity risk is estimated on the same sample period as currency risk for each country. Likewise, adjusted  $R^2$ s on bond markets are derived from:

$$r_{t+1}^{b,\$} = \alpha + \beta r_{t+1}^{m,world,\$} + \gamma r_{t+1}^{b,world,\$} + \varepsilon_{t+1},$$

where  $r_{t+1}^{b,\$}$  denotes the returns in U.S. dollars on a foreign country's 10-year bond return index and  $r_{t+1}^{b,world,\$}$  is the world bond return, obtained as the average of all the 10-year bond returns. On equity (bond) markets, adjusted  $R^2$ s range from 27% (32%) to 87% (89%).

**Comovement of GDP and consumption** In the international real business cycle literature, pricing kernels respond to shocks to fundamental macroeconomic variables. Thus, shares of systematic currency risk could be related to measures of comovement based on consumption or output growth for example. Adjusted  $R^2$ s on consumption growth rates are derived from:

$$\Delta y_{t+1} = \alpha + \beta \Delta y_{t+1}^{world} + \varepsilon_{t+1},$$

where  $\Delta y_{t+1}$  denotes the annual growth rate of real foreign consumption and  $\Delta y_{t+1}^{world}$  corresponds to the annual growth rate of the world consumption (measured as the sum of all OECD consumptions). Adjusted  $R^2$ s on GDP growth rates, which are available for more countries and on longer samples, are derived similarly. The GDP and consumption series, measured in U.S. dollars or at purchasing power parity, come from the World Bank. They are only available at an annual frequency.

## 4.2 Systematic Currency Risk and World Comovement

The different shares of systematic currency risk across countries appear related to measures of world equity, fixed income, and macroeconomic comovement. A simple cross-country regression confirms the findings:

$$R_i^{2,FX} = \alpha + \beta R_i^{2,X} + \varepsilon_i,$$

where  $R_i^{2,FX}$  denotes the share of systematic variation in the exchange rate of country  $i$ , obtained as in the previous sections using the carry and dollar factors.  $R_i^{2,X}$  denotes the share of systematic variation measured with either equity returns, bond returns, or macroeconomic variables.

Panel I of Table 5 reports the slope coefficients ( $\beta$ ) on this cross-country regression. On the full sample period (1983–2010) as on the post-1999 sample, all the slope coefficients are positive and significant. A large share of systematic risk in equity returns, bond returns, output growth, or consumption growth is associated with a large share of systematic currency risk. Slope coefficients range between 0.6 and 1.1 across asset and macroeconomic measures of world comovement.

[Table 5 about here.]

As an example, Figure 3 reports the share of systematic *currency* risk as a function of the share of systematic *equity* risk, as well as one-standard error confidence intervals (obtained by bootstrapping) on both sides of each point estimate. The figure focuses on the 1999.1–2010.12 sub-sample, after the introduction of the euro. Over this sub-sample, the data set offers an (almost) balanced panel of both developed and emerging economies. Over the last 10 years, the higher the share of systematic risk on equity markets, the higher the share of systematic risk on currency markets.

Saudi Arabia (SA) and Hong Kong (HK) appear as outliers because their share of systematic equity risk is much higher than their share of systematic currency risk, which is close to zero. This finding is not surprising: again, both countries have pegged their currencies to the U.S. dollar throughout the sample. Among the developed countries, the U.K., for example, also exhibits higher shares of systematic equity risk than their currencies would suggest. How to interpret this finding? In any no-arbitrage model, exchange rates are informative about SDFs, whereas equity returns depend on both SDFs and the properties of the dividend processes. Thus, one potential hypothesis is that British firms (and thus their dividends) are more highly exposed to world shocks than their corresponding SDF would suggest, because of a higher leverage for example. I leave a thorough study of outliers for subsequent research and focus on the main finding, i.e., a strong link between measures of systematic equity and currency risk.

[Figure 3 about here.]

Part of this link is mechanical because asset returns and macroeconomic variables are converted into U.S. dollars.<sup>9</sup> If returns and growth rates in local currencies are much less volatile than currency movements, then tests of world integration boil down to measures of systematic currency risk using the dollar risk factor, and thus the link between systematic risk measured on equity, bond, output, and consumption on the one hand and currency markets on the other hand. The factor regressions in Section 2 thus offer new insight on the large literature on asset market integration, showing that part of its results are driven by exchange rates.

But the link between measures of systematic risk is not purely mechanical: it also exists for series converted at purchasing power parities (PPP). PPP series are much smoother than realized exchange rates and do not drive asset or macroeconomic measures of integration. Panel II of Table 5 reports similar slope coefficients ( $\beta$ ) to those of Panel I, but the underlying asset and macroeconomic measures of integration are now based on series converted using PPP. Again, all the slope coefficients are positive — ranging from 0.4 to 1.1. — and significant. Standard errors, however, are certainly underestimated. They do not take into account the uncertainty stemming from the estimation of each  $R^2$  on bond, equity, currency, consumption, or output time series, and the number of observations (one per country, subject to data availability) is limited.

To push the comparison further and to increase the sample size, I compare time-varying equity and currency  $R^2$ s. The methodology remains the same, but equity and currency regressions are now run on rolling windows of 60 months. Panels III and IV of Table 5 report the results obtained on series in either U.S. dollars or PPP dollars, on either the full sample or the last ten years, with country fixed effects and a time trend. Again, in all cases, the  $R^2$ s on currency markets are

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<sup>9</sup>As a first approximation, the world equity return corresponds to the average of the country-specific returns; likewise, the dollar risk factor is approximately the average change in exchange rates of the U.S dollar (ignoring capitalization weights for equity indices and portfolios for currencies). Tests of stock return comovement thus correspond to:

$$r_{t+1}^m - \Delta s = \alpha + \beta \left( \frac{1}{N} \sum_{i=1}^N r_{t+1}^{m,i} + Dollar_{t+1} \right) + \gamma \left( \frac{1}{N} \sum_{i=1}^N r_{t+1}^{hml,i} + Dollar_{t+1} \right) + \varepsilon_{t+1}.$$

positively and significantly related to the  $R^2$ s on equity and bond markets.<sup>10</sup>

To take a precise example, the integration of the Australian stock market appears highly correlated through time to the integration of the Australian dollar. But for Switzerland, this relationship does not exist, suggesting that trading on the Swiss franc is not linked to trading on the Swiss stock market. Data on capital flows collected by the U.S. Treasury are consistent with this finding. Capital flows linked to the sales and purchases of Swiss stocks and bonds account for less than 20% of the capital flows between the U.S. and Switzerland. But capital flows linked to the sales and purchases of Australian stocks and bonds account for more than 40% of the capital flows between the U.S. and Australia.

Overall, the cross-country regression results reported in Table 5 highlight a clear link across measures of systematic risk obtained on different asset and good markets. These results, combined with the asset pricing power of the dollar factor, point to a risk-based approach to exchange rates. The carry and dollar factors thus not only measure the share of systematic variation, but also the share of systematic risk in bilateral exchange rates. Pursuing this risk perspective on exchange rates, the next section highlights the key theoretical implications of the empirical findings.

## 5 Implications and Interpretations

In this section, I first highlight the preference-free implications of the cross-country differences in currency systematic risk, before reviewing a reduced-form and a macro-finance model to propose an interpretation of the carry and dollar factors.

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<sup>10</sup>All slope coefficients are significant (even after clustering by country), except for bond markets in PPP dollars over the last ten years: there, systematic risks in currency and bond markets share a common strong upward trend, although there is not enough variation around this trend to measure an additional link between shares of systematic risk. The common trend, however, implies a strong link between the two markets.



## 5.1 Preference-free Results

Let us start again from the definition of the change in the nominal exchange rate  $\Delta s^i$ :

$$\Delta s_{t+1}^i = m_{t+1} - m_{t+1}^i,$$

where  $m$  and  $m^i$  denote the nominal pricing kernels of the domestic and country  $i$  investors. Without loss of generality, each pricing kernel can be decomposed into country-specific and world shocks:

$$\Delta s_{t+1}^i = \overbrace{m_{t+1}^{US-spec.} + m_{t+1}^W}^{m_{t+1}} - \overbrace{m_{t+1}^{i-spec.} + m_{t+1}^{W,i}}^{m_{t+1}^i},$$

where  $m_{t+1}^{US-spec.}$  and  $m_{t+1}^{i-spec.}$  denote the U.S.- and country  $i$ -specific components of the U.S. and country  $i$  pricing kernels, while  $m_{t+1}^W$  and  $m_{t+1}^{W,i}$  denote their respective global components. Each component can be a vector of several shocks. By construction, the country-specific shocks are orthogonal across countries and orthogonal to the world shocks:  $cov(m^{i-spec.}, m^{j-spec.}) = 0$  and  $cov(m^{i-spec.}, m^{W,j}) = 0$ , for any  $i, j$ . Bilateral exchange rates depend on (home minus foreign) differences in country-specific and world shocks. Foreign country-specific shocks drive part of the exchange rate movements, but since they can be diversified away, they are not part of a risk factor built from the perspective of the representative U.S. investor. In this general framework, I consider the implications of the dollar and carry factors and the dollar beta-sorted portfolios, before turning to the share of global shocks in exchange rates.

**Dollar Risk Factor** In large baskets of currencies, foreign country-specific shocks average out (assuming that there are enough currencies in the baskets for the law of large number to apply). The dollar risk factor is the average exchange rate of all currencies defined in terms of U.S. dollars and thus corresponds to:

$$Dollar_{t+1} = \frac{1}{N} \sum_i \Delta s_{t+1}^i = m_{t+1}^{US-spec.} + m_{t+1}^W - \frac{1}{N} \sum_i m_{t+1}^{W,i},$$

where  $N$  denotes the number of currencies in the sample. As a result, the dollar risk factor depends on both U.S.-specific and world shocks, but not on foreign-specific shocks. The covariance between the change in exchange rate and the dollar risk factor is:

$$\text{cov}(\Delta s^i, \text{Dollar}) = \text{Var}(m^{US-spec.}) + \text{Cov}(m^W, m^W - \frac{1}{N} \sum_i m^{W,i}) - \text{cov}(m^{W,i}, m^W - \frac{1}{N} \sum_i m^{W,i}).$$

The first two terms are the same for all currencies. In the data, there are large differences in the slope coefficients of bilateral exchange rates on the dollar risk factor. Thus the third term in the covariance decomposition must be non-zero. In other words, the dollar factor must depend on world shocks, not only on U.S.-specific shocks, and foreign pricing kernels must differ in their loading on those global shocks.

**Carry Risk Factor** The carry risk factor is the average exchange rate of high- versus low-interest rate currencies:

$$\text{Carry}_{t+1} = \frac{1}{N_H} \sum_{i \in H} \Delta s_{t+1}^i - \frac{1}{N_L} \sum_{i \in L} \Delta s_{t+1}^i = \frac{1}{N_H} \sum_{i \in H} m_{t+1}^{W,i} - \frac{1}{N_L} \sum_{i \in L} m_{t+1}^{W,i},$$

where  $N_H$  ( $N_L$ ) denotes the number of high (low) interest rate currencies in the sample. The carry factor, defined as a difference in baskets of exchange rates, is dollar-neutral. As Lustig et al. (2011) show, the carry factor accounts for world shocks, provided that pricing kernels differ in their exposure to world shocks — without heterogeneity, the carry factor would not exist. Lustig et al. (2011) show that these world shocks are priced globally. The intuition is the following. Borrowing in low interest rate currencies and lending in high interest rate currencies provides the same return to any investor in the world. Since these returns are the same from the perspective of any investor, they must compensate them for taking on global risk that is priced globally (i.e. with the same price across countries). Total returns to carry trades vary across investors only because they need to convert their gains expressed in low interest rate currencies back into their own currency.

**Dollar-beta Portfolios** The dollar-beta portfolios corresponds to sorts on the covariances between some world shocks in each country’s pricing kernel and the corresponding world shocks in the U.S. pricing kernel. The correlation between the dollar and carry factors is less than 0.1. Moreover, recall that the carry factor does not price the cross-section of dollar-beta portfolios. In other words, the exchange rate of the high dollar beta portfolio minus the exchange rate of the low dollar beta portfolio has a low correlation with the carry trade factor. Thus the dollar factor must reflect world shocks that are different from the world shocks captured by the carry factor. Any model in international finance must then feature two kinds of global shocks, one of them being priced globally.

**Global Systematic Shocks in Bilateral Exchange Rates** The carry factor depends only on global shocks priced globally, while the dollar factor depends on both U.S. and global shocks. The factor regressions are thus informative about the share of global shocks in the pricing kernels. In order to determine the share of global risk factors in exchange rates, I regress the changes in bilateral exchange rates on the conditional and unconditional carry factors and the global component of the dollar factor. This global component is measured via the following long-short investment strategy: long the high dollar beta portfolio and short the low dollar beta portfolio. This long-short difference cancels out the U.S.-specific component of the U.S. pricing kernel and focuses on its global component.

Table 6 reports the results. They are obtained on a smaller number of observations than in the previous tables because building the dollar-beta portfolios uses 60 observations. Loadings on the global component of the dollar factor are significant for all developed currencies. Unsurprisingly, the global component of the dollar factor accounts for a lower share of the exchange rate variations than the dollar factor itself. The difference in  $R^2$ s range from 2 to 15 percentage points. Overall, global shocks account for a large share of the exchange rate changes, with  $R^2$ s ranging from 17% to 82%.

[Table 6 about here.]

**The Return of an Old Puzzle** The large share of global shocks in exchange rates constitute a new challenge for models in international economics. Backus and Smith (1993) note that constant relative risk-aversion implies that changes in real exchange rates should be perfectly correlated to relative consumption growth rates. In the data, however, the unconditional correlation is close to zero, if not negative. Generalizing this point in a preference-free setting, Brandt, Cochrane and Santa-Clara (2006) point that volatile pricing kernels (as implied by the Hansen and Jagannathan (1991) bounds for example) imply very volatile exchange rates, unless pricing kernels are highly correlated across countries.<sup>11</sup> In the data, this correlation must be above 0.9, in strong contrast to the low cross-country correlation of consumption growth rates.

Colacito and Croce (2011) propose a solution to the Backus and Smith (1993) and Brandt et al. (2006) quandaries. As noted in the introduction, the solution relies on global long-run risk shocks driving the pricing kernels, but not affecting their differences and thus exchange rates. A crucial assumption here is that countries are symmetric such that global shocks cancel out. This solution to the puzzle needs to be refined in light of the new evidence reported in this paper: global factors account for a significant share of the exchange rate variations. As a result, the Backus and Smith (1993) and Brandt et al. (2006) puzzles are back.

The empirical findings clearly indicate that countries must differ in their pricing kernels, more precisely in how their pricing kernels respond to global shocks. The sources and impact of heterogeneity in international finance is the subject of ongoing research. For example, Hassan (2012) considers the impact of different country sizes on asset returns, Gourinchas, Rey and Govillot (2011) entertain different risk-aversion coefficients, while Maggiori (2012a) studies the consequence of different levels of financial development.

I end this paper with a brief overview of a reduced-form model and a long-run risk model. These models aid the interpretation of the dollar and carry factors and the empirical results. The detailed presentations of the models, the proofs, and the simulations are in the separate appendix;

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<sup>11</sup>In complete markets (or for the projection of SDFs on the space of traded assets), the change in real exchange rates is  $\Delta q_{t+1}^i = m_{t+1}^{real} - m_{t+1}^{real,i}$ . Thus the variance of exchange rates is:  $Var[\Delta q_{t+1}^i] = Var[m_{t+1}^{real}] + Var[m_{t+1}^{real,i}] - 2cov(m_{t+1}^{real}, m_{t+1}^{real,i})$ .

the paper focuses only on the intuition derived from the models.

## 5.2 A Reduced-Form Model

Building on Cox, Ingersoll and Ross (1985) and Backus, Foresi and Telmer (2001), Lustig, Rousanov and Verdelhan (2012) start from the law of motion of the log pricing kernel or stochastic discount factor (SDF)  $m^i$ .

**Stochastic Discount Factors** Lustig et al. (2012) assume that each SDF responds to country-specific shocks (denoted  $u^i$ , uncorrelated across countries) and two global shocks (denoted  $u^w$  and  $u^g$ ). The SDFs are heteroscedastic because, as Bekaert (1996), Bansal (1997), and Backus et al. (2001) have shown, expected currency log excess returns depend on the conditional variances of the home and foreign (lognormal) SDFs.<sup>12</sup> If SDFs were homoscedastic, expected excess returns would be constant and the UIP condition would be valid. The time-varying volatilities of the country-specific shocks are also country-specific, while those of global shocks are either country-specific (for the  $u^g$  shocks) or global (for the  $u^w$  shocks). In the language of finance, the  $u^w$  shocks are priced globally, while the  $u^g$  shocks are priced locally. To be parsimonious, all the parameters that govern the SDFs are the same across countries, except for the exposure of each SDF to the shocks priced globally ( $\delta^i$ ).

In the model, bilateral exchange rates, interest rates, average forward discounts, and the carry and dollar risk factors can all be derived in closed forms. Particularly simple expressions arise in the following special case: assuming that (i) the U.S. SDF's exposure to world shocks priced globally ( $u^w$ ) is equal to the average exposure to world shocks across countries ( $\bar{\delta}^i = \delta$ ), and (ii) baskets of high and low interest rate currencies exhibit the same level of country-specific volatilities. The first assumption implies that the time-series mean of the average forward discount among developed countries is zero, a close approximation to the data. The second assumption implies that the conditional correlation between the dollar and carry factors is zero; in the data, the unconditional

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<sup>12</sup>If SDFs are not lognormal, expected currency excess returns depend on the higher cumulants, some of which must be time-varying.

correlation is less than 0.1. In that special case, the carry factor captures world shocks that are priced globally, while the dollar factor is driven by U.S.-specific shocks and world shocks that are priced locally. The average forward discount rate measures the gap between the average local volatility and its current value in the U.S. The change in exchange rate of country  $i$  responds to country  $i$ 's specific shocks, U.S.-specific shocks, and world shocks, in proportion to the differential exposure to these world shocks in the foreign country and in the U.S. SDFs.

**Intuitions for the Cross-Sections of Currency Excess Returns** The reduced-form model provides a rationale for the two risk factors and their associated cross-sections of currency returns.

Sorting countries by their interest rates leads to a large cross-section of currency excess returns in the data. Low (high) interest rate currencies offer low (high) average excess returns. In the model, as Lustig et al. (2012) show, sorting countries by their interest rates is similar to sorting by the exposure ( $\delta^i$ ) to global shocks priced globally ( $u^w$ ); high interest rate countries are low  $\delta^i$  countries. During a bad global shock,  $u^w < 0$ , these currencies depreciate: carry trades are risky because high (low) interest rate currencies depreciate (appreciate) in bad times.

In the model, sorting countries by their dollar betas is similar to sorting by the level of the country-specific price of risk, which is relevant for global shocks priced locally ( $u^g$ ). On the one hand, when the average forward discount rate is positive, the U.S. interest rate is lower than the world average, and the U.S.-specific market price of risk is above its long-run mean. Currencies with large loadings on the dollar factor are currencies with low country-specific prices of risk. They offer large excess returns on average to compensate investors for taking on global risks that are priced locally: during periods of a bad global shock ( $u^g < 0$ ), these currencies depreciate. On the other hand, when the average forward discount is negative, currencies with large loadings on the dollar factor are currencies with large country-specific prices of risk. They tend to appreciate during a bad global shock ( $u^g < 0$ ): again, this pattern is a source of risk since investors are short those currencies (and long the U.S. dollar) when the average forward discount rate is negative.

**Shares of Systematic Variation** Simulated data from the model, using the parameters reported in Lustig et al. (2012), reproduce the stark contrast between UIP and factor regressions. The interest rate difference accounts for only 0.2% to 0.4% of the bilateral exchange rate dynamics. The conditional and unconditional carry factors explain up to 25% of the exchange rate variation, thus slightly more than in the data. Adding the dollar factor delivers an average  $R^2$  close to 60%, in line with the average value in the data.  $R^2$ s and slope coefficients are time-varying, as they are in the data. The model also implies a large cross-section of interest rate-sorted portfolio returns, from  $-4.8\%$  to  $2.3\%$  per annum, thus delivering a high-minus-low carry trade excess return of  $7.1\%$ . The model, however, does not produce a large enough cross-section of dollar-sorted portfolio returns. This shortcoming is due to the low value predictive power of the average forward discount in the model (cf Lustig, Roussanov, and Verdelhan, 2012) and the low dispersion in dollar loadings.

Overall, the reduced-form model offers a simple interpretation for the dollar and carry factors and their associated cross-section of returns. I turn now to a long-run risk example that illustrates the theoretical challenges brought by the empirical findings.

### 5.3 Long-Run Risk Example

The long-run risk literature works off the class of preferences due to Kreps and Porteus (1978) and Epstein and Zin (1989). Following Bansal and Yaron (2004), real consumption growth in each country exhibits a persistent long-run expected growth component and is subject to temporary shocks. The dynamics of consumption growth are key because they influence the properties of the SDF: shocks to the SDF are ultimately consumption shocks, while the market prices of risk are derived in equilibrium from the preference parameters and the properties of the consumption process.

**Preferences and Endowments** I start from the symmetric, two-country, long-run risk model of Bansal and Shaliastovich (2012) and extend it to  $N$  different countries. Following the insights of the reduced-form model, consumption growth should respond to local shocks, global shocks with a time-varying volatility that is common across countries, and global shocks with a country-specific

volatility. The first set of global shocks is priced globally, while the second set is priced locally.

As in Bansal and Shaliastovich (2012), the long-run risk process is common across countries. Colacito and Croce (2011) show that the long-run risk components must be highly correlated across countries in order to deliver exchange rates that are as smooth as in the data. Bansal and Shaliastovich (2012) thus assume that the foreign economy shares the same long-run risk component as the domestic economy. This assumption is in line with empirical evidence: Nakamura, Sergeyev and Steinsson (2012) find a common long-run risk component across countries in the Barro and Ursua (2008) dataset.

Unlike in Bansal and Shaliastovich (2012), some temporary shocks to consumption growth are common across countries: yet, their volatilities are country-specific. In the data, consumption growth processes exhibit a low correlation across countries. To take into account this empirical fact, consumption growth also responds to country-specific temporary shocks; those shocks are homoscedastic.

The description of consumption growth is then rich enough to produce different kinds of local and global shocks. The key remaining question is the source of heterogeneity across countries. To narrow it down to the structural parameters, it is useful to derive the exchange rates and carry and dollar factors in the model.

**Sources of Cross-country Differences** For the carry factor to exist, high- and low-interest rate currencies must differ in their loadings on the common components of the SDF [see Lustig et al. (2011)]. Thus, countries must differ in their market price of short-run consumption growth risk, in their market price of long-run risk, and/or in their market prices of long-run volatility risk. Those market prices of risk depend on the risk-aversion coefficient, the EIS, the persistence of the long-run risk component, and the average wealth-consumption ratio (which is determined in equilibrium as a function of the preference and endowment parameters).

The dollar factor loads on global shocks only if the U.S. economy exhibits market prices of risk that differ from their cross-country average counterparts. It implies that the representative U.S. investor must differ in at least one of the dimensions above from the average world investor.



The model thus links the necessary heterogeneity to differences in structural parameters, but the empirical relevance of each one of them is an open question.

**Simulation Results** As an illustration, I consider two sets of simulations: countries differ either by their risk-aversion coefficient or by the share of their country-specific volatility. There is no reason for countries to differ only along one dimension; the dichotomy is only there for clarity. There are 36 countries in the simulation and the first country is the domestic country (i.e., the U.S.).<sup>13</sup> The model starts from simple consumption dynamics (consistent with the cross-country correlations) and produces reasonable risk-free rates, as well as equity and bond risk premia; it is thus an interesting laboratory to study exchange rates.

When countries differ by their shares of country-specific volatility, exchange rates are driven almost exclusively by short-term shocks; there is no role for long-term shocks as market prices of risk do not vary significantly across countries. As a result, the carry factors account for only 0.2% to 0.7% of the monthly exchange rate movements.

Differences in risk-aversion coefficients have a higher impact. The model produces a cross-section of currency excess returns when countries are sorted by their short-term nominal interest rates (or dollar betas). In the model, as in the data, interest rate differences explain almost none of the exchange rate time-series. When risk-aversion differences are small, long-run risk shocks play no role in exchange rate variation because of the absence of cross-country differences in their market prices of risk. In this case, most of the exchange rate variation is due to country-specific shocks. When a country's risk-aversion differs (e.g., 4 to 6), common long-run risk shocks explain up to 36% of the exchange rate changes. This large increase in the role of common shocks accounts for the increase in the share of systematic risk measured by the dollar and carry factors: the carry factors account for 0.8% to 8.4% of the monthly changes in exchange rates, while the carry and dollar factors jointly deliver  $R^2$ s between 47% and 75% across countries, in line with the empirical

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<sup>13</sup>Most preference and endowment parameters are those proposed in Bansal and Shaliastovich (2012). The calibration differs in risk-aversion, which varies from 4 to 6, instead of 10 in Bansal and Shaliastovich (2012). The inflation mean, the inflation shock volatility, and the expected inflation volatility are lower than in Bansal and Shaliastovich (2012) in order to reproduce the characteristics of aggregate inflation in the sample.

results in Section 2. Likewise, systematic risk in equity markets ranges from 70% to 88%. As in the data,  $R^2$ s increase on both equity and currency markets across countries. This is the main achievement of the model, showing that the cross-asset and cross-country results of Section 4 can be reproduced in a simple model.

In light of the general discussion earlier in this section, some weaknesses of the simulations are unsurprising: the correlation between changes in exchange rates and relative consumption growth (cf. the Backus and Smith (1993) puzzle) is too high, and the changes in exchange rates are too volatile, with standard deviations ranging between 17% and 19% (cf. the Brandt et al. (2006) puzzle). Moreover, the cross-sections of carry and dollar beta-sorted excess returns are small, the cross-sectional variation in equity  $R^2$ s measured in local currencies is small, and the carry factors explain a large share of the dollar beta-sorted portfolios; all of these outcomes are at odds with the data. Reproducing all the empirical results reported in this paper constitute a challenge for models in international economics.

## 6 Conclusion

The carry and dollar factors account for 20% to 90% of bilateral exchange rates at daily, monthly, quarterly, and annual frequencies. Both factors are *risk* factors. Cross-currency differences in the shares of systematic variation appear related to financial and macroeconomic measures of comovement. The findings point to a risk-based approach to currency markets, suggesting that a large share of exchange rate movement is due to global risk factors.

The findings are important for both academics and practitioners. For practitioners, they imply the need for global currency risk management. Many international mutual fund managers do not hedge their currency exposures. If changes in exchange rates were independent and random, then buying assets in many different currencies would offer a simple diversification mechanism of currency risk. This paper, however, shows that such diversification is wishful thinking. The decomposition of exchange rates in two risk factors also simplifies optimal portfolio allocation and hedging. As an example, a mean-variance investor allocating resources among, for example, 12

currencies would need to estimate the inverse of a  $3 \times 3$  (instead of a  $12 \times 12$ ) covariance matrix.

For researchers, the role of the carry and dollar factors motivate the study of systematic components in exchange rates, since those components account for a large share of bilateral exchange rate movements. Unlike many covariances between exchange rates and macroeconomic variables, the loadings on the risk factors and  $R^2$ s are precisely estimated. They offer a new source of cross-country differences and thus new potential targets for future models in macroeconomics that seek to link the deep characteristics of each economy to the behavior of its exchange rate.

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## Data Appendix

The main data set contains at most 37 different currencies of the following countries: Australia, Austria, Belgium, Canada, China (Hong Kong), Czech Republic, Denmark, Finland, France, Germany, Greece, Hungary, India, Indonesia, Ireland, Italy, Japan, Kuwait, Malaysia, Mexico, Netherlands, New Zealand, Norway, Philippines, Poland, Portugal, Saudi Arabia, Singapore, South Africa, South Korea, Spain, Sweden, Switzerland, Taiwan, Thailand, United Kingdom, as well as the euro area. The euro series start in January 1999. Euro area countries are excluded after this date; only the euro series remains.<sup>14</sup>

Some of these currencies have pegged their exchange rate partly or completely to the U.S. dollar over the course of the sample. They are in the sample because forward contracts were easily accessible to investors and their forward prices are not inconsistent with covered interest rate parity. Based on large failures of covered interest rate parity, however, the following observations are deleted from the sample: South Africa from the end of July 1985 to the end of August 1985; Malaysia from the end of August 1998 to the end of June 2005; Indonesia from the end of December 2000 to the end of May 2007; Turkey from the end of October 2000 to the end of November 2001; United Arab Emirates from the end of June 2006 to the end of November 2006.

Two important points need to be highlighted. First, note that for each currency inserted on the left-hand side of a regression, that currency is excluded from any portfolio that appears on the right-hand side. The objective is to prevent some purely mechanical correlation to arise. Excluding

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<sup>14</sup>All the countries are included in the carry and dollar factors. To save space, some country-level results are not reported. Finland, Portugal, and Spain are omitted because there are only 22 months of data for these countries. Belgium and the Netherlands are also omitted because the sample already contains many European countries (results on these countries are very similar to those of France and Germany).

or not a single currency pair, however, has little impact on the properties of the factors because a large sample of countries is used to build them. Excluding one currency does not mean that all relevant information is dropped. Assume that two foreign countries  $A$  and  $B$  decide to peg their currency to each other, then excluding  $A$  from the dollar and carry portfolios does not matter much since the same information is available in the exchange rate between country  $B$  and the U.S. For this reason, all the countries in the euro area are excluded after January 1999, keeping only the euro. But the objective of this paper is to highlight common components across currencies, so there would be no point in trying to exclude all the countries whose exchange rates might be correlated.

Second, portfolios always use the largest available sample of countries. Even when studying the bilateral changes in exchange rates of developed countries, portfolios and thus risk factors are derived from the large sample of developed and emerging countries. The average forward discount is obtained using all developed countries in the sample: Australia, Austria, Belgium, Canada, Denmark, Euro Area, Finland, France, Germany, Greece, Ireland, Italy, Japan, Netherlands, New Zealand, Norway, Portugal, Spain, Sweden, Switzerland, and United Kingdom.

Table 1: Interest Rates, Industrial Production, and Inflation

Country	Panel I: Interest Rates			Panel II: Industrial Production and Inflation					
	$\alpha$	$\beta$	$R^2$	$\alpha$	$\beta$	$\gamma$	$\delta$	$R^2$	$N$
Australia	0.16 (0.25)	-0.88 (0.68)	<b>0.08</b> [0.73]	0.24 (0.27)	-3.31 (1.83)	-5.37 (10.62)	81.59 (79.86)	<b>2.04</b> [3.27]	312
Canada	-0.03 (0.14)	-1.02 (0.74)	<b>0.21</b> [0.78]	-0.25 (0.23)	-2.16 (1.40)	-12.48 (8.85)	-67.84 (59.33)	<b>0.75</b> [1.71]	312
Denmark	-0.21 (0.19)	-0.26 (0.71)	<b>-0.27</b> [0.53]	-0.18 (0.17)	-0.33 (0.63)	-3.65 (6.16)	-91.06 (47.73)	<b>1.54</b> [2.12]	312
Euro Area	-0.18 (0.27)	-2.66 (2.22)	<b>0.55</b> [2.31]	-0.27 (0.33)	-0.65 (1.80)	-0.93 (8.34)	-192.36 (91.11)	<b>8.00</b> [6.10]	143
France	-0.19 (0.26)	-0.17 (1.14)	<b>-0.54</b> [1.20]	-0.09 (0.28)	-0.91 (0.85)	0.78 (10.56)	40.50 (88.83)	<b>-1.64</b> [2.96]	107
Germany	-0.23 (0.28)	0.43 (1.12)	<b>-0.43</b> [1.12]	-0.18 (0.36)	-0.00 (1.03)	-9.65 (9.05)	-15.22 (60.34)	<b>-1.05</b> [1.87]	181
Italy	-0.12 (0.34)	0.40 (1.01)	<b>-0.45</b> [2.07]	0.02 (0.24)	0.44 (1.23)	-6.55 (8.88)	35.18 (80.02)	<b>-0.93</b> [3.24]	177
Japan	-0.91 (0.26)	-2.42 (0.90)	<b>1.71</b> [1.59]	-1.40 (0.59)	-2.10 (1.15)	-13.46 (7.91)	-133.39 (64.78)	<b>3.07</b> [2.37]	325
New Zealand	0.21 (0.25)	-0.99 (0.49)	<b>0.91</b> [1.44]	-0.13 (0.22)	0.65 (0.78)	-11.39 (9.40)	-40.94 (44.19)	<b>-0.22</b> [1.51]	312
Norway	-0.23 (0.22)	0.44 (0.89)	<b>-0.17</b> [0.93]	-0.02 (0.14)	-0.41 (0.76)	-6.77 (7.86)	-39.77 (37.35)	<b>0.10</b> [1.68]	312
Sweden	-0.04 (0.21)	-0.35 (0.99)	<b>-0.23</b> [0.78]	-0.02 (0.16)	-0.92 (0.73)	-8.86 (7.16)	-84.85 (34.31)	<b>2.76</b> [2.31]	312
Switzerland	-0.35 (0.28)	-0.53 (1.07)	<b>-0.19</b> [0.77]	-0.59 (0.30)	-0.87 (0.71)	-7.98 (6.52)	-68.13 (53.33)	<b>0.82</b> [1.66]	325
United Kingdom	0.22 (0.23)	-1.39 (1.24)	<b>0.43</b> [1.43]	-0.19 (0.23)	-1.38 (1.27)	-11.68 (7.89)	-6.94 (32.75)	<b>0.42</b> [1.65]	325

Notes: Panel I reports country-level results from the following regression:

$$\Delta s_{t+1} = \alpha + \beta(i_t^* - i_t) + \varepsilon_{t+1},$$

where  $\Delta s_{t+1}$  denotes the bilateral exchange rate in foreign currency per U.S. dollar, and  $i_t^* - i_t$  is the interest rate difference between the foreign country and the U.S. Panel II reports country-level results from the following regression:

$$\Delta s_{t+1} = \alpha + \beta(i_t^* - i_t) + \gamma \Delta ip_{t+1} + \delta(\Delta cpi_{t+1}^* - \Delta cpi_{t+1}) + \varepsilon_{t+1},$$

where  $\Delta ip_{t+1}$  denotes the log difference in the industrial production index and  $\Delta cpi_{t+1}^* - \Delta cpi_{t+1}$  denotes the foreign minus domestic log difference in the consumer price indices. The table reports the constant  $\alpha$ , the slope coefficients  $\beta$ ,  $\gamma$ , and  $\delta$ , as well as the  $R^2$  of this regression (in percentage points), and the number  $N$  of observations. Standard errors in parentheses are Newey and West (1987) standard errors computed with the optimal number of lags according to Andrews (1991). The standard errors for the  $R^2$ s are reported in brackets; they are obtained by bootstrapping. Data are monthly, from Barclays and Reuters (Datastream). All variables are in percentage points. The sample period is 11/1983–12/2010.

Table 2: Carry and Dollar Factors: Monthly Tests in Developed Countries

Country	$\alpha$	$\beta$	$\gamma$	$\delta$	$\tau$	$R^2$	$R^2_{\$}$	$R^2_{no \$}$	$W$	$N$
Australia	0.07 (0.23)	-0.44 (0.60)	0.77 (0.49)	0.16 (0.13)	0.74 (0.13)	<b>25.59</b> (5.77)	20.05 [5.72]	7.71 [4.31]	***	312
Canada	-0.11 (0.11)	-0.02 (0.63)	-0.61 (0.42)	0.21 (0.06)	0.34 (0.07)	<b>19.38</b> (6.94)	13.11 [4.34]	8.14 [4.97]	***	312
Denmark	-0.01 (0.07)	-0.20 (0.38)	0.53 (0.13)	-0.16 (0.03)	1.51 (0.04)	<b>86.08</b> (1.67)	83.63 [2.03]	3.97 [3.99]	***	312
Euro Area	0.07 (0.11)	-0.52 (0.86)	0.10 (0.23)	-0.28 (0.05)	1.62 (0.08)	<b>80.60</b> (3.58)	76.22 [3.99]	-0.05 [4.81]	***	143
France	-0.15 (0.07)	-0.10 (0.34)	0.80 (0.14)	-0.13 (0.03)	1.38 (0.04)	<b>90.97</b> (1.48)	87.58 [1.93]	12.30 [5.90]	***	181
Germany	-0.21 (0.09)	-0.03 (0.34)	0.79 (0.17)	-0.03 (0.04)	1.42 (0.04)	<b>91.00</b> (1.36)	88.35 [1.75]	22.83 [6.20]	***	181
Italy	-0.03 (0.22)	0.26 (0.69)	0.68 (0.20)	-0.07 (0.11)	1.24 (0.10)	<b>68.97</b> (5.25)	64.59 [6.92]	2.16 [6.13]	***	177
Japan	-0.44 (0.24)	-1.13 (0.86)	-0.10 (0.45)	-0.39 (0.11)	0.83 (0.12)	<b>29.52</b> (5.51)	23.58 [5.45]	5.34 [3.47]	***	325
New Zealand	0.10 (0.20)	-0.58 (0.39)	0.76 (0.38)	-0.11 (0.11)	0.95 (0.11)	<b>29.80</b> (5.31)	26.96 [5.78]	3.43 [2.85]	*	312
Norway	-0.07 (0.12)	0.29 (0.37)	0.48 (0.11)	-0.06 (0.05)	1.35 (0.08)	<b>71.23</b> (3.99)	69.87 [3.98]	3.13 [3.36]	***	312
Sweden	0.06 (0.10)	-0.28 (0.35)	0.99 (0.16)	-0.06 (0.04)	1.39 (0.06)	<b>72.42</b> (2.90)	67.65 [3.41]	5.94 [3.46]	***	312
Switzerland	-0.14 (0.11)	-0.19 (0.41)	0.94 (0.19)	-0.11 (0.06)	1.46 (0.06)	<b>74.61</b> (2.45)	69.03 [2.98]	12.09 [3.70]	***	325
United Kingdom	0.06 (0.15)	-0.15 (0.71)	0.63 (0.47)	-0.03 (0.09)	1.06 (0.09)	<b>50.76</b> (5.09)	49.90 [5.29]	2.13 [3.01]		325

Notes: This table reports country-level results from the following regression:

$$\Delta s_{t+1} = \alpha + \beta(i_t^* - i_t) + \gamma(i_t^* - i_t)Carry_{t+1} + \delta Carry_{t+1} + \tau Dollar_{t+1} + \varepsilon_{t+1},$$

where  $\Delta s_{t+1}$  denotes the bilateral exchange rate in foreign currency per U.S. dollar, and  $i_t^* - i_t$  is the interest rate difference between the foreign country and the U.S.,  $Carry_{t+1}$  denotes the dollar-neutral average change in exchange rates obtained by going long a basket of high interest rate currencies and short a basket of low interest rate currencies, and  $Dollar_{t+1}$  corresponds to the average change in exchange rates against the U.S. dollar. The table reports the constant  $\alpha$ , the slope coefficients  $\beta$ ,  $\gamma$ ,  $\delta$ , and  $\tau$ , as well as the  $R^2$  of this regression (in percentage points) and the number of observations  $N$ . Standard errors in parentheses are Newey and West (1987) standard errors computed with the optimal number of lags according to Andrews (1991). The standard errors for the  $R^2$ s are reported in brackets; they are obtained by bootstrapping.  $R^2_{\$}$  denotes the  $R^2$  of a similar regression with only the  $Dollar$  factor (i.e., without the conditional and unconditional  $Carry$  factors).  $R^2_{no \$}$  denotes the  $R^2$  of a similar regression without the  $Dollar$  factor.  $W$  denotes the result of a Wald test: the null hypothesis is that the loadings  $\gamma$  and  $\delta$  on the conditional and unconditional carry factors are jointly zero. Three asterisks (\*\*\*) correspond to a rejection of the null hypothesis at the 1% confidence level; two asterisks and one asterisk correspond to the 5% and 10% confidence levels. Data are monthly, from Barclays and Reuters (Datastream). All variables are in percentage points. The sample period is 11/1983–12/2010.

Table 3: Carry and Dollar Factors: Monthly Tests in Emerging and Developing Countries

Country	$\alpha$	$\beta$	$\gamma$	$\delta$	$\tau$	$R^2$	$R^2_{\$}$	$R^2_{no \$}$	$W$	$N$
Hong Kong	-0.00 (0.01)	-0.15 (0.09)	0.06 (0.05)	0.00 (0.00)	0.02 (0.01)	<b>5.40</b> (3.32)	4.85 [3.10]	1.29 [2.29]		325
Czech Republic	-0.14 (0.17)	-0.11 (0.35)	-0.04 (0.16)	-0.21 (0.09)	1.76 (0.09)	<b>64.09</b> (4.71)	62.28 [4.64]	-0.62 [2.34]	**	167
Hungary	0.39 (0.38)	-0.35 (0.57)	-0.40 (0.18)	0.18 (0.15)	1.86 (0.14)	<b>67.69</b> (5.09)	67.14 [4.89]	1.17 [4.54]	**	158
India	0.31 (0.24)	-0.57 (0.66)	0.22 (0.29)	0.03 (0.11)	0.49 (0.07)	<b>31.38</b> (7.05)	30.72 [6.59]	7.61 [5.80]		158
Indonesia	1.93 (1.31)	-1.21 (1.41)	0.21 (0.44)	0.22 (0.44)	1.75 (0.50)	<b>9.75</b> (7.14)	10.80 [5.88]	1.72 [6.22]		90
Kuwait	-0.16 (0.03)	2.17 (0.19)	0.53 (0.10)	-0.09 (0.02)	0.22 (0.04)	<b>52.24</b> (11.14)	44.45 [10.00]	25.66 [14.37]	***	167
Malaysia	0.09 (0.13)	0.10 (0.53)	0.10 (0.23)	0.19 (0.10)	0.42 (0.07)	<b>23.04</b> (5.19)	18.17 [4.57]	6.40 [5.22]		230
Mexico	0.40 (0.28)	-0.36 (0.36)	-0.29 (0.15)	0.68 (0.16)	0.22 (0.15)	<b>26.09</b> (8.44)	9.11 [6.94]	24.48 [8.19]	***	167
Philippines	0.13 (0.37)	-0.02 (0.88)	0.63 (0.21)	-0.01 (0.10)	0.47 (0.10)	<b>32.59</b> (7.79)	19.48 [6.35]	23.92 [8.63]	***	167
Poland	-0.08 (0.20)	1.09 (0.71)	1.13 (0.30)	0.10 (0.08)	1.89 (0.11)	<b>74.77</b> (5.43)	70.73 [6.09]	18.44 [8.37]	***	106
Saudi Arabia	0.00 (0.01)	-0.39 (0.35)	0.18 (0.10)	-0.00 (0.00)	0.00 (0.00)	<b>8.57</b> (11.24)	2.83 [8.18]	8.84 [10.84]		167
Singapore	-0.17 (0.11)	-0.29 (0.60)	0.12 (0.15)	0.08 (0.03)	0.50 (0.04)	<b>48.19</b> (4.19)	47.19 [4.38]	6.29 [4.05]	*	312
South Africa	0.87 (0.51)	-0.58 (0.79)	0.04 (0.37)	0.18 (0.28)	1.07 (0.14)	<b>24.87</b> (5.50)	24.14 [5.66]	2.36 [2.44]		324
South Korea	0.27 (0.27)	0.60 (1.71)	0.62 (0.49)	0.14 (0.11)	1.38 (0.27)	<b>51.83</b> (6.21)	51.30 [5.99]	13.63 [9.19]		106
Taiwan	0.05 (0.12)	0.45 (0.31)	0.29 (0.13)	0.08 (0.06)	0.50 (0.06)	<b>35.77</b> (5.41)	34.39 [6.11]	6.94 [5.19]	**	167
Thailand	-0.07 (0.18)	-0.36 (1.16)	0.88 (0.43)	-0.01 (0.12)	0.79 (0.17)	<b>27.98</b> (5.82)	19.20 [5.63]	13.50 [7.29]		167
Turkey	-0.71 (0.39)	0.69 (0.11)	-0.19 (0.04)	1.12 (0.25)	0.65 (0.17)	<b>39.03</b> (8.08)	27.34 [8.00]	32.80 [7.26]	***	154
United Arab Emirates	-0.00 (0.00)	-0.22 (0.14)	0.10 (0.07)	-0.00 (0.00)	0.00 (0.00)	<b>15.10</b> (19.30)	3.32 [12.36]	15.39 [19.27]		162

Notes: This table reports country-level results from the following regression:

$$\Delta s_{t+1} = \alpha + \beta(i_t^* - i_t) + \gamma(i_t^* - i_t)Carry_{t+1} + \delta Carry_{t+1} + \tau Dollar_{t+1} + \varepsilon_{t+1},$$

where  $\Delta s_{t+1}$  denotes the bilateral exchange rate in foreign currency per U.S. dollar, and  $i_t^* - i_t$  is the interest rate difference between the foreign country and the U.S.,  $Carry_{t+1}$  denotes the dollar-neutral average change in exchange rates obtained by going long a basket of high interest rate currencies and short a basket of low interest rate currencies, and  $Dollar_{t+1}$  corresponds to the average change in exchange rates against the U.S. dollar. The table reports the constant  $\alpha$ , the slope coefficients  $\beta$ ,  $\gamma$ ,  $\delta$ , and  $\tau$ , as well as the  $R^2$  of this regression and the number of observations  $N$ .  $R^2_{\$}$  ( $R^2_{no \$}$ ) denotes the  $R^2$  of a similar regression with only (without) the  $Dollar$  factor.  $W$  denotes the result of a Wald test on the joint significance of  $\gamma$  and  $\delta$ . See Table 2 for additional information.

Table 4: Portfolios of Countries Sorted By Dollar Exposures

Panel I: Summary Statistics						
<i>Portfolio</i>	1	2	3	4	5	6
Spot change: $\Delta s$						
<i>Mean</i>	−0.97	−2.12	−2.88	−3.66	−2.99	−5.07
<i>Std</i>	3.29	5.31	6.70	7.72	10.19	10.68
Forward Discount: $f - s$						
<i>Mean</i>	0.34	0.74	0.99	1.47	2.00	2.07
<i>Std</i>	0.54	1.11	1.24	1.44	0.70	0.55
Excess Return: $rx$						
<i>Mean</i>	1.31	2.86	3.87	5.13	4.99	7.14
	[0.70]	[1.17]	[1.41]	[1.61]	[2.16]	[2.18]
<i>Std</i>	3.34	5.38	6.68	7.62	10.20	10.64
<i>SR</i>	0.39	0.53	0.58	0.67	0.49	0.67
Excess Return: $rx$ (with bid-ask spreads)						
<i>Mean</i>	0.58	1.43	2.11	3.73	3.73	5.84
	[0.72]	[1.11]	[1.40]	[1.61]	[2.05]	[2.37]
Panel II: Risk Prices						
	$\lambda_{Cond.Dollar}$	$b_{Cond.Dollar}$	$R^2$	$RMSE$	$\chi^2$	
<i>GMM</i> <sub>1</sub>	4.73	0.94	83.06	0.80		
	[1.54]	[0.31]			66.57	
<i>GMM</i> <sub>2</sub>	4.51	0.90	81.74	0.83		
	[1.50]	[0.30]			66.91	
<i>FMB</i>	4.73	0.94	85.22	0.80		
	[1.41]	[0.28]			50.40	
	[1.41]	[0.28]			52.96	
<i>Mean</i>	4.61					
Panel III: Factor Betas						
<i>Portfolio</i>	1	2	3	4	5	6
$\alpha$	0.81	0.87	0.64	0.76	−1.17	0.44
	[0.90]	[1.00]	[1.06]	[0.91]	[0.99]	[0.90]
$\beta$	0.11	0.44	0.71	0.99	1.40	1.52
	[0.03]	[0.06]	[0.06]	[0.06]	[0.06]	[0.05]
$R^2$	4.40	28.98	48.00	71.64	78.97	86.39

*Notes:* Panel I reports summary statistics on portfolios of currencies sorted on their exposure to the dollar factor. See Section 3 for details on the construction of these portfolios. The table reports, for each portfolio, the mean and standard deviations of the average change in log spot exchange rates  $\Delta s$ , the average log forward discount  $f - s$ , and the average log excess return  $rx$  without bid-ask spreads. All moments are annualized and reported in percentage points. For excess returns, the table also reports Sharpe ratios, computed as ratios of annualized means to annualized standard deviations and the mean excess returns net of bid-ask spreads. Panel II reports results from GMM and Fama-McBeth asset pricing procedures. The market price of risk  $\lambda$ , the adjusted  $R^2$ , the square-root of mean-squared errors  $RMSE$  and the  $p$ -values of  $\chi^2$  tests on pricing errors are reported in percentage points.  $b$  denotes the vector of factor loadings ( $m_{t+1} = 1 - b_{Cond.Dollar}_{t+1}$ ). The last row reports the mean of the risk factor. Excess returns used as test assets and risk factors do not take into account bid-ask spreads. All excess returns are multiplied by 12 (annualized). Shanken (1992)-corrected standard errors are reported in parentheses. The second step of the FMB procedure does not include a constant. Panel III reports OLS estimates of the factor betas.  $R^2$ s and  $p$ -values are reported in percentage points. The standard errors in brackets are Newey and West (1987) standard errors computed with the optimal number of lags according to Andrews (1991). The alphas are annualized and in percentage points. Data are monthly, from Barclays and Reuters in Datastream. The sample period is 12/1988–12/2010.

Table 5: Shares of Currency Systematic Variation and World Comovement

	Equity 11/1983–12/2010	Equity 1/1999–12/2010	Bonds 11/1983–12/2010	Bonds 1/1999–12/2010	GDP 1983–2010	Consumption 1983–2010
Panel I: Constant $R^2$ s, series in U.S. dollars						
$\beta$	0.77	0.65	1.10	0.66	0.73	0.63
s.e	[0.28]	[0.19]	[0.16]	[0.12]	[0.15]	[0.29]
$R^2$	19.55	31.86	76.28	72.54	47.82	16.26
$N$	33	28	17	13	29	27
Panel II: Constant $R^2$ s, series in PPP dollars						
$\beta$	0.56	0.50	1.11	0.38	0.73	0.67
s.e	[0.35]	[0.23]	[0.28]	[0.21]	[0.22]	[0.28]
$R^2$	7.61	15.54	50.84	22.04	27.98	18.36
$N$	33	28	17	13	30	27
Panel III: Time-Varying $R^2$ s, series in U.S. dollars						
$\beta$	0.47	0.30	0.81	0.55		
s.e	[0.07]	[0.07]	[0.07]	[0.11]		
s.e	(0.07)	(0.12)	(0.10)	(0.13)		
<i>trend</i>	2.44	0.17	1.22	-1.03		
$R^2$	79.75	82.72	90.80	90.97		
$N$	5253	2915	3085	1578		
Panel IV: Time-Varying $R^2$ s, series in PPP dollars						
$\beta$	0.25	0.28	0.41	0.04		
s.e	[0.08]	[0.08]	[0.11]	[0.10]		
s.e	(0.10)	(0.11)	(0.14)	(0.12)		
<i>trend</i>	3.57	0.38	2.51	1.02		
$R^2$	76.74	82.35	77.41	86.12		
$N$	5253	2915	3085	1578		

Notes: Panels I and II report results from the following cross-country regressions:

$$R_i^{2,FX} = \alpha + \beta R_i^{2,X} + \varepsilon_i,$$

where  $R_i^{2,FX}$  denotes the share of systematic variation in the exchange rate of country  $i$ .  $R_i^{2,X}$  are obtained in tests of world integration, using either equity returns, bond returns, GDP growth, or consumption growth. Panels III and IV report results from the following cross-country panel regressions with country and time fixed effects:

$$R_{i,t}^{2,FX} = \alpha^i + \beta R_{i,t}^{2,X} + \gamma trend + \varepsilon_{i,t},$$

where  $R_{i,t}^{2,FX}$  and  $R_{i,t}^{2,X}$  are estimated on rolling windows of 60 months. Panels I and III use financial and macroeconomic series expressed in U.S. dollars. Panels II and IV use series expressed in purchasing power parity (PPP) dollars. The table reports the slope coefficient  $\beta$ , the standard errors, the cross-sectional  $R^2$  (in percentage points), as well as the number of observations  $N$  (i.e., countries, or country-month pairs). For the panel regressions, the table also reports the  $t$ -statistic on the common time trend.  $R^2$ s on financial markets (equity, bonds, and currencies) are obtained on monthly series. The sample periods are 11/1983–12/2010 and 1/1999–12/2010.  $R^2$ s on macroeconomic variables (consumption and output) are obtained on annual series. The sample period is 1983–2010. The standard errors (s.e.) in the panel regressions are obtained from the Newey-West autocorrelation consistent covariance estimator computed with a horizon of 60 months (in brackets) and from clustering by country (in parenthesis).



Table 6: Monthly Shares of Global Shocks in Bilateral Exchange Rates

Country	$\alpha$	$\gamma$	$\delta$	$\tau$	$R^2$	$R^2_{\$}$	$R^2_{Global\$}$	$W$	$N$
Australia	-0.08 (0.18)	0.07 (0.64)	0.32 (0.15)	0.36 (0.10)	<b>18.61</b> (7.33)	34.87 [7.05]	13.42 [5.99]	***	266
Canada	-0.09 (0.11)	-1.01 (0.52)	0.23 (0.06)	0.21 (0.06)	<b>17.68</b> (8.50)	25.74 [7.53]	8.58 [4.76]	***	266
Denmark	0.11 (0.08)	-0.00 (0.12)	0.03 (0.04)	0.87 (0.03)	<b>80.55</b> (3.21)	84.56 [2.00]	80.65 [3.08]		266
Euro Area	0.14 (0.10)	-0.35 (0.20)	-0.15 (0.06)	0.89 (0.04)	<b>82.62</b> (3.34)	82.80 [3.42]	81.72 [3.74]	**	143
France	0.05 (0.10)	-0.04 (0.15)	0.15 (0.04)	0.88 (0.06)	<b>81.56</b> (4.83)	88.87 [2.44]	80.31 [4.92]	***	122
Germany	0.08 (0.11)	-0.15 (0.18)	0.13 (0.04)	0.92 (0.06)	<b>81.92</b> (4.66)	88.89 [2.16]	80.85 [4.80]	***	122
Italy	0.26 (0.18)	0.66 (0.20)	0.19 (0.07)	0.69 (0.06)	<b>64.55</b> (5.54)	64.93 [6.56]	51.34 [9.01]	***	122
Japan	-0.02 (0.19)	-0.41 (0.58)	-0.32 (0.15)	0.42 (0.11)	<b>14.92</b> (5.65)	26.78 [6.44]	13.17 [5.23]	**	266
New Zealand	-0.01 (0.17)	-0.17 (0.60)	0.20 (0.17)	0.50 (0.08)	<b>25.38</b> (6.23)	43.64 [5.27]	24.45 [5.99]		266
Norway	0.07 (0.10)	0.20 (0.13)	0.15 (0.05)	0.78 (0.06)	<b>67.64</b> (5.86)	72.08 [4.43]	65.82 [6.07]	***	266
Sweden	0.16 (0.11)	0.33 (0.23)	0.17 (0.05)	0.83 (0.04)	<b>66.96</b> (4.73)	73.31 [3.05]	64.63 [5.34]	***	266
Switzerland	0.03 (0.11)	0.29 (0.29)	-0.05 (0.07)	0.85 (0.06)	<b>68.11</b> (3.94)	74.85 [2.75]	67.61 [4.16]		266
United Kingdom	0.13 (0.12)	0.67 (0.53)	0.09 (0.08)	0.56 (0.06)	<b>44.92</b> (5.66)	48.47 [6.62]	41.84 [6.27]	***	266

Notes: This table reports country-level results from the following regression:

$$\Delta s_{t+1} = \alpha + \gamma(i_t^* - i_t)Carry_{t+1} + \delta Carry_{t+1} + \tau Dollar_{t+1}^{global} + \varepsilon_{t+1},$$

where  $\Delta s_{t+1}$  denotes the bilateral exchange rate in foreign currency per U.S. dollar, and  $i_t^* - i_t$  is the interest rate difference between the foreign country and the U.S.,  $Carry_{t+1}$  denotes the dollar-neutral average change in exchange rates obtained by going long a basket of high interest rate currencies and short a basket of low interest rate currencies, and  $Dollar_{t+1}^{global}$  corresponds to the change in exchange rates in a high dollar-beta portfolio minus the change in exchange rates in a low dollar-beta portfolio. See Section 3 for details on the construction of these portfolios. The table reports the constant  $\alpha$ , the slope coefficients  $\beta$ ,  $\gamma$ ,  $\delta$ , and  $\tau$ , as well as the  $R^2$  of this regression (in percentage points) and the number of observations  $N$ . Standard errors in parentheses are Newey and West (1987) standard errors computed with the optimal number of lags according to Andrews (1991). The standard errors for the  $R^2$ s are reported in brackets; they are obtained by bootstrapping.  $R^2_{Global\$}$  denotes the  $R^2$  of a similar regression with only the  $Dollar^{global}$  factor (i.e., without the conditional and unconditional  $Carry$  factors).  $R^2_{\$}$  denotes the  $R^2$  of a similar regression using the same  $Carry$  factors, along with the  $Dollar$  factor.  $W$  denotes the result of a Wald test: the null hypothesis is that the loadings  $\gamma$  and  $\delta$  on the conditional and unconditional carry factors are jointly zero. Three asterisks (\*\*\*) correspond to a rejection of the null hypothesis at the 1% confidence level; two asterisks and one asterisk correspond to the 5% and 10% confidence levels. Data are monthly, from Barclays and Reuters (Datastream). All variables are in percentage points. The sample period is 11/1983–12/2010.

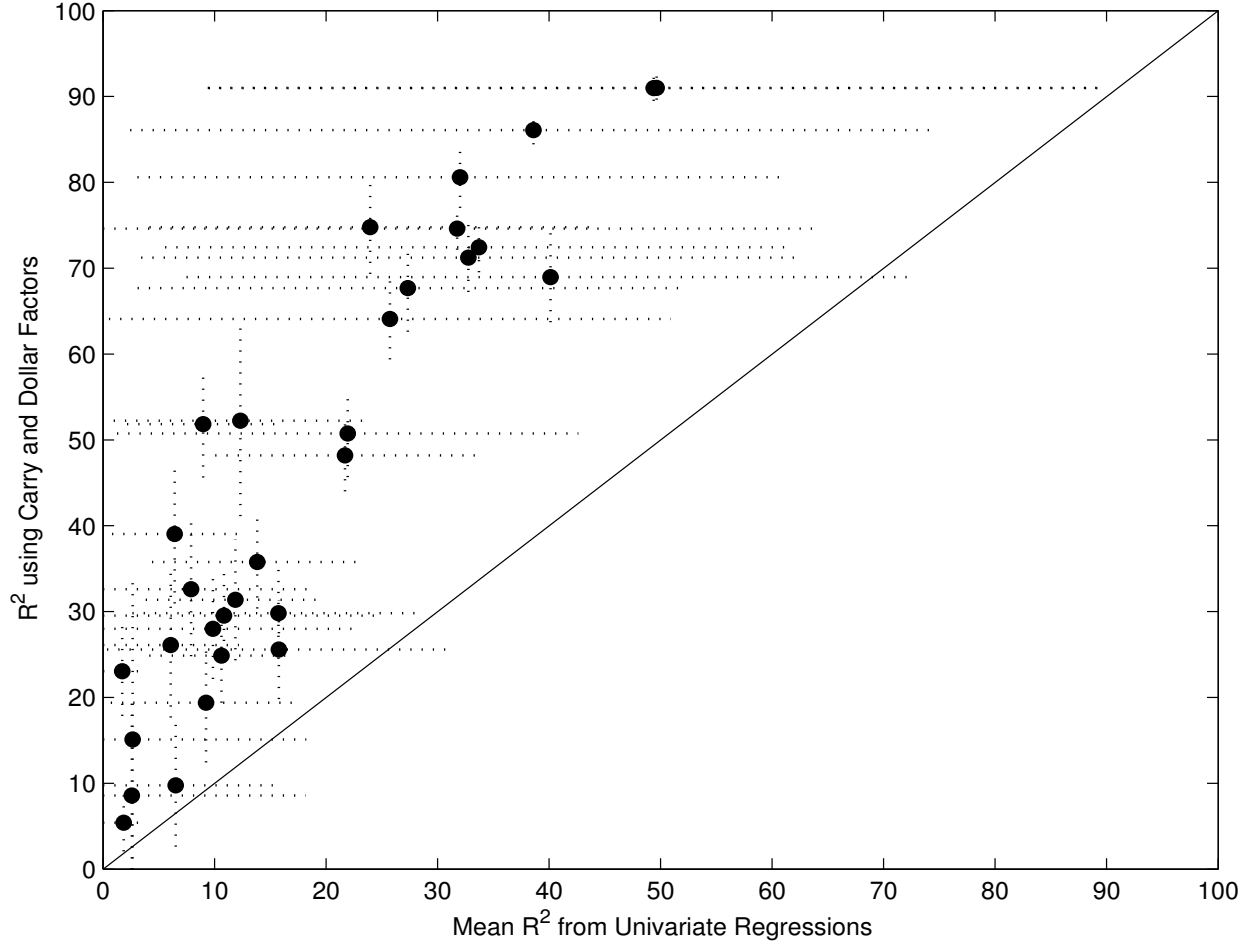


Figure 1: Measuring Systematic Risk with Factors vs. Individual Exchange Rates

The figure compares  $R^2$ s obtained with the carry and dollar factors (vertical axis) to those obtained from random univariate regressions of one exchange rate changes on others (horizontal axis).

Adjusted  $R^2$ s on the vertical axis correspond to the following regressions:

$$\Delta s_{t+1} = \alpha + \beta(i_t^* - i_t) + \gamma(i_t^* - i_t)Carry_{t+1} + \delta Carry_{t+1} + \tau Dollar_{t+1} + \varepsilon_{t+1},$$

where  $\Delta s_{t+1}$  denotes the bilateral exchange rate in foreign currency per U.S. dollar,  $i_t^* - i_t$  denotes the interest rate difference,  $Carry_{t+1}$  denotes the dollar-neutral average change in exchange rates obtained by going long a basket of high interest rate currencies and short a basket of low interest rate currencies, and  $Dollar_{t+1}$  corresponds to the average change in exchange rates against the U.S. dollar. Dots correspond to point estimates, while dotted lines represent confidence intervals (defined as one-standard error above and below the point estimates). Standard errors are obtained by bootstrapping.

Adjusted  $R^2$ s on the horizontal axis correspond to the following experiment. Each currency  $i$  is regressed (in log changes) on a constant and another currency  $j$  ( $i \neq j$ ):

$$\Delta s_{t+1}^i = \alpha + \beta \Delta s_{t+1}^j + \varepsilon_{t+1}^{i,j}.$$

Each regression delivers a pair-specific adjusted R-square, denoted  $R^{2^{i,j}}$ . For each currency  $i$ , the figure reports the mean of all  $R^{2^{i,j}}$  for  $j \neq i$ . Dots correspond to the mean estimates, while dotted lines represent confidence intervals (defined as one-standard deviation above and below the mean estimates). Data are monthly. The sample period is 11/1983–12/2010.

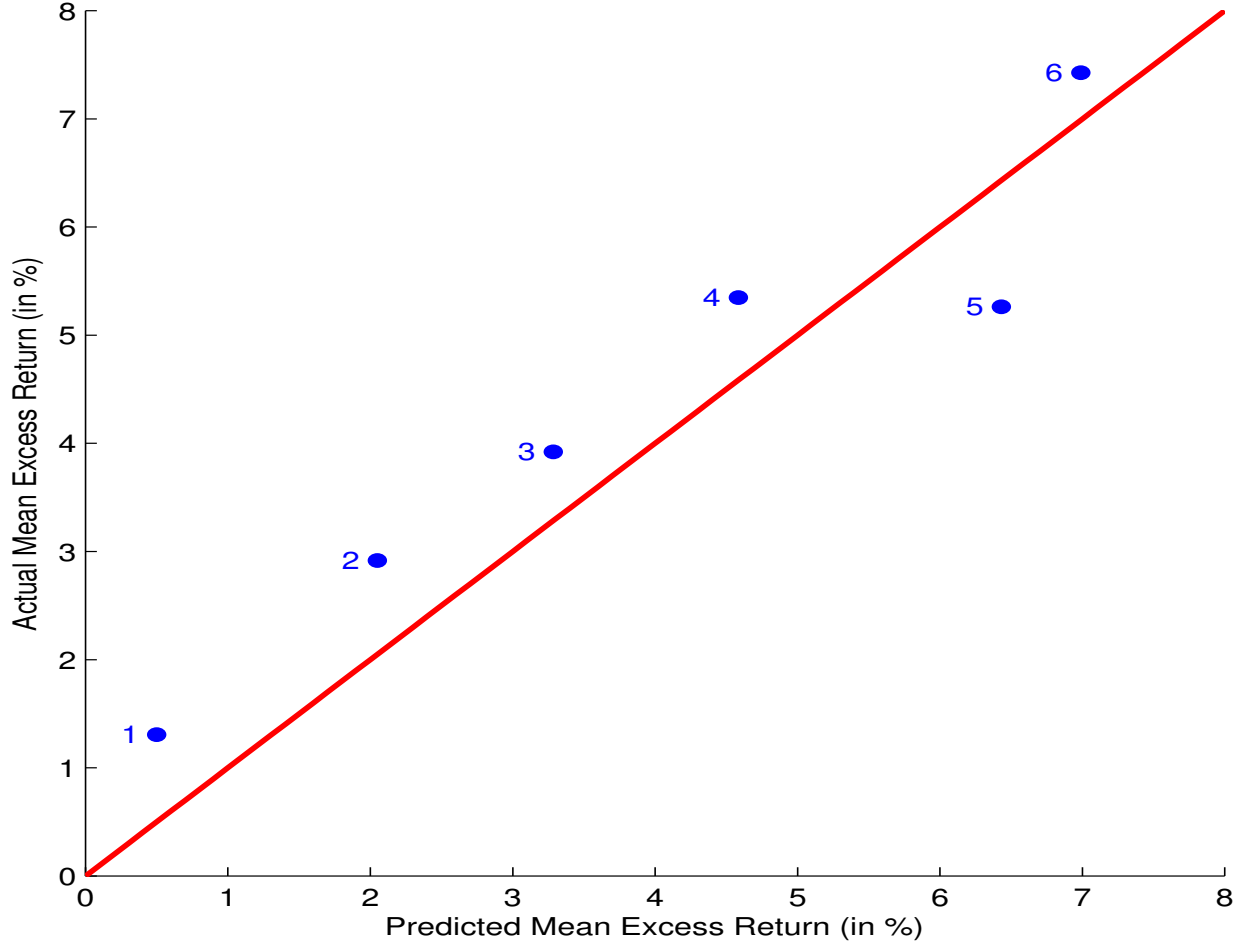


Figure 2: Realized vs. Predicted Excess Returns: Portfolios of Countries Sorted on Dollar Exposures

The figure plots realized average excess returns on the vertical axis against predicted average excess returns on the horizontal axis. The portfolios are based on each currency's exposure to the dollar factor. At each date  $t$ , each currency  $i$  change in exchange rate is regressed on a constant and the dollar and carry factors using a 60-month rolling window that ends in period  $t - 1$ . Currency  $i$ 's exposure to the *Dollar* factor is denoted  $\tau_t^i$ . Currencies are then sorted into six groups at time  $t$  based on the slope coefficients  $\tau_t^i$ . Portfolio 1 contains currencies with the lowest taus. Portfolio 6 contains currencies with the highest taus. At each date  $t$  and for each portfolio, the investor goes long if the average forward discount is positive and short otherwise. Each portfolio  $j$ 's actual excess return is regressed on a constant and the conditional dollar excess return to obtain the slope coefficient  $\beta^j$ . Each predicted excess return then corresponds to the OLS estimate  $\beta^j$  multiplied by the mean of the conditional dollar excess return. All returns are annualized. Data are monthly. The sample period is 12/1988–12/2010.

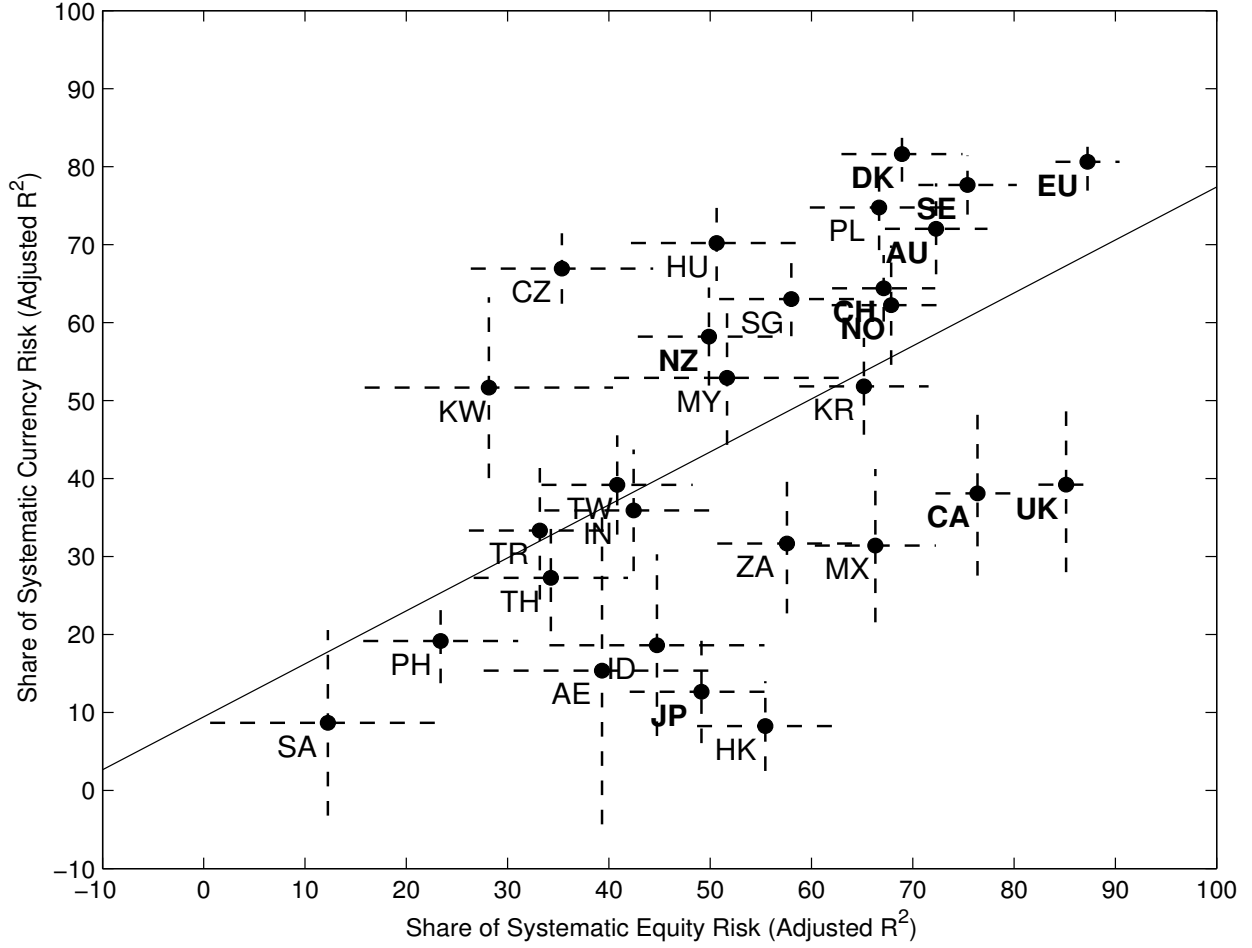


Figure 3: Systematic Equity and Currency Risk

The figure plots adjusted  $R^2$ s on currency markets as a function of adjusted  $R^2$ s on equity markets. Dots correspond to point estimates, while dotted lines represent confidence intervals (defined as one-standard error above and below the point estimates). Standard errors are obtained by bootstrapping. Adjusted  $R^2$ s on currency markets are obtained from the following regressions:

$$\Delta s_{t+1} = \alpha + \beta(i_t^* - i_t) + \gamma(i_t^* - i_t)Carry_{t+1} + \delta Carry_{t+1} + \tau Dollar_{t+1} + \varepsilon_{t+1},$$

where  $\Delta s_{t+1}$  denotes the bilateral exchange rate in foreign currency per U.S. dollar,  $i_t^* - i_t$  denotes the interest rate difference,  $Carry_{t+1}$  denotes the dollar-neutral average change in exchange rates obtained by going long a basket of high interest rate currencies and short a basket of low interest rate currencies, and  $Dollar_{t+1}$  corresponds to the average change in exchange rates against the U.S. dollar. Adjusted  $R^2$ s on equity markets are derived from:

$$r_{t+1}^{m,\$} = \alpha + \beta r_{t+1}^{m,world,\$} + \gamma r_{t+1}^{hml,world,\$} + \varepsilon_{t+1},$$

where  $r_{t+1}^{m,\$}$  denotes the returns on a foreign country's MSCI stock market index in U.S. dollars,  $r_{t+1}^{m,world}$  corresponds to returns on the MSCI world equity index in U.S. dollars, and  $r_{t+1}^{hml,world,\$}$  is the difference between returns on the world MSCI value equity index and the world MSCI growth equity index in U.S. dollars (i.e., high minus low book-to-market equity returns). Data are monthly. The sample is 1/1999–12/2010. The country codes correspond to the international standard: Australia (AU), Canada (CA), Hong Kong (HK), Czech Republic (CZ), Denmark (DK), Finland (FI), Hungary (HU), India (IN), Indonesia (ID), Japan (JP), Kuwait (KW), Malaysia (MY), Mexico (MX), New Zealand (NZ), Norway (NO), Philippines (PH), Poland (PL), Saudi Arabia (SA), Singapore (SG), South Africa (ZA), South Korea (KR), Sweden (SE), Switzerland (CH), Taiwan (TW), Thailand (TH), Turkey (TR), United Kingdom (UK), as well as the euro area (EU).